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# Walleye Pollock (Theragra chalcogramma) stock assessment for British Columbia in 2017 

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## Foreword

This series documents the scientific basis for the evaluation of aquatic resources and ecosystems in Canada. As such, it addresses the issues of the day in the time frames required and the documents it contains are not intended as definitive statements on the subjects addressed but rather as progress reports on ongoing investigations.

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#### Abstract

A new stock assessment is presented for two British Columbia (BC) stocks of Walleye Pollock (Theragra chalcogramma, WAP), with the BC North stock encompassing the three most northerly Pacific Marine Fisheries Commission (PMFC) major areas (5C, 5D, 5E) and the BC South stock including the remaining four outside PMFC major areas (3C, 3D, 5A, 5B plus minor areas $12 \& 20)$. These stock definitions were selected on the basis of a difference in observed mean weights, with BC North mean weights estimated near $1.0 \mathrm{~kg} / \mathrm{fish}$ while the equivalent BC South mean weights averaged near $0.5 \mathrm{~kg} /$ fish. A delay-difference production model was used to assess each stock, using data from fishery-independent surveys, a CPUE series derived from commercial bottom trawl catch rates, and an annual mean weight series derived from unsorted commercial catch samples. Because there are no useable BC ageing data, we used survey age samples from the Gulf of Alaska (GoA) to specify growth for the BC North stock. The BC South stock proved more problematic, with the GoA growth model unable to fit the BC South observed mean weights, eventually requiring us to use a published WAP growth model from the Asian Sea of Okhotsk. Each stock assessment explored a range of plausible natural mortality values as well as a range of ages for the knife-edge selectivity assumption because the biomass indices and the mean weight data used in the delaydifference model were not informative for these parameters. The stock assessment was conducted in a Bayesian framework, where the best fit to the data was used as the starting point for a search across the joint posterior parameter distributions using the Monte Carlo Markov Chain (MCMC) procedure. Twelve runs were made for the BC North stock and 11 for the BC South stock, with each run consisting of 60 million MCMC iterations, sampling every $50,000^{\text {th }}$ iteration, discarding the first 200 draws for burn-in, leaving 1,000 draws to comprise the posterior. Composite reference (model averaged) scenarios were used to represent each stock, with the model average for both stocks consisting of eight model runs selected on the basis of a subjective evaluation of the quality of the MCMC posterior. Each composite reference scenario included three values for instantaneous natural mortality ( $M=0.25,0.30,0.35$ ) and covered two or three ages at which knife-edge recruitment $(k)$ to the fishery occurred ( $k=3,4$ in BC North and including $k=5$ in BC South). The MCMC posteriors for the two composite scenarios were constructed by pooling the 1000 MCMC samples from each of the selected runs to give a posterior of 3,000 samples for $B C$ North and 6,000 samples for $B C$ South, thus giving equal weight to each run. The composite reference scenario was evaluated against historical reference points (HRPs) based on the reconstructed spawning biomass trajectory due to concerns about the stability of estimating $B_{0}$ and $B_{2017}$. The HRP $B_{\text {avg }}$, the average spawning biomass from 1967-2016, was used as a proxy for $B_{\mathrm{MSY}}$, and $B_{\text {min }}$, the minimum spawning biomass from which it subsequently recovered to $B_{\text {avg }}$, was used in place of $0.4 B_{\text {msry }}$. Twice $B_{\text {min }}$ was used in place of $0.8 B_{\text {msy }}$. The average exploitation rate over the period 1967-2016 ( $u_{\text {avg }}$ ) was used in place of $u_{\text {Msr }}$. The biomass at the beginning of 2017 for the model average BC North stock was evaluated as being primarily above the USR while the 2017 beginning year biomass for the BC South stock was evaluated as being entirely above the USR. For each stock, the assessment provides a decision table which evaluates the probability of the model average case staying above five reference points across a wide range of 22 constant catches. However, the paper warns that the probabilities in these decision tables should be viewed cautiously as the delay-difference model used in this stock assessment is not capable of making reliable multi-year projections because it has no latent age structure to inform predictions and the stock-recruitment function is poorly determined.


## 1. INTRODUCTION

Walleye Pollock (Theragra chalcogramma, acronym WAP) shares an evolutionary lineage with the Pacific, Atlantic, and Greenland Cods (Gadus macrocephalus, G. morhua, and G. ogac, respectively). The Alaskan stock assessment group currently uses the 2013 taxonomic nomenclature change for this species, adopting the new name Gadus chalcogrammus (masculine genus name) over Theragra chalcogramma (feminine). In this assessment, we use Theragra chalcogramma.

Walleye Pollock's appearance features an olive green to brown colour dorsally, often mottled or blotched; silvery on its sides, and pale ventrally; the fins are darker (Cohen et al. 1990). In Alaska, the species has been observed to reach a maximum length of 105 cm , a maximum weight of 6.05 kg , and a maximum age of 22 years (NOAA 2010). In British Columbia (BC), the observed maximum length was also 105 cm for a male specimen caught off the southeast coast of Moresby Island (PMFC major 5C, minor 6, locality 2) at 94 m in June 2015 (DFO GFBio database, accessed Dec 13, 2016).

This species is a semi-demersal ( $30-400 \mathrm{~m}$ ) schooling fish that performs diurnal vertical migrations and becomes increasingly associated with the bottom as the fish ages (Cohen et al. 1990, NOAA 2010). It reaches first maturity at $3-4$ years ( 30 to 38 cm total length); fecundity varies with age (Cohen et al. 1990). Fish congregate in dense schools to spawn, usually at 50250 m depth; spawning in the Strait of Georgia occurs from January to March (Cohen et al. 1990). In BC, Walleye Pollock prey upon shrimps, sand lance, and herring (Cohen et al. 1990); predation on juvenile pollock (cannibalism) and other bony fish is reported in the eastern Bering Sea (NOAA 2010).

The Alaskan Walleye Pollock fisheries support the largest landings of any single species in the USA, and possibly in the world. The highest concentrations of pollock occur in the Eastern Bering Sea (annual catches > 1 million tonnes, lanelli et al. 2015) with lesser amounts caught in the Gulf of Alaska (annual catches $>100,000$ tonnes, Dorn et al. 2015). These Alaskan fisheries are supported by complex age-structured stock assessment models, reflecting a strong investment by government in acoustic and trawl surveys, catch sampling and an intensive ageing program. In BC, the annual 2016 coastwide TAC was 4225 t comprising 1320 t in PMFC 5CDE, 1790 t in 5AB, and 1115 t in the Strait of Georgia. No TAC has ever been set for 3CD other than a 270 t quota set in 1997 only Appendix A. In this assessment, we consider two stocks: "BC North" (PMFC 5CDE) and "BC South" (PMFCs 5AB+3CD + minor areas 12 \& 20), for reasons outlined below.
The last stock assessment of Walleye Pollock (WAP) was conducted in 1997 (Saunders and Andrews $1998{ }^{1}$ ). This was a qualitative assessment that did not estimate stock status with respect to established reference points, partially because the DFO Precautionary Approach reference points had not been set. Harvest advice is required for this species to determine if current harvest levels are sustainable and are compliant with the DFO Fishery Decision-Making Framework Incorporating the Precautionary Approach (DFO 2006, 2009). A request was submitted to Science by staff in the Fisheries and Aquaculture Management's Groundfish Management Unit (GMU) in 2013 but was considered unachievable due to data limitations. In 2016, a review of the available data suggested that Science could potentially assess Walleye Pollock using a delay-difference model similar to the approach used to assess Shortspine

[^0]Thornyhead (Sebastolobus alascanus) in 2016 (Starr and Haigh 2017). Initial runs treated the entire offshore region of BC as a single stock (excluding the Strait of Georgia, Figure 1); however, after some exploration, it became clear that a more tenable hypothesis was that there were at least two separate stocks, given the large difference in observed mean weight in the respective fisheries: BC North ( $1.056 \mathrm{~kg} / \mathrm{fish}$ ) vs. BC South ( $0.521 \mathrm{~kg} / \mathrm{fish}$ ).


Figure 1. Walleye Pollock assessment areas comprising Pacific Marine Fisheries Commission (PMFC) major and minor areas - green for 5CDE, orange for 5AB + minor area 12, blue for 3CD + minor area 20, and red for 4B less minor areas 12 and 20. The Groundfish Management Unit area boundaries, which differ from PMFC area, are superimposed for comparison. This assessment is for areas called 'BC North' (5CDE, green) and 'BC South' (5AB3CD, orange + blue).

The modelling of these stocks presented many difficulties, primarily associated with a scarcity of usable data. In particular, there were almost no age data for BC pollock with the small amount of available data determined using an ageing methodology known to be biased. We were forced to go outside of BC to obtain usable age-length data, using survey derived data from the eastern Gulf of Alaska (Martin Dorn, pers. comm.) for BC North (see Appendix D). Suitable growth data for BC South pollock proved harder to find, requiring us to use a growth model from the Sea of Okhotsk (Janusz and Horbowy 1997), which lies between the eastern Russian
mainland and Kamchatka Peninsula, just west of the Bering Sea. A further problem was the generation of unstable reference points based on maximum sustainable yield (MSY) calculations. This issue necessitated the adoption of historical-based reference points (HRPs), specifically the average estimated spawning biomass over the full time period of the model (1967-2016), denoted $B_{\text {avg }}$ and a limit reference point (LRP) set to the minimum estimated spawning biomass in a year from which the biomass trajectory recovered to above $B_{\text {avg. }}$. The mode of the posterior distribution (MPD) for this year was $B_{2001}$ for BC North and $B_{2008}$ for BC South. The upper stock reference (USR) was set to twice the LRP. The advice is provided in decision tables that give the probability that $B_{2018,2019}$ (1- and 2-year projected spawning biomass) will be greater than the various reference points (Appendix F).

### 1.1. RANGE AND DISTRIBUTION

Walleye Pollock occur along the North Pacific rim, ranging from the Sea of Japan, extending north into Russian and Alaskan waters, then down through BC and southward to southern California (FAO Aquatic Species Distribution Map Viewer, accessed May 8, 2017). The species primarily hugs the coastline in this range, but form two large population centres in the Sea of Okhotsk and the Bering Sea. In BC, there are four primary spawning grounds - Dixon Entrance/northern Hecate Strait, Queen Charlotte Sound, SW Vancouver Island, and the Strait of Georgia (Saunders et al. 1989). These are reflected by high CPUE density in Figure 2, which also shows an additional hotspot in upper Moresby Gully. The bulk (98\%) of the commercial captures from the BC North population lies between depths 55 m and 457 m while those from BC South lie between 90 m and 401 m in 5AB and between 64 m and 470 m in 3CD (Appendix D).

### 1.2. ASSESSMENT BOUNDARIES

This assessment includes Pacific Marine Fisheries Commission (PMFC) major areas 5CDE for the BC North stock and PMFC major areas 5AB and 3CD plus minor areas 12 (Queen Charlotte Strait) and 20 (entrance to Juan de Fuca Strait) for the BC South stock (Figure 1). These assessed areas are similar to the GMU TAC (total allowable catch) areas, so managers can divide any recommended catch, delimited by brackets [ ], using simple TAC ratios:

$$
\begin{aligned}
& \mathrm{TAC}_{5 \mathrm{CDE}}=[\mathrm{BC} \text { North }] \\
& \mathrm{TAC}_{5 \mathrm{AB}}=[\mathrm{BCSouth}]\left(\frac{\mathrm{TAC}_{5 \mathrm{AB}}}{\mathrm{TAC}_{5 \mathrm{AB}}+\mathrm{TAC}_{3 \mathrm{CD}}}\right) \\
& \mathrm{TAC}_{3 \mathrm{CD}}=[\mathrm{BCSouth}]\left(\frac{\mathrm{TAC}_{3 \mathrm{CD}}}{\mathrm{TAC}_{5 \mathrm{AB}}+\mathrm{TAC}_{3 \mathrm{CD}}}\right)
\end{aligned}
$$



Figure 2. Aerial distribution of Walleye Pollock mean trawl (combined bottom and midwater trawl) tow catch per unit effort (kg/hour) from Feb 17, 1996 to Dec 31, 2016 (accessed May 3, 2017). Isobaths show the 100, 200, and 500 m depth contours. Note that cells with $<3$ fishing vessels are not displayed. Each cell represents, on average, $32 \mathrm{~km}^{2}$.

## 2. CATCH DATA

The preparation methods and a full catch history for the assessment of the two Walleye Pollock stocks are presented in detail in Appendix A. The catch used in the model appears in Table A.7. The 5-year average catch (over 2011-15) was 992 t in BC North and $3,256 \mathrm{t}$ for BC South. Information about species caught concurrently with WAP commercial catches are presented in Appendix D.

## 3. FISHERIES MANAGEMENT

Table A. 2 summarises all management actions taken for Walleye Pollock since 1981. The last assessment of Walleye Pollock was conducted in 1997 (Saunders and Andrews 1998¹) - an assessment that did not estimate stock status with respect to DFO Precautionary Approach reference points. Harvest advice is required for this species to determine if current harvest levels are sustainable and are compliant with the DFO Sustainable Fisheries Framework's Decision-making Framework incorporating the Precautionary Approach. A request was submitted to Science by staff in the Fisheries and Aquaculture Management's Groundfish Management Unit in 2013 but was considered unachievable due to data limitations.

One of the conditions from the 2014 Marine Stewardship Council (MSC) Canadian hake fishery certification was that a Pollock assessment be completed by 2018. In response to this MSC condition, the December 5, 2014 GTAC (Groundfish Trawl Advisory Committee) Summary states: "GTAC recommended that DFO ensure that redstripe, pollock and rougheye be assessed by 2018 to ensure compliance with the MSC conditions".

## 4. SURVEY DESCRIPTIONS

Four sets of fishery independent survey indices have been used to track changes in the biomass of the BC North population (Appendix B):

1. Goose Island Gully (GIG) Historical ${ }^{2}$ - an early series of 9 indices extending from 1967 to 1995. Most of these surveys were performed by the research vessel G.B. Reed, but two commercial vessels (Eastward Ho and Ocean Selector) were used in 1984 and 1994 respectively. The 1995 survey used both the Ocean Selector and the Frosti. Only tows located in GIG were used to ensure continuity across all surveys.
2. HS Assemblage - a series of 11 indices spanning 1984 to 2003 . The original design of this survey assigned fishing locations by 10 -fathom depth intervals within a 10 nautical mile grid of Hecate Strait (HS). The HS Assemblage survey was designed as a systematic fixedstation survey. In 2004, this survey series was discontinued in favour of the HS Synoptic survey.
3. HS Synoptic - a random-stratified "synoptic" (species comprehensive) trawl survey covering all of Hecate Strait and extending into Dixon Entrance and across the top of Graham Island. This survey has been repeated 6 times between 2005 to 2015 using two vessels and a consistent design, including targeting a wide range of finfish species.
4. WCHG Synoptic - a random-stratified synoptic trawl survey covering the west coast (WC) of Graham Island in Haida Gwaii (HG) and western part of Dixon Entrance. This survey has been repeated 6 times between 2006 to 2016 using three vessels and a consistent design, including targeting a wide range of finfish species.
Three sets of fishery independent survey indices have been used to track changes in the biomass of the BC South population (Appendix B):
5. Goose Island Gully (GIG) Historical- an early series of 9 indices extending from 1967 to 1995. Most of these surveys were performed by the research vessel G.B. Reed, but two commercial vessels (Eastward Ho and Ocean Selector) were used in 1984 and 1994

[^1]respectively. The 1995 survey used both the Ocean Selector and the Frosti. Only tows located in GIG were used to ensure continuity across all surveys.
2. WCVI Synoptic - a random-stratified synoptic trawl survey covering the west coast of Vancouver Island (WCVI). This survey has been repeated 7 times between 2004 to 2016 using the same vessel and a consistent design, including targeting a wide range of finfish species.
3. QCS Synoptic - a random-stratified synoptic trawl survey covering all of Queen Charlotte Sound (QCS) and targeting a wide range of finfish species. This survey has been repeated 8 times between 2003 to 2015, using three different vessels (see Table B.10) but with a consistent design.
The relative biomass survey indices were used as data in the model to index abundance, along with their associated relative errors to weight each index value inversely proportional to its variance. All surveys in the Synoptic series use the same net, an Atlantic Western IIA box trawl net.

## 5. COMMERCIAL TRAWL CPUE

Commercial catch and effort data from the bottom trawl fishery were used to generate indices of abundance for Walleye Pollock in this stock assessment (Figure 3, Figure 4). This was done for several reasons, with the primary reason being the lack of long-term abundance information for use in this model. In addition, it was hoped that the nature of the bottom trawl fishery, with all of the WAP catches taken while targeting other groundfish species, would result in an index series that was not unduly affected by economic considerations.

The theoretical basis for the analysis is described in Appendix C, Section C.2. The analysis (Sections C. 3 and C.4) is based on tow-by-tow data which reported Walleye Pollock landings or discards or which operated in a depth range where WAP would be expected to be caught. The period analysed was from 1996, when compulsory onboard observer coverage began, to 2015, the last complete year of data. Three analyses were performed for each stock:

- a regression analysis on all positive catch records which assumed a log-normal distribution, where the effect on catch rates by DFO locality, fishing depth, month, $0.1^{\circ}$ latitude bands, hours fished and vessel were estimated and removed from the trend, leaving a standardised annual abundance trend;
- a similar analysis using the presence/absence of WAP in the data set, which assumed a binomial distribution and which removed the effects of DFO locality, fishing depth, month, $0.1^{\circ}$ latitude bands and hours fished, resulting in an alternative annual abundance trend; and
- an analysis which combined the log-normal and binomial series using the delta-lognormal method of Vignaux (1994; see Equation C.4).
The combined series was used as input to the WAP stock assessment models for the BC North (Figure 3) and the BC South (Figure 4) stocks. The initial decline in the index may have been partly influenced by the adjustment to a new regulatory regime (onboard observers, transferrable individual vessel quotas). Model sensitivity runs which dropped this series were made to investigate the impact of the CPUE series in the stock assessment.


Figure 3. Combined index series for the BC North bottom trawl fishery along with the contributing lognormal and binomial index series, all normalised on their respective geometric means. The 95\% confidence bounds are based on 500 bootstrap replicates.


Figure 4. Combined index series for the BC South bottom trawl fishery. Details in Figure 3 caption.

## 6. BIOLOGICAL INFORMATION

### 6.1. NATURAL MORTALITY

Although the Alaskan fisheries stock assessments use age-specific mortality rates for Walleye Pollock, the underlying assumption is that $M=0.30$ at full maturity (Dorn et al. 2015). Agespecific $M$ values of $0.90,0.45$, and 0.30 for ages 1,2 and $3+$, respectively, have been used in Alaskan Eastern Bering Sea catch-age models since 1982 (lanelli et al. 2015). The delaydifference model (Section 7) assumes that maturity matches selectivity, i.e., all recruited fish are mature with a single natural mortality rate. Runs were made with $M=0.25, M=0.30$ and $M=0.35$ to bracket plausible values for this parameter.

### 6.2. KNIFE-EDGE SELECTIVITY

Dorn et al. (2012) provide a range of selectivity ogives for the Gulf of Alaska (GoA) fisheries and surveys, with the median age selected to these commercial fisheries ranging from age 3 to age 5 (see columns 5 to 7 in Table D.8). Based on the ogives in this table, ages 3 and 4 were selected as the most likely ages to use for the age of knife-edge recruitment in the WAP delaydifference model. Knife-edge selectivity at age 5 was also run as an additional sensitivity for the BC South stock because of the lower maximum size in the growth model used for this stock.

### 6.3. GROWTH PARAMETERS

Growth parameters were estimated from WAP length and age data from DFO biological samples collected from 1976 to 1995 (Appendix D); however, ages determined by the break and burn method (MacLellan 1997) existed for only 17 specimens from commercial trips (16 in BC North, 1 in BC South). Otoliths from commercial sources and aged by surface readings were more abundant (230 in BC North, 399 in BC South) but are known to be biased (Stanley 1987), at least for Pacific rockfish species. The remaining processed otoliths came from research surveys but were aged using pectoral fin ray counts ( 638 in BC South) or an unknown method (210 in BC South). Unfortunately, pectoral fin ageing is thought to be biased, especially at older ages (MacLellan et al. $1990^{3}$ ), because fin ray deposition slows down or ceases at older ages.
As the DFO age data were insufficient to derive growth parameters not biased by ageing methodology, we approached an Alaskan colleague, Martin Dorn (Research Fish Biologist, NOAA Fisheries, Sand Point, Seattle), who supplied 8,882 age-length pairs randomly selected from six biannual surveys conducted in the Gulf of Alaska (GoA) between 2005 and 2015. These fish had all been aged from otoliths prepared using the "break \& burn" method and he advised us to use the samples from the Eastern GoA as growth varied across the GoA. We used these data to estimate a growth function ( $L_{\neq}=66.944 \mathrm{~cm}, K=0.212, t_{0}=-1.136$, Figure 5) for the BC North stock that adequately fit the mean weight data for three knife-edge ages 3,4 , and 5 (Appendix D). However, we could not use this function for the BC South stock because fish sampled from Dixon Entrance were, on average, twice as large as those sampled from more southern BC waters. This North stock likely belongs to a larger stock that includes Dixon Entrance, northern Hecate Strait, and waters off of Southeast Alaska (Gustafson et al. 2000).

For the BC South stock, other sources were explored. Growth functions based on fin ray ages published by Saunders et al. (1989) for the west coast of Vancouver Island (3CD) and the Strait

[^2]of Georgia (apparently derived from age-length pairs not available in the general DFO database) were tested but featured growth rate coefficients that were so steep that neither function could fit the South mean weight data satisfactorily. NOAA colleagues, who work on pelagic fish off the west coasts of Washington and Oregon, indicated that age-length pairs from fisheries off Washington or Oregon were non-existent. In the end, we found a growth function published by Janusz and Horbowy (1997) for Walleye Pollock in the central Sea of Okhotsk $\left(L_{*}=50.827 \mathrm{~cm}, K=0.199, t_{0}=-1.790\right.$, Figure 5$)$, which provided satisfactory fits to the observed BC South mean weight data for knife-edge ages 3 and 4 (Appendix D). We have no reason to believe that the Sea of Okhotsk (OS) relationship represents BC South (BCS) other than the estimated OS growth (1991-94) was consistent with our mean weight data. Although the study region in the Sea of Okhotsk sits at higher latitudes (54-55 ${ }^{\circ} \mathrm{N}$ ) than BC South (48$52^{\circ} \mathrm{N}$ ) and experiences cooler temperatures ( $8-12^{\circ} \mathrm{C}$ in summer vs. $\sim 15^{\circ} \mathrm{C}$ ), the two populations of pollock share comparable length compositions: mean $=40 \mathrm{~cm}$ (OS) vs. 37 cm (BCS), maximum $=75 \mathrm{~cm}$ (OS) vs. 78 cm (BCS), predominant lengths $=35-45 \mathrm{~cm}$ (OS) vs. 23-50 cm (BCS) for $\sim 76 \%$ of the lengths.


Figure 5. Interpolated combined-sex growth models used to estimate the Walford parameters used in the WAP delay-difference stock assessment model. [left panel]: Eastern Gulf of Alaska model; [right panel]: Okhotsk Sea model.

Growth and length-weight parameters (Section D.1.1) appropriate for each stock were used to prepare Walford plots (Figure D.16) which provide the growth parameter values used as input to the WAP delay-difference model. The Walford parameters are calculated from the knife-edge recruitment age to 30 y for the each growth model. The Walford parameters will vary slightly with changing age assumptions at knife-edge recruitment for both growth models (GoA in the North, Sea of Okhotsk in the South). Table D. 9 presents the Walford parameters used in the stock assessment along with the mean length and mean weight associated with each of the knife-edge age at recruitment assumptions.

### 6.4. MEAN WEIGHT

In excess of 50,000 WAP length observations, taken from unsorted samples, were available from the combined offshore trawl fisheries (see Appendix D), 18,873 in the North (1973-2016) and 32,125 in the South (1972-2016). All lengths were converted to weights using combinedsex, stock-specific parameters in Section D.1.1. Although females attain larger sizes than males
(Figure 5), the allometric relationship between length and weight remains similar between the sexes (Figure D.4). To remove some of the variance due to influential factors in the data, an additive lognormal model (Schnute et al. 2004) was used to adjust the annual index of fish weight for each BC pollock stock (Section D.2.1). The only explanatory variable that had an observable effect on the series was the minor PMFC area. Additional explanatory variables (e.g., season, depth, sex) were explored but their effects on the annual index series were minimal and consequently not used. The most striking feature between the North and South is that fish are roughly twice as large in the North with a geometric mean weight of $1.056 \mathrm{~kg} / \mathrm{fish}$ vs. $0.521 \mathrm{~kg} /$ fish in the South (see Appendix D). This observation may be due to a number of possible hypotheses:
A. True differences in growth exist between two discrete populations;
B. Older (and larger fish) migrate north;
C. Exploitation rates in the south are higher, cropping off all the big fish.

Hypothesis C seems unlikely, given the spotty exploitation history in the south. We prefer hypothesis A; the modal age in the GoA survey data, which we use to represent the BC North stock, can be either age 1 or age 3 (i.e., many young fish); however, these data do not rule out hypothesis B. To eliminate migration, we would need age samples from the BC South stock.

### 6.5. MATURITY

Maturity is not used as input into a delay-difference model. Instead, the model assumes that the age at knife-edge recruitment also defines maturity, resulting in a population where all recruited fish are fully mature. Maturity ogives based on the available DFO age and length data were constructed to test the consistency of the available data with the assumption that all fish from the knife-edge age were mature (see Section D.1.3).

The resulting analysis indicated that the median age at full maturity for WAP is 3.4 y using pectoral fins and 4.6 y using surface-read otoliths (Figures D. 5 \& D.6). Median lengths at full maturity using unsorted research, unsorted commercial, and sorted commercial are 49.7 cm , $56.6 \mathrm{~cm}, 57.3 \mathrm{~cm}$, respectively (Figure D.15). These values are consistent with the ages (3 to 5 y) used in this stock assessment as candidates for knife-edge recruitment. These analyses indicate that the assumption that vulnerable fish are fully mature is not wholly satisfied, particularly when knife-edge selectivity equals age three. Dorn et al. (2012) report that estimates of the $50 \%$ maturity age in the Gulf of Alaska "..are highly variable and range from 3.5 years in 1983 to 6.1 years in 1991, with an average of 4.9 years".

### 6.6. STEEPNESS

A Beverton-Holt (BH) stock-recruitment function was used to generate average recruitment estimates in each year, based on the biomass of recruited WAP (Equation E.22). Recruitment deviations from this average (Equation E.23) were estimated to improve the fit to the model data and to introduce variability in the Bayesian estimation phase. The BH function was parameterised using a steepness parameter, $h$, which specifies the proportion of the maximum recruitment available at $0.2 B_{0}$ (Mace and Doonan 1988). The parameter $h$ was estimated in the model, constrained by a prior that took the form of a beta distribution with mean 0.7 and standard deviation 0.15. This prior was the same one used by Forrest et al. (2015) for Pacific Cod (Gadus macrocephalus), and is very similar to the prior developed for west coast rockfish by Forrest et al. (2010; mean=0.674; standard deviation=0.168). Myers et al. (1999) reported median $h=0.55$ for $T$. chalcogramma (with 20 and 80 percentiles of 0.53 and 0.58 ); however,
this estimate is based on only two fish. Estimated $h$ for G. morhua was 0.84 (0.76. 0.90) based on 21 fish, and $h=0.79(0.67,0.87)$ for the family Gadidae based on 49 fish (Myers et al. 1999).

## 7. DELAY-DIFFERENCE MODEL

Delay difference models represent an intermediate approach between aggregated surplus production models and age-structured models. The delay-difference structure tracks the effects of recruitment, survival and growth on biomass, without requiring an explicit age-structured framework, and can perform well as long as its major assumptions are met (Hilborn and Walters 1992). Difference equations, which allow for a time-delay between spawning and recruitment, are used to build population models in discrete time-steps (generally 1 year), in which the surviving biomass for next year is predicted from the surviving biomass from the previous year, after adjusting for growth and mortality and adding next year's recruitment. An advantage of delay difference models over simpler production models is that they do not assume constant recruitment over time.

The key assumptions of the delay difference model are:

- Growth in mean body weight follows the linear relationship described by the Ford-Walford equation (E.1).
- Knife-edge selectivity, i.e., all fish aged $k$ and older are equally vulnerable to the fishing gear. A corollary to the assumption of knife-edge selectivity is that maturity is also knifeedge and the same as selectivity. All fish in the model are mature and fully selected.
- Constant natural mortality at age, i.e., all fish aged $k$ and older have the same natural mortality rate.
This model is described with equations in Appendix E (see also Forrest et al. 2015). The model was fit to the annual catch data (Table A.7), three or four survey series in the North and three in the South (Appendix B), a series of CPUE biomass indices from commercial bottom trawls (Appendix C), and a series of fishery mean weights described in Section D.2. We did not attempt to alter the relative weighting of the component data series, instead using the observation error CVs estimated by the surveys without modification. An arbitrary CV=0.3 was used for the CPUE data and $\mathrm{CV}=0.15$ for the mean weight data. We decided not to explore sensitivity to the variance components of the model because, in our previous work with a Shortspine Thornyhead delay-difference model, we had not found much effect on model estimates other than to raise or lower model uncertainty (Starr and Haigh 2017). Given limited resources, we opted to concentrate on sensitivity runs which would directly affect model conclusions.

We describe an "example case" run in detail for each of the BC North and BC South stocks. We then present a range of alternative runs for each stock which explore the effect of key model assumptions (age at knife-edge selectivity, $M$ and the effect of some of the series indices). Each example case is not more likely than the other runs ( 12 in total for the BC North and 11 for the BC South). It is simply one run in the group of 12 or 11 runs presented for each BC stock. These alternative analyses were run (with full MCMC simulations) to see how model results differed when input assumptions were changed. Specifications for these runs are given in Table 1 for both Walleye Pollock stocks.

Table 1. Summary of the analyses performed to test the sensitivity of the delay-difference model to variations in natural mortality $M$, knife-edge recruitment age $k$ and included series indices. All runs for the BC North stock use the eastern Gulf of Alaska growth function (Martin Dorn, pers. comm.); all runs for the BC South stock use the Okhotsk Sea growth function (Janusz and Horbowy 1997). The column marked 'MCMC rank' provides a subjective ranking of the MCMCs, where $1=$ good, $2=$ acceptable, and $3=$ poor.

|  | Case Run ID |  | Run \# | M |  | MC rank |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| BC North example 4 surveys + CPUE | S00 | M.30+k3 | 1 | 0.30 | 3 | 1.25 |
|  | S01 | M.25+k3 | 16 | 0.25 | 3 | 3 |
|  | S02 | M. $30+\mathrm{k} 4$ | 4 | 0.30 | 4 | 2.5 |
|  | S03 | M. $35+\mathrm{k} 3$ | 2 | 0.35 | 3 | 1.5 |
|  | S04 | M. $35+\mathrm{k} 4$ | 12 | 0.35 | 4 | 1.5 |
| No GIG survey | S05 | M. $25+\mathrm{k} 3-\mathrm{GIG}$ | 9 | 0.25 | 3 | 3 |
|  | S06 | M.25+k4-GIG | 7 | 0.25 | 4 | 2 |
|  | S07 | M.30+k3-GIG | 3 | 0.30 | 3 | 3 |
|  | S08 | M.30+k4-GIG | 5 | 0.30 | 4 | 1.5 |
|  | S09 | M.35+k4-GIG | 10 | 0.35 | 4 | 1.25 |
| No GIG or CPUE | S10 | M.30+k3-GIG-CPUE | 8 | 0.30 | 3 | 2 |
|  | S11 | M.30+k4-GIG-CPUE | 6 | 0.30 | 4 | 2 |
| BC South example $k$ at $M=0.30$ | S00 | M.30+k3 | 4 | 0.30 | 3 | 2 |
|  | S01 | M.30+k4 | 5 | 0.30 | 4 | 2 |
|  | S02 | M. $30+\mathrm{k} 5$ | 15 | 0.30 | 5 | 2 |
| k at $\mathrm{M}=0.25$ | S03 | M.25+k3 | 11 | 0.25 | 3 | 1 |
|  | S04 | M.25+k4 | 9 | 0.25 | 4 | 1.5 |
|  | S05 | M.25+k5 | 14 | 0.25 | 5 | 2 |
| k at M=0.35 | S06 | M. $35+\mathrm{k} 3$ | 12 | 0.35 | 3 | 2 |
|  | S07 | M.35+k4 | 13 | 0.35 | 4 | 2 |
|  | S08 | M.35+k5 | 16 | 0.35 | 5 | 2 |
| No CPUE | S09 | M.30+k3-CPUE | 10 | 0.30 | 3 | 1 |
|  | S10 | M.30+k4-CPUE | 8 | 0.30 | 4 | 3 |

The MPD (mode of the posterior distribution) "best fit" was used as the starting point for a Bayesian search across the joint posterior distributions of the parameters using the Monte Carlo Markov Chain (MCMC) procedure. All models were run for 60,000,000 iterations, sampling every $50,000^{\text {th }}$, to give 1,200 draws ( 1,000 samples after dropping the first 200 as a "burn-in").
The range of model exploration represented in Table 1 was undertaken because there is substantial uncertainty in specifying the productivity of these stocks (represented by $M$ and $h$ in the growth model), as well as selecting the age at full knife-edge recruitment ( $k$ ), when making the assumptions that are mandatory when using a delay difference model. Because the available data are not informative with respect to these key model parameters, it is not possible to objectively rule out most of these alternative hypotheses. Instead, after covering a range of plausible values for the key parameter assumptions, we chose model runs for advice based on a subjective ranking of the MCMC diagnostics (e.g., autocorrelation, stability of traces) using a simple ranking system ( 1 = good, 2 = acceptable, 3 = poor), selecting only those runs that ranked $\leq 2$ (using the mean ranking assigned by each author independently). This stock assessment adopted a "Model Averaging" approach, using selected model runs for each stock that represent a range of hypotheses based on plausible F values (see Sections 8.1.2 and 8.2.2) to construct a "Model Average Composite" for providing advice to managers (Section 9).

## 8. MODEL RESULTS

### 8.1. BC NORTH

### 8.1.1. Model Example

Results for an example model run for BC North (Case S00, Table 1), which assumed $M=0.30$ and $k=3$ years, are presented to illustrate model behaviour, particularly in how these models fit the available data, the shape of the biomass trajectory and the predictions of stock status. This example case was chosen because the model estimated fishing mortality rates $(F)$ less than 2 ( $\approx$ maximum harvest rate $u<0.86$ ) in the MCMC runs (Table 2). However, there are other plausible model runs that could be used as the example case, regardless of the assumptions made regarding key productivity parameters, because the available data are not informative with respect to productivity and age at knife-edge selectivity.

Table 2. Median values for select MCMC-derived parameters and quantities for the 12 runs described for BC North in Table 1. The historical reference points use averages from 1967-2016. Projections to 2018 were made assuming TAC=1000 t, a value near to the 2011-2015 average catch of $992 t$. Shaded rows highlight runs contributing to the Model Average.

| Case | $h$ | $\begin{array}{r} B_{\text {avg }} \\ \text { (tonnes) } \end{array}$ | $\begin{gathered} B_{2017 /} \\ B_{\text {avg }} \end{gathered}$ | $\begin{gathered} B_{2018} / \\ B_{\text {avg }} \end{gathered}$ | Yrmin | $\begin{gathered} B_{2017} / \mathrm{l} \\ B_{\text {min }} \end{gathered}$ | $\begin{gathered} {\left[B_{2018}>\right.} \\ \left.B_{2017}\right] \end{gathered}$ | Uavg | $\begin{gathered} u_{2016} \\ u_{\text {avg }} \end{gathered}$ | Median $F_{\text {max }}$ by MCMC | \# Years median $F_{t}>2$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| S00: M.30+k3 | 0.74 | 7,568 | 0.57 | 0.42 | 2001 | 1.9 | 0 | 0.14 | 2.0 | 0.71 | 0 |
| S01: M. $25+\mathrm{k} 3$ | 0.75 | 4,962 | 0.48 | 0.28 | 1986 | 2.3 | 0 | 0.21 | 2.0 | 1.69 | 0 |
| S02: M.30+k4 | 0.78 | 3,377 | 0.44 | 0.21 | 1986 | 3.1 | 0.06 | 0.35 | 1.7 | 19.4 | 7 |
| S03: M.35+k3 | 0.73 | 10,270 | 0.60 | 0.47 | 2001 | 1.9 | 0.02 | 0.11 | 2.0 | 0.51 | 0 |
| S04: M.35+k4 | 0.77 | 3,695 | 0.46 | 0.24 | 1986 | 3.1 | 0.05 | 0.33 | 1.6 | 16.1 | 6 |
| S05: M.25+k3-GIG | 0.78 | 4,628 | 0.54 | 0.39 | 1986 | 2.7 | 0.09 | 0.26 | 1.8 | 9.97 | 1 |
| S06: M.25+k4-GIG | 0.81 | 3,272 | 0.51 | 0.35 | 1986 | 2.9 | 0.12 | 0.36 | 1.6 | 19.0 | 6 |
| S07: M.30+k3-GIG | 0.77 | 5,325 | 0.65 | 0.53 | 1986 | 3.7 | 0.10 | 0.24 | 1.6 | 8.01 | 0 |
| S08: M.30+k4-GIG | 0.80 | 3,438 | 0.58 | 0.43 | 1986 | 3.5 | 0.14 | 0.35 | 1.5 | 18.9 | 6 |
| S09: M.35+k4-GIG | 0.80 | 3,725 | 0.62 | 0.49 | 1986 | 3.5 | 0.16 | 0.34 | 1.5 | 17.5 | 5 |
| S10: M.30+k3-GIG-CPUE | 0.75 | 6,986 | 1.3 | 1.1 | 2000 | 9.9 | 0.05 | 0.20 | 0.80 | 18.4 | 1 |
| S11: M.30+k4-GIG-CPUE | 0.80 | 4,248 | 1.1 | 0.92 | 1986 | 8.9 | 0.07 | 0.31 | 0.87 | 19.9 | 8 |

The MPD fits in the example model run to the survey and CPUE indices are generally acceptable although the model is incapable of fitting the abrupt changes in some series like the HS assemblage (Figure 6). The model is also not capable of fitting the mean weights near the end of the time series; generally, the fit fluctuates without trend (Figure 7). The fit to the catch series is tight (Figure 7). The MCMC spawning biomass trajectory in relation to $B_{\text {avg }}$ appears in Figure 7 and shows the median limit reference point $B_{2001}=0.30 B_{\text {avg }}$. The MCMC recruitment of age 3 fish shows that 10 recruitment-year medians exceed the mean recruitment Figure 7.

Observing the alternative runs for BC North, the MPD fits to the data series (Figure F.13: mean weight, Figure F.14: HS assemblage, Figure F.15: HS synoptic, Figure F. 16 WCHG synoptic, and Figure F.17: North CPUE), demonstrate that this model has little power to distinguish among hypotheses. Figure F. 17 suggests that some combinations of $M$ and $k$ are better able to fit the high 1996 CPUE index point. In general, model runs with $k=4$ fit the mean weight data better than the models with $k=3$ or $k=5$. However, models with $k=3$ tend to have lower maximum exploitation rates because the estimated stock size is larger. Models with $k=4$ tend to reach the model's maximum fishing mortality rate (constrained at $F=20$ ) in some years because there are too few fish available to match the observed catch. This is likely due to a failure in the model assumption of knife-edge selectivity at a specified age.


Figure 6. BC North MPD fits (triangles) to the four survey abundance series and to the CPUE index series for the example model Case SOO.

When the MCMC traces of the alternative run hypotheses are examined (Figure F.18), they appear to be relatively well-behaved with smooth running quantiles; however, the autocorrelation plots (Figure F.19) point to roughly half of the MCMC chains containing important lagged correlation effects. The example case (S00) shows some serial correlation at the beginning, but this dissipates over time. Some of the cases (e.g., S02 and S11) show periodic correlation and two cases (S01 and S07) show high positive serial correlation through most of the time series. Visually, the best cases (S06, S08, S09) all feature $k=4$, regardless of $M$, and no GIG survey.


Figure 7. BC North example model case S00: [top left] MPD fit to the mean weight data; [top right] MPD fit to the catch data; [bottom left] MCMC time trend of $B_{t} / B_{\text {avg, }}$, showing the median (heavy black line), $5 \%$ and $95 \%$ quantiles (dashed lines) from the posterior distribution, as well as showing the MPD trend and the historical reference points $L R P=B_{\text {min }} / B_{\text {avg }}$ and $U S R=2 L R P$; [bottom right] MCMC time trend of recruitment showing $90 \%$ credibility intervals.

Table F.3, which gives the negative log likelihoods (NLLs) for alternative model fits, provides a more quantitative basis to make comparisons. This table shows that there are some big differences in the component NLLs, indicating how well the models fit the various data components. These likelihoods can be use to select models among similar alternative runs, depending on which components are deemed important. For instance, case S10 had the lowest NLL for mean weight, but suffered from periodic autocorrelation in the $\log R_{0}$ MCMC traces (Figure F.19). The objective function value (OFV) can be compared within three clusters where model components are symmetrical - \{S00, ... S04\}, \{S05, ..., S09\}, and \{S10, S11\}. Within the
second cluster, S08 has the lowest OFV, the best fit to mean weight, and the lowest NLLs for all survey components. These results suggest that a model without the GIG survey featuring $M=0.35$ and $k=4$ offers the best fit to the data with little autocorrelation. Unfortunately, choosing $k>3$ results in unrealistically high fishing mortality rates in some years (see $F_{\text {max }}$ in Table 2). These high fishing mortality rates are likely a by-product of the assumed knife-edge recruitment made by the delay-difference model, which sometimes results in insufficient biomass available to accommodate catch levels in some years when $k>3$.

Median estimates for current biomass lie well above the LRP $=B_{\text {min }}$ reference point across the full range of alternative runs (Table 2). Additionally, current (2017) spawning biomass only falls below $0.5 B_{\text {avg }}$ three times (cases S01, S02, and S04). Projected spawning biomass has a high probability of being lower than the current spawning biomass under a catch strategy of $1000 \mathrm{t} / \mathrm{y}$. Generally, assumptions of higher knife-edge selectivity produce higher rates of fishing mortality and lower estimates of standing stock. Regardless of the alternative run, the BC North stock is not large and most likely represents the southern extreme of a larger SE Alaska population. If this is correct, the existence of this larger (and mostly unfished) population of Pollock may provide a rescue effect for the BC portion of the population in situations of low abundance.

Seven of the alternative BC North runs removed the GIG survey abundance index estimates, resulting in model runs with a lower average biomass ( $B_{\text {avg }}$ ) and higher average removal rates ( $u_{\text {avg }} B_{\text {avg }}$, Figures F21 \& F.22); however, estimated stock status ( $B_{2017} / B_{\text {avg }}$ ) for the first five cases remains similar across these runs and to the example case (Figure F21). Removing the commercial CPUE series and the GIG survey increases stock status ( $B_{2017} / B_{\text {avg }}$ ) and reduces current exploitation rate ( $u_{2016} / U_{\text {avg }}$ ) relative to the example run (Figure F22).

### 8.1.2. Model Average

Initially, model runs used for a Model Average were selected based on MCMC diagnostics alone. However, the RPR participants further restricted the initial selection based on aspects of the estimated fishing mortality in the MCMC samples (Figure 8). The following criteria (also see Table 1 and Table 2) were used for selecting model runs for inclusion to the Model Average posterior for this assessment:

- use model runs where the median $F_{\text {max }}$ for MCMC samples was $<2$;
- add model runs where median annual $F_{t}$ was $>2$ only once;
- remove model runs with poor MCMC diagnostics (rank > 2).

Three alternative BC North runs were selected for inclusion to the Model Average posterior based on the above criteria. Table 3 gives the model-based and HRP-based quantities ( 0.05 , $0.50,0.95$ quantiles) from the Model Average posterior based on the 3000 pooled MCMC samples (runs S00, S03, S10). This table shows that the BC North composite stock is evaluated to be $68 \%$ of $B_{\text {avg }}$ ( $90 \%$ credibility range: $38 \%$ to $162 \%$ ) and to be $231 \%$ above $B_{\text {min }}(90 \%$ credibility range: $129 \%$ to $1610 \%$ ). There is a high degree of positive skewness in the estimates of current and average stock size because, while the lower bounds of stock size are determined by the catch history and the assumed model maximum exploitation rate, there is little information in the data to constrain the upper bounds of stock size.


Figure 8. BC North: Fishing mortality rate ( $F_{t}$ ) for the alternative model runs (see Table 1). Annual boxplots show the 5, 25, 50, 75 and $95 \%$ quantiles. The horizontal dashed line indicates $F=2$, the horizontal solid line with number above/below indicates the median of $1000 F_{\text {max }}$ estimates, the vertical bar to the right of the $F_{\text {max }}$ line shows the $90 \%$ credibility interval for $F_{\text {max }}$, and the annual circles along the $F_{\text {max }}$ line indicate the frequency of MCMC runs that reached $F_{\text {max }}$ in that year.

Table 3. BC North: the 0.05, 0.50, and 0.95 quantiles of MCMC-derived quantities from 3000 MCMC samples comprising the Model Average Composite scenario. Definitions: $B_{2017}$ : current beginning year spawning biomass, $B_{\text {avg }}$ : average biomass from 1967 to $2016, B_{\text {min }}$ : minimum biomass that acts as the LRP (USR = 2LRP), $u_{2016: ~}$ harvest rate (ratio of total catch to vulnerable biomass) in the middle of 2016, and $u_{\text {avg: }}$ : average harvest rate from 1967 to 2016. All biomass values are in tonnes. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

|  | Quantile |  |  |
| :--- | ---: | ---: | ---: |
|  | $5 \%$ | $50 \%$ | $95 \%$ |
| Model based |  |  |  |
| $B_{2017}$ | 2,621 | 6,185 | 13,927 |
| $B_{\text {avg }}$ | 5,634 | 7,837 | 14,626 |
| $B_{2017} / B_{\text {avg }}$ | 0.385 | 0.683 | 1.62 |
| $U_{2016}$ | 0.106 | 0.214 | 0.406 |
| HRP-based |  |  |  |
| $B_{\text {min }}$ | 654 | 2,051 | 4,818 |
| $2 B_{\text {min }}$ | 1,307 | 4,101 | 9,636 |
| $B_{\text {min }} / B_{\text {avg }}$ | 0.0921 | 0.27 | 0.388 |
| $2 B_{\text {min }} / B_{\text {avg }}$ | 0.184 | 0.54 | 0.775 |
| $B_{2017} / B_{\text {min }}$ | 1.29 | 2.31 | 16.1 |
| $U_{\text {avg }}$ | 0.0744 | 0.15 | 0.234 |
| $U_{2016} / U_{\text {avg }}$ | 0.602 | 1.79 | 2.52 |

### 8.2. BC SOUTH

### 8.2.1. Model Example

Results for an example model run for BC South (Case S00, Table 1), which assumed $M=0.30$ and $k=3$ years, are presented to illustrate model behaviour, particularly in how these models fit the available data, the shape of the biomass trajectory and the predictions of stock status. This example case was chosen because the model estimated low fishing mortality rates $(F)$ in the MCMC runs (Table 4).

Table 4. Median values for select MCMC-derived parameters and quantities for the 11 runs described for BC South in Table 1. The historical reference points use averages from 1967-2016. Projections to 2020 were made assuming TAC=3250 $t$, a value below the 2011-2015 average catch of $3,256 t$. Shaded rows highlight runs contributing to the Model Average.

| Case | $h$ | $\begin{array}{r} B \text { avg } \\ \text { (tonnes) } \end{array}$ | $\begin{gathered} B_{2017 /} \\ B_{\text {avg }} \end{gathered}$ | $\begin{gathered} \text { B2020/ } \\ \text { Bavg } \end{gathered}$ | Yrmin | $\begin{gathered} B_{2017} / \\ B_{\text {min }} \end{gathered}$ | $\begin{gathered} P\left[B_{2020}\right. \\ \left.>B_{2017}\right] \end{gathered}$ | Uavg | $\begin{gathered} u_{2016} \\ \text { Uavg } \end{gathered}$ | Median <br> $F_{\text {max }}$ by <br> MCMC | \# Years median $F_{t}>2$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| S00: M.30+k3 | 0.75 | 89,549 | 1.0 | 0.88 | 2008 | 5.8 | 0.01 | 0.04 | 0.73 | 0.28 | 0 |
| S01: M.30+k4 | 0.77 | 21,257 | 0.79 | 0.64 | 2008 | 8 | 0.02 | 0.16 | 0.82 | 18.3 | 1 |
| S02: M.30+k5 | 0.79 | 14,835 | 0.72 | 0.53 | 2008 | 8.3 | 0.01 | 0.24 | 0.82 | 17.9 | 4 |
| S03: M.25+k3 | 0.76 | 54,998 | 1.0 | 0.87 | 2008 | 6.4 | 0 | 0.06 | 0.74 | 0.49 | 0 |
| S04: M. $25+\mathrm{k} 4$ | 0.78 | 20,412 | 0.85 | 0.64 | 2008 | 8.5 | 0 | 0.16 | 0.80 | 18.3 | 1 |
| S05: M.25+k5 | 0.81 | 13,022 | 0.84 | 0.64 | 2008 | 8.7 | 0.01 | 0.24 | 0.82 | 19.4 | 5 |
| S06: M.35+k3 | 0.74 | 183,563 | 1.0 | 0.86 | 2008 | 5.5 | 0.01 | 0.02 | 0.74 | 0.12 | 0 |
| S07: M. $35+\mathrm{k} 4$ | 0.78 | 21,814 | 0.75 | 0.54 | 2008 | 7.6 | 0.01 | 0.17 | 0.82 | 14.2 | 1 |
| S08: M. $35+\mathrm{k} 5$ | 0.79 | 14,623 | 0.72 | 0.45 | 2008 | 7.9 | 0.01 | 0.23 | 0.85 | 19.7 | 3 |
| S09: M.30+k3-CPUE | 0.75 | 33,336 | 0.62 | 0.48 | 2008 | 18 | 0.01 | 0.14 | 0.81 | 19.2 | 2 |
| S10: M.30+k4-CPUE | 0.76 | 19,971 | 0.90 | 0.65 | 2008 | 17 | 0.01 | 0.18 | 0.69 | 18.4 | 2 |

The MPD fits in the example case to the survey and CPUE indices are generally acceptable, although the model is incapable of fitting the abrupt changes in some series like the

WCVI synoptic in 2010 (Figure 9). The model is also not capable of fitting the high mean weights in the first half of the time series (Figure 10). It is not possible to know if the early (pre1980) mean weight samples, which seem to be inconsistent with the later mean weights in the series, are truly representative the BC South population. Apart from the poor fit to these early data, the fit to the remainder of the mean weight series fluctuates with a slight dome-shaped trend. The fit to the catch series is tight (Figure 10). The MCMC spawning biomass trajectory in relation to $B_{\text {avg }}$ appears in Figure 10 and shows the median limit reference point $B_{2008}=0.18 B_{\text {avg }}$. The MCMC recruitment of age 3 fish shows that only four recruitment-year medians exceed the mean recruitment Figure 10.

Observing the alternative runs for BC South, the MPD fits to the data series (Figure F.36: mean weight, Figure F.37: GB Reed, Figure F.38: WCVI synoptic, Figure F. 39 QCS synoptic, and Figure F.40: South CPUE) demonstrate that this model has little power to distinguish among hypotheses. The MCMC traces of the alternative hypotheses (Figure F.41) appear to be generally well-behaved with smooth running quantiles for most cases.


Figure 9. BC South MPD fits (triangles) to the three survey abundance series and to the CPUE index series for the example model Case S0O.


Figure 10. BC South example model case S00: [top left] MPD fit to the mean weight data; [top right] MPD fit to the catch data; [bottom left] MCMC time trend of $B_{t} / B_{\text {avg, }}$, showing the median (heavy black line), $5 \%$ and $95 \%$ quantiles (dashed lines) from the posterior distribution, as well as showing the MPD trend and the historical reference points $\angle R P=B_{\text {min }} / B_{\text {avg }}$ and $U S R=2 L R P$; [bottom right] MCMC time trend of recruitment showing $90 \%$ credibility intervals.

Table F.22, which gives the negative log likelihoods (NLLs) for the alternative model fits, provides a more quantitative basis to make comparisons. This table shows that there are some large differences in the component NLLs, indicating how well the models fit the various data components. These likelihoods can be used to select models among the alternative runs, depending on which components are deemed important and as long as the comparisons are made for models with the same data components. Excluding the cases that drop CPUE, almost all components show the smallest NLL for case S08, which suggests that this case provides the best fit to the data. However, case S 08 features a high $k$ of $5 y$ and the MCMC chains have
unacceptable levels of autocorrelation (Figure F.42). Also, at $k=5$, the median $F_{\text {max }}$ is hitting a model constraint of $F=20$ (Table 4), i.e., the complete removal of spawning biomass. Generally, the higher the $k$, the better the fit to mean weight, but the overall stock size is estimated to be smaller, resulting in very high levels of $F$ in some years. This latter result is caused by the assumption of knife-edge recruitment at a specified age that is made by the delay-difference model, sometimes resulting in insufficient biomass to accommodate catch levels in some years when $k>3$.

The example case (S00) demonstrates a step-wise increase in the $95^{\text {th }}$ quantile and S05 appears to have a downward trend in median $\log R_{0}$. The autocorrelation plots (Figure F.42) point to substantial lagged correlation effects in many of the MCMC chains. The example case shows serial correlation for at least 20 lags before dissipating, but it re-appears at later lags. Case S06, a variation on the example case with $M=0.35$, demonstrates substantial periodic correlation. The only cases to show low auto-correlation are S01 ( $M=0.30, k=4$ ), S03 ( $M=0.25$, $k=3)$, S 04 ( $M=0.25, k=4$ ), and S 09 ( $M=0.30, k=3$, no CPUE). It appears that the model is sensitive to combinations of $M$ and $k$ with a tendency for better MCMC diagnostics when $M$ is low with low $k$. When the CPUE series is removed, the example combination of moderate $M$ and $k$ yields MCMC chains with low autocorrelation.
Median estimates for current biomass lie well above the LRP $=B_{\text {min }}$ reference point across the full range of alternative runs (Table 4). Additionally, the minimum median current spawning biomass depletion is only $0.62 B_{\text {avg }}$ (case S09). Projected spawning biomass has a high probability of being lower than the current spawning biomass under a catch strategy of $3250 \mathrm{t} / \mathrm{y}$. Unlike the BC North stock, all cases find the same year for a biomass minimum (in 2008). The southern stock is at least an order of magnitude larger in population than the northern one. The two runs which discard the CPUE series also estimate levels of biomass which result in very high estimates for $F_{\text {max }}$, even when $k=3$.

### 8.2.2. Model Average

Initially, model runs used for a Model Average were selected based on MCMC diagnostics alone. However, the RPR participants further restricted the initial selection based on aspects of the estimated fishing mortality in the MCMC samples (Figure 11). The following criteria (also see Table 1 and Table 4) were used for selecting model runs for inclusion to the Model Average posterior for this assessment:

- use model runs where the median $F_{\text {max }}$ for MCMC samples was < 2 ;
- add model runs where median annual $F_{t}$ was $>2$ only once;
- remove model runs with poor MCMC diagnostics (rank > 2 ).

Six alternative BC South runs were selected for inclusion to the Model Average posterior based on the above criteria. Table 5 gives the model-based and HRP-based quantities ( $0.05,0.50$, 0.95 quantiles) from the Model Average posterior based on the 6000 pooled MCMC samples (runs S00, S01, S03, S04, S06, S07). This table shows that the composite stock is evaluated to be $90 \%$ of $B_{\text {avg }}\left(90 \%\right.$ credibility range: $59 \%$ to $135 \%$ ) and to be $677 \%$ above $B_{\text {min }}(90 \%$ credibility range: $233 \%$ to $1080 \%$ ). There is a high degree of positive skewness in the estimates of current and average stock size because, while the lower bounds of stock size are determined by the catch history and the assumed model maximum exploitation rate, there is little information in the data to constrain the upper bounds of stock size.


Figure 11. BC South: Fishing mortality rate $\left(F_{t}\right)$ for the alternative model runs (see Table 1). Annual boxplots show the 5, 25, 50, 75 and $95 \%$ quantiles. The horizontal dashed line indicates $F=2$, the horizontal solid line with number above/below indicates the median of $1000 F_{\text {max }}$ estimates, the vertical bar to the right of the $F_{\text {max }}$ line shows the $90 \%$ credibility interval for $F_{\text {max }}$, and the annual circles along the $F_{\text {max }}$ line indicate the frequency of MCMC runs that reached $F_{\text {max }}$ in that year.

Table 5. BC South: The 0.05, 0.50, and 0.95 quantiles of MCMC-derived quantities from 6000 MCMC samples comprising the Model Average Composite scenario. Definitions: $B_{2017}$ : current beginning year spawning biomass, $B_{\text {avg }}$ : average biomass from 1967 to $2016, B_{\text {min }}$ : minimum biomass that acts as the $L R P(U S R=2 * L R P), u_{2016}$ : harvest rate (ratio of total catch to vulnerable biomass) in the middle of 2016, and $u_{\text {avg: }}$ : average harvest rate from 1967 to 2016. All biomass values are in tonnes. For reference, the average catch over the last 5 years (2011-2015) is $3,256 t$.

|  | Quantiles |  |  |
| :--- | ---: | ---: | ---: |
|  | $5 \%$ | $50 \%$ | $95 \%$ |
| Model based |  |  |  |
| $B_{2017}$ | 12,737 | 28,923 | 317,629 |
| $B_{\text {avg }}$ | 16,938 | 33,487 | 292,976 |
| $B_{2017} / B_{\text {avg }}$ | 0.589 | 0.899 | 1.35 |
| $U_{2016}$ | 0.00787 | 0.0829 | 0.171 |
| HRP-based |  |  |  |
| $B_{\text {min }}$ | 1,543 | 6,520 | 58,110 |
| $2 B_{\text {min }}$ | 3,086 | 13,041 | 11,219 |
| $B_{\text {min }} / B_{\text {avg }}$ | 0.0753 | 0.138 | 0.296 |
| $2 B_{\text {min }} / B_{\text {avg }}$ | 0.150 | 0.277 | 0.593 |
| $B_{2017} / B_{\text {min }}$ | 2.33 | 6.77 | 10.8 |
| $U_{\text {avg }}$ | 0.0113 | 0.119 | 0.195 |
| $U_{2016} / U_{\text {avg }}$ | 0.589 | 0.772 | 1.00 |

## 9. ADVICE FOR MANAGERS

### 9.1. MANAGEMENT TARGETS

The Sustainable Fisheries Framework (SFF, DFO 2009) established provisional reference points to guide management and assess harvest in relation to sustainability. These reference points are the Limit Reference Point (LRP) of $0.4 B_{\text {MSY }}$ and the upper stock reference point (USR) of $0.8 B_{\mathrm{MSY}}$, which have not been adopted in this assessment due to concerns about the stability of estimating $B_{0}$ and $B_{2017}$ using the iSCAM delay-difference model (see Appendix E for discussion). In their stead, this assessment adopted historical reference points (HRPs): $B_{\text {avg }}$ (average spawning biomass from 1967-2016) as a proxy for $B_{\text {MSY }}$, and $B_{\text {min }}$ (spawning biomass in the year when the reconstructed biomass reached a minimum from which it subsequently recovered to $B_{\text {avg }}$ ) in place of $0.4 B_{\text {MSY }}$. The current ( $B_{2017}$ ) spawning biomass is used as an indicator for the probability of an increase or decrease. Therefore, the following reference points are used:

- Current spawning biomass: $B_{2017}$
- Limit Reference Point (LRP): $B_{\text {min }}$
- Upper Stock Reference (USR): $2 B_{\text {min }}$
- $B_{\text {MSY }}$ proxy: $B_{\text {avg }}$ (average over the years 1967-2016)
- Average removal rate: $u_{\text {avg }}$ (average over the years 1967-2016)


### 9.2. HARVEST ADVICE

### 9.2.1. BC North

Figure 12 shows the current stock status ( $B_{2017} / B_{\text {avg }}$ ) relative to two historical-based reference points $\left(B_{\text {min }} / B_{\text {avg }}=\mathrm{LRP}\right.$ and $\left.2 B_{\text {min }} / B_{\text {avg }}=\mathrm{USR}\right)$ for the BC North Model Average Composite posterior and for each of the component runs comprising the Model Average Composite posterior. This plot shows that the 2017 biomass for the Model Average run is evaluated to be primarily above the USR.

A decision table of probabilities, based on Model Average Composite posterior (Table 6), forms the basis of the advice to managers. Note that the probabilities for 2017 cannot change as the 2016 catches have already been taken. The probability that the estimated spawning biomass at the beginning of $2017\left(B_{2017}\right)$ is greater than the LRP $\left(B_{\min }\right)$ is 0.99 , greater than the USR $\left(2 B_{\text {min }}\right)$ is 0.62 and greater than $B_{\text {avg }}$ is 0.27 . The estimated harvest rate $u_{2016}$ has a probability of 0.74 of being greater than the average removal rate ( uavg ), indicating that the 2016 harvest rate is likely above this indicator.

Table 6. BC North: Decision table for the Model Average Composite scenario for 5 reference points: the current year spawning biomass, the limit reference point $B_{\text {min }}$, the upper stock reference $2 B_{\text {min }}$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period; for projection-year biomass $B_{2018}$ and mid-year harvest rate $u_{2017}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities are the proportion of MCMC samples from three pooled scenarios chosen for their acceptable MCMC diagnostics. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\text {min }}\right)=0.99, P\left(B_{2017}>2 B_{\text {min }}\right)=$ $0.62, P\left(B_{2017}>B_{\text {avg }}\right)=0.27$, and $P\left(u_{2016}>u_{\text {avg }}\right)=0.74$. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

| Catch $(\mathrm{t})$ | $\mathrm{P}\left(B_{2018}>B_{2017}\right)$ | $\mathrm{P}\left(B_{2018}>B_{\text {min }}\right)$ | $\mathrm{P}\left(B_{2018}>2 B_{\text {min }}\right)$ | $\mathrm{P}\left(B_{2018}>B_{\text {avg }}\right)$ | $\mathrm{P}\left(u_{2017}>u_{\text {avg }}\right)$ |
| :---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.23 | 0.99 | 0.58 | 0.26 | 0 |
| 100 | 0.17 | 0.99 | 0.56 | 0.26 | 0 |
| 200 | 0.13 | 0.98 | 0.54 | 0.25 | 0 |
| 300 | 0.10 | 0.98 | 0.53 | 0.25 | 0.01 |
| 400 | 0.08 | 0.97 | 0.51 | 0.24 | 0.08 |
| 500 | 0.07 | 0.96 | 0.50 | 0.23 | 0.24 |
| 600 | 0.06 | 0.95 | 0.49 | 0.22 | 0.45 |
| 700 | 0.05 | 0.94 | 0.47 | 0.22 | 0.59 |
| 800 | 0.04 | 0.93 | 0.46 | 0.21 | 0.65 |
| 900 | 0.03 | 0.91 | 0.45 | 0.21 | 0.68 |
| 1000 | 0.03 | 0.90 | 0.43 | 0.20 | 0.70 |
| 1200 | 0.02 | 0.87 | 0.42 | 0.19 | 0.74 |
| 1400 | 0.01 | 0.84 | 0.40 | 0.18 | 0.80 |
| 1600 | 0.01 | 0.80 | 0.39 | 0.16 | 0.85 |
| 1800 | 0.01 | 0.76 | 0.38 | 0.15 | 0.90 |
| 2000 | 0.01 | 0.71 | 0.37 | 0.13 | 0.93 |
| 2500 | 0 | 0.62 | 0.35 | 0.11 | 0.98 |
| 300 | 0 | 0.54 | 0.34 | 0.09 | 0.99 |
| 3500 | 0 | 0.48 | 0.32 | 0.07 | 1 |
| 4000 | 0 | 0.43 | 0.30 | 0.05 | 1 |
| 4500 | 0 | 0.40 | 0.28 | 0.04 | 1 |
| 5000 | 0 | 0.37 | 0.26 | 0.03 | 1 |

The average annual removals by the trawl fishery over the last five years (2011-2015) from the BC North stock were 992 t . This value can guide a manager by locating a similar catch strategy
(1000 t) in Table 6 (shaded row). If this annual catch were maintained over one year, the probability that $B_{2018}$ will be greater than $B_{2017}$ is 0.03 , i.e., a decline in stock abundance is expected with a high probability. Additionally, the probability that $B_{2018}$ will be greater than the LRP and USR is 0.90 and 0.43 , respectively, which should be seen as a cautionary indicator. However, managers should keep in mind that projections (Figure 13) from this simple delaydifference model are uncertain because it has no latent age structure to inform predictions and the stock-recruitment function is poorly determined. Appendix F supplies an additional table for projections to 2019.


Figure 12. $B C$ North: Status of the current stock $B_{2017}$ relative to $B_{\text {avg }}$ with the circles showing median biomass reference points ( $B_{\text {min }} / B_{\text {avg }}$ [red], $2 B_{\text {min }} / B_{\text {avg }}$ [green]), where $B_{\text {avg }}$ is a proxy for $B_{M S Y}, B_{\text {min }}$ is the limit reference point (LRP), and $2 B_{\text {min }}$ is the upper stock reference point (USR). The $90 \%$ credibility range is shown for the LRP and USR. Stock status is shown for the Model Average Composite scenario comprising three pooled model runs (see Table 1 for definitions of these model runs). Box plots show the $5,25,50,75$ and 95 percentiles from the MCMC posteriors. $M=$ instantaneous natural mortality $\left(y^{-1}\right) ; k=$ age (y) at knife-edge recruitment.

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Figure 13. BC North. Median estimates (solid black line) and 90\% credibility intervals (black dashed lines, grey fill) for the model-average $B_{t}$ (biomass in year $t$ in tonnes) for Walleye Pollock. Also shown are the initial biomass $B_{1967}$ (green circle), current biomass $B_{2017}$ (yellow circle), two-year projections $B_{2018-19}$ (pink fill), the median of average biomass $B_{\text {avg }}$ (blue dotted line), the historical catch (red bars) and the catch strategy (pink bars, $1000 t$ ).

### 9.2.2. BC South

Figure 14 shows the current stock status ( $B_{2017} / B_{\text {avg }}$ ) relative to two historical-based reference points $\left(B_{\text {min }} / B_{\text {avg }}=\mathrm{LRP}\right.$ and $\left.2 B_{\text {min }} / B_{\text {avg }}=\mathrm{USR}\right)$ for the BC South Model Average Composite posterior and for each of the component runs comprising the Model Average Composite posterior. This plot shows that the Model Average run is evaluated to be above the $90 \%$ credibility interval of the USR, as are all of the component runs where $k=3$.

A decision table of probabilities, based on the Model Average Composite posterior (Table 7), forms the basis of the advice to managers for this stock. Note that the probabilities for 2017 cannot change as the 2016 catches have already been taken. The probability that the estimated spawning biomass at the beginning of 2017 ( $B_{2017}$ ) is greater than the LRP $\left(B_{\text {min }}\right)$ is 1 , greater than the USR $\left(2 B_{\min }\right)$ is 0.96 and greater than $B_{\text {avg }}$ is 0.34 . The estimated harvest rate $u_{2016}$ has a probability of 0.05 of being greater than the estimated average removal rate ( $u_{\text {avg }}$ ), indicating that the current harvest rate is likely below this indicator.

Table 7. BC South: Decision table for the Model Average scenario for 5 reference points: the current year spawning biomass, the limit reference point $B_{\text {min }}$, the upper stock reference $2 B_{\text {min }}$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period; for projectionyear biomass $B_{2018}$ and mid-year harvest rate $u_{2017}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities are the proportion of MCMC samples from six pooled scenarios chosen for their acceptable MCMC diagnostics. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\text {min }}\right)=1, P\left(B_{2017}>2 B_{\text {min }}\right)=0.96, P\left(B_{2017}>B_{\text {avg }}\right)=0.34$, and $P\left(u_{2016}>u_{\text {avg }}\right)=0.05$. For reference, the average catch over the last 5 years (2011-2015) is $3,256 t$.

| Catch $(\mathrm{t})$ | $\mathrm{P}\left(B_{2018}>B_{2017}\right)$ | $\mathrm{P}\left(B_{2018}>B_{\min }\right)$ | $\mathrm{P}\left(B_{2018}>2 B_{\min }\right)$ | $\mathrm{P}\left(B_{2018}>B_{\text {avg }}\right)$ | $\mathrm{P}\left(u_{2017}>u_{\text {avg }}\right)$ |
| :---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.05 | 1 | 0.96 | 0.20 | 0 |
| 500 | 0.03 | 1 | 0.95 | 0.19 | 0 |
| 1000 | 0.02 | 1 | 0.95 | 0.18 | 0 |
| 1500 | 0.01 | 1 | 0.95 | 0.17 | 0 |
| 1750 | 0.01 | 1 | 0.95 | 0.16 | 0.02 |
| 2000 | 0.01 | 1 | 0.95 | 0.16 | 0.07 |
| 2250 | 0.01 | 1 | 0.95 | 0.15 | 0.20 |
| 2500 | 0.01 | 1 | 0.95 | 0.15 | 0.38 |
| 2750 | 0.01 | 0.99 | 0.95 | 0.15 | 0.56 |
| 3000 | 0 | 0.99 | 0.95 | 0.14 | 0.73 |
| 3250 | 0 | 0.99 | 0.95 | 0.14 | 0.85 |
| 3500 | 0 | 0.99 | 0.95 | 0.13 | 0.93 |
| 4000 | 0 | 0.99 | 0.95 | 0.13 | 0.99 |
| 4500 | 0 | 0.99 | 0.94 | 0.12 | 1 |
| 5000 | 0 | 0.98 | 0.94 | 0.11 | 1 |
| 5500 | 0 | 0.98 | 0.94 | 0.11 | 1 |
| 6000 | 0 | 0.98 | 0.94 | 0.11 | 1 |
| 6500 | 0 | 0.98 | 0.93 | 0.10 | 1 |
| 7000 | 0 | 0.97 | 0.93 | 0.10 | 1 |
| 8000 | 0 | 0.97 | 0.91 | 0.09 | 1 |
| 9000 | 0 | 0.96 | 0.87 | 0.09 | 1 |
| 10000 | 0 | 0.94 | 0.82 | 0.08 | 1 |

The average annual removals by the trawl fishery over the last five years (2011-2015) from the BC South stock were $3,256 \mathrm{t}$. A catch strategy of 3250 t in Table 7 (shaded row) indicates that if this annual catch were maintained over one year, the probability that $B_{2018}$ will be greater than $B_{2017}$ is 0 , i.e., a decline in stock abundance is expected (Figure 14). The probability that $B_{2018}$ will be greater than the LRP and USR is 0.99 and 0.95 , which indicates that the stock should remain above these reference points (Figure 14), lying well above the two $B_{\text {min }}$ reference points. Again it should be noted that the projections (Figure 15) by this model are uncertain because it
has no latent age structure to inform predictions and the stock-recruitment function is poorly determined. Appendix F supplies an additional table for projections to 2019.


Figure 14. BC South: Status of the current stock $B_{2017}$ relative to $B_{\text {avg }}$ with the circles showing median biomass reference points ( $B_{\text {min }} / B_{\text {avg }}[r e d], 2 B_{\text {min }} / B_{\text {avg }}$ [green]), where $B_{\text {avg }}$ is a proxy for $B_{M S} \gamma, B_{\text {min }}$ is the limit reference point (LRP), and $2 B_{\text {min }}$ is the upper stock reference point (USR). The $90 \%$ credibility range is shown for the LRP and USR. Stock status is shown for the Model Average Composite scenario comprising six pooled model runs (see Table 1 for definitions of these model runs). Box plots show the 5, $25,50,75$ and 95 percentiles from the MCMC posteriors. $M=$ instantaneous natural mortality $\left(y^{-1}\right) ; k=$ age (y) at knife-edge recruitment

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Figure 15. BC South. Median estimates (solid black line) and 90\% credibility intervals (black dashed lines, grey fill) for the model-average $B_{t}$ (biomass in year $t$ in tonnes) for Walleye Pollock. Also shown are the initial biomass $B_{1967}$ (green circle), current biomass $B_{2017}$ (yellow circle), two-year projections $B_{2018-19}$ (pink fill), the median of average biomass Bavg (blue dotted line), the historical catch (red bars) and the catch strategy (pink bars, $3250 t$ ).

### 9.3. ASSESSMENT SCHEDULE

Advice was also requested concerning the appropriate time interval between future assessments and, for the interim years between assessments, potential values of indicators that could trigger a full assessment earlier than usual (as per DFO, 2016). We suggest the next full stock assessment be scheduled for 2022, such that there will be new indices available from the synoptic surveys (three for HS, two for WCHG, three for QCS, and two for WCVI). By then, otoliths aged by break-and burn should be available to determine growth functions for the

BC North and BC South stocks of Walleye Pollock. Having considered the possible indicators that could be monitored in the interim years, we conclude that none are suitable for triggering an earlier-than-scheduled full assessment. Although advice for the interim years is explicitly included in this assessment in the form of decision tables (see Table 6 and Table 7), the predictions in these tables should be viewed cautiously as the delay-difference model used in this stock assessment has no latent age structure to inform predictions and the stockrecruitment function is poorly determined.

## 10. DISCUSSION AND CONCLUSIONS

The median estimate for current stock status ( $B_{2017} / B_{\text {avg }}$ ) for the Model Average Composite stock is estimated to be 0.68 for BC North (Table 3) and 0.90 for BC South (Table 5). Both stocks are expected to decline over the next one to two years at a level of catch consistent with the 20112015 average catch in each region (North = 1000 t/year, South = 3250 t/year). These declines will affect the two stocks similarly with respect to stock status - in the North there is a high probability ( 0.90 ) that $B_{2018}$ will be greater than the LRP $B_{\min }$ (Table 6) while in the South the probability is almost certain (0.99) that $B_{2018}$ will be greater than the LRP $B_{\text {min }}$ (Table 7)

This stock assessment is not capable of giving advice on equilibrium levels of yield, nor does it provide confidence in the absolute stock size, given that the available data can be fit reasonably well across a wide range of stock production hypotheses and that equilibrium calculations can vary depending on the definition of the equilibrium biomass. Following the examples of Pacific Cod (Forrest et al. 2015) and Rock Sole (Holt et al. 2016), this assessment uses historical reference points to guide managers on the sustainability of the Walleye Pollock removals by the trawl fleets (bottom and midwater). There was no simulation performed to determine the suitability of these reference points, but $B_{\text {min }}$ as a limit reference point is thought to be a reasonable benchmark because the stock has declined to this level in the past and has recovered from it.

The stock assessment projections indicate that recent catches will reduce the biomass over the next three years once the information from biomass indices is no longer available. This drop indicates that stock abundance has been maintained in the past through good recruitment or possibly stock productivity is too low. Projections are always less reliable than stock reconstruction results, but these limitations should increase the caution when considering the projection probabilities presented in Table 6 and Table 7 compared to the stock status results presented in Figure 12 and Figure 14.
We are aware that the assessment of this species has a number of important limitations. These are mainly related to the lack of reliable ageing of this species from BC waters. It is possible that this lack could be remedied in advance of the next stock assessment for this species as there exists a large number of unprocessed ageing structures in storage at the Pacific Biological Station.

It is suspected that the BC North stock may be part of a larger SE Alaska stock (Gustafson et al. 2000, and references therein), which suggests that it should not be evaluated as a unit stock. This possibility should be further explored, because concepts such as stock status and longterm yields need to be evaluated for the total stock, not just the part of it that appears in BC waters. The SE Gulf of Alaska stock is lightly exploited, given the long-term prohibition of trawling in the SE Alaskan panhandle, which may possibly provide a "rescue effect" (i.e. replenishment of the BC part of the population) from this larger, less exploited parent population.

Within BC waters, there is uncertainty with respect to the stock structure adopted by this stock assessment. While there is a clear difference in mean size between northern and southern Pollock, it is unclear how this large difference has been maintained. The simplest explanation is that the stocks are distinct, which is the hypothesis adopted here. However, there are other processes which could cause this observation, including migration of older fish from south to north and differential exploitation rates. Without reliable ageing, it is not possible to rule out the first alternative hypothesis, although the NOAA survey data indicate the presence of younger age classes in the Eastern Gulf of Alaska. The hypothesis of differential harvest rates seems less likely, given the relative equivalence of the catch histories over much of the recent 25 years.
This assessment of Walleye Pollock touches on a number of issues and concerns. Table 8 attempts to summarise them.

Table 8. Summary table of issues encountered in the stock assessment of Walleye Pollock.
$\left.\begin{array}{lll}\hline \text { Concern } & \text { Issue } & \text { Solution } \\ \hline \text { Boundary } & \begin{array}{l}\text { BC North may only represent the southern } \\ \text { stocks }\end{array} & \begin{array}{l}\text { Acknowledge that perceived changes in } \\ \text { therefore, sustainable fishing may be } \\ \text { irrelevant if there is an ongoing rescue effect. }\end{array} \\ \text { stock may be due to factors other than } \\ \text { population dynamics. }\end{array}\right]$

## 11. FUTURE RESEARCH AND DATA REQUIREMENT

The following issues should be considered when planning future stock assessments and management evaluations for Walleye Pollock.

1. Determine the most reliable method for ageing this species in BC. While the Alaskan Pollock are aged using otolith break and burn methodology, there is uncertainty whether this procedure is suitable for BC Pollock.
2. Available BC Pollock ageing data (currently only available on paper) from the 1980s should be entered into the DFO data system.
3. Review the existing otolith repository and begin break-and-burn ageing if this ageing method is deemed reliable and where there exist sufficient samples/specimens to yield useful stock assessment data.
4. If otolith ageing is deemed reliable, review the otolith sampling plans for Walleye Pollock in the commercial fishery and in the synoptic surveys, and adjust as needed to ensure that stocks are adequately represented.
5. Collect length-stratified biological samples from the commercial fishery and from research surveys to ensure that age structures represent the full size range of WAP in BC.
6. Reassess the growth curves for BC WAP stocks when reliable ages become available.

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## APPENDIX A. CATCH DATA

## A.1. BRIEF HISTORY OF THE FISHERY

The early history of the British Columbia (BC) trawl fleet is discussed by Forrester and Smith (1972); however, the document focuses on major stocks of flatfish, cod, and rockfish with no mention of Walleye Pollock. Saunders et al. (1989) noted that aside from market demand for Walleye Pollock, Canadian landings of this species followed its abundance in the northeast Pacific Ocean.

A Pollock fishery in Queen Charlotte Strait (PMFC minor area 12) reportedly started in 1992 to target spawning fish in the first quarter of the year (Saunders and Andrews 19984). Saunders et al. (1989) identify four primary spawning grounds for Walleye Pollock in BC waters:

- Dixon Entrance/northern Hecate Strait,
- Queen Charlotte Sound,
- South West Vancouver Island, and
- the Strait of Georgia,
which are broadly outlined in Figure A. 1 as the adopted areas for assessing this species. A summary of BC spawning areas appears in Gustafson et al. (2000).
The highest capture rates by the bottom trawl commercial fishery (averaged over 1996-2017, Figure A.2) occur in:
- Dixon Entrance (perhaps as part of a larger South East Alaska population, Thompson 1981),
- upper Moresby Gully,
- lower Queen Charlotte Sound outside Queen Charlotte Strait, and
- off Juan de Fuca Strait (gyre in summer).

The midwater trawl data (Figure A.3) show catch rates 10 times higher than those for the bottom trawl data, and more likely correspond to the spawning grounds (e.g., Queen Charlotte Strait). The Strait of Georgia (or Gulf) spawning stock shows up as two discrete patches (north and south) in Figure A.3.

Starting in April, 2012, groundfish bottom trawl activities were confined to a trawl footprint agreed to by the commercial fishing industry, DFO management, and the David Suzuki Foundation (Wallace et al. 2015; DFO 2016).

[^3]

Figure A.1. Walleye Pollock assessment areas comprising Pacific Marine Fisheries Commission (PMFC) major and minor areas - green for 5CDE, orange for 5AB + minor area 12, blue for $3 C D+$ minor area 20 , and red for 4B less minor areas 12 and 20. The Groundfish Management Unit area boundaries, which differ from PMFC area, are superimposed for comparison. This assessment is for areas called 'North' (5CDE, green) and 'South' (5AB3CD, orange + blue).

Table A.1. Annual trawl Total Allowable Catches (TACs) in tonnes for Walleye Pollock in groundfish management areas. Entries '---' denote no TAC set; empty cells denote unknown TAC. Note: year can either be calendar year (1994-1996) or fishing year (1997 on).

| Year | Start | End | 5CDE | 5AB | 3CD | Gulf | Coast | Outside | Notes |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1981 | 1/1/1981 | 12/31/1981 | - | - | - | 3400 | 3400 | - | a |
| 1982 | 1/1/1982 | 12/31/1982 | - | - | - | 3400 | 3400 | - | - |
| 1983 | 1/1/1983 | 12/31/1983 | - | - | - | 3400 | 3400 | - | - |
| 1984 | 1/1/1984 | 12/31/1984 | - | - | - | 3400 | 3400 | - | - |
| 1985 | 1/1/1985 | 12/31/1985 | - | - | - | 3400 | 3400 | - | - |
| 1986 | 1/1/1986 | 12/31/1986 | - | - | - | 3400 | 3400 | - | - |
| 1987 | 1/1/1987 | 12/31/1987 | - | - | - | 3400 | 3400 | - | - |
| 1988 | 1/1/1988 | 12/31/1988 | - | - | - | 3400 | 3400 | - | - |
| 1989 | 1/1/1989 | 12/31/1989 | - | - | - | 3400 | 3400 | - | - |
| 1990 | 1/1/1990 | 12/31/1990 | - | - | - | 3400 | 3400 | - | - |
| 1991 | 1/1/1991 | 12/31/1991 | - | - | - | 3700 | 3700 | - | - |
| 1992 | 1/1/1992 | 12/31/1992 | - | - | - | 3700 | 3700 | - | - |
| 1993 | 1/1/1993 | 12/31/1993 | - | - | - | 3700 | 3700 | - | - |
| 1994 | 1/1/1994 | 12/31/1994 | - | - | - | 2000 | 2000 | - | b |
| 1995 | 1/1/1995 | 12/31/1995 | 2900 | 1750 | - | 2260 | 6910 | 4650 | c |
| 1996 | 2/6/1996 | 3/31/1997 | 3190 | 1898 | - | 1490 | 6578 | 5088 | d, e |
| 1997 | 4/1/1997 | 3/31/1998 | 825 | 1790 | 270 | 1115 | 4000 | 2885 | f |
| 1998 | 4/1/1998 | 3/31/1999 | 825 | 1790 | - | 1115 | 3730 | 2615 | - |
| 1999 | 4/1/1999 | 3/31/2000 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2000 | 4/1/2000 | 3/31/2001 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2001 | 4/1/2001 | 3/31/2002 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2002 | 4/1/2002 | 3/31/2003 | 1320 | 1790 | - | 1115 | 4225 | 3110 | g,h |
| 2003 | 4/1/2003 | 3/31/2004 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2004 | 4/1/2004 | 3/31/2005 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2005 | 4/1/2005 | 3/31/2006 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2006 | 4/1/2006 | 3/31/2007 | 1320 | 1790 | - | 1115 | 4225 | 3110 | i |
| 2007 | 3/10/2007 | 3/31/2008 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2008 | 3/8/2008 | 2/20/2009 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2009 | 2/21/2009 | 2/20/2010 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2010 | 2/21/2010 | 2/20/2011 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2011 | 2/21/2011 | 2/20/2013 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2012 | 2/21/2011 | 2/20/2013 | 1320 | 1790 | - | 1115 | 4225 | 3110 | j |
| 2013 | 2/21/2013 | 2/20/2014 | 1320 | 1790 | - | 1115 | 4225 | 3110 | k |
| 2014 | 2/21/2014 | 2/20/2015 | 1320 | 1790 | - | 1115 | 4225 | 3110 | - |
| 2015 | 2/21/2015 | 2/20/2016 | 1320 | 1790 | - | 1115 | 4225 | 3110 | I |
| 2016 | 2/21/2016 | 2/20/2017 | 1320 | 1790 | - | 1115 | 4225 | 3110 | m |

[^4]Table A.2. Codes to notes on management actions and quota adjustments that appear in Table A.1. Abbreviations under 'Management Actions': WAP = Walleye Pollock, DMP = dockside monitoring program, $I V Q=$ individual vessel quota, lbs = pounds ( $0.4536 \mathrm{~kg} / \mathrm{lb}$ ).

|  | Year | Management Actions* |
| :---: | :---: | :---: |
| a | 1981 | Pollock TAC (1981-1994): only 4B=Areas 13-18, 29 |
| b | 1994 | Started DMP for Trawl fleet. |
| c | 1995 | Pollock TAC areas: $5 \mathrm{CDE}=5 \mathrm{CD}$; 5AB=Area 12; 4B=Areas 13-18, 29 |
| d | 1996 | Started 100\% onboard observer program for offshore Trawl fleet. |
| e | 1996 | Pollock TAC areas: $5 \mathrm{CDE}=5 \mathrm{CD}$; 5AB=Areas 11,12; 4B=Areas 13-18, 29 |
| f | 1997 | Started IVQ system for Trawl Total Allowable Catch (TAC) species (April 1, 1997) |
| g | 2002 | Established the inshore rockfish conservation strategy. |
| h | 2002 | Closed areas to preserve four hexactinellid (glassy) sponge reefs. |
| i | 2006 | Introduced an Integrated Fisheries Management Plan (IFMP) for most groundfish fisheries. |
| j | 2012 | Freeze the footprint of where groundfish bottom trawl activities can occur (all vessels under the authority of a valid Category "T" commercial groundfish trawl license selecting Option A as identified in the IFMP). |
| k | 2013 | To support groundfish research the Groundfish Trawl Industry agreed to the trawl TAC offsets to account for unavoidable mortality incurred in during the 2013 DFO and Trawl industry agreed upon Groundfish Trawl Multi-species surveys: WAP in 5CDE $=2.2 \mathrm{t}$, $5 \mathrm{AB}=1.2 \mathrm{t}$. |
| I | 2015 | Research allocations for 2015 to account for the mortalities associated with survey catches within TACs: WAP=4.3 t. |
| m | 2016 | Research allocations for 2016 to account for the mortalities associated with survey catches within TACs: WAP $=0.3 \mathrm{t}$. |



Figure A.2. Aerial distribution of Walleye Pollock bottom trawl tow mean catch per unit effort (kg/hour) from Feb 17, 1996 to Sep 24, 2017. Isobaths show the 100, 200, and 500 m depth contours. Note that cells with $<3$ fishing vessels are not displayed. Each cell represents, on average, 32 km².


Figure A.3. Aerial distribution of Walleye Pollock midwater trawl tow mean catch per unit effort (kg/hour) from Feb 17, 1996 to Sep 24, 2017. Isobaths show the 100, 200, and 500 m depth contours. Note that cells with $<3$ fishing vessels are not displayed. Each cell represents, on average, 32 km².


Figure A.4. Aerial distribution of accumulated Walleye Pollock catch (tonnes) from bottom trawls before the introduction of the trawl footprint (green polygons) in April 2012, limiting areas in which trawl vessels can operate. Note that cells with $<3$ fishing vessels are not displayed. Each cell represents, on average, 32 km².


Figure A.5. Aerial distribution of accumulated Walleye Pollock catch (tonnes) from bottom trawls after the introduction of the trawl footprint (green polygons) in April 2012. See Figure A. 2 captions for details.

## A.2. CATCH HISTORY

This assessment collects reported catches back to 1954 but considers the start of the fishery to be 1967 (Figure A.6, Table A.3) when the Canadian domestic fleet started to increase the capture of Walleye Pollock.

Starting in 2015, all official catch tables from the databases below have been merged into one table called "GF_MERGED_CATCH", which is available in DFO's GFFOS database. All groundfish DFO databases are now housed on the DFBCV9TWVASP001 server (formerly on the SVBCPBSGFIIS server). Walleye Pollock catch by fishery sector ultimately comes from the following four DFO databases:

- GFCatch (1954-1995) - trawl and trap (Rutherford 1999);
- PacHarvest observer trawl (1996-2007) - trawl;
- GFFOS groundfish subset from Fishery Operation System (2006-2016) - all fisheries, gear types, and modern surveys; and
- GFBio joint-venture hake and research survey catches (1947-2016) - multiple gear types.

However, all these data sources are superseded by GFFOS from 2007 on because this latter repository was designed to record all landings and discards from commercial fisheries and research activities. Prior to this, midwater landings of pollock only appeared in either GFCatch or GFBio. The latter was designed primarily to hold biological information from fish samples but also served to record catch from J-V hake fishing events and research survey activity.

## A.2.1. Coastwide Stock

Total annual catches used in the delay-difference population model comprised those from bottom and midwater trawls in three offshore areas - 5CDE (around Haida Gwaii), 5AB (Queen Charlotte Sound and Strait), and 3CD (west coast Vancouver Island plus the entrance to Juan de Fuca Strait) - and areas unknown (Table A.3, Figure A.6). Only catches in the Strait of Georgia (4B) were excluded from this assessment. Additionally, catches not in 4B are presented by fishing gear (Table A.4) and by DFO database (Table A.5). These include records from nontrawl fisheries; however, the amounts from these sources are so minor that the population assessment only considered removals by trawl gear.

Table A.3. Reported catches (in tonnes, landings + releases) of WAP in PMFC 5CDE +5AB+3CD by coastal region. Catches from the trawl fishery explicitly excluding the Gulf region (4B) were used in the population model (see Table A.7). Catch for 2016 is not complete.

| Year | 5CDE | 5AB | 3CD | UNK | Total | Year | 5CDE | 5AB | 3CD | UNK | Total |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1954 | 0 | 14.9 | 5.93 | 0 | 21 | 1986 | 99.7 | 8.39 | 3.74 | 0 | 112 |
| 1955 | 2.77 | 1.4 | 5.55 | 0 | 10 | 1987 | 34.7 | 35.3 | 3.62 | 0 | 74 |
| 1956 | 14.5 | 20.6 | 52.8 | 0 | 88 | 1988 | 52.1 | 13.3 | 265 | 0 | 330 |
| 1957 | 7.26 | 3.16 | 4.35 | 0 | 15 | 1989 | 42.5 | 27.6 | 944 | 0 | 1014 |
| 1958 | 14.4 | 3.38 | 15.9 | 0 | 34 | 1990 | 422 | 142 | 625 | 0 | 1189 |
| 1959 | 1.72 | 2.24 | 18.6 | 0 | 23 | 1991 | 529 | 48.5 | 454 | 0 | 1032 |
| 1960 | 9.47 | 5.87 | 12.1 | 0 | 27 | 1992 | 1372 | 1772 | 1769 | 0 | 4913 |
| 1961 | 6.2 | 1 | 3.46 | 0 | 11 | 1993 | 4447 | 3828 | 671 | 0 | 8946 |
| 1962 | 11.8 | 0 | 20.3 | 0 | 32 | 1994 | 1344 | 3279 | 192 | 0.005 | 4815 |
| 1963 | 3.54 | 5.87 | 11.9 | 0 | 21 | 1995 | 1685 | 2574 | 16.3 | 0 | 4275 |
| 1964 | 2.08 | 5.56 | 22.9 | 0 | 31 | 1996 | 887 | 685 | 2812 | 52.9 | 4437 |
| 1965 | 9.21 | 0 | 30.2 | 0 | 39 | 1997 | 612 | 128 | 268 | 11.4 | 1019 |
| 1966 | 134 | 1.27 | 26.5 | 0 | 162 | 1998 | 819 | 61 | 3.14 | 5.46 | 889 |
| 1967 | 68 | 2.38 | 4.27 | 1 | 76 | 1999 | 1213 | 34.7 | 6.22 | 4.52 | 1258 |
| 1968 | 17.6 | 6.67 | 3.68 | 0 | 28 | 2000 | 987 | 57.5 | 99.6 | 4.57 | 1149 |
| 1969 | 47.2 | 33.2 | 32.5 | 0 | 113 | 2001 | 122 | 6.8 | 2854 | 50 | 3033 |
| 1970 | 7.5 | 0 | 34.9 | 0 | 42 | 2002 | 84 | 19.5 | 2726 | 246 | 3076 |
| 1971 | 0 | 0 | 47.4 | 0 | 47 | 2003 | 625 | 1723 | 2516 | 65 | 4929 |
| 1972 | 1.03 | 172 | 70.1 | 0 | 243 | 2004 | 1036 | 590 | 675 | 71.4 | 2372 |
| 1973 | 23.5 | 70.4 | 4.98 | 0 | 99 | 2005 | 501 | 851 | 154 | 277 | 1783 |
| 1974 | 49.8 | 19.4 | 0 | 0 | 69 | 2006 | 543 | 2863 | 105 | 113 | 3624 |
| 1975 | 132 | 34.4 | 18.1 | 0 | 185 | 2007 | 1302 | 2095 | 2.24 | 30.7 | 3430 |
| 1976 | 818 | 469 | 17.4 | 0 | 1304 | 2008 | 354 | 1090 | 4.96 | 21.6 | 1471 |
| 1977 | 659 | 244 | 55.3 | 0 | 958 | 2009 | 1430 | 2004 | 283 | 139 | 3856 |
| 1978 | 1776 | 324 | 51.6 | 0 | 2152 | 2010 | 1702 | 1976 | 112 | 26.5 | 3817 |
| 1979 | 1923 | 164 | 63.6 | 0 | 2151 | 2011 | 831 | 1935 | 1341 | 31 | 4138 |
| 1980 | 1285 | 41 | 27.2 | 0 | 1353 | 2012 | 1129 | 1751 | 1226 | 41.6 | 4148 |
| 1981 | 693 | 15 | 184 | 0 | 892 | 2013 | 824 | 1345 | 154 | 25.7 | 2349 |
| 1982 | 824 | 7.38 | 105 | 0 | 936 | 2014 | 643 | 2013 | 4092 | 33 | 6781 |
| 1983 | 1084 | 20.8 | 6.99 | 0 | 1112 | 2015 | 1532 | 1858 | 563 | 121 | 4074 |
| 1884 | 699 | 18.7 | 7.52 | 0 | 725 | 2016 | 170 | 714 | 20.5 | 52 | 957 |
| 1985 | 1180 | 1.38 | 8.56 | 0 | 1190 | - | - | - | - | - | - |
|  |  |  |  |  |  |  |  |  |  |  |  |

Table A.4. Reported catches (in tonnes, landings + releases) of WAP in PMFC 5CDE + 5AB+3CD by fishing gear. Catch for 2016 is not complete.

| Year | Bottom Trawl | Midw. Trawl | Trap | Hook\& Line | Total | Year | $\begin{gathered} \text { Bottom } \\ \text { Trawl } \end{gathered}$ | Midw. Trawl | Trap | Hook\& Line | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1954 | 20.9 | 0 | 0 | 0 | 21 | 1986 | 43.9 | 67.9 | 0 | 0 | 112 |
| 1955 | 9.71 | 0 | 0 | 0 | 10 | 1987 | 72.4 | 1.17 | 0 | 0 | 74 |
| 1956 | 87.8 | 0 | 0 | 0 | 88 | 1988 | 75.9 | 254 | 0 | 0 | 330 |
| 1957 | 14.8 | 0 | 0 | 0 | 15 | 1989 | 76.5 | 938 | 0 | 0 | 1015 |
| 1958 | 33.6 | 0 | 0 | 0 | 34 | 1990 | 556 | 633 | 0 | 0 | 1189 |
| 1959 | 22.5 | 0 | 0 | 0 | 23 | 1991 | 538 | 493 | 0 | 0 | 1031 |
| 1960 | 27.5 | 0 | 0 | 0 | 28 | 1992 | 1089 | 3825 | 0 | 0 | 4914 |
| 1961 | 10.5 | 0 | 0 | 0 | 11 | 1993 | 2753 | 6192 | 0 | 0 | 8945 |
| 1962 | 32.1 | 0 | 0 | 0 | 32 | 1994 | 1102 | 3713 | 0 | 0 | 4815 |
| 1963 | 21.3 | 0 | 0 | 0 | 21 | 1995 | 815 | 3460 | 0 | 0.008 | 4275 |
| 1964 | 30.6 | 0 | 0 | 0 | 31 | 1996 | 1026 | 3410 | 0 | 0 | 4436 |
| 1965 | 39.4 | 0 | 0 | 0 | 39 | 1997 | 311 | 708 | 0 | 0.001 | 1019 |
| 1966 | 162 | 0 | 0 | 0 | 162 | 1998 | 203 | 686 | 0 | 0.023 | 889 |
| 1967 | 75.6 | 0 | 0 | 0 | 76 | 1999 | 278 | 980 | 0 | 0.004 | 1258 |
| 1968 | 28 | 0 | 0 | 0 | 28 | 2000 | 172 | 977 | 0 | 0.009 | 1149 |
| 1969 | 113 | 0 | 0 | 0 | 113 | 2001 | 194 | 2840 | 0 | 0.004 | 3034 |
| 1970 | 42.4 | 0 | 0 | 0 | 42 | 2002 | 135 | 2940 | 0 | 0 | 3075 |
| 1971 | 47.4 | 0 | 0 | 0 | 47 | 2003 | 186 | 4742 | 0 | 0 | 4928 |
| 1972 | 243 | 0.02 | 0 | 0 | 243 | 2004 | 145 | 2227 | 0 | 0.005 | 2372 |
| 1973 | 97.5 | 1.42 | 0 | 0 | 99 | 2005 | 446 | 1336 | 0 | 0.006 | 1782 |
| 1974 | 66.7 | 2.53 | 0 | 0 | 69 | 2006 | 143 | 3481 | 0 | 0.049 | 3624 |
| 1975 | 123 | 61.5 | 0 | 0 | 185 | 2007 | 142 | 3288 | 0 | 0 | 3430 |
| 1976 | 899 | 406 | 0 | 0 | 1305 | 2008 | 74.6 | 1396 | 0 | 0 | 1471 |
| 1977 | 904 | 54.2 | 0 | 0 | 958 | 2009 | 110 | 3746 | 0 | 0 | 3856 |
| 1978 | 1313 | 839 | 0 | 0 | 2152 | 2010 | 117 | 3699 | 0 | 0 | 3816 |
| 1979 | 1378 | 772 | 0.005 | 0 | 2150 | 2011 | 175 | 3963 | 0 | 0.059 | 4138 |
| 1980 | 831 | 522 | 0 | 0 | 1353 | 2012 | 155 | 3993 | 0 | 0.041 | 4148 |
| 1981 | 576 | 316 | 0 | 0 | 892 | 2013 | 162 | 2186 | 0 | 0.053 | 2348 |
| 1982 | 254 | 682 | 0 | 0 | 936 | 2014 | 153 | 6628 | 0 | 0.054 | 6781 |
| 1983 | 239 | 872 | 0 | 0 | 1111 | 2015 | 284 | 3790 | 0 | 0.046 | 4074 |
| 1984 | 158 | 567 | 0 | 0 | 725 | 2016 | 96.6 | 860 | 0 | 0.013 | 957 |
| 1985 | 65.5 | 1125 | 0 | 0 | 1191 | - | - | - | - | - | - |

Table A.5. Reported catches (in tonnes, landings + releases) of WAP in PMFC 5CDE + 5AB+3CD reported by DFO database. Catch for 2016 is not complete.
\(\left.$$
\begin{array}{lrrrrr|rrrrrr}\hline \text { Year } & \begin{array}{r}\text { GF } \\
\text { Catch }\end{array} & \text { GFBio } & \begin{array}{r}\text { Pac } \\
\text { Harvest }\end{array} & \begin{array}{r}\text { GF } \\
\text { FOS }\end{array} & \text { Total } & \text { Year } & \begin{array}{r}\text { GF } \\
\text { Catch }\end{array}
$$ \& GFBio \& \begin{array}{r}Pac <br>

Harvest\end{array} \& $$
\begin{array}{r}\text { GF }\end{array}
$$ \& Total\end{array}\right]\)| FOS |
| :--- |



Figure A.6. Reported total (landed + released) catch (t) for Walleye Pollock in PMFC major areas 5CDE + $5 A B+3 C D$ from (A) all gear types, (B) DFO databases, and (C) the three regions and areas unknown (trawl fishery only). Note that catches in panels A and B include those from non-trawl fisheries; however, these are negligible.

## A.2.2. North vs. South Stocks

Initially, the assessment focused on a coastwide stock; however, after some exploration it became evident that Walleye Pollock in the North were twice as big, on average, as those in the South ( $1.04 \mathrm{~kg} /$ fish vs. $0.51 \mathrm{~kg} /$ fish, respectively, Table A.6, Figure A.7.). A meeting of the Technical Working Group on Mar 23, 2016, facilitated the decision to model the North (PMFC 5CDE) separately from the South (PMFC 5AB3CD + minor areas 12 \& 20). The catch inputs to the model appear in Table A.7. We note that catch by gear type differs by region (Figure A.8).

Table A.6. Annual mean weight (kg) of Walleye Pollock by PMFC minor areas roughly arrange in a North-to-South direction. The final row gives the geometric mean of the fish by minor area (see Figure A.7). Headers in blue indicate areas for the North stock, those in pink indicate areas for the South stock.

| Year | 35 | 3 | 4 |  | 5 | 6 |  | 27 | 8 | 11 | 12 | 23 | 21 | 20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1972 | --- | --- | --- | --- | --- | --- | --- | ----- | 0.611 | 1.227 | --- | --- | --- | --- |
| 1973 | --- | --- | --- | --- | 0.643 | 0.300 | 0.302 | 2 | --- | 1.229 | --- | --- | --- | 1.227 |
| 1974 | --- | --- | --- | --- | - |  | 0.832 | 2 | --- | --- | --- | --- | --- | --- |
| 1976 | --- | --- | 1.157 | --- | 1.293 | --- | 1.727 | 7 | --- | --- | --- | --- | --- | --- |
| 1977 | 1.562 | 1.058 | 1.323 | 1.478 | 1.037 | --- | --- | - | 1.679 | --- | --- | --- | 0.868 | --- |
| 1978 | --- | 1.348 | 1.263 | - | --- | 1.369 | 0.425 | 5 | 1.740 | 0.907 | --- | --- | --- | --- |
| 1979 | 1.481 | 1.563 | 1.423 | 1.482 | 0.981 | 1.167 | --- | -- --- | --- | --- | --- | --- | --- | 1.021 |
| 1980 | --- | --- | 0.827 | --- | --- | --- | --- | -- --- | --- | --- | --- | --- | --- | --- |
| 1981 | --- | --- | 1.100 | --- | --- | --- | --- | -- --- | --- | --- | --- | --- | --- | --- |
| 1985 | --- | --- | 1.434 | --- | --- | --- | --- | -- --- | --- | --- | --- | --- | --- | --- |
| 1988 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- | 0.709 | --- | --- |
| 1989 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- | 0.395 | 0.820 | --- |
| 1990 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- | 0.536 | 0.585 | --- |
| 1991 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- | 0.481 | --- | --- |
| 1992 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- | 0.332 | 0.356 | --- |
| 1993 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- | 0.285 | 0.271 | --- |
| 1994 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- |  | 0.374 | --- |
| 1996 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- | 0.271 | 0.294 | --- |
| 1997 | --- | 0.994 | 1.179 | --- | --- | --- | --- | - | --- | --- | --- | 0.377 | 0.401 | --- |
| 1998 | --- | 1.024 | 1.148 | --- | --- | --- | 0.720 | 0 | --- | --- | --- | --- | --- |  |
| 1999 | --- | 1.236 | 0.930 | --- | --- | --- | --- | -- --- | --- | --- | --- | --- | --- | --- |
| 2000 | --- | 1.269 | 0.833 | --- | --- | --- | --- | ----- | --- | --- | --- |  | 0.352 | --- |
| 2001 | --- | --- | --- | --- | --- | --- | --- | -- --- | --- | --- | --- | 0.230 | 0.271 | --- |
| 2002 | --- | --- | --- | --- | --- | --- | --- | ----- | --- | --- |  | 0.057 | 0.479 | --- |
| 2003 | --- | 1.008 | 1.014 | --- | --- | --- | --- | -- --- | --- | --- | 0.698 | 0.394 | --- | --- |
| 2004 | --- | --- | 0.856 | --- | --- | --- | --- | -- --- | --- | --- | 0.700 | --- | 0.497 | --- |
| 2005 | --- | --- | 1.067 | --- | --- | --- | --- | ---- | --- | 0.633 | 0.857 | --- | --- | --- |
| 2006 | --- | --- | 0.822 | --- | --- | --- | --- | ---- | --- | 0.608 | 0.632 | --- | --- | --- |
| 2007 | --- |  | 0.878 | --- | --- | --- | --- | --- --- | --- | 0.782 | 0.732 | --- | --- | --- |
| 2008 | --- | 1.225 | --- | --- | --- | --- | --- | -- --- | --- |  | 0.282 | --- | --- | --- |
| 2009 | - | 1.409 | 1.279 | --- | --- | --- | --- | -- --- | --- | --- | 0.394 | 0.596 | --- | 0.322 |
| 2010 | --- | 1.018 | 1.212 | --- | --- | --- | --- | -- --- | --- | --- | 0.508 | --- | --- | --- |
| 2011 | --- |  | --- | --- | --- | --- | --- | -- --- | 0.336 | --- | 0.527 | 0.592 | 0.448 | 0.503 |
| 2012 | --- | --- | 0.914 | 1.066 | --- | --- | --- | -- --- | --- | 0.601 | 0.575 | , |  | 0.276 |
| 2013 | --- | 0.908 | 0.974 | --- | --- | --- | --- | -- --- | --- | 0.853 | 0.609 | --- | --- | 0.243 |
| 2014 | --- | --- | --- | --- | --- | 1.067 | --- | - | --- | 0.550 | 0.571 | --- | 0.414 | 0.286 |
| 2015 | --- | --- | 1.084 | --- | --- | 0.974 |  | -0.95426 | 0.501 | --- | --- | --- | --- | --- |
| 2016 | --- | --- | 1.606 | --- | --- | --- | -- | -- | --- | --- | 0.570 | --- | --- | --- |
| Geomean | 1.521 | 1.157 | 1.085 | 1.327 | 0.959 | 0.870 | 0.668 | 80.954 | 0.786 | 0.787 | 0.569 | 0.352 | 0.430 | 0.453 |

Table A.7. Catches used in the 2016 Walleye Pollock delay-difference model. Columns labelled 'North' include catches from PMFC area 5CDE, 'South' includes catches from PMFC areas 5AB (+ minor area 12) and $3 C D$ (+ minor area 20), and 'Coast' includes catches from 'North', 'South', and unknown areas (see Table A.3). The coastwide equivalent catch from Saunders and Andrews (19984), excluding 4B, is reported in the column labelled 'WP G07-7' for comparison purposes only.

| Year | North | South | Coast | W9P | Year | North | South | Coast |
| :--- | ---: | ---: | ---: | ---: | :--- | ---: | ---: | ---: |
| 1967 | 68 | 7 | 76 | 56 | 1997 | 612 | 396 | 1019 |
| 1968 | 18 | 10 | 28 | 26 | 1998 | 819 | 64 | 889 |
| 1969 | 47 | 66 | 113 | 94 | 1999 | 1213 | 41 | 1259 |
| 1970 | 8 | 35 | 42 | 8 | 2000 | 987 | 157 | 1149 |
| 1971 | 0 | 47 | 47 | 5 | 2001 | 122 | 2861 | 3034 |
| 1972 | 1 | 242 | 243 | 173 | 2002 | 84 | 2746 | 3075 |
| 1973 | 23 | 75 | 99 | 85 | 2003 | 625 | 4239 | 4928 |
| 1974 | 50 | 19 | 69 | 61 | 2004 | 1036 | 1265 | 2373 |
| 1975 | 132 | 52 | 184 | 102 | 2005 | 501 | 1004 | 1782 |
| 1976 | 818 | 487 | 1304 | 1296 | 2006 | 543 | 2968 | 3624 |
| 1977 | 659 | 299 | 958 | 841 | 2007 | 1302 | 2097 | 3430 |
| 1978 | 1776 | 376 | 2152 | 2031 | 2008 | 354 | 1095 | 1470 |
| 1979 | 1923 | 227 | 2150 | 2045 | 2009 | 1430 | 2287 | 3856 |
| 1980 | 1285 | 68 | 1353 | 2932 | 2010 | 1702 | 2088 | 3817 |
| 1981 | 693 | 199 | 892 | 1640 | 2011 | 831 | 3276 | 4138 |
| 1982 | 824 | 112 | 936 | 1706 | 2012 | 1129 | 2977 | 4148 |
| 1983 | 1084 | 28 | 1111 | 1064 | 2013 | 824 | 1499 | 2349 |
| 1984 | 699 | 26 | 725 | 758 | 2014 | 643 | 6106 | 6781 |
| 1985 | 1180 | 10 | 1190 | 1263 | 2015 | 1532 | 2420 | 4074 |
| 1986 | 100 | 12 | 112 | 195 | 2016 | 1532 | 2420 | 4074 |
| 1987 | 35 | 39 | 74 | 1389 | - | - | - | - |
| 1988 | 52 | 278 | 330 | 269 | - | - | - | - |
| 1989 | 42 | 972 | 1014 | 975 | - | - | - | - |
| 1990 | 422 | 767 | 1189 | 1086 | - | - | - | - |
| 1991 | 529 | 503 | 1031 | 948 | - | - | - | - |
| 1992 | 1372 | 3541 | 4913 | 3501 | - | - | - | - |
| 1993 | 4447 | 4498 | 8945 | 5410 | - | - | - | - |
| 1994 | 1344 | 3471 | 4815 | 1717 | - | - | - | - |
| 1995 | 1685 | 2590 | 4275 | 2390 | - | - | - | - |
| 1996 | 887 | 3497 | 4436 | 3907 | - | - | - | - |
|  |  |  |  |  |  |  |  |  |



Figure A.7. Trend in mean weight of Walleye Pollock by PMFC minor area arranged from North to South (see Table A.6).


Figure A.8. Annual catch (t) of Walleye Pollock by gear type in the North and South areas.

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## APPENDIX B. TRAWL SURVEYS

## B.1. INTRODUCTION

This appendix summarizes the derivation of relative Walleye Pollock (WAP) abundance indices from the following bottom trawl surveys:

- a set of historical surveys operated in the Goose Island Gully (GIG) of Queen Charlotte Sound (Section B.3);
- Hecate Strait assemblage or multi-species survey (Section B.4);
- Hecate Strait synoptic survey (Section B.5);
- Queen Charlotte Sound synoptic survey (Section B.6);
- west coast Vancouver Island synoptic survey (Section B.7);
- west coast Haida Gwaii synoptic survey (Section B.8).

Only surveys which were used in the WAP stock assessment are presented. The NMFS Triennial, WCVI shrimp and QC Sound shrimp surveys have been omitted because the presence of WAP in these surveys has been sporadic, rendering these surveys poor candidates to provide reliable abundance series for this species.

## B.2. ANALYTICAL METHODS

Catch and effort data for strata $i$ in year $y$ yield catch per unit effort (CPUE) values $U_{y i}$. Given a set of data $\left\{C_{y i j}, E_{y i j}\right\}$ for tows $j=1, \ldots, n_{y i}$,

Eq. B. $1 \quad U_{y i}=\frac{1}{n_{y i}} \sum_{j=1}^{n_{v i}} \frac{C_{y i j}}{E_{y i j}}$,
where $C_{y i j}=$ catch $(\mathrm{kg})$ in tow $j$, stratum $i$, year $y$;
$E_{y i j}=$ effort (h) in tow $j$, stratum $i$, year $y$;
$n_{y i}=$ number of tows in stratum $i$, year $y$.
CPUE values $U_{y i}$ convert to CPUE densities $\delta_{y i}\left(\mathrm{~kg} / \mathrm{km}^{2}\right)$ using:
Eq. B. $2 \quad \delta_{y i}=\frac{1}{v w} U_{y i}$,
where $v=$ average vessel speed $(\mathrm{km} / \mathrm{h})$;
$w=$ average net width (km).
Alternatively, if vessel information exists for every tow, CPUE density can be expressed
Eq. B. $3 \quad \delta_{y i}=\frac{1}{n_{y i}} \sum_{j=1}^{n_{y i}} \frac{C_{y i j}}{D_{y i j} w_{y i j}}$,
where $C_{y i j}=$ catch weight $(\mathrm{kg})$ for tow $j$, stratum $i$, year $y$;

```
D Dij}=\mathrm{ distance travelled (km) for tow j, stratumi}i\mathrm{ , year y;
w
nyi}=\mathrm{ number of tows in stratum }i\mathrm{ , year }y\mathrm{ .
```

The annual biomass estimate is then the sum of the product of CPUE densities and bottom areas across $m$ strata:

Eq. B. $4 \quad B_{y}=\sum_{i=1}^{m} \delta_{y i} A_{i}=\sum_{i=1}^{m} B_{y i}$,
where $\delta_{y i}=$ mean CPUE density $\left(\mathrm{kg} / \mathrm{km}^{2}\right)$ for stratum $i$, year $y$;
$A_{i}=$ area $\left(\mathrm{km}^{2}\right)$ of stratum $i ;$
$B_{y i}=$ biomass (kg) for stratum $i$, year $y$;
$m=$ number of strata.
The variance of the survey biomass estimate $V_{y}\left(\mathrm{~kg}^{2}\right)$ follows:
Eq. B. $5 \quad V_{y}=\sum_{i=1}^{m} \frac{\sigma_{y i}^{2} A_{i}^{2}}{n_{y i}}=\sum_{i=1}^{m} V_{y i}$,
where $\sigma_{y i}^{2}=$ variance of CPUE density $\left(\mathrm{kg}^{2} / \mathrm{km}^{4}\right)$ for stratum $i$, year $y$;
$V_{y i}=$ variance of the biomass estimate $\left(\mathrm{kg}^{2}\right)$ for stratum $i$, year $y$.
The coefficient of variation (CV) of the annual biomass estimate for year $y$ is
Eq. B. $6 \quad C V_{y}=\frac{\sqrt{V_{y}}}{B_{y}}$.

## B.3. EARLY GIG SURVEYS IN QUEEN CHARLOTTE SOUND

## B.3.1. Data selection

Tow-by-tow data from a series of historical trawl surveys were available for 12 years spanning the period from 1965 to 1995. The first two surveys, in 1965 and 1966, were wide-ranging, with the 1965 survey extending from near San Francisco to halfway up the Alaskan Panhandle (Westrheim 1966a, 1967b). The 1966 survey was only slightly less ambitious, ranging from the southern US-Canada border in Juan de Fuca Strait into the Alaskan Panhandle (Westrheim 1966b, 1967b). It was apparent that the design of these two early surveys was exploratory and that these surveys would not be comparable to the subsequent Queen Charlotte Sound (QCS) surveys which were much narrower in terms of area covered and which had a much higher density of tows in the Goose Island Gully (GIG). This can be seen in the small number of tows used by the first two surveys in GIG (Table B.1). As a consequence, these surveys are not included in this series.

The 1967 ([left panel]: Figure B.1) and 1969 ([left panel]: Figure B.2) surveys (Westrheim 1967a, 1969; Westrheim et al. 1968) also performed tows on the west coast of Vancouver Island, the west coast of Haida Gwaii and SE Alaska, but both of these surveys had a reasonable number of tows in the GIG grounds (Table B.1). The 1971 survey ([left panel]:

Figure B.3) was entirely confined to GIG (Harling et al. 1971) while the 1973 ([left panel]: Figure B.4), 1976 ([left panel]: Figure B.5) and 1977 ([left panel]: Figure B.6) surveys covered both Goose Island and Mitchell Gullies in QCS (Harling et al. 1973; Westrheim et al. 1976; Harling and Davenport 1977).

A 1979 survey (Nagtegaal and Farlinger 1980) was conducted by a commercial fishing vessel (Southward Ho, Table B.1), with the distribution of tows being very different from the preceding and succeeding surveys (plot not provided; see Figure C5 in Edwards et al. 2012). As well, the distribution of tows by depth was also different from the other surveys (Table B.2). These observations imply a substantially different survey design and consequently this survey was not included in the time series.

The 1984 survey was conducted by two vessels: the G.B. Reed and the Eastward Ho (Nagtegaal et al. 1986). Part of the design of this survey was to compare the catch rates of the two vessels (one was a commercial fishing vessel and the other a government research vessel - Greg Workman, DFO, pers. comm.), thus they both followed similar design specifications, including the configuration of the net. Unfortunately, the tows were not distributed similarly in all areas, with the G.B. Reed fishing mainly in the shallower portions of the GIG, while the Eastward Ho fished more in the deeper and seaward parts of the GIG ([left panel]: Figure B.7) although the two vessels fished more contiguously in Mitchell Gully (immediately to the north). When the depth-stratified catch rates for POP (the main design species of the surveys) of the two vessels were compared within the GIG only (using a simple ANOVA), the Eastward Ho catch rates were significantly higher ( $\mathrm{p}=0.049$ ) than those observed for the G.B. Reed. However, the difference in catch rates was no longer significant when tows from Mitchell's Gully were added to the analysis ( $p=0.12$ ). Given the lack of significance when the full suite of available tows were compared, along with the uneven spatial distribution of tows among vessels within the GIG (although the ANOVA was depth-stratified, it is possible that the depth categories were too coarse), the most parsimonious conclusion was that there was no detectable difference between the two vessels. Consequently, all the GIG tows from both vessels were pooled for this survey year.
The 1994 survey (Hand et al. 1995), also conducted by a commercial vessel (the Ocean Selector, Table B.2) ([left panel]: Figure B.8), was modified by the removal of 19 tows which were part of an acoustic experiment and therefore were not considered appropriate for biomass estimation (they were tows used to estimate species composition for ensonified schools). Although this survey was designed to emulate as closely as possible the previous G.B. Reed surveys in terms of tow location selection (same fixed tow locations, G. Workman, DFO, pers. comm.), the timing of this survey was about two to three months earlier than the previous surveys (starting in mid-June rather than August or September, Table B.3).
The 1995 survey (Yamanaka et al, 1996), conducted by two commercial fishing vessels: the Ocean Selector and the Frosti (Table B.2), used a random stratified design with each vessel duplicating every tow (G. Workman, DFO, pers. comm.). This type of design was entirely different from that used in the previous surveys. As well, the focus of this survey was on Pacific Ocean Perch (POP), with tows optimised to capture this species. Given the difference in design (random stations rather than fixed locations), this survey was not used in the stock assessment.

Table B.1. Number of tows in GIG and in other areas (Other) by survey year and vessel conducting the survey for the 12 historical (1965 to 1995) surveys. Survey years in grey were not used in the assessment.

| Survey | GB Reed |  |  |  | Southward Ho |  | Eastward Ho |  | Ocean Selector |  | Frosti |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Year | Other | GIG | Other | GIG | Other | GIG | Other | GIG | Other | GIG |
| 1965 | 76 | 8 | - | - | - | - | - | - | - | - |  |
| 1966 | 49 | 15 | - | - | - | - | - | - | - | - |  |
| 1967 | 17 | 33 | - | - | - | - | - | - | - | - |  |
| 1969 | 3 | 32 | - | - | - | - | - | - | - | - |  |
| 1971 | 3 | 36 | - | - | - | - | - | - | - | - |  |
| 1973 | 13 | 33 | - | - | - | - | - | - | - | - |  |
| 1976 | 23 | 33 | - | - | - | - | - | - | - | - |  |
| 1977 | 15 | 47 | - | - | - | - | - | - | - | - |  |
| 1979 | - | - | 20 | 59 |  | - | - | - | - | - |  |
| 1984 | 19 | 42 | - | - | 15 | 27 | - | - | - | - |  |
| 1994 | - | - | - | - | - | - | 2 | 69 | - | - |  |
| 1995 | - | - | - | - | - | - | 2 | 55 | 1 | 57 |  |

Table B.2. Total number of tows by 20 fathom depth interval (in metres) in GIG and in other areas (Other) by survey year for the 12 historical (1965 to 1995) surveys. Survey years in grey were not used in the assessment. There are more tows in the GIG portion of this table than in Table B. 3 because some of the tows in this table were not suitable for index calculations.

| Areas other than GIG |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Survey year |  |  |  |  |  | 20 fathom depth interval (m) |  |  |  | Total Tows |
|  | 66-146 | 147-183 | 184-219 | 220-256 | 257-292 | 293-329 | 330-366 | 367-402 | 440-549 |  |
| 1965 | 3 | 15 | 26 | 17 | 6 | 6 | 1 | 1 | 1 | 76 |
| 1966 | 3 | 11 | 18 | 8 | 2 | 1 | 3 | 2 | 1 | 49 |
| 1967 | 1 | - | 6 | 1 | 2 | 1 | 1 | 4 | - | 16 |
| 1969 | - | 1 | - | 1 | - | 1 | - | - | - | 3 |
| 1971 | - | - | - | - | - | - | - | - | - | - |
| 1973 | - | - | 4 | 3 | 2 | 2 | 2 | - | - | 13 |
| 1976 | - | - | 4 | 4 | 4 | 4 | 4 | - | - | 20 |
| 1977 | - | - | 3 | 2 | 2 | 3 | 2 | - | - | 12 |
| 1979 | 11 | 2 | 1 | 5 | 1 | - | - | - | - | 20 |
| 1984 | - | - | 4 | 10 | 7 | 7 | 6 | - | - | 34 |
| 1994 | - | - | - | - | - | - | - | - | - | - |
| 1995 | - | - | - | - | - | - | - | - | - | - |
| GIG |  |  |  |  |  |  |  |  |  |  |
| Survey year |  |  |  |  |  | 20 fathom depth interval (m) |  |  |  | Total |
|  | 66-146 | 147-183 | 184-219 | 220-256 | 257-292 | 293-329 | 330-366 | 367-402 | 440-549 | Tows |
| 1965 | - | 2 | 4 | 1 | 1 | - | - | - | - | 8 |
| 1966 | 3 | 2 | 3 | 5 | 2 | - | - | - | - | 15 |
| 1967 | 1 | 6 | 11 | 6 | 10 | - | - | - | - | 34 |
| 1969 | - | 9 | 11 | 6 | 6 | - | - | - | - | 32 |
| 1971 | - | 5 | 15 | 9 | 10 | - | - | - | - | 39 |
| 1973 | - | 7 | 11 | 7 | 8 | - | - | - | - | 33 |
| 1976 | - | 7 | 15 | 8 | 6 | - | - | - | - | 36 |
| 1977 | 1 | 12 | 14 | 14 | 9 | - | - | - | - | 50 |
| 1979 | 23 | 12 | 18 | 6 | - | - | - | - | - | 59 |
| 1984 | - | 13 | 25 | 17 | 13 | 1 | - | - | - | 69 |
| 1994 | - | 15 | 18 | 20 | 18 | - | - | - | - | 71 |
| 1995 | 2 | 23 | 47 | 22 | 15 | 6 | - | - | - | 115 |

Given that the only area that was consistently monitored by these surveys was the GIG grounds, tows lying between $50.9^{\circ} \mathrm{N}$ and $51.6^{\circ} \mathrm{N}$ latitude from the seven acceptable survey years, covering the period from 1967 to 1984, were used to index the WAP population (Table B.1).

The original depth stratification of these surveys was in 20 fathom ( 36.1 m ) intervals, with the important strata for WAP ranging from 70 fathoms ( 183 m ) to 160 fathoms ( 300 m ). For the GIG survey series, the shallowest tow capturing WAP was 121 m . Similarly, the deepest tow capturing WAP was 282 m . These depth strata were combined for analysis into three ranges: $70-100 \mathrm{fm}, 100-120 \mathrm{fm}$ and 120-160 fm, for a total of 333 tows from the eight accepted survey years (Table B.3).

Table B.3. Number of tows available by survey year and depth stratum for the analysis of the historical GIG trawl survey series. Survey year in grey was not used in the WAP stock assessment.

| Depth stratum |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Survey Year | $\begin{array}{r} 120-183 \mathrm{~m} \\ (70-100 \mathrm{fm}) \\ \hline \end{array}$ | $\begin{array}{r} 184-218 \mathrm{~m} \\ (100-120 \mathrm{fm}) \\ \hline \end{array}$ | $\begin{array}{r} 219-300 \mathrm{~m} \\ (120-160 \mathrm{fm}) \\ \hline \end{array}$ | Total | Start Date | End Date |
| 1967 | 7 | 11 | 15 | 33 | 07-Sep-67 | 03-Oct-67 |
| 1969 | 9 | 11 | 12 | 32 | 14-Sep-69 | 24-Sep-69 |
| 1971 | 4 | 15 | 17 | 36 | 14-Oct-71 | 28-Oct-71 |
| 1973 | 7 | 11 | 15 | 33 | 07-Sep-73 | 24-Sep-73 |
| 1976 | 7 | 13 | 13 | 33 | 09-Sep-76 | 26-Sep-76 |
| 1977 | 13 | 14 | 20 | 47 | 24-Aug-77 | 07-Sep-77 |
| 1984 | 13 | 23 | 33 | 69 | 05-Aug-84 | 08-Sep-84 |
| 1994 | 10 | 16 | 24 | 50 | 21-Jun-94 | 06-Jul-94 |
| 1995 | 22 | 45 | 45 | 112 | 11-Sep-95 | 22-Sep-95 |

Table B.4. Biomass estimates for Walleye Pollock from the historical Goose Island Gully trawl surveys for the years 1967 to 1995. Biomass estimates are based on three depth strata (Table B.3), assuming that the survey tows were randomly selected within these areas. Bootstrap bias corrected confidence intervals and CVs are based on 1000 random draws with replacement.

| Survey <br> Year | Biomass (t) <br> (Eq. B.4) | Mean bootstrap <br> biomass (t) | Lower bound <br> biomass (t) | Upper bound <br> biomass (t) | Bootstrap <br> CV | Analytic CV <br> (Eq. B.6) |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1967 | 141 | 143 | 32 | 289 | 0.467 | 0.465 |
| 1969 | 792 | 779 | 367 | 1,495 | 0.358 | 0.359 |
| 1971 | 1,432 | 1,417 | 109 | 4,283 | 0.735 | 0.740 |
| 1973 | 2,628 | 2,553 | 544 | 6,386 | 0.584 | 0.588 |
| 1976 | 5,669 | 5,678 | 2,220 | 9,918 | 0.336 | 0.346 |
| 1977 | 2,486 | 2,532 | 1,226 | 4,083 | 0.286 | 0.292 |
| 1984 | 1,284 | 1,290 | 522 | 2,356 | 0.345 | 0.346 |
| 1994 | 1,312 | 1,324 | 634 | 2,186 | 0.304 | 0.305 |

A doorspread density (Eq. B.3) was calculated for each tow based on the catch of WAP, using a fixed doorspread value of 61.6 m (Yamanaka et al. 1996) for every tow and the recorded distance travelled. Unfortunately, the speed, effort and distance travelled fields were not well populated for these surveys. Therefore, missing values for these fields were filled in with the mean values for the survey year. This resulted in the majority of the tows having distances towed near 3 km , which was the expected result given the design specification of $1 / 2$ hour tows at an approximate speed of $6 \mathrm{~km} / \mathrm{h}$ (about 3.2 knots).

## B.3.2. Results

Maps showing the locations where WAP were caught in the Goose Island Gully (GIG) indicate that this species is found throughout the GIG in most years, excepting 1967 and 1969 (see Figure B. 1 to Figure B.8). WAP was taken relatively frequently in small amounts, with 277 of the 444 valid tows capturing WAP with a median catch weight of 16 kg . The largest WAP tow in terms of catch weight was 2370 kg in 1971. WAP were mainly taken at depths from 159 to $276 \mathrm{~m}(5 \%$ and $95 \%$ quantiles of the starting depth empirical distribution), with the minimum and maximum observed depths at 148 and 296 m respectively (Figure B.9).

Estimated biomass levels in the GIG for Walleye Pollock from the historical GIG trawl surveys were variable, with the maximum biomass recorded in 1976 (at 5669 t ) and the minimum biomass in 1967 (at 141 t ) (Figure B.10; Table B.4). Survey relative errors are moderate to high for this species, ranging from a low of 0.29 in 1977 to 0.74 in 1971 (Table B.4). The proportion of tows which caught WAP was variable between years, ranging between $33 \%$ and $96 \%$ of the tows (Figure B.11). Overall, 277 tows from a total 444 valid tows (62\%) contained WAP.


Figure B.1. Valid tow locations and density plots for the historic 1967 Goose Island Gully (GIG) survey. Tow locations are colour-coded by depth range: black=120-183m; red=184-218m; grey=219-300m. Circle sizes in the right-hand density plot scaled across all years (1967, 1969, 1971, 1973, 1976, 1977, 1984, and 1994), with the largest circle $=12,825 \mathrm{~kg} / \mathrm{km}^{2}$ in 1971. Black boundary lines show the extent of the modern Queen Charlotte Sound synoptic survey and the red solid lines indicate the boundaries between PMFC areas 5A, 5B and 5C.


Figure B.2. Tow locations and density plots for the historic 1969 Goose Island Gully (GIG) survey (see Figure B. 1 caption).


Figure B.3. Tow locations and density plots for the historic 1971 Goose Island Gully (GIG) survey (see Figure B. 1 caption).


Figure B.4. Tow locations and density plots for the historic 1973 Goose Island Gully (GIG) survey (see Figure B. 1 caption).


Figure B.5. Tow locations and density plots for the historic 1976 Goose Island Gully (GIG) survey (see Figure B. 1 caption).


Figure B.6. Tow locations and density plots for the historic 1977 Goose Island Gully (GIG) survey (see Figure B. 1 caption).


Figure B.7. [left panel]: Tow location colours indicate the vessel fishing rather than depth: black=G.B. Reed; red=Eastward Ho. Additional locations fished by vessel in Mitchell Gully are also shown; [right panel]: density plot for the historic 1984 Goose Island Gully (GIG) survey (see Figure B. 1 caption).


Figure B.8. Tow locations and density plots for the historic 1994 Goose Island Gully (GIG) survey (see Figure B. 1 caption).


Figure B.9. Distribution of observed catch weights of Walleye Pollock (WAP) for the historic Goose Island Gully (GIG) surveys (Table B.3) by survey year and 25 m depth zone. Depth zones are indicated by the mid point of the depth interval and circles in the panel are scaled to the maximum value ( 2658 kg ) in the 150-175 m interval in 1976. The 1\% and 99\% quantiles for the WAP empirical start of tow depth distribution $=148 \mathrm{~m}$ and 282 m respectively.


Figure B.10. Plot of biomass estimates for the WAP historic Goose Island Gully (GIG) surveys: 1967 to 1994 (values provided in Table B.4). Bias corrected 95\% confidence intervals from 1000 bootstrap replicates are plotted.


Figure B.11. Proportion of tows by year which contain WAP from the historic Goose Island Gully (GIG) surveys: 1967 to 1995.

## B.4. HECATE STRAIT ASSEMBLAGE SURVEY

## B.4.1. Data selection

This survey was conducted 11 times over the period 1984 to 2003 in Hecate Strait (HS) between Moresby and Graham Islands and the mainland (all valid tow starting positions are shown by survey year in Figure B. 12 to Figure B. 22 (Sinclair 1999). The design overlaid a 10 nm square grid over Hecate Strait and placed one tow per grid square in each 10-fathom depth interval over the range of 10 to 80 fathoms ( 18 to 146 m ). Strata deeper than 80 fathoms were sampled in some survey years, but these were excluded from the analysis because the deeper strata were not sampled in all survey years. Tow positions were selected non-randomly by substrate type and were fixed after the first survey, although there was some variation in how tow positions were revisited, and new tow positions were added over the years. There were 85 to 105 valid tows in each survey year after the initial year, which had over 140 tows (Table B.5). Sinclair (1999) chose to analyze these data using the 10 fathom depth intervals as depth strata, without reference to the overlaid grid pattern by assuming that the tow locations had been selected randomly.

Table B.5. Number of usable tows for biomass estimation by year and depth stratum for the Hecate Strait assemblage survey over the period 1984 to 2003. Also shown is the area of each depth stratum and the vessel conducting the survey by survey year.

| Year | Vessel |  |  |  |  |  | Depth stratum |  | Total tows |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 10-19fm | 20-29fm | 30-39fm | 40-49fm | 50-59fm | 60-69fm | 70-79fm |  |
|  | G.B. Reed/ |  |  |  |  |  |  |  |  |
| 1984 | Arctic Ocean | 19 | 19 | 23 | 25 | 23 | 23 | 14 | 146 |
| 1987 | Southward Ho | 15 | 12 | 12 | 11 | 16 | 10 | 9 | 85 |
| 1989 | Southward Ho | 17 | 12 | 12 | 15 | 12 | 9 | 13 | 90 |
| 1991 | Southward Ho | 18 | 12 | 15 | 10 | 21 | 15 | 7 | 98 |
| 1993 | W.E. Ricker | 16 | 20 | 11 | 15 | 10 | 15 | 7 | 94 |
| 1995 | W.E. Ricker | 17 | 19 | 15 | 16 | 14 | 14 | 7 | 102 |
| 1996 | W.E. Ricker | 25 | 24 | 21 | 10 | 11 | 10 | 4 | 105 |
| 1998 | W.E. Ricker | 14 | 11 | 17 | 13 | 13 | 14 | 4 | 86 |
| 2000 | W.E. Ricker | 18 | 22 | 19 | 14 | 15 | 11 | 6 | 105 |
| 2002 | Viking Storm | 17 | 17 | 15 | 16 | 11 | 10 | 6 | 92 |
| 2003 | W.E. Ricker | 15 | 17 | 16 | 18 | 15 | 9 | 5 | 95 |
| Area (km ${ }^{\text {2 }}$ ) | - | 2,657 | 1,651 | 908 | 828 | 912 | 792 | 612 | 8,360 ${ }^{1}$ |

Table B.6. Biomass estimates for Walleye Pollock from the Hecate Strait assemblage trawl survey for the survey years 1984 to 2003, using the method of Starr et al. (20065) (see text for explanation). Bootstrap bias corrected confidence intervals and CVs are based on 1000 random draws with replacement.

| Survey <br> Year | Biomass (t) <br> (Eq. B.4) | Mean bootstrap <br> biomass (t) | Lower bound <br> biomass (t) | Upper bound <br> biomass (t) | Bootstrap <br> CV | Analytic CV <br> (Eq. B.6) |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1984 | 383 | 382 | 229 | 569 | 0.216 | 0.227 |
| 1987 | 514 | 505 | 114 | 1,268 | 0.611 | 0.618 |
| 1989 | 410 | 907 | 817 | 114 | 807 | 0.420 |
| 1991 | 1,193 | 1,227 | 429 | 1,602 | 0.337 | 0.430 |
| 1993 | 425 | 429 | 303 | 3,111 | 0.579 | 0.340 |
| 1995 | 1,684 | 1,684 | 251 | 703 | 0.260 | 0.265 |
| 1996 | 2,022 | 62 | 2,031 | 647 | 3,240 | 0.379 |
| 1998 | 1,253 | 1,251 | 348 | 4,058 | 0.383 | 0.372 |
| 2000 | 549 | 536 | 736 | 114 | 0.335 | 0.393 |
| 2002 |  | 243 | 1,925 | 0.247 | 0.322 |  |
| 2003 |  | 970 | 0.348 | 0.242 |  |  |

Two methods have been used to generate a doorspread density value (Eq. B.3) for each survey tow, given that there are no estimates of doorspread or wingspread for this survey and there only exist estimates of [distance_travelled] and [speed ] for the final three survey years. The method proposed by Sinclair (1999) was to calculate a CPUE (kg/h) for each tow and to convert this value to biomass per area swept ( $\mathrm{kg} / \mathrm{km}^{2}$ ) by assuming a constant area swept by each tow, with $0.0486 \mathrm{~km}^{2} / \mathrm{h}$ as the constant. A second method was proposed by Starr et al. (2006), who assumed a constant doorspread value of 43 m and a constant speed of $5.1 \mathrm{~km} / \mathrm{h}$ (Eq. B.2). There was little practical difference between these methods when the resulting biomass indices are treated as relative, as was demonstrated by Starr et al. (2006).

## B.4.2. Results

Catch densities of WAP from this survey were generally highest in the northern half of Hecate Strait, rarely extending to the top of Graham Island or the southern part of Hecate Strait (Figure B. 12 to Figure B.22). WAP were mainly taken at depths from 57 to 139 m ( $5 \%$ and $95 \%$ quantiles of the empirical depth distribution), with observations down to depths deeper than 200 m in the years that the deep strata were sampled (Figure B.23).
Estimated WAP biomass indices from this trawl survey showed no trend from 1984 to 2003 (Table B.6; Figure B.24). The estimated relative errors were moderate to high, ranging from 23 to $62 \%$ (Table B.6). These estimates of variability may be biased low, given the non-random selection of tow locations. On average, one third of the survey tows captured WAP (ranging from 0.20 to 0.45 ) (Figure B.25). Overall, 362 of the 1,098 valid survey tows contained WAP.

[^5]

Figure B.12. Valid tow locations and density plots for the 1984 Hecate Strait assemblage survey with strata in fathoms: 10-19 (black), 20-29 (red), 30-39 (grey), 40-49 (blue), 50-59 (sienna), 60-69 (cyan), 7079 (turquoise); and density plots for the 1984 Hecate Strait assemblage survey. Circle sizes in the righthand density plot scaled across all years (1984, 1987, 1989, 1991, 1993, 1995, 1996, 1998, 2000, 2002, 2003), with the largest circle $=12,777 \mathrm{~kg} / \mathrm{km}^{2}$ in 1996.



Figure B.13. Tow locations and density plots for the 1987 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.14. Tow locations and density plots for the 1989 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.15. Tow locations and density plots for the 1991 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.16. Tow locations and density plots for the 1993 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.17. Tow locations and density plots for the 1995 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.18. Tow locations and density plots for the 1996 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.19. Tow locations and density plots for the 1998 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.20. Tow locations and density plots for the 2000 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.21. Tow locations and density plots for the 2002 Hecate Strait assemblage survey (see Figure B. 12 caption).


Figure B.22. Tow locations and density plots for the 2003 Hecate Strait assemblage survey (see Figure B. 12 caption).

Hecate Strait assemblage survey


## Survey year

Maximum circle size $=2261 \mathrm{~kg}$
Figure B.23. Distribution of observed catch weights of Walleye Pollock for the Hecate Strait assemblage survey (Table B.5) by survey year and 25 m depth zone. Depth zones are indicated by the mid point of the depth interval and circles in the panel are scaled to the maximum value $(2,261 \mathrm{~kg})$ in the $50-75 \mathrm{~m}$ interval in 1996. The 1\% and 99\% quantiles for the WAP empirical start of tow depth distribution= 37 m and 203 m respectively. Note that tows deeper than $148 \mathrm{~m}(80 \mathrm{fm})$ are presented here when available but were excluded from the biomass analysis because the deeper strata were not sampled consistently in all survey years.

Pollock: Hecate Strait assemblage survey


Figure B.24. Plot of biomass estimates for WAP (values provided in Table B.6) from the Hecate Strait assemblage survey over the period 1984 to 2003 . Bias corrected $95 \%$ confidence intervals from 1000 bootstrap replicates are plotted.


Figure B.25. Proportion of tows by year which contain WAP from the Hecate Strait assemblage survey over the period 1984 to 2003.

## B.5. HECATE STRAIT SYNOPTIC SURVEY

## B.5.1. Data selection

This survey has been conducted in six alternating years over the period 2005 to 2015 in Hecate Strait (HS) between Moresby and Graham Islands and the mainland and in Dixon Entrance at the top of Graham Island (all valid tow starting positions by survey year are shown in Figure B. 26 to Figure B.31). This survey treats the full spatial coverage as a single areal stratum divided into four depth strata: $10-70 \mathrm{~m} ; 70-130 \mathrm{~m} ; 130-220 \mathrm{~m}$; and $220-500 \mathrm{~m}$ (Table B.7).

A doorspread density value (Eq. B.3) was generated for each tow based on the catch of Walleye Pollock (WAP), the mean doorspread for the tow and the distance travelled; [distance_travelled ] is a database field which is calculated directly from the tow track. This field is used preferentially for the variable $D_{y i j}$ in Eq. B.3. A calculated value ([vessel_speed ] $X$ [tow_duration]) is used for this variable if [distance_travelled] is missing, but there were no instances of this occurring in the 6 trawl surveys. Missing values for the [doorspread] field were filled in with the mean doorspread for the survey year ( 217 values over all years: Table B.8).

Table B.7. Number of usable tows for biomass estimation by year and depth stratum for the Hecate Strait synoptic survey over the period 2005 to 2015. Also shown is the area of each depth stratum and the vessel conducting the survey by survey year.

|  |  | Depth stratum |  |  |  | Total |
| :---: | :---: | ---: | ---: | ---: | ---: | ---: |
| Year | Vessel | $\mathbf{1 0 - 7 0}$ | $\mathbf{7 0 - 1 3 0}$ | $\mathbf{1 3 0 - 2 2 0}$ | $\mathbf{2 2 0 - 5 0 0}$ | tows |
| 2005 | Frosti | 79 | 88 | 26 | 9 | 202 |
| 2007 | W.E. Ricker | 48 | 43 | 36 | 7 | 134 |
| 2009 | W.E. Ricker | 53 | 43 | 48 | 12 | 156 |
| 2011 | W.E. Ricker | 70 | 51 | 50 | 14 | 185 |
| 2013 | W.E. Ricker | 74 | 42 | 43 | 16 | 175 |
| 2015 | W.E. Ricker | 47 | 46 | 40 | 15 | 148 |
| Area $\left(\mathrm{km}^{2}\right)$ | - | 5,958 | 3,011 | 2,432 | 1,858 | $13,259^{1}$ |
| total area for survey |  |  |  |  |  |  |

Table B.8. Number of missing doorspread values by year for the Hecate Strait synoptic survey over the period 2005 to 2015 as well as showing the number of available doorspread observations and the mean doorspread value for the survey year.

| Year | Number tows <br> with missing <br> doorspread | Number tows with <br> doorspread <br> observations ${ }^{2}$ | Mean doorspread (m) <br> used for tows with <br> missing values $^{2}$ |
| :---: | ---: | ---: | ---: |
| 2005 | 7 | 217 | 64.4 |
| 2007 | 98 | 37 | 59.0 |
| 2009 | 93 | 70 | 54.0 |
| 2011 | 13 | 186 | 54.8 |
| 2013 | 6 | 169 | 51.7 |
| 2015 | 0 | 151 | 59.4 |
| Total | 217 | 830 | 57.6 |
| valid biomass estimation tows only |  |  |  |
| ${ }^{2}$ includes tows not used for biomass estimation |  |  |  |

Table B.9. Biomass estimates for Walleye Pollock from the Hecate Strait synoptic trawl survey for the survey years 2005 to 2015. Bootstrap bias corrected confidence intervals and CVs are based on 1000 random draws with replacement.

| Survey <br> Year | Biomass (t) <br> $($ Eq. B.4) | Mean bootstrap <br> biomass (t) | Lower bound <br> biomass (t) | Upper bound <br> biomass (t) | Bootstrap <br> CV | Analytic CV <br> $($ Eq. B.6) |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2005 | 1750 | 1741 | 919 | 3293 | 0.330 | 0.321 |
| 2007 | 1394 | 1422 | 568 | 2776 | 0.391 | 0.391 |
| 2009 | 1028 | 1032 | 604 | 1798 | 0.275 | 0.267 |
| 2011 | 1073 | 1069 | 561 | 1856 | 0.308 | 0.308 |
| 2013 | 1828 | 1816 | 1113 | 2851 | 0.234 | 0.232 |
| 2015 | 1972 | 1974 | 1062 | 3293 | 0.285 | 0.273 |

## B.5.2. Results

Catches of WAP from this survey are concentrated along the 100 m depth contour in Dixon Entrance and then follow that contour into upper Hecate Strait (Figure B. 26 to Figure B.31). WAP are mainly taken at depths from 43 to 239 m ( $5-95 \%$ quantiles), but there are sporadic observations to depths up to about 330 m and down to about 20 m (Figure B.32).
Estimated WAP doorspread biomass from this trawl survey showed no overall trend over the period 2005 to 2015, with the highest estimates recorded in 2005, 2013 and 2015 and the lowest estimate in 2009 (Table B.9; Figure B.33). The estimated relative errors were moderate, ranging from 23 to $39 \%$ (Table B.9). On average, $55 \%$ of the survey tows captured WAP (ranging from 0.52 to 0.60 by year) (Figure B.34). Overall, 549 of the 1000 valid survey tows contained WAP with a low median catch weight for positive tows ( $3.3 \mathrm{~kg} / \mathrm{tow}$ ) and a maximum catch weight across all six surveys of 1622 kg (in 2005).


Figure B.26. Valid tow locations and density plots for the 2005 Hecate Strait synoptic survey. Circle sizes in the right-hand density plot scaled across all years (2005, 2007, 2009, 2011, 2013, 2015), with the largest circle $=14,373 \mathrm{~kg} / \mathrm{km}^{2}$ in 2005. Red lines indicate boundaries for PMFC major statistical areas 5C, $5 D$ and $5 E$.


Figure B.27. Tow locations and density plots for the 2007 Hecate Strait synoptic survey (see Figure B. 26 caption).


Figure B.28. Tow locations and density plots for the 2009 Hecate Strait synoptic survey (see Figure B. 26 caption).


Figure B.29. Tow locations and density plots for the 2011 Hecate Strait synoptic survey (see Figure B. 26 caption).


Figure B.30. Tow locations and density plots for the 2013 Hecate Strait synoptic survey (see Figure B. 26 caption).


Figure B.31. Tow locations and density plots for the 2015 Hecate Strait synoptic survey (see Figure B. 26 caption).


Figure B.32. Distribution of observed catch weights of Walleye Pollock for the Hecate Strait synoptic survey (Table B.7) by survey year and 25 m depth zone. Catches are plotted at the mid-point of the interval and circles in the panel are scaled to the maximum value ( 2491 kg ) in the 75-100 m interval in 2005. The $1 \%$ and $99 \%$ quantiles for the WAP empirical start of tow depth distribution= 25 m and 288 m respectively.


Figure B.33. Plot of biomass estimates for Walleye Pollock values provided in Table B. 9 from the Hecate Strait synoptic survey over the period 2005 to 2015 . Bias corrected 95\% confidence intervals from 1000 bootstrap replicates are plotted.


Figure B.34. Proportion of tows by year which contain Walleye Pollock from the Hecate Strait synoptic survey over the period 2005 to 2015.

## B.6. QUEEN CHARLOTTE SOUND SYNOPTIC TRAWL SURVEY

## B.6.1. Data selection

This survey has been conducted in eight years over the period 2003 to 2015 in Queen Charlotte Sound (QCS), which lies between the top of Vancouver Island and the southern portion of Moresby Island and extends into the lower part of Hecate Strait between Moresby Island and the mainland. The design divided the survey into two large areal strata which roughly correspond to the PMFC regions 5A and 5B while also incorporating part of 5C (all valid tow starting positions are shown by survey year in Figure B. 35 to Figure B.42). Each of these two areas was divided into four depth strata: 50-125 m; 125-200 m; 200-330 m; and 330-500 m (Table B.10).
A doorspread density value (Eq. B.3) was generated for each tow based on the catch of Walleye Pollock (WAP), the mean doorspread for the tow and the distance travelled. [distance_travelled] is a database field which is calculated directly from the tow track. This field is used preferentially for the variable $D_{y i j}$ in Eq. B.3. A calculated value ([vessel_speed] $X$ [tow_duration]) is used for this variable if [distance_travelled ] is missing, but there were only two instances of this occurring in the 8 trawl surveys. Missing values for the [doorspread] field were filled in with the mean doorspread for the survey year (101 values over all years, Table B.11).

Table B.10. Number of usable tows for biomass estimation by year and depth stratum for the Queen Charlotte Sound synoptic survey over the period 2003 to 2015. Also shown is the area of each stratum for the 2015 survey and the vessel conducting the survey by survey year.

| Year | Vessel | South depth strata |  |  |  | North depth strata |  |  |  | Total tows |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 50-125 | 125-200 | 200-330 | 330-500 | 50-125 | 125-200 | 200-330 | 330-500 |  |
| 2003 | Viking Storm | 29 | 56 | 29 | 6 | 5 | 39 | 50 | 19 | 233 |
| 2004 | Viking Storm | 42 | 48 | 31 | 8 | 20 | 38 | 37 | 6 | 230 |
| 2005 | Viking Storm | 29 | 60 | 29 | 8 | 8 | 45 | 37 | 8 | 224 |
| 2007 | Viking Storm | 33 | 62 | 24 | 7 | 19 | 57 | 48 | 7 | 257 |
| 2009 | Viking Storm | 34 | 60 | 28 | 8 | 10 | 44 | 43 | 6 | 233 |
| 2011 | Nordic Pearl | 38 | 67 | 25 | 8 | 10 | 51 | 45 | 8 | 252 |
| 2013 | Nordic Pearl | 32 | 66 | 29 | 10 | 9 | 46 | 44 | 5 | 241 |
| 2015 | Frosti | 30 | 65 | 26 | 4 | 12 | 50 | 44 | 8 | 239 |
| Area (km $\left.{ }^{2}\right)^{2}$ |  | 5,072 | 5,432 | 2,712 | 548 | 1,804 | 4,060 | 3,748 | 1,252 | 24,628 |

Table B.11. Number of missing doorspread values by year for the Queen Charlotte Sound synoptic survey over the period 2003 to 2015 as well as showing the number of available doorspread observations and the mean doorspread value for the survey year.

| Year | Number tows with <br> missing doorspread ${ }^{\mathbf{1}}$ | Number tows with <br> doorspread observations ${ }^{2}$ | Mean doorspread (m) used for <br> tows with missing values ${ }^{2}$ |
| :---: | ---: | ---: | ---: |
| 2003 | 13 | 236 | 72.1 |
| 2004 | 8 | 267 | 72.8 |
| 2005 | 1 | 258 | 74.5 |
| 2007 | 5 | 262 | 71.8 |
| 2009 | 2 | 248 | 71.3 |
| 2011 | 30 | 242 | 67.0 |
| 2013 | 42 | 226 | 69.5 |
| 2015 | 0 | 249 | 70.5 |
| Total | 101 | 1,988 | 71.2 |

[^6]Table B.12. Biomass estimates for Walleye Pollock from the Queen Charlotte Sound synoptic trawl survey for the survey years 2003 to 2015. Bootstrap bias corrected confidence intervals and CVs are based on 1000 random draws with replacement.

| Survey <br> Year | Biomass (t) <br> $($ Eq. B.4) | Mean bootstrap <br> biomass (t) | Lower bound <br> biomass (t) | Upper bound <br> biomass (t) | Bootstrap <br> CV | Analytic CV <br> $($ Eq. B.6) |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2003 | 254 | 254 | 176 | 370 | 0.183 | 0.188 |
| 2004 | 335 | 335 | 224 | 492 | 0.198 | 0.194 |
| 2005 | 648 | 649 | 423 | 1020 | 0.221 | 0.219 |
| 2007 | 518 | 518 | 305 | 796 | 0.240 | 0.243 |
| 2009 | 372 | 375 | 209 | 563 | 0.239 | 0.239 |
| 2011 | 2671 | 2639 | 926 | 6745 | 0.541 | 0.539 |
| 2013 | 1667 | 1689 | 874 | 2605 | 0.266 | 0.265 |
| 2015 | 2114 | 2116 | 1066 | 3669 | 0.313 | 0.315 |

## B.6.2. Results

Catch densities of WAP were very low for the first five surveys, resulting in low incidence for this species (Figure B. 35 to Figure B.39). However, the incidence of this species in this survey increased after 2009, particularly in the 2011 survey (Figure B.40), but was also higher in the 2013 (Figure B.41) and 2015 surveys (Figure B.42). The tows that captured WAP in these three surveys were generally located in Queen Charlotte Strait near the northern end of Vancouver Island, which is also the location of one of the three targeted midwater fisheries for this species. WAP were mainly taken at depths from 112 to 291 m (5-95\% quantiles), but there were sporadic observations up to depths near 400 m and down to about 60 m (Figure B.43).


Figure B.35. Valid tow locations (50-125m stratum: black; 126-200m stratum: red; 201-330m stratum: grey; 331-500m stratum: blue) and density plots for the 2003 QC Sound synoptic survey. Circle sizes in the right-hand density plot scaled across all years (2003-2005, 2007, 2009, 2011, 2013, 2015), with the largest circle $=17,403 \mathrm{~kg} / \mathrm{km}^{2}$ in 2011. Boundaries delineate the North and South areal strata.


Figure B.36. Tow locations and density plots for the 2004 Queen Charlotte Sound synoptic survey (see Figure B. 35 caption).


Figure B.37. Tow locations and density plots for the 2005 Queen Charlotte Sound synoptic survey (see Figure B. 35 caption).


Figure B.38. Tow locations and density plots for the 2007 Queen Charlotte Sound synoptic survey (see Figure B. 35 caption).


Figure B.39. Tow locations and density plots for the 2009 Queen Charlotte Sound synoptic survey (see Figure B. 35 caption).


Figure B.40. Tow locations and density plots for the 2011 Queen Charlotte Sound synoptic survey (see Figure B. 35 caption).


Figure B.41. Tow locations and density plots for the 2013 Queen Charlotte Sound synoptic survey (see Figure B. 35 caption).


Figure B.42. Tow locations and density plots for the 2015 Queen Charlotte Sound synoptic survey (see Figure B. 35 caption).


Survey year
Maximum circle size $=2651 \mathrm{~kg}$
Figure B.43. Distribution of observed catch weights of Walleye Pollock for the two main Queen Charlotte Sound synoptic survey areal strata (Table B.10) by survey year and 25 m depth zone. Catches are plotted at the mid-point of the interval and circles in the panel are scaled to the maximum value (2651 kg) in the 125-150 m interval in the 2011 southern stratum. The $1 \%$ and $99 \%$ quantiles for the WAP empirical start of tow depth distribution $=76 \mathrm{~m}$ and 353 m respectively.

Estimated WAP doorspread biomass from this trawl survey were low from 2003 to 2009, followed by a step up in biomass beginning with the 2011 survey (Table B.12; Figure B.44). The estimated relative errors were variable for this species, lying between 19 and $24 \%$ when the survey biomass estimates were low, but jumping to higher values after 2009 (Table B.12). Between 23 and 62\% of the South stratum tows and 39 to $62 \%$ of the North stratum tows captured some WAP (Figure B.45). Overall, 858 of the 1909 valid survey tows ( $45 \%$ ) contained WAP, with the North stratum having a $53 \%$ average proportion non-zero tows while the equivalent South stratum proportion was $38 \%$. Although this species occurs frequently in this survey, catch weights tend to be low, with the median catch weight for positive tows around $1.8 \mathrm{~kg} /$ tow across all 8 surveys, but the maximum catch weight was 2127 kg in the 2011 survey.


Figure B.44. Plot of biomass estimates for WAP (values provided in Table B.12) from the Queen Charlotte Sound synoptic survey over the period 2003 to 2015. Bias corrected 95\% confidence intervals from 1000 bootstrap replicates are plotted.


Figure B.45. Proportion of tows by stratum and year which contain WAP from the Queen Charlotte Sound synoptic survey over the period 2003 to 2015.

## B.7. WEST COAST VANCOUVER ISLAND SYNOPTIC TRAWL SURVEY

## B.7.1. Data selection

This survey has been conducted seven times in the period 2004 to 2016 off the west coast of Vancouver Island by RV W.E. Ricker. It comprises a single areal stratum, separated into four depth strata: 50-125 m; 125-200 m; 200-330 m; and 330-500 m (Table B.13). Approximately 150 to $1802-\mathrm{km}^{2}$ blocks are selected randomly among the four depth strata when conducting each survey (Olsen et. al. 2008).

A doorspread density value was generated for each tow based on the catch of Walleye Pollock, the mean doorspread for the tow and the distance travelled (Eq. B.3). The distance travelled was provided as a data field, determined directly from vessel track information collected during the tow. There were only two missing values in this field which were filled in by multiplying the vessel speed by the time that the net was towed. There were a large number of missing values for the doorspread field, which were filled in using the mean doorspread for the survey year or a default value of 64.7 m for the three years with no doorspread data (Table B.14). The default value is based on the mean of the observed doorspread from the net mensuration equipment, averaged across the years with doorspread estimates.

Table B.13. Stratum designations, number of usable and unusable tows, for each year of the west coast Vancouver Island synoptic survey. Also shown is the area of each stratum in 2016 and the start and end dates for each survey.

| Survey year | Stratum depth zone |  |  |  | Total Tows ${ }^{1}$ | Unusable tows | Start date | End date |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 50-125 m | 125-200 m | 200-330 m | 330-500 m |  |  |  |  |
| 2004 | 34 | 34 | 13 | 8 | 89 | 17 | 26-May-04 | 09-Jun-04 |
| 2006 | 61 | 62 | 28 | 13 | 164 | 12 | 24-May-06 | 18-Jun-06 |
| 2008 | 54 | 50 | 32 | 23 | 159 | 19 | 27-May-08 | 21-Jun-08 |
| 2010 | 58 | 47 | 22 | 9 | 136 | 8 | 08-Jun-10 | 28-Jun-10 |
| 2012 | 61 | 46 | 26 | 20 | 153 | 4 | 23-May-12 | 15-Jun-12 |
| 2014 | 55 | 49 | 29 | 14 | 147 | 6 | 29-May-14 | 20-Jun-14 |
| 2016 | 54 | 41 | 26 | 19 | 140 | 7 | 25-May-16 | 15-Jun-16 |
| Area (km²) | 5804 | 3796 | 708 | 608 | 10,916 ${ }^{2}$ | - |  |  |

${ }^{1}$ GFBio usability codes $=0,1,2,6$
${ }^{2}$ Total area (km ${ }^{2}$ ) for 2016 synoptic survey
Table B.14. Number of tows with and without doorspread measurements by survey year for the WCVI synoptic survey. Mean doorspread values for those tows with measurements are provided.

|  | Without | Number tows <br>  <br>  <br>  <br> doorspread <br> doorspread | With <br> doorspread <br> $(\mathbf{m})$ |
| :---: | ---: | ---: | ---: |
| 2004 | 89 | 0 | - |
| 2006 | 96 | 69 | 64.3 |
| 2008 | 58 | 107 | 64.5 |
| 2010 | 136 | 0 | - |
| 2012 | 153 | 0 | - |
| 2014 | 14 | 139 | 64.3 |
| 2016 | 0 | 147 | 65.5 |
| All surveys | 546 | 462 | 64.7 |



Figure B.46. Valid tow locations (50-125m stratum: black; 126-200m stratum: red; 201-330m stratum: grey; 331-500m stratum: blue) and density plots for the 2004 west coast Vancouver Island synoptic survey. Circle sizes in the right-hand density plot scaled across all years (2004, 2006, 2008, 2010, 2012, 2014), with the largest circle $=8962 \mathrm{~kg} / \mathrm{km}^{2}$ in 2010. The red solid lines indicate the boundaries for PMFC areas $3 C, 3 D$ and $5 A$.


Figure B.47. Tow locations and density plots for the 2006 west coast Vancouver Island synoptic survey (see Figure B. 46 caption).


Figure B.48. Tow locations and density plots for the 2008 west coast Vancouver Island synoptic survey (see Figure B. 46 caption).


Figure B.49. Tow locations and density plots for the 2010 west coast Vancouver Island synoptic survey (see Figure B. 46 caption).


Figure B.50. Tow locations and density plots for the 2012 west coast Vancouver Island synoptic survey (see Figure B. 46 caption).


Figure B.51. Tow locations and density plots for the 2014 west coast Vancouver Island synoptic survey (see Figure B. 46 caption).


Figure B.52. Tow locations and density plots for the 2016 west coast Vancouver Island synoptic survey (see Figure B. 46 caption).


Figure B.53. Distribution of observed weights of Walleye Pollock by survey year and 25 m depth zone. Catches are plotted at the mid-point of the interval and circles in the panel are scaled to the maximum value (1060 kg) in the $50-75$ m interval in 2010. The $1 \%$ and $99 \%$ quantiles for the WAP empirical start of tow depth distribution $=61 \mathrm{~m}$ and 270 m respectively.

## B.7.2. Results

Walleye Pollock are mainly taken near the entrance of Juan de Fuca Strait, although the incidence of these encounters appears to vary between years (Figure B. 46 to Figure B.51). There is an important midwater trawl fishery for this species in Juan de Fuca Strait, but much of it occurs further in the Strait, although there are also catches in the same location where WAP are taken in this survey (see Figure A.2). Walleye Pollock were mainly taken at depths from 75 to 213 m (5-95\% quantiles) and there were almost no observations at depths greater than 300 m (Figure B.53). Estimated biomass levels for Walleye Pollock from this trawl survey show elevated biomass levels beginning with 2010, but there is no apparent trend after that year. Relative errors are high, ranging from 36 to $66 \%$ across the seven surveys (Figure B. 54 ; Table B.15).
The proportion of tows capturing Walleye Pollock ranged between 9 and $41 \%$ for the seven surveys, with a mean value of $25 \%$ (Figure B.55). About one quarter of the tows from this survey contain WAP, but as in the QC Sound synoptic survey, the median catch weight for positive tows was low (around $1.3 \mathrm{~kg} / \mathrm{tow}$ ) and the maximum catch weight across all six surveys was 1060 kg (in 2010).


Figure B.54. Plot of biomass estimates for Walleye Pollock from the 2004 to 2016 west coast Vancouver Island synoptic trawl surveys (Table B.15). Bias-corrected 95\% confidence intervals from 1000 bootstrap replicates are plotted.


Figure B.55. Proportion of tows by stratum and year capturing Walleye Pollock in the WCVI synoptic trawl surveys, 2004-2016.

Table B.15. Biomass estimates for Walleye Pollock from the WCVI synoptic trawl survey for the survey years 2004 to 2016. Bootstrap bias-corrected confidence intervals and CVs are based on 1000 random draws with replacement.

| Survey <br> Year | Biomass (t) <br> (Eq. B.4) | Mean bootstrap <br> biomass (t) | Lower bound <br> biomass (t) | Upper bound <br> biomass (t) | Bootstrap <br> CV | Analytic CV <br> (Eq. B.6) |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2004 | 569 | 281 | 92 | 1684 | 0.662 | 0.669 |
| 2006 | 7 | 28 | 8 | 61 | 0.473 | 0.490 |
| 2008 | 7 | 7 | 3 | 14 | 0.408 | 0.419 |
| 2010 | 1598 | 489 | 481 | 274 | 4493 | 0.648 |
| 2012 | 1121 | 1112 | 221 | 968 | 0.380 | 0.609 |
| 2014 | 237 | 231 | 365 | 2542 | 0.486 | 0.378 |
| 2016 |  | 104 | 438 | 0.364 | 0.490 |  |

## B.8. WEST COAST HAIDA GWAII SYNOPTIC TRAWL SURVEY

## B.8.1. Data selection

The west coast Haida Gwaii (WCHG) survey has been conducted six times in the period 2006 to 2016 off the west coast of Haida Gwaii. A survey conducted in 2014 did not complete a sufficient number of tows for it to be considered completed. The survey comprises a single areal stratum extending from about $53^{\circ} \mathrm{N}$ to the BC-Alaska border and east to $133^{\circ} \mathrm{W}$ (e.g., Olsen et al. 2008). The 2006 survey used a different depth stratification scheme compared to the later synoptic surveys: 150-200 m, 200-330 m, 330-500 m, 500-800 m, and 800-1300 m (Workman et al. 2007). All tows from this survey were re-stratified into the four depth strata used from 2007 onwards: 180-330 m; 330-500 m; 500-800 m; and 800-1300 m, based on the mean of the beginning and end depths of each tow (Table B.16). Plots of the locations of all valid tows by year and stratum are presented in Figure B. 56 (2006), Figure B. 57 (2007), Figure B. 58 (2008), Figure B. 59 (2010), Figure B. 60 (2012) and Figure B. 61 (2016). Note that the depth stratum boundaries for this survey differ from those used for the Queen Charlotte Sound (Edwards et al., 2012) and west coast Vancouver Island (Edwards et al., 2014) synoptic surveys due to the considerable difference in the seabed topography of the area being surveyed. The deepest stratum ( $800-1300 \mathrm{~m}$ ) was omitted from this analysis because of lack of coverage in 2007.

Table B.16. Stratum designations, vessel name, number of usable and unusable tows, for each year of the west coast Haida Gwaii synoptic survey. Also shown are the area of each stratum and the dates of the first and last survey tow in each year.

| Survey year | Vessel | Depth stratum |  |  |  | Total tows ${ }^{1}$ | Unusable tows | Start date | End date |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{array}{r} 180- \\ 330 \mathrm{~m} \end{array}$ | $\begin{array}{r} 330- \\ 500 \mathrm{~m} \\ \hline \end{array}$ | $\begin{array}{r} 500- \\ 800 \mathrm{~m} \end{array}$ | $\begin{array}{r} 800- \\ 1300 \mathrm{~m} \end{array}$ |  |  |  |  |
| 2006 | Viking Storm | 55 | 26 | 16 | 13 | $110^{2}$ | 13 | 30-Aug-06 | 22-Sep-06 |
| 2007 | Nemesis | 68 | 34 | 9 | 0 | 111 | 5 | 14-Sep-07 | 12-Oct-07 |
| 2008 | Frosti | 71 | 31 | 8 | 8 | 118 | 9 | 28-Aug-08 | 18-Sep-08 |
| 2010 | Viking Storm | 82 | 29 | 12 | 6 | 129 | 2 | 28-Aug-10 | 16-Sep-10 |
| 2012 | Nordic Pearl | 75 | 29 | 10 | 16 | 130 | 11 | 27-Aug-12 | 16-Sep-12 |
| 2016 | Frosti | 69 | 28 | 5 | 10 | 112 | 8 | 28-Aug-16 | 24-Sep-16 |
| Area (km²) | - | 1104 | 1024 | 956 | 2248 | $5332{ }^{3}$ | - | - | - |

${ }^{1}$ GFBio usability codes $=0,1,2,6 ;{ }^{2}$ excludes 2 tows S of $53^{\circ} \mathrm{N} ;{ }^{3}$ Total area in $2016\left(\mathrm{~km}^{2}\right)$

Table B.17. Number of valid tows with doorspread measurements, the mean doorspread values (in m) from these tows for each survey year and the number of valid tows without doorspread measurements.

| Year | Tows with doorspread Tows missing doorspread | Mean doorspread (m) |  |
| :---: | ---: | ---: | ---: |
| 2006 | 93 | 30 | 77.7 |
| 2007 | 113 | 3 | 68.5 |
| 2008 | 123 | 4 | 80.7 |
| 2010 | 129 | 2 | 79.1 |
| 2012 | 92 | 49 | 73.8 |
| 2016 | 105 | 15 | 74.1 |
| Total/Average | 655 | 103 | $75.8^{1}$ |

${ }^{1}$ average 2006-2016: all observations
A doorspread density (Eq. B.3) was generated for each tow based on the catch of Walleye Pollock (WAP), the mean doorspread for the tow and the distance travelled.
[distance_travelled ] is a database field which is calculated directly from the tow track. This field is used preferentially for the variable $D_{y i j}$ in Eq. B.3. A calculated value ([vessel_speed] X [tow_duration]) is used for this variable if [distance_travelled ] is missing, but there were no instances of this occurring in the six trawl surveys. Missing values for the [doorspread] field were filled in with the mean doorspread for the survey year (103 values over all years, Table B.17).



Figure B.56. Valid tow locations by stratum (180-330m: black; 330-500m: red; 500-800m: grey; 8001300m: blue) and density plots for the 2006 Viking Storm synoptic survey. Circle sizes in the right-hand density plot scaled across all years (2006-2016), with the largest circle $=2384{\mathrm{~kg} / \mathrm{km}^{2}}^{2}$ in 2016. The red lines show the Pacific Marine Fisheries Commission 5E and 5D major area boundaries.


Figure B.57. Tow locations and density plots for the 2007 Nemesis synoptic survey (see Figure B. 56 caption).


Figure B.58. Tow locations and density plots for the 2008 Frosti synoptic survey (see Figure B. 56 caption).


Figure B.59. Tow locations and density plots for the 2010 Viking Storm synoptic survey (see Figure B. 56 caption).


Figure B.60. Tow locations and density plots for the 2012 Viking Storm synoptic survey (see Figure B. 56 caption).


Figure B.61. Tow locations and density plots for the 2016 Frosti synoptic survey (see Figure B. 56 caption).

## B.8.2. Results

WAP was taken in small amounts off the north coast of Graham Island, into the western part of Dixon Entrance and down the west side of Graham in most of the first five surveys [Figure B. 56 (2006), Figure B. 57 (2007), Figure B. 58 (2008), Figure B. 59 (2010), Figure B. 60 (2012)]. WAP catches in the 2016 survey (Figure B.61) were distributed down the west coast of Graham Island down to 53 N (the southern limit of the survey) and there was one large tow west of Langara Island. Walleye Pollock were mainly taken at depths from 210 to 394 m ( 5 to $95 \%$ quantiles), with the majority of the observations lying between 220 and 370 m depth
(Figure B.62).
Table B.18. Biomass estimates for Walleye Pollock from the six west coast Haida Gwaii synoptic surveys. Bootstrap bias-corrected confidence intervals and coefficients of variation (CVs) are based on 1000 random draws with replacement.

| Survey <br> Year | Biomass (t) <br> (Eq. B.4) | Mean bootstrap <br> biomass (t) | Lower bound <br> biomass (t) | Upper bound <br> biomass (t) | Bootstrap <br> CV | Analytic CV <br> (Eq. B.6) |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2006 | 31.0 | 31.0 | 21.8 | 45.9 | 0.194 | 0.191 |
| 2007 | 15.7 | 15.9 | 9.4 | 24.2 | 0.236 | 0.237 |
| 2008 | 34.3 | 34.5 | 20.0 | 54.7 | 0.242 | 0.238 |
| 2010 | 38.4 | 38.6 | 26.8 | 53.8 | 0.183 | 0.191 |
| 2012 | 67.2 | 67.3 | 41.7 | 96.9 | 0.210 | 0.212 |
| 2016 | 61.8 | 62.0 | 18.5 | 159.3 | 0.600 | 0.618 |

Estimated biomass levels for Walleye Pollock from these trawl surveys may show a weak ascending trend (ranging from 31 t in 2010 to 60+ t in 2012 and 2016) (Figure B.63;
Table B.18). The estimated relative errors (RE) for these surveys were low (compared to other WAP surveys) up to 2016, ranging from 18 to $24 \%$, but the 2016 RE was $60 \%$ (Table B.18).
The proportion of tows that captured Walleye Pollock ranged from 32 to $53 \%$ of tows over the six synoptic survey years, with an overall mean of $43 \%$ (Figure B.64). The median WAP catch weight for positive tows was low ( $<2 \mathrm{~kg} / \mathrm{tow}$ ) and the maximum catch weight across all six
surveys was 308 kg (in 2016). This tow was the largest of the series, with the next largest tow being 69 kg , which explains the high variance observed in the 2016 biomass estimate.


Figure B.62. Distribution of observed weights of Walleye Pollock by survey year and 50 m depth zone intervals. Catches are plotted at the mid-point of the interval and circles in the each panel are scaled to the maximum value ( $397 \mathrm{~kg}-200-250 \mathrm{~m}$ interval in 2016). Minimum and maximum depths observed for WAP: 157 m and 473 m , respectively. Depth is taken at the start position for each tow.


Figure B.63. Biomass estimates for Walleye Pollock from the six west coast Haida Gwaii synoptic surveys (Table B.18). Bias-corrected 95\% confidence intervals from 1000 bootstrap replicates are plotted.


Figure B.64. Proportion of tows by year that contain Walleye Pollock for the six west coast Haida Gwaii synoptic surveys.

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## APPENDIX C. COMMERCIAL TRAWL CPUE

## C.1. INTRODUCTION

Commercial catch and effort data have been used to generate indices of abundance in several ways. The simplest indices are derived from the arithmetic mean or geometric mean of catch divided by an appropriate measure of effort (Catch Per Unit Effort or CPUE) but such indices make no adjustments for changes in fishing practices or other non-abundance factors which may affect catch rates. Consequently, methods to standardise for changes to vessel configuration, the timing or location of catch and other possible effects have been developed to remove potential biases to CPUE that may result from such changes. In these models, abundance is represented as a "year effect" and the dependent variable is either an explicitly calculated CPUE represented as catch divided by effort, or an implicit CPUE represented as catch per tow or catch per record. In the latter case, additional effort terms can be offered as explanatory variables, allowing the model to select the effort term with the greatest explanatory power. It is always preferable to standardise for as many factors as possible when using CPUE as a proxy for abundance. Unfortunately, it is often not possible to adjust for factors that might affect the behaviour of fishers, particularly economic factors, resulting in indices that may not entirely reflect the underlying stock abundance.

## C.2. METHODS

## C.2.1. Arithmetic and Unstandardised CPUE

Arithmetic and unstandardised CPUE indices provide potential measures of relative abundance, but are generally considered unreliable because they fail to take into account changes in the fishery, including spatial and temporal changes as well as behavioural and gear changes. They are frequently calculated because they provide a measure of the overall effect of the standardisation procedure.

Arithmetic CPUE ( $A_{y}$ ) in year $y$ was calculated as the total catch for the year divided by the total effort in the year using Eq. C.1:
Eq. C. $1 \quad A_{y}=\sum_{i=1}^{n_{\nu}} C_{i, y} / \sum_{i=1}^{n_{\nu}} E_{i, y}$
where $C_{i, y}$ is the [catch], $E_{i, y}=T_{i, y}$ ([tows]) or $E_{i, y}=H_{i, y}$ ([hours_fished]) for record $i$ in year $y$ , and $n_{y}$ is the number of records in year $y$.

Unstandardised (geometric) CPUE assumes a log-normal error distribution. An unstandardised index of CPUE $\left(G_{y}\right)$ in year $y$ was calculated as the geometric mean of the ratio of catch to effort for each record $i$ in year $y$, using Eq. C.2:

Eq. C. 2

$$
G_{y}=\exp \left[\sum_{i=1}^{n_{y}} \ln \left(\frac{C_{i, y}}{E_{i, y}}\right) / n_{y}\right]
$$

where $C_{i, y}, E_{i, y}$ and $n_{y}$ are as defined for Eq. C. 1

## C.2.2. Standardised CPUE

These models are preferred over the unstandardised models described above because they can account for changes in fishing behaviour and other factors which may affect the estimated abundance trend, as long as the models are provided with adequate data. In the models described below, catch per record is used as the dependent variable and the associated effort is treated as an explanatory variable.

## C.2.2.1. Lognormal Model

Standardised CPUE assumes a lognormal error distribution, with explanatory variables used to represent changes in the fishery. A standardised CPUE index (Eq. C.3) is calculated from a generalised linear model (GLM) (Quinn and Deriso 1999) using a range of explanatory variables including [year], [month], [depth], [vesse/] and other available factors:
Eq. C. $3 \quad \ln \left(I_{i}\right)=B+Y_{y_{i}}+\alpha_{a_{i}}+\beta_{b_{i}}+\ldots+f\left(\chi_{i}\right)+f\left(\delta_{i}\right)+\ldots+\varepsilon_{i}$
where $I_{i}=C_{i}$ or catch;
$B=$ the intercept;
$Y_{y_{i}}=$ year coefficient for the year corresponding to record $i$;
$\alpha_{a_{i}}$ and $\beta_{b_{i}}=$ coefficients for factorial variables $a$ and $b$ corresponding to record $i$; $f\left(\chi_{i}\right)$ and $f\left(\delta_{i}\right)$ are polynomial functions (to the 3rd order) of the continuous variables $\chi_{i}$ and $\delta_{i}$ corresponding to record $i$;
$\varepsilon_{i}=$ an error term.
The actual number of factorial and continuous explanatory variables in each model depends on the model selection criteria. Because each record represents a single tow, $C_{i}$ has an implicit associated effort of one tow. Hours fished for the tow is represented on the right-hand side of the equation, usually as a continuous (polynomial) variable.

Note that calculating standardised CPUE with Eq. C. 3 without additional explanatory variables is equivalent to using Eq. C.2, provided the same definition for $E_{i, y}$ is used.

Canonical coefficients and standard errors were calculated for each categorical variable (Francis 1999). Standardised analyses typically set one of the coefficients to 1.0 without an error term and estimate the remaining coefficients and the associated error relative to the fixed coefficient. This is required because of parameter confounding. The Francis (1999) procedure rescales all coefficients so that the geometric mean of the coefficients is equal to 1.0 and calculates a standard error for each coefficient, including the fixed coefficient.

Coefficient-distribution-influence (CDI) plots are visual tools to facilitate understanding of patterns which may exist in the combination of coefficient values, distributional changes, and annual influence (Bentley et al. 2012). CDI plots were used to illustrate each explanatory variable added to the model.

## C.2.2.2. Binomial Logit Model

The procedure described by Eq. C. 3 is necessarily confined to the positive catch observations in the data set because the logarithm of zero is undefined. Observations with zero catch were modelled by fitting a logit regression model based on a binomial distribution and using the presence/absence of Walleye Pollock as the dependent variable (where 1 is substituted for
$\ln \left(I_{i}\right)$ in Eq. C. 3 if it is a successful catch record and 0 if it is not successful) and using the same data set. Explanatory factors are estimated in the model in the same manner as described in Eq. C.3. Such a model provides an alternative series of standardised coefficients of relative annual changes that is analogous to the series estimated from the lognormal regression.

## C.2.2.3. Combined Model

A combined model, integrating the two sets of relative annual changes estimated by the lognormal and binomial models, can be estimated using the delta distribution, which allows zero and positive observations (Vignaux 1994). Such a model provides a single index of abundance which integrates the signals from the positive (lognormal) and binomial series. This approach uses the following equation to calculate an index based on the two contributing indices:

Eq. C. $4 \quad{ }^{C} Y_{y}={ }^{L} Y_{y} /\left(1-P_{0}\left[1-\frac{1}{{ }^{B} Y_{y}}\right]\right)$
where ${ }^{C} Y_{y}=$ combined index for year $y$,
${ }^{L} Y_{y}=$ lognormal index for year $y$,
${ }^{B} Y_{y}=$ binomial index for year $y$, and
$P_{0} \quad=$ proportion zero for base year 0.
Francis (2001) suggests that a bootstrap procedure is the appropriate way to estimate the variability of the combined index. Therefore, confidence bounds for the combined model were estimated using a bootstrap procedure based on 500 replicates, drawn with replacement.
The index series plots below present normalised values, i.e., each series is divided by its geometric mean so that the series is centred on 1 . This facilitates comparison among series.

## C.3. PRELIMINARY INSPECTION OF THE DATA

The analyses reported in this Appendix is based on tow-by-tow total catch (landings + discards) data collected over the period 1996-2015 for which detailed positional data for every tow are available and there is an estimate of discarded catch for the tow because of the presence of an observer on board the vessel. These data are held in the DFO PacHarvTrawl (PacHarvest) and GFFOS databases (Fisheries and Oceans Canada, Pacific Region, Groundfish Data Unit).

Tow-by-tow catch and effort data for Walleye Pollock (WAP) from the BC trawl fishery operating from Juan de Fuca Strait to the Dixon Entrance from 1996 to 2015 were selected using the following criteria:

- Tow start date between 1 January 1996 and 31 December 2015;
- Bottom trawl type (includes 'unknown' gear) or Midwater trawl type (includes 'unspecified trawl' for vessels fishing in the 'HAKE' fishing sector);
- Fished in PMFC regions: 3C, 3D, 4B, 5A, 5B, 5C, 5D or 5E;
- Fishing success code $<=1$ (code $0=$ unknown; code $1=$ useable);
- Catch of at least one fish or invertebrate species (no water hauls or inanimate object tows);
- Valid depth field;
- Valid latitude and longitude co-ordinates;
- Valid estimate of time towed that was > 0 hours and <= 24 hours.

Each record represents a single tow, which results in equivalency between the number of records and number of tows. Catch per record can therefore be used to represent CPUE, because each record (tow) has an implicit effort component.
The catch and effort data for WAP were sub-divided into two areas (BC North and BC South) based on the localised distribution of midwater trawl catches (see Appendix A). The data were further sub-divided into two gear types for each area: midwater trawl and bottom trawl, on the premise that fishing for this species will differ, with the midwater fishery largely targeting WAP while the bottom trawl fishery catches WAP while targeting other species (Thompson 1981). Furthermore, the catch rates for the two gear types differ greatly and it would be inappropriate to combine the two capture methods in a single analysis.
Figure C. 1 plots the distribution of depth for all successful WAP tows for each of two areas and two gear types per area. A depth range for each analysis was selected from these plots and are summarised in Table C.1.

Table C.1. Depth bins used in CPUE analyses of stock by gear.

| Analysis | Gear | First <br> year | Depth <br> range <br> $(\mathrm{m})$ | Upper <br> bound <br> effort $(\mathrm{h})$ | Minimum <br> bin <br> records | N <br> depth <br> bins | N <br> latitude <br> bins | N <br> locality <br> bins |
| :--- | :--- | :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Area 5CDE | Midwater trawl | 1996 | $50-550$ | 10 | 35 | 20 | 5 | 9 |
|  | Bottom trawl | 1996 | $50-450$ | 8 | 150 | 16 | 18 | 15 |
| Area 3CD5AB | Midwater trawl | 2003 | $50-450$ | 10 | 30 | 16 | 16 | 25 |
| + Minors 12\&20 | Bottom trawl | 1996 | $50-400$ | 8 | 75 | 14 | 28 | 31 |

Vessel qualification criteria for the bottom trawl fisheries were based on number of trips per year and number of years fishing to avoid including vessels which only occasionally captured Walleye Pollock. However, this was not possible to do for the equivalent midwater trawl fisheries because of the small number of participating vessels and the relatively small amount of available catch information (see table below). For the 5CDE fishery, only two midwater trawl vessels provided data in 2014 and there were only four in 2015 ([upper left panel] Figure C.2). Vessel numbers are higher for the southern fishery, but the overlap of vessels over time is poor in this fishery ([lower left panel] Figure C.2). Consequently the [vessel] explanatory variable was not offered to the midwater models. The vessel qualification criteria used in each analysis appear in Table C. 2 and the distribution of tows by vessel and fishery is presented in Figure C.2. Once a vessel was selected, all data for the qualifying vessel were included, regardless of the number of trips in a year. All vessels were included (i.e., no year restriction) in the 5CDE and 5AB3CD analyses but [vessel] was not selected as a categorical explanatory variable in these analyses.

Table C.2. Vessel qualification criteria used in CPUE analyses of stock by gear.

| Analysis | Trawl Gear | Vessel selection criteria |  |  | Data set characteristics |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{gathered} \mathrm{N} \\ \text { years } \end{gathered}$ | $\begin{gathered} \mathrm{N} \\ \text { trips } \end{gathered}$ | Minimum Records | $\begin{gathered} \mathrm{N} \\ \text { vessels } \end{gathered}$ | $\begin{gathered} \% \\ \text { catch } \end{gathered}$ | catch <br> (t) | Total records | Positive records |
| Area 5CDE | Midwater | NA | NA | NA | NA | 100 | 12,218 | 2,771 | 1,722 |
|  | Bottom | 4 | 4 | 0 | 22 | 82 | 2,347 | 54,836 | 16,344 |
| Area 3CD5AB | Midwater | NA | NA | NA | NA | 100 | 27,448 | 25,139 | 3,492 |
| +Minors12\&20 | Bottom | 5 | 5 | 100 | 40 | 77 | 1,317 | 125,538 | 16,393 |

Table C. 2 shows the number of vessels used in each analysis and the fraction of the total catch represented in each core fleet. There were considerably more vessels participating in the
bottom trawl analyses, with good vessel overlap across years ([right column] Figure C.2). Table C. 3 reports the explanatory variables offered to the models, based on the tow-by-tow information in each record, with the number of available categories varying as indicated in Table C. 1 and Table C.2Table C.4, Table C.5, Table C.6, and Table C. 7 summarise the data used in each analysis by calendar year, including the number of records, the total hours fished and the associated WAP catch.

Table C.3. Explanatory variables offered to the CPUE model, based on the tow-by-tow information.

| Variable | Data type |
| :--- | :--- |
| Year | 20 categories (calendar years) |
| Hours_fished | continuous: 3rd order polynomial |
| Month | Fishing locality areas identified by Rutherford (1995) <br> (includes a final aggregated category) |
| DFO_locality | Latitude aggregated by $0.1^{\circ}$ bands starting at $48^{\circ} \mathrm{N}$ <br> (includes a final aggregated category) |
| Latitude_bands | See Table C.2 for number of categories by analysis <br> (no final aggregated category) |
| Vessel | See Table C.1 for number of categories by analysis <br> (no final aggregated category) |
| Depth_bands |  |

Table C.4. Summary data for the Walleye Pollock midwater trawl fishery in 5CDE by year for the core data set (after applying all data filters).

| Year | Number vessels ${ }^{1}$ | Number trips ${ }^{1}$ | Number tows ${ }^{1}$ | Number records ${ }^{1}$ | Number records ${ }^{2}$ | $\begin{array}{r} \text { \% zero } \\ \text { records }^{2} \end{array}$ | Total catch (t) $)^{1}$ | Total hours ${ }^{1}$ | CPUE $(\mathrm{kg} / \mathrm{h})$ (Eq. C.1) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1996 | 18 | 37 | 94 | 94 | 131 | 28.2 | 156.5 | 281 | 557.1 |
| 1997 | 10 | 20 | 54 | 54 | 60 | 10.0 | 395.8 | 162 | 2,445.7 |
| 1998 | 7 | 12 | 49 | 49 | 59 | 16.9 | 238.8 | 135 | 1,765.1 |
| 1999 | 10 | 24 | 85 | 85 | 95 | 10.5 | 617.5 | 284 | 2,170.8 |
| 2000 | 6 | 8 | 43 | 43 | 59 | 27.1 | 258.0 | 138 | 1,867.8 |
| 2001 | 6 | 11 | 28 | 28 | 144 | 80.6 | 56.7 | 115 | 492.2 |
| 2002 | 7 | 14 | 38 | 38 | 145 | 73.8 | 20.1 | 97 | 208.4 |
| 2003 | 7 | 16 | 60 | 60 | 140 | 57.1 | 566.8 | 201 | 2,824.5 |
| 2004 | 5 | 26 | 112 | 112 | 185 | 39.5 | 920.8 | 321 | 2,868.3 |
| 2005 | 6 | 11 | 90 | 90 | 151 | 40.4 | 127.5 | 151 | 844.8 |
| 2006 | 6 | 23 | 90 | 90 | 147 | 38.8 | 440.1 | 267 | 1,648.1 |
| 2007 | 8 | 45 | 230 | 230 | 341 | 32.6 | 1,172.4 | 765 | 1,533.5 |
| 2008 | 10 | 25 | 115 | 115 | 266 | 56.8 | 286.2 | 275 | 1,041.4 |
| 2009 | 6 | 24 | 116 | 116 | 181 | 35.9 | 1,238.5 | 402 | 3,084.0 |
| 2010 | 5 | 25 | 153 | 153 | 205 | 25.4 | 1,627.9 | 488 | 3,338.6 |
| 2011 | 5 | 17 | 70 | 70 | 106 | 34.0 | 742.1 | 222 | 3,338.9 |
| 2012 | 5 | 18 | 82 | 82 | 95 | 13.7 | 1,033.1 | 307 | 3,366.9 |
| 2013 | 5 | 15 | 76 | 76 | 88 | 13.6 | 705.5 | 320 | 2,206.2 |
| 2014 | 3 | 10 | 42 | 42 | 59 | 28.8 | 470.7 | 182 | 2,588.2 |
| 2015 | 4 | 16 | 95 | 95 | 114 | 16.7 | 1,143.2 | 352 | 3,251.2 |

${ }^{1}$ calculated for tows with Walleye Pollock catch >0
${ }^{2}$ calculated for all tows

Table C.5. Summary data for the Walleye Pollock bottom trawl fishery in 5CDE by year for the core data set (after applying all data filters and selection of core vessels).

| Year | Number vessels ${ }^{1}$ | Number trips ${ }^{1}$ | Number tows ${ }^{1}$ | Number records ${ }^{1}$ | Number records ${ }^{2}$ | $\begin{array}{r} \% \text { zero } \\ \text { records }^{2} \end{array}$ | Total catch $(t)^{1}$ | Tota hours | CPUE $(\mathrm{kg} / \mathrm{h})$ (Eq. C.1) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1996 | 20 | 139 | 1,091 | 1,091 | 2,434 | 55.2 | 388.4 | 2,407 | 161.4 |
| 1997 | 19 | 162 | 909 | 909 | 2,538 | 64.2 | 131.9 | 1,870 | 70.5 |
| 1998 | 18 | 189 | 1,002 | 1,002 | 3,520 | 71.5 | 122.2 | 2,195 | 55.7 |
| 1999 | 20 | 221 | 1,299 | 1,299 | 3,908 | 66.8 | 177.7 | 2,736 | 64.9 |
| 2000 | 19 | 189 | 926 | 926 | 4,202 | 78.0 | 121.2 | 2,010 | 60.3 |
| 2001 | 20 | 163 | 837 | 837 | 3,310 | 74.7 | 56.3 | 1,831 | 30.8 |
| 2002 | 19 | 176 | 988 | 988 | 3,593 | 72.5 | 60.1 | 1,764 | 34.1 |
| 2003 | 17 | 158 | 730 | 730 | 2,851 | 74.4 | 29.2 | 1,334 | 21.9 |
| 2004 | 17 | 185 | 838 | 838 | 3,152 | 73.4 | 89.2 | 1,539 | 57.9 |
| 2005 | 17 | 228 | 1,092 | 1,092 | 3,206 | 65.9 | 344.9 | 1,803 | 191.3 |
| 2006 | 15 | 157 | 763 | 763 | 2,493 | 69.4 | 65.7 | 1,584 | 41.5 |
| 2007 | 15 | 123 | 578 | 578 | 2,106 | 72.6 | 65.4 | 1,076 | 60.8 |
| 2008 | 13 | 129 | 587 | 587 | 2,151 | 72.7 | 62.1 | 1,036 | 60.0 |
| 2009 | 14 | 107 | 407 | 407 | 2,465 | 83.5 | 75.5 | 746 | 101.2 |
| 2010 | 12 | 122 | 473 | 473 | 2,265 | 79.1 | 49.9 | 873 | 57.1 |
| 2011 | 13 | 168 | 829 | 829 | 2,474 | 66.5 | 88.5 | 1,419 | 62.3 |
| 2012 | 11 | 139 | 916 | 916 | 2,359 | 61.2 | 93.1 | 1,604 | 58.0 |
| 2013 | 12 | 146 | 841 | 841 | 2,409 | 65.1 | 112.9 | 1,628 | 69.4 |
| 2014 | 10 | 129 | 686 | 686 | 1,721 | 60.1 | 81.2 | 1,364 | 59.6 |
| 2015 | 10 | 107 | 552 | 552 | 1,679 | 67.1 | 131.9 | 932 | 141.5 |

${ }^{1}$ calculated for tows with Walleye Pollock catch >0
${ }^{2}$ calculated for all tows
Table C.6. Summary data for the Walleye Pollock midwater trawl fishery in 3CD5AB+Minor Areas 12\&20 by year for the core data set (after applying all data filters, including vessel filter).

| Year | Number vessels ${ }^{1}$ | Number trips ${ }^{1}$ | Number tows ${ }^{1}$ | Number records ${ }^{1}$ | Number records ${ }^{2}$ | $\begin{array}{r} \% \text { zero } \\ \text { records } \end{array}$ | Total catch ( $\mathrm{t}^{1}$ | Total hours | CPUE (kg/h) (Eq. C.1) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2003 | 35 | 110 | 293 | 293 | 1,135 | 74.2 | 1,772.1 | 753 | 2,354.7 |
| 2004 | 22 | 42 | 100 | 100 | 818 | 87.8 | 611.3 | 254 | 2,408.4 |
| 2005 | 26 | 59 | 417 | 417 | 1,431 | 70.9 | 828.6 | 1,015 | 816.4 |
| 2006 | 16 | 64 | 578 | 578 | 1,675 | 65.5 | 2,444.8 | 1,502 | 1,628.1 |
| 2007 | 20 | 56 | 358 | 358 | 2,591 | 86.2 | 1,980.0 | 1,024 | 1,933.8 |
| 2008 | 12 | 35 | 143 | 143 | 2,669 | 94.6 | 1,050.5 | 571 | 1,840.4 |
| 2009 | 18 | 58 | 266 | 266 | 2,099 | 87.3 | 2,186.8 | 672 | 3,255.1 |
| 2010 | 9 | 37 | 179 | 179 | 2,125 | 91.6 | 1,832.0 | 839 | 2,184.8 |
| 2011 | 18 | 58 | 228 | 228 | 2,123 | 89.3 | 2,835.3 | 737 | 3,846.4 |
| 2012 | 17 | 54 | 224 | 224 | 2,115 | 89.4 | 2,772.2 | 644 | 4,302.8 |
| 2013 | 14 | 37 | 116 | 116 | 2,319 | 95.0 | 1,433.3 | 418 | 3,427.2 |
| 2014 | 23 | 101 | 338 | 338 | 2,125 | 84.1 | 5,557.0 | 1,008 | 5,514.1 |
| 2015 | 21 | 75 | 252 | 252 | 1,914 | 86.8 | 2,143.9 | 528 | 4,059.4 |

${ }^{1}$ calculated for tows with Walleye Pollock catch >0
${ }^{2}$ calculated for all tows

Table C.7. Summary data for the Walleye Pollock bottom trawl fishery in 3CD5AB+Minor Areas $12 \& 20$ by year for the core data set (after applying all data filters and selection of core vessels).

| Year | Number vessels ${ }^{1}$ | Number trips ${ }^{1}$ | Number tows ${ }^{1}$ | Number records ${ }^{1}$ | Number records ${ }^{2}$ | $\begin{aligned} & \text { \% zero } \\ & \text { records } \end{aligned}$ | Total catch <br> (t) ${ }^{1}$ | Total hours ${ }^{1}$ | CPUE (kg/h) (Eq. C.1) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1996 | 34 | 102 | 237 | 237 | 4,111 | 94.2 | 81.8 | 626 | 130.5 |
| 1997 | 36 | 238 | 988 | 988 | 5,157 | 80.8 | 68.9 | 2,294 | 30.0 |
| 1998 | 36 | 253 | 969 | 969 | 5,885 | 83.5 | 40.9 | 2,297 | 17.8 |
| 1999 | 36 | 310 | 1,041 | 1,041 | 6,892 | 84.9 | 33.3 | 2,539 | 13.1 |
| 2000 | 37 | 311 | 958 | 958 | 8,407 | 88.6 | 34.1 | 2,128 | 16.0 |
| 2001 | 37 | 273 | 728 | 728 | 7,825 | 90.7 | 93.2 | 1,430 | 65.2 |
| 2002 | 37 | 276 | 834 | 834 | 8,983 | 90.7 | 55.1 | 1,763 | 31.3 |
| 2003 | 37 | 313 | 856 | 856 | 9,054 | 90.5 | 55.0 | 1,677 | 32.8 |
| 2004 | 37 | 340 | 1,057 | 1,057 | 8,424 | 87.5 | 38.9 | 2,340 | 16.6 |
| 2005 | 35 | 319 | 1,128 | 1,128 | 8,462 | 86.7 | 68.9 | 2,474 | 27.8 |
| 2006 | 29 | 260 | 963 | 963 | 6,871 | 86.0 | 42.1 | 2,318 | 18.1 |
| 2007 | 29 | 221 | 852 | 852 | 5,871 | 85.5 | 62.9 | 2,068 | 30.4 |
| 2008 | 27 | 163 | 479 | 479 | 4,834 | 90.1 | 11.3 | 1,117 | 10.1 |
| 2009 | 27 | 208 | 602 | 602 | 5,685 | 89.4 | 28.2 | 1,340 | 21.1 |
| 2010 | 25 | 227 | 762 | 762 | 5,925 | 87.1 | 50.7 | 1,914 | 26.5 |
| 2011 | 25 | 198 | 780 | 780 | 5,709 | 86.3 | 86.3 | 1,892 | 45.6 |
| 2012 | 26 | 205 | 713 | 713 | 4,719 | 84.9 | 47.9 | 1,610 | 29.8 |
| 2013 | 22 | 189 | 691 | 691 | 4,529 | 84.7 | 45.5 | 1,603 | 28.4 |
| 2014 | 22 | 202 | 750 | 750 | 4,077 | 81.6 | 285.5 | 1,699 | 168.1 |
| 2015 | 21 | 219 | 1,005 | 1,005 | 4,118 | 75.6 | 86.1 | 2,240 | 38.4 |

${ }^{1}$ calculated for tows with Walleye Pollock catch $>0 ;{ }^{2}$ calculated for all tows


Figure C.1. Depth distribution of tows capturing WAP for the four GLM analyses from 1996 to 2015 using $25 m$ intervals (each bin is labelled with the upper bound of the interval). Vertical lines indicate the $1 \%$ and 99\% quantiles.


Figure C.2. Bubble plot showing vessel participation (number tows) by the core fleets in the four GLM analyses. Vessels are coded in ascending order total catch by year.

## C.4. RESULTS

## C.4.1. Areas 5CDE (North)

## C.4.1.1. Midwater trawl fishery

A standardised lognormal General Linear Model (GLM) analysis was performed on positive catch records from the midwater trawl tow-by-tow data set generated as described in Section C.3. Six explanatory variables (described in Section C. 3 above) were offered to the model and $\operatorname{In}$ (catch) was used as the dependent variable, where catch is the total by weight of landed plus discarded Walleye Pollock in each record (tow) (Eq. C.3). The resulting CPUE index series is presented in Figure C.3.
The [Year] categorical variable was forced as the first variable in the model without regard to its effect on the model deviance. The remaining five variables were offered sequentially, with a stepwise acceptance of the remaining variables with the best AIC. This process was continued until the improvement in the model $\mathrm{R}^{2}$ was less than $1 \%$ (Table C.8). This model selected three of the five remaining explanatory variables, including [DFO_locality], [Month], and
[Hours_fished] in addition to [Year]. The final lognormal model accounted for 61\% of the total model deviance (Table C.8), with the year variable explaining $28 \%$ of the model deviance.

Model residuals do not fit the underlying lognormal distributional assumption very well, with deviations at the tails and in the body of the residual distribution (Figure C.4).
A stepwise plot showing the effect on the year indices as each explanatory variable was introduced into the model shows that the standardisation procedure brought down the unstandardised year indices from peaks observed in 1997, 1999, 2004, 2010, 2012 and 2015 (Figure C.5).

CDI plots of the three explanatory variables introduced to the model in addition to [Year] show a strong standardisation effect with the addition of the [DFO_locality] variable (Figure C.6). Unfortunately, two locations, which have no previous history in the model, dominate in 2014 and 2015. This means that these location coefficients will be poorly estimated and confounded with the year effect. The variables [Month] (Figure C.7) and [Hours_fished] (Figure C.8) show high annual variability, with strong shifts in the resulting indices. This is evidence of poor consistency in the underlying data.

The year indices show a rising index from the mid- to late-2000s (Figure C.3) but this model has poor residual diagnostics, strong abrupt annual shifts in data and is consequently not reliable.

Table C.8. Order of acceptance of variables into the lognormal model of positive total mortalities (verified landings plus discards) of Walleye Pollock 5CDE midwater trawl fishery with the amount of explained deviance $\left(R^{2}\right)$ for each variable. Variables accepted into the model are marked in bold with an *. Year was forced as the first variable and vessel was not offered to this model.

| Variable | 1 | 2 | 3 | 4 | 5 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Year* $^{*}$ | $\mathbf{0 . 2 8 4 6}$ | - | - | - | - |
| DFO_locality* $^{*}$ | 0.4278 | $\mathbf{0 . 5 3 1 3}$ | - | - | - |
| Month* $^{*}$ | 0.2004 | 0.4082 | $\mathbf{0 . 5 8 8 7}$ | - | - |
| Hours_fished* $^{\text {Latitude_bands }}$ | 0.1463 | 0.3380 | 0.5611 | $\mathbf{0 . 6 0 6 5}$ | - |
| Depth_bands | 0.1003 | 0.3731 | 0.5344 | 0.5906 | 0.6087 |
| Improvement in deviance | 0.0466 | 0.3094 | 0.5377 | 0.5944 | 0.6115 |



Figure C.3. Three CPUE series for WAP from 1996 to 2015 in 5CDE midwater trawl fishery. The solid line is the standardised CPUE series from the lognormal model (Eq. C.3). The arithmetic (Eq. C.1) and the unstandardised series (Eq. C.2) are also presented. All series are scaled to same geometric mean.


Figure C.4. Residual diagnostic plots for the GLM lognormal analysis for Walleye Pollock in 5CDE midwater trawl fishery. Upper left: histogram of the standardised residuals with overlaid lognormal distribution (SDNR = standard deviation of normalised residuals. MASR $=$ median of absolute standardised residuals). Lower left: Q-Q plot of the standardised residuals with the outside horizontal and vertical lines representing the $5^{\text {th }}$ and $95^{\text {th }}$ percentiles of the theoretical and observed distributions. Upper right: standardised residuals plotted against the predicted CPUE. Lower right: observed CPUE plotted against the predicted CPUE.


Figure C.5. Plot showing the year coefficients after adding each successive term of the standardised lognormal regression analysis for Walleye Pollock in the 5CDE midwater trawl fishery. The final model is shown with a thick solid black line. Each line has been scaled so that the geometric mean equals 1.0.


Figure C.6. CDI plot showing the effect of introducing the categorical variable [DFO_locality] to the lognormal regression model for Walleye Pollock in the 5CDE midwater trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.7. CDI plot showing the effect of introducing the categorical variable [Month] to the lognormal regression model for Walleye Pollock in the 5CDE midwater trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution of variable records by year (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.8. CDI plot showing the effect of introducing the continuous variable [Hours_fished] to the lognormal regression model for Walleye Pollock in the 5CDE midwater trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).

## C.4.1.2. Bottom trawl fishery: positive lognormal model

A standardised lognormal General Linear Model (GLM) analysis was performed on positive catch records from the bottom trawl tow-by-tow data set generated as described in Section C.3. Seven explanatory variables (described in Section C. 3 above) were offered to the model and $\ln$ (catch) was used as the dependent variable, where catch is the total by weight of landed plus discarded Walleye Pollock in each record (tow) (Eq. C.3). The resulting CPUE index series is presented in Figure C.9.

The [Year] categorical variable was forced as the first variable in the model without regard to its effect on the model deviance. The remaining six variables were offered sequentially, with a stepwise acceptance of the remaining variables with the best AIC. This process was continued until the improvement in the model $\mathrm{R}^{2}$ was less than $1 \%$ (Table C.9). This model selected all six of the remaining explanatory variables, including [DFO_locality], [Depth_bands]
[Month], [Vessel], [Hours_fished] and [Latitude_bands] in addition to [Year]. The final lognormal model accounted for $29 \%$ of the total model deviance (Table C.9), with the year variable explaining $5 \%$ of the model deviance.
Model residuals show an excellent fit with the underlying lognormal distributional assumption, with no deviation at the tails or in the body of the residual distribution (Figure C.10).
A stepwise plot showing the effect on the year indices as each explanatory variable was introduced into the model shows that the standardisation procedure did not modify the underlying annual indices, except for a small drop at the beginning of the series (Figure C.11). This is indicative of stability in the underlying data, with few annual shifts.
CDI plots of the six explanatory variables introduced to the model in addition to [Year] show a strong standardisation effect at the beginning of the series with the addition of [DFO_locality] variable (Figure C.12). There is a trend in the standardisation effect from the addition of the [Depth_bands] variable (Figure C.13), but the effect is relatively small. The variables [Month] (Figure C.14), [Vesssel] (Figure C.15), [Hours_fished] (Figure C.16) and [Latitude_bands] (Figure C.17) all show some standardisation effects, but the overall model impact is small.
The year indices show a decline at the beginning of the series, followed by no change to 2015 (Figure C.9). This model has excellent diagnostics and shows little change from the unstandardised series.

Table C.9. Order of acceptance of variables into the lognormal model of positive total mortalities (verified landings plus discards) of Walleye Pollock 5CDE bottom trawl fishery with the amount of explained deviance $\left(R^{2}\right)$ for each variable. Variables accepted into the model are marked in bold with an *. Year was forced as the first variable.

| Variable $^{\|c\|} 1$ | 2 | 3 | 4 | 5 | 6 | 7 |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Year $^{*}$ | $\mathbf{0 . 0 5 1 9}$ | - | - | - | - | - | - |
| DFO_locality* $^{*}$ | 0.1237 | $\mathbf{0 . 1 7 7 4}$ | - | - | - | - | - |
| Depth_bands* $^{\text {Month* }}$ | 0.0635 | 0.1114 | $\mathbf{0 . 2 1 4 7}$ | - | - | - | - |
| Vessel* $^{*}$ | 0.0504 | 0.1018 | 0.2048 | $\mathbf{0 . 2 4 7 4}$ | - | - | - |
| Hours_fished* $^{\text {Latitude_bands* }}$ | 0.0225 | 0.0746 | 0.2043 | 0.2383 | $\mathbf{0 . 2 6 9 2}$ | - | - |
| Improvement in deviance | 0.0137 | 0.0630 | 0.1946 | 0.2304 | 0.2623 | $\mathbf{0 . 2 8 1 0}$ | - |



Figure C.9. Three CPUE series for Walleye Pollock from 1996 to 2015 in 5CDE bottom trawl fishery. The solid line is the standardised CPUE series from the lognormal model (Eq. C.3). The arithmetic series (Eq. C.1) and the unstandardised series (Eq. C.2) are also presented. All three series have been scaled to same geometric mean.


Figure C.10. Residual diagnostic plots for the GLM lognormal analysis for Walleye Pollock in 5CDE bottom trawl fishery. Upper left: histogram of the standardised residuals with overlaid lognormal distribution (SDNR = standard deviation of normalised residuals. MASR $=$ median of absolute standardised residuals). Lower left: Q-Q plot of the standardised residuals with the outside horizontal and vertical lines representing the $5^{\text {th }}$ and $95^{\text {th }}$ percentiles of the theoretical and observed distributions. Upper right: standardised residuals plotted against the predicted CPUE. Lower right: observed CPUE plotted against the predicted CPUE.


Figure C.11. Plot showing the year coefficients after adding each successive term of the standardised lognormal regression analysis for Walleye Pollock in the 5CDE bottom trawl fishery. The final model is shown with a thick solid black line. Each line has been scaled so that the geometric mean equals 1.0.


Figure C.12. CDI plot showing the effect of introducing the categorical variable [DFO_locality] to the lognormal regression model for Walleye Pollock in the 5CDE bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.13. CDI plot showing the effect of introducing the categorical variable [Depth_bands] to the lognormal regression model for Walleye Pollock in the 5CDE bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution of variable records by year (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.14. CDI plot showing the effect of introducing the categorical variable [Month] to the lognormal regression model for Walleye Pollock in the 5CDE bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.15. CDI plot showing the effect of introducing the categorical variable [Vessel] to the lognormal regression model for Walleye Pollock in the 5CDE bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.16. CDI plot showing the effect of introducing the categorical variable [Hours_fished] to the lognormal regression model for Walleye Pollock in the 5CDE bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.17. CDI plot showing the effect of introducing the categorical variable [Latitude_bands] to the lognormal regression model for Walleye Pollock in the 5CDE bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).

## C.4.1.3. Bottom trawl fishery: binomial logit model

The same variables used in the lognormal model were offered sequentially to this model, beginning with the year categorical variable, until the improvement in the model $R^{2}$ was less than $1 \%$ (Table C.10). This model drops in the first four years, followed by a period of little trend up to 2009, then shifts to a strong increase that peaks in 2014 (Figure C.18).

Table C.10. Order of acceptance of variables into the binomial model of presence/absence of verified landings plus discards of Walleye Pollock in 5CDE bottom trawl fishery with the amount of explained deviance $\left(R^{2}\right)$ for each variable. Variables accepted into the model are marked in bold with an *. Year was forced as the first variable.

| Variable | 1 | 2 | 3 | 4 | 5 | 6 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Year* $^{*}$ | $\mathbf{0 . 0 1 6 6}$ | - | - | - | - | - |
| DFO_locality* $^{*}$ | 0.1414 | $\mathbf{0 . 1 5 6 5}$ | - | - | - | - |
| Depth_bands* $^{\text {Month* }}$ | 0.0829 | 0.0968 | $\mathbf{0 . 2 0 1 5}$ | - | - | - |
| Latitude_bands* $^{\text {Hours_fished }}$ | 0.0273 | 0.0433 | 0.1803 | $\mathbf{0 . 2 2 2 4}$ | - | - |
| Vessel | 0.1396 | 0.1547 | 0.1788 | 0.2143 | $\mathbf{0 . 2 3 4 9}$ | - |
| Improvement in deviance | 0.0090 | 0.0250 | 0.1618 | 0.2120 | 0.2327 | 0.2447 |



Figure C.18. Binomial index series for the 5CDE bottom trawl fishery also showing the trend in proportion of zero tows from the same data set.

## C.4.1.4. Bottom trawl fishery: combined model

The combined model (Eq. C.4) initially drops, as seen in both the lognormal and binomial series, but then takes on a gradually increasing trend that more closely resembles the binomial series (Figure C.19), giving the series a slight "U" shape.

$95 \%$ bias corrected error bars for combined index based on 500 bootstrap replicates
Figure C.19. Combined index series (Eq. C.4) for the 5CDE bottom trawl fishery also showing the contributing lognormal and binomial index series. Confidence bounds based on 500 bootstrap replicates.

## C.4.2. Areas 3CD5AB + Minor Areas 12\&20 (South)

## C.4.2.1. Midwater trawl fishery

A standardised lognormal General Linear Model (GLM) analysis was performed on positive catch records from the midwater trawl tow-by-tow data set generated as described in Section C.3. Six explanatory variables (described in Section C. 3 above) were offered to the model and $\ln$ (catch) was used as the dependent variable, where catch is the total by weight of landed plus discarded Walleye Pollock in each record (tow) (Eq. C.3). The resulting CPUE index series is presented in Figure C. 20.

The [Year] categorical variable was forced as the first variable in the model without regard to its effect on the model deviance. The remaining five variables were offered sequentially, with a stepwise acceptance of the remaining variables with the best AIC. This process was continued until the improvement in the model $\mathrm{R}^{2}$ was less than $1 \%$ (Table C.11). This model selected four of the five remaining explanatory variables, including [ $0.1^{\circ}$ Latitude_bands],
[DFO_locality], [Month], and [Hours_fished] in addition to [Year]. The final lognormal model accounted for $66 \%$ of the total model deviance (Table C.11), with the year variable explaining $11 \%$ of the model deviance.
As seen in the 5CDE model, residuals do not fit the underlying lognormal distributional assumption very well, with considerable deviation at the tails and in the body of the residual distribution (Figure C.21).
A stepwise plot showing how the year indices change as each explanatory variable was introduced into the model shows a strong standardisation effect near the end of the series, beginning around 2010, with the main effect occurring with the addition of the initial [0.1${ }^{\circ}$ Latitude_bands] variable (Figure C.22).
CDI plots of the four explanatory variables introduced to the model in addition to [Year] show a strong, but variable with large swings between years, standardisation effect with the addition of the [ $0.1^{\circ}$ Latitude_bans] variable (Figure C.23). This type of sharp shifts between years is not a desirable feature in this type of analysis because it indicates lack of stability in the underlying data. The explanatory variables [DFO_Locality] (Figure C.24), [Month] (Figure C.25) and [Hours_fished] (Figure C.26) all show similar behaviour in their contributions to the standardisation model. These substantial year-to-year changes indicate that this model is likely to be unstable.

The year indices show a rising index from the late-2000s (Figure C.20), but this model has poor residual diagnostics, abrupt annual shifts in the underlying data and is consequently not reliable.

Table C.11. Order of acceptance of variables into the lognormal model of positive total mortalities (verified landings plus discards) of Walleye Pollock 3CD5AB+Minor Areas 12\&20 midwater trawl fishery with the amount of explained deviance $\left(R^{2}\right)$ for each variable. Variables accepted into the model are marked with an *. Year was forced as the first variable and vessel was not offered to this model.

| Variable | 1 | 2 | 3 | 4 | 5 | 6 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Year* $^{*}$ | $\mathbf{0 . 1 0 8 9}$ | - | - | - | - | - |
| Latitude_bands* $^{\text {DFO_locality* }}$ | 0.4481 | $\mathbf{0 . 5 9 9 6}$ | - | - | - | - |
| Month* $^{*}$ | 0.4628 | 0.5938 | $\mathbf{0 . 6 2 9 3}$ | - | - | - |
| Hours_fished | 0.2775 | 0.4054 | 0.6144 | $\mathbf{0 . 6 4 5 7}$ | - | - |
| Depth_bands* | 0.1159 | 0.2081 | 0.6131 | 0.6401 | $\mathbf{0 . 6 5 7 9}$ | - |
| Improvement in deviance | 0.2945 | 0.3620 | 0.6070 | 0.6375 | 0.6523 | 0.6656 |



Figure C.20. Three CPUE series for Walleye Pollock from 2003 to 2015 in 3CD5AB+Minor Areas 12\&20 midwater trawl fishery. The solid line is the standardised CPUE series from the lognormal model (Eq. C.3). The arithmetic series (Eq. C.1) and the unstandardised series (Eq. C.2) are also presented. All three series have been scaled to same geometric mean.


Figure C.21. Residual diagnostic plots for the GLM lognormal analysis for Walleye Pollock in 3CD5AB+Minor Areas 12\&20 midwater trawl fishery. Upper left: histogram of the standardised residuals with overlaid lognormal distribution (SDNR = standard deviation of normalised residuals. MASR $=$ median of absolute standardised residuals). Lower left: Q-Q plot of the standardised residuals with the outside horizontal and vertical lines representing the $5^{\text {th }}$ and $95^{\text {th }}$ percentiles of the theoretical and observed distributions. Upper right: standardised residuals plotted against the predicted CPUE. Lower right: observed CPUE plotted against the predicted CPUE.


Figure C.22. Plot showing the year coefficients after adding each successive term of the standardised lognormal regression analysis for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 midwater trawl fishery. The final model is shown with a thick solid black line. Each line has been scaled so that the geometric mean equals 1.0.


Figure C.23. CDI plot showing the effect of introducing the categorical variable [Latitude_bands] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 midwater trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.24. CDI plot showing the effect of introducing the categorical variable [DFO_locality] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 midwater trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.25. CDI plot showing the effect of introducing the categorical variable [Month] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 midwater trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution of variable records by year (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.26. CDI plot showing the effect of introducing the categorical variable [Hours_fished] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 midwater trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).

## C.4.2.2. Bottom trawl fishery: positive lognormal model

A standardised lognormal General Linear Model (GLM) analysis was performed on positive catch records from the bottom trawl tow-by-tow data set generated as described in Section C.3. Seven explanatory variables (described in Section C. 3 above) were offered to the model and $\ln$ (catch) was used as the dependent variable, where catch is the total by weight of landed plus discarded Walleye Pollock in each record (tow) (Eq. C.3). The resulting CPUE index series is presented in Figure C. 27.
The [Year] categorical variable was forced as the first variable in the model without regard to its effect on the model deviance. The remaining six variables were offered sequentially, with a stepwise acceptance of the remaining variables with the best AIC. This process was continued until the improvement in the model $\mathrm{R}^{2}$ was less than 1\% (Table C.12). This model selected five of the six remaining explanatory variables, including [DFO_locality], [Month], [Vessel],
[Hours_fished] and [Latitude_bands] in addition to [Year]. The final lognormal model accounted for $24 \%$ of the total model deviance (Table C.12), with the year variable explaining only $3 \%$ of the model deviance.

Model residuals show an excellent fit with the underlying lognormal distributional assumption, with only a small deviation at the upper tail of the distribution and none in the lower tail or in the body of the residual distribution (Figure C.28).
A stepwise plot showing the effect on the year indices as each explanatory variable was introduced into the model shows that the standardisation procedure made small adjustments to the unstandardised series in 1996, 1997, 2000, 2005 and 2010, resulting in a relatively smooth annual trend (Figure C.29).

CDI plots of the five explanatory variables introduced to the model in addition to [Year] show reasonably strong standardisation effects at the beginning, the middle and the end of the series with the addition of [DFO_locality] variable (Figure C.30). The variables [Month]
(Figure C.31), [Vessel] (Figure C.32), [Hours_fished] (Figure C.33) and
[Latitude_bands] (Figure C.34) also impact the standardisation model, but the effects are relatively minor. A possible exception to this is a strong deviation in the vessel effect in 2001 (Figure C.32), but an examination of the shift in Figure C. 29 shows little change in the 2001 year index when [Vessel] was added to the model.

The lognormal year indices show a declining trend at the beginning of the series, followed by a flat or slightly increasing trend towards the end of the series (Figure C.27). This model has excellent diagnostics and shows only small changes from the unstandardised series.

Table C.12. Order of acceptance of variables into the lognormal model of positive total mortalities (verified landings plus discards) of Walleye Pollock 3CD5AB+Minor Areas 12\&20 bottom trawl fishery with the amount of explained deviance ( $R^{2}$ ) for each variable. Variables accepted into the model are marked in bold with an *. Year was forced as the first variable.

| Variable | 1 | 2 | 3 | 4 | 5 | 6 | 7 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Year* $^{*}$ | $\mathbf{0 . 0 3 1 8}$ | - | - | - | - | - | - |
| DFO_Iocality* $^{\text {* }}$ | 0.1366 | $\mathbf{0 . 1 6 7 5}$ | - | - | - | - | - |
| Month $^{*}$ | 0.0486 | 0.0751 | $\mathbf{0 . 1 8 9 8}$ | - | - | - | - |
| Vessel | 0.0412 | 0.0754 | 0.1900 | $\mathbf{0 . 2 1 2 1}$ | - | - | - |
| Hours_fished* | 0.0155 | 0.0456 | 0.1828 | 0.2041 | $\mathbf{0 . 2 2 9 6}$ | - |  |
| Latitude_bands | 0.1151 | 0.1449 | 0.1813 | 0.2025 | 0.2248 | $\mathbf{0 . 2 4 1 6}$ | 0.2379 |
| Depth_bands | 0.0185 | 0.0482 | 0.1740 | 0.1965 | 0.2194 | 0.2361 | 0.2479 |
| Improvement in deviance | 0 | 0.1358 | 0.0223 | 0.0224 | 0.0174 | 0.0120 | 0.0063 |



Figure C.27. Three CPUE series for Walleye Pollock from 1996 to 2015 in 3CD5AB+Minor Areas 12\&20 bottom trawl fishery. The solid line is the standardised CPUE series from the lognormal model (Eq. C.3). The arithmetic series (Eq. C.1) and the unstandardised series (Eq. C.2) are also presented. All three series have been scaled to same geometric mean.


Figure C.28. Residual diagnostic plots for the GLM lognormal analysis for Walleye Pollock in 3CD5AB+Minor Areas 12\&20 bottom trawl fishery. Upper left: histogram of the standardised residuals with overlaid lognormal distribution (SDNR = standard deviation of normalised residuals. MASR = median of absolute standardised residuals). Lower left: Q-Q plot of the standardised residuals with the outside horizontal and vertical lines representing the $5^{\text {th }}$ and $95^{\text {th }}$ percentiles of the theoretical and observed distributions. Upper right: standardised residuals plotted against the predicted CPUE. Lower right: observed CPUE plotted against the predicted CPUE.


Figure C.29. Plot showing the year coefficients after adding each successive term of the standardised lognormal regression analysis for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 bottom trawl fishery. The final model is shown with a thick solid black line. Each line has been scaled so that the geometric mean equals 1.0.


Figure C.30. CDI plot showing the effect of introducing the categorical variable [DFO_locality] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.31. CDI plot showing the effect of introducing the categorical variable [Month] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution of variable records by year (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.32. CDI plot showing the effect of introducing the continuous variable [Vessel] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.33. CDI plot showing the effect of introducing the categorical variable [Hours_fished] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).


Figure C.34. CDI plot showing the effect of introducing the categorical variable [Latitude_bands] to the lognormal regression model for Walleye Pollock in the 3CD5AB+Minor Areas 12\&20 bottom trawl fishery. Each plot consists of subplots showing the effect by level of variable (top left), the relative distribution by year of variable records (bottom left), and the cumulative effect of variable by year (bottom right).

## C.4.2.3. Bottom trawl fishery: binomial logit model

The same variables used in the lognormal model were offered sequentially to this model, beginning with the year categorical variable, until the improvement in the model $R^{2}$ was less than $1 \%$ (Table C.13). This model shows no little trend up to about 2009 or 2010, but is then followed by a strong increase up to 2015 (Figure C.35).

Table C.13. Order of acceptance of variables into the binomial model of presence/absence of verified landings plus discards of Walleye Pollock in 3CD5AB+Minor Areas 12\&20 bottom trawl fishery with the amount of explained deviance $\left(R^{2}\right)$ for each variable. Variables accepted into the model are marked in bold with an *. Year was forced as the first variable.

| Variable | 1 | 2 | 3 | 4 | 5 | 6 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Year* $^{*}$ | $\mathbf{0 . 0 1 4 9}$ |  | - | - | - | - |
| Latitude_bands | 0.0921 | $\mathbf{0 . 1 0 7 5}$ | - | - | - | - |
| Depth_bands* | 0.0541 | 0.0675 | $\mathbf{0 . 1 5 1 0}$ | - | - | - |
| DFO_locality*** | 0.0865 | 0.1042 | 0.1415 | $\mathbf{0 . 1 8 5 4}$ | - | - |
| Hours_fished * | 0.0135 | 0.0277 | 0.1167 | 0.1673 | $\mathbf{0 . 2 0 0 0}$ | - |
| Vessel | 0.0166 | 0.0306 | 0.1162 | 0.1591 | 0.1918 | 0.2072 |
| Month | 0.0098 | 0.0259 | 0.1111 | 0.1565 | 0.1890 | 0.2029 |
| Improvement in deviance | 0 | 0.0927 | 0.0435 | 0.0344 | 0.0146 | 0.0071 |



Error bars=+/-1.96*SE
Figure C.35. Binomial index series for the 3CD5AB + Minor Areas 12\&20 bottom trawl fishery also showing the trend in proportion of zero tows from the same data set.

## C.4.2.4. Bottom trawl fishery: combined model

The combined model (Eq. C.4) shows an even stronger "U" pattern than did the 5CDE (North) model, with a declining trend to early 2000s, followed by a period of little change and increasing strongly to 2015, taking its signal from the binomial series (Figure C.36).

$95 \%$ bias corrected error bars for combined index based on 500 bootstrap replicates
Figure C.36. Combined index series (Eq. C.4) for the 3CD5AB+Minor Areas $12 \& 20$ bottom trawl fishery also showing the contributing lognormal and binomial index series. Confidence bounds based on 500 bootstrap replicates.

## C.5. COMPARISONS WITHIN AND AMONG STOCKS

Figure C. 37 compares the lognormal midwater and bottom trawl indices for the defined 5CDE (North) WAP area, showing a strong increase in the midwater trawl CPUE beginning from 2008 while there was only a moderate increase in the bottom trawl CPUE in the mid-2000s, followed by a stable index over the last decade.
Figure C. 38 compares the lognormal midwater and bottom trawl indices for the defined 3CD5AB+Minor Areas 12\&20 (South) WAP area, showing an increasing trend in midwater trawl CPUE from the mid-2000s while the bottom trawl index is mainly stable over the same period, after dropping from a high level observed in the late 1990s.
Both lognormal midwater trawl indices (Figure C.39) increase, starting at the end of the 2000s. The overall level of increase is much greater for the South series, but this series also shows strong annual variations which are unlikely to be abundance driven. Given the poor residual diagnostics for both series (see Figure C. 4 and Figure C.21) and the strong year-to-year variations observed in both series, neither can be considered sufficiently reliable for use as an indicator of abundance trends.

The two combined bottom trawl CPUE series show less annual variability than the midwater trawl series (Figure C.40) and show similar patterns up to the early 2010s, with both series declining from high levels in 1996 or 1997 and then maintaining a consistent level just below the series mean up to around 2012/2013. The large jump from 1996 to 1997 in the South series is unlikely to be abundance driven, especially given the small amount of data in 1996 (see Table C.7: 1996 has about one-quarter the number tows and less than half the trips of subsequent years). The North and South series diverge after about 2013, with the South series
showing a strong increase to over three times the series mean while the North maintains a level just above the series mean (Figure C.40).


Figure C.37. Comparison of the lognormal bottom trawl and midwater CPUE index series for the 5CDE (North) WAP area.


Each relative series scaled so that the geometric mean=1.0 from 2003 to 2015
Figure C.38. Comparison of the lognormal bottom trawl and midwater CPUE index series for the 3CD5AB+Minor Areas 12\&20 (South) WAP area.


Each relative series scaled so that the geometric mean=1.0 from 2003 to 2015
Figure C.39. Comparison of the lognormal midwater trawl CPUE index series for the North and South WAP area definitions.


Figure C.40. Comparison of the combined (Eq. C.4) bottom trawl CPUE index series for the North and South WAP area definitions.

## C.6. RELATIVE INDICES OF ABUNDANCE

Table C. 14 to Table C. 16 summarise the relative indices of abundance derived from the CPUE analyses for the two Walleye Pollock stocks. The midwater gear yields indices from 1996 to 2015 for BC North, but only from 2003 to 2015 for BC South (Table C.14). These indices were not used in the model because they rely on directed effort that is heavily influenced by fisher behaviour and management regulation.

CPUE indices used in the delay-difference model appear as the delta-lognormal (combined) indices from the bottom trawl data (BC North: Table C.15, BC South: Table C.16).
A fixed CV of 0.3 was applied to each CPUE index in the model, which iSCAM accepts as the inverse of CV: $v_{j t}=1 / c_{j t}=1 / 0.3=3.333$ (see Eqn E.29).

Table C.14. Relative indices of annual CPUE (kg/tow) from the arithmetic, geometric and lognormal models of non-zero catches of midwater trawl fishery for Walleye Pollock in the indicated sub-regions with associated standard error (SE) for the lognormal model. -‘‘i indices not available.

|  | 5CDE |  |  |  | 3CD5AB+Minor Areas 12\&20 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Arithmetic (Eq. C.1) | Geometric (Eq. C.2) | Lognormal (Eq. C.3) | Standard error | Arithmetic (Eq. C.1) | Geometric (Eq. C.2) | Lognormal (Eq. C.3) | Standard error |
| 1996 | 557 | 528 | 1,048 | 0.171 | - | - | - | - |
| 1997 | 2,446 | 3,402 | 2,286 | 0.206 | - | - | - | - |
| 1998 | 1,765 | 2,765 | 1,659 | 0.211 | - | - | - | - |
| 1999 | 2,171 | 3,901 | 2,086 | 0.164 | - | - | - | - |
| 2000 | 1,868 | 2,304 | 1,825 | 0.226 | - | - | - | - |
| 2001 | 492 | 422 | 597 | 0.276 | - | - | - | - |
| 2002 | 208 | 67 | 301 | 0.252 | - | - | - | - |
| 2003 | 2,825 | 1,638 | 1,425 | 0.195 | 2,355 | 4,119 | 2,742 | 0.116 |
| 2004 | 2,868 | 3,018 | 1,838 | 0.150 | 2,408 | 2,499 | 1,691 | 0.166 |
| 2005 | 845 | 278 | 328 | 0.187 | 816 | 387 | 395 | 0.096 |
| 2006 | 1,648 | 1,310 | 1,278 | 0.168 | 1,628 | 1,895 | 933 | 0.085 |
| 2007 | 1,534 | 1,363 | 1,334 | 0.114 | 1,934 | 1,356 | 1,654 | 0.110 |
| 2008 | 1,041 | 348 | 1,118 | 0.175 | 1,840 | 935 | 1,247 | 0.146 |
| 2009 | 3,084 | 3,391 | 2,961 | 0.143 | 3,255 | 2,185 | 1,730 | 0.106 |
| 2010 | 3,339 | 5,541 | 3,509 | 0.127 | 2,185 | 3,000 | 3,040 | 0.135 |
| 2011 | 3,339 | 2,906 | 3,621 | 0.183 | 3,846 | 4,336 | 5,575 | 0.116 |
| 2012 | 3,367 | 7,401 | 4,430 | 0.169 | 4,303 | 4,791 | 7,442 | 0.117 |
| 2013 | 2,206 | 4,152 | 4,525 | 0.176 | 3,427 | 4,476 | 3,618 | 0.150 |
| 2014 | 2,588 | 2,996 | 3,217 | 0.276 | 5,514 | 7,090 | 11,208 | 0.101 |
| 2015 | 3,251 | 7,179 | 3,596 | 0.254 | 4,059 | 5,861 | 7,777 | 0.117 |

Table C.15. Relative indices of annual CPUE from the arithmetic, unstandardised, lognormal models of non-zero bottom trawl catches of Walleye Pollock in 5CDE. Also shown are the indices from the binomial model of presence/absence in this fishery and the combined delta-lognormal model (Eq. C.4). All indices are scaled so that their geometric means equal 1.0. Upper and lower 95\% confidence bounds and associated standard error (SE) are presented for the lognormal model, while bootstrapped upper and lower 95\% confidence bounds are presented for the combined model.

| Year | Arithmetic Index (Eq. C.1) | Geometric Index (Eq. C.2) | Lognormal (Eq. C.3) |  |  |  | Binomial Index (Eq. C.3) | Combined (Eq. C.4) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Index | Lower bound | Upper bound | SE |  | Index | Lower bound | Upper bound |
| 1996 | 2.538 | 2.341 | 2.399 | 2.172 | 2.650 | 0.0497 | 1.522 | 3.005 | 2.721 | 3.336 |
| 1997 | 1.109 | 1.108 | 1.242 | 1.122 | 1.375 | 0.0509 | 1.258 | 1.423 | 1.269 | 1.593 |
| 1998 | 0.875 | 0.830 | 0.961 | 0.871 | 1.061 | 0.0495 | 0.817 | 0.869 | 0.778 | 0.974 |
| 1999 | 1.021 | 0.951 | 1.040 | 0.954 | 1.135 | 0.0436 | 1.074 | 1.098 | 0.995 | 1.197 |
| 2000 | 0.948 | 0.877 | 0.904 | 0.818 | 0.998 | 0.0498 | 0.564 | 0.644 | 0.574 | 0.729 |
| 2001 | 0.484 | 0.653 | 0.525 | 0.473 | 0.582 | 0.0520 | 0.622 | 0.399 | 0.358 | 0.440 |
| 2002 | 0.536 | 0.553 | 0.567 | 0.516 | 0.624 | 0.0474 | 0.835 | 0.520 | 0.462 | 0.564 |
| 2003 | 0.344 | 0.530 | 0.632 | 0.567 | 0.705 | 0.0546 | 0.915 | 0.611 | 0.549 | 0.686 |
| 2004 | 0.911 | 0.707 | 0.726 | 0.655 | 0.804 | 0.0511 | 0.704 | 0.598 | 0.534 | 0.667 |
| 2005 | 3.009 | 1.914 | 1.422 | 1.296 | 1.559 | 0.0463 | 0.912 | 1.370 | 1.223 | 1.510 |
| 2006 | 0.652 | 0.799 | 0.940 | 0.845 | 1.046 | 0.0534 | 1.167 | 1.036 | 0.933 | 1.165 |
| 2007 | 0.957 | 1.301 | 1.286 | 1.140 | 1.451 | 0.0604 | 0.918 | 1.245 | 1.102 | 1.424 |
| 2008 | 0.943 | 1.280 | 1.046 | 0.927 | 1.180 | 0.0603 | 0.846 | 0.966 | 0.862 | 1.098 |
| 2009 | 1.592 | 1.264 | 1.113 | 0.965 | 1.285 | 0.0716 | 0.550 | 0.779 | 0.641 | 0.894 |
| 2010 | 0.899 | 1.018 | 1.089 | 0.952 | 1.245 | 0.0670 | 0.792 | 0.966 | 0.843 | 1.108 |
| 2011 | 0.980 | 1.083 | 0.942 | 0.850 | 1.046 | 0.0519 | 1.327 | 1.108 | 0.961 | 1.220 |
| 2012 | 0.913 | 1.120 | 1.224 | 1.108 | 1.352 | 0.0497 | 1.770 | 1.636 | 1.470 | 1.819 |
| 2013 | 1.091 | 1.123 | 1.052 | 0.950 | 1.165 | 0.0510 | 1.421 | 1.277 | 1.143 | 1.426 |
| 2014 | 0.937 | 0.919 | 1.083 | 0.966 | 1.215 | 0.0574 | 1.973 | 1.512 | 1.332 | 1.716 |
| 2015 | 2.225 | 1.006 | 0.950 | 0.838 | 1.078 | 0.0630 | 1.331 | 1.119 | 0.969 | 1.307 |

Table C.16. Relative indices of annual CPUE from the arithmetic, unstandardised, lognormal models of non-zero bottom trawl catches of Walleye Pollock in 3CD5AB+Minor Areas 12\&20. Also shown are the indices from the binomial model of presence/absence in this fishery and the combined delta-lognormal model (Eq. C.4). All indices are scaled so that their geometric means equal 1.0. Upper and lower 95\% confidence bounds and associated standard error (SE) are presented for the lognormal model, while bootstrapped upper and lower 95\% confidence bounds are presented for the combined model.

| Year | Arithmetic Index (Eq. C.1) | Geometric Index (Eq. C.2) | Lognormal (Eq. C.3) |  |  |  | Binomial Index (Eq. C.3) | Combined (Eq. C.4) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Index | Lower bound | Upper bound | SE |  | Index | Lower bound | Upper bound |
| 1996 | 2.308 | 1.415 | 1.816 | 1.497 | 2.202 | 0.0985 | 0.399 | 0.753 | 0.548 | 1.021 |
| 1997 | 1.573 | 2.206 | 2.005 | 1.817 | 2.213 | 0.0503 | 1.424 | 2.803 | 2.427 | 3.105 |
| 1998 | 0.760 | 1.247 | 1.305 | 1.183 | 1.440 | 0.0500 | 1.118 | 1.457 | 1.297 | 1.614 |
| 1999 | 0.516 | 0.907 | 1.058 | 0.963 | 1.163 | 0.0483 | 1.074 | 1.138 | 1.008 | 1.260 |
| 2000 | 0.464 | 0.849 | 1.072 | 0.972 | 1.183 | 0.0500 | 0.833 | 0.906 | 0.811 | 1.028 |
| 2001 | 1.413 | 0.904 | 0.722 | 0.645 | 0.809 | 0.0577 | 0.684 | 0.506 | 0.439 | 0.577 |
| 2002 | 0.695 | 0.880 | 0.704 | 0.634 | 0.782 | 0.0533 | 0.693 | 0.499 | 0.439 | 0.570 |
| 2003 | 0.698 | 1.063 | 0.951 | 0.858 | 1.055 | 0.0528 | 0.813 | 0.785 | 0.691 | 0.896 |
| 2004 | 0.552 | 0.790 | 0.771 | 0.704 | 0.846 | 0.0470 | 1.058 | 0.817 | 0.723 | 0.910 |
| 2005 | 0.939 | 1.251 | 0.990 | 0.904 | 1.083 | 0.0460 | 1.166 | 1.149 | 1.025 | 1.267 |
| 2006 | 0.670 | 0.822 | 0.858 | 0.779 | 0.945 | 0.0491 | 0.958 | 0.829 | 0.737 | 0.924 |
| 2007 | 1.147 | 0.806 | 0.822 | 0.742 | 0.910 | 0.0521 | 1.005 | 0.830 | 0.728 | 0.945 |
| 2008 | 0.266 | 0.512 | 0.579 | 0.507 | 0.662 | 0.0681 | 0.638 | 0.379 | 0.327 | 0.451 |
| 2009 | 0.571 | 0.663 | 0.656 | 0.582 | 0.739 | 0.0609 | 0.710 | 0.476 | 0.406 | 0.546 |
| 2010 | 0.934 | 0.928 | 0.922 | 0.827 | 1.028 | 0.0554 | 0.776 | 0.728 | 0.639 | 0.829 |
| 2011 | 1.649 | 1.305 | 1.204 | 1.080 | 1.341 | 0.0552 | 1.085 | 1.306 | 1.163 | 1.511 |
| 2012 | 1.135 | 0.901 | 0.851 | 0.761 | 0.952 | 0.0572 | 1.305 | 1.097 | 0.962 | 1.258 |
| 2013 | 1.123 | 1.034 | 1.076 | 0.958 | 1.207 | 0.0588 | 1.361 | 1.441 | 1.256 | 1.680 |
| 2014 | 7.772 | 1.377 | 1.348 | 1.207 | 1.506 | 0.0565 | 2.010 | 2.574 | 2.194 | 2.989 |
| 2015 | 2.401 | 1.139 | 1.369 | 1.241 | 1.509 | 0.0500 | 2.833 | 3.525 | 3.075 | 3.945 |

## C.7. REFERENCS - CPUE

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## APPENDIX D. BIOLOGICAL DATA

## D.1. GROWTH AND MATURITY

This appendix describes the derivation of the length-weight relationship, von Bertalanffy growth relationship, maturity schedule, natural mortality and Walford parameters used in the Walleye Pollock (WAP) delay-difference stock assessment model. These analyses are based on Walleye Pollock biological data extracted from the Fisheries and Oceans Canada (DFO) Groundfish database "GFBioSQL" on Dec 13, 2016 (345,429 records). General data selection criteria for many analyses are summarized in Table D.1, though each analysis can vary the selection.

Table D.1. Data selection criteria for analyses of Walleye Pollock biological data for growth and lengthweight analysis.

| Field | Criterion | Notes |
| :---: | :---: | :---: |
| Trip type | [trip_type] == c(2, 3) | Definition of research observations. |
|  | [trip_type] $==\mathrm{c}(1,4,5)$ | Definition of commercial observations |
| Sample type | [sample_type] $==\mathrm{c}(1,2,6,7)$ | Only random or total samples. |
| Ageing method | [agemeth]==3 \| (==0 \& [year]>=1980) or $==7$ (fin rays) | Break \& burn method, or unknown from 1980 onwards, assumed to be B\&B. |
| Species category code | [SPECIES_CATEGORY_CODE] == 1 (or 3) | 1 = Unsorted samples <br> 3 = Sorted (keeper) samples |
| Sex code | [sex] $==\mathrm{c}(1,2)$ | Clearly identified sex <br> (1=male or 2=female). |
|  |  | North (5CDE): |
|  |  | PMFC major area codes 7:9 |
| Area code | [stock] select valid stock area observations | South (5AB3CD): <br> PMFC major areas 5:6 (5AB) + minor 12 <br> (Queen Charlotte Strait) + majors 3:4 <br> (3CD) + minor 20 (Juan de Fuca Strait) |
| Tow status | select [Not_available_reason_code] == NULL | Not rejected, valid tow. |

## D.1.1. Length-Weight

A log-linear relationship with additive errors was fitted to females, males, and combined to all valid weight and length data pairs $i,\left\{W_{i}, L_{i}\right\}$ :

$$
\begin{equation*}
\ln \left(W_{i}\right)=b \ln \left(L_{i}\right)+a+\varepsilon \tag{D.1}
\end{equation*}
$$

where $a$ and $b$ are the intercept and slope parameters.
Samples from gear types bottom trawl (Table D.2) and midwater trawl (Table D.3) were analysed separately, but not used in the assessment. As the delay-difference model was not sex-specific, males and females were combined for allometric relationships for stocks North (Figure D.1), South (Figure D.2), and coastwide (Figure D.3).

Table D.2. Length-weight parameter estimates, standard errors (SE) and number of observations (n) for Walleye Pollock (females, males and combined) for all research/charter and commercial samples from bottom trawls operating in stock areas $5 C D E, 5 A B, 3 C D$, and $4 B$ from 1973 to $2015 . \bar{W}_{\text {pred }}=$ mean weight (in kg ) from the fitted data set.

| Stock | $n$ | $\ln (\mathrm{a})$ | $\begin{gathered} \mathrm{SE} \\ \ln (a) \end{gathered}$ | b | SE(b) | $\bar{W}_{i}$ | $\begin{gathered} \hline \mathbf{S D} \\ W_{i} \end{gathered}$ | $\begin{array}{r} \min \\ W_{i} \end{array}$ | $\begin{gathered} \max \\ W_{i} \end{gathered}$ | $\bar{W}_{\text {pred }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $5 C D E+5 A B+3 C D$ (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 5,209 | -11.782 | 0.0178 | 3.0153 | 0.0050 | 0.5257 | 0.4457 | 0.016 | 2.744 | 0.6322 |
| Males | 3,363 | -11.841 | 0.0227 | 3.0333 | 0.0067 | 0.3167 | 0.3065 | 0.016 | 1.927 | 0.4216 |
| F+M | 8,575 | -11.801 | 0.0135 | 3.0210 | 0.0039 | 0.4437 | 0.4099 | 0.016 | 2.744 | 0.5433 |
| $5 C D E+5 A B+3 C D$ (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 253 | -10.359 | 0.2081 | 2.6726 | 0.0534 | 1.1168 | 0.3703 | 0.447 | 2.052 | 1.4237 |
| Males | 258 | -10.373 | 0.2130 | 2.6741 | 0.0557 | 0.8995 | 0.2524 | 0.249 | 1.745 | 1.2031 |
| F+M | 511 | -10.407 | 0.1413 | 2.6839 | 0.0366 | 1.0071 | 0.3342 | 0.249 | 2.052 | 1.3461 |
| 5CDE (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 2,589 | -11.691 | 0.0217 | 2.9889 | 0.0061 | 0.5689 | 0.5011 | 0.027 | 2.418 | 0.6395 |
| Males | 1,491 | -11.852 | 0.0286 | 3.0362 | 0.0084 | 0.3565 | 0.3682 | 0.023 | 1.927 | 0.4820 |
| F+M | 4,080 | -11.745 | 0.0168 | 3.0043 | 0.0048 | 0.4913 | 0.4683 | 0.023 | 2.418 | 0.5717 |
| 5CDE (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 181 | -9.835 | 0.2172 | 2.5323 | 0.0562 | 1.0037 | 0.3423 | 0.447 | 2.052 | 1.3805 |
| Males | 214 | -10.221 | 0.2159 | 2.6322 | 0.0567 | 0.8569 | 0.2427 | 0.249 | 1.745 | 1.2024 |
| F+M | 398 | -9.981 | 0.1518 | 2.5695 | 0.0396 | 0.9277 | 0.3072 | 0.249 | 2.052 | 1.3139 |
| $5 A B$ (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 1,531 | -12.010 | 0.0304 | 3.0881 | 0.0083 | 0.6254 | 0.4087 | 0.046 | 2.364 | 0.9322 |
| Males | 813 | -12.214 | 0.0369 | 3.1484 | 0.0107 | 0.3738 | 0.3183 | 0.024 | 1.446 | 0.6126 |
| F+M | 2,345 | -12.073 | 0.0225 | 3.1059 | 0.0063 | 0.5380 | 0.3982 | 0.024 | 2.364 | 0.8158 |
| 5AB (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 69 | -10.176 | 0.7024 | 2.6399 | 0.1766 | 1.4060 | 0.2529 | 0.737 | 2.041 | 1.5211 |
| Males | 43 | -10.681 | 1.1595 | 2.7638 | 0.2977 | 1.1056 | 0.1914 | 0.709 | 1.559 | 1.1787 |
| F+M | 113 | -10.392 | 0.5231 | 2.6917 | 0.1325 | 1.2896 | 0.2721 | 0.709 | 2.041 | 1.4182 |
| 3CD (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 1,093 | -11.204 | 0.0553 | 2.8332 | 0.0161 | 0.2812 | 0.2002 | 0.016 | 2.744 | 0.3087 |
| Males | 1,053 | -11.125 | 0.0564 | 2.8144 | 0.0168 | 0.2162 | 0.1217 | 0.016 | 0.82 | 0.1940 |
| F+M | 2,144 | -11.137 | 0.0389 | 2.8159 | 0.0115 | 0.2495 | 0.1695 | 0.016 | 2.744 | 0.2554 |
| 4B (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 413 | -11.911 | 0.0509 | 3.0228 | 0.0150 | 0.3189 | 0.2657 | 0.009 | 1.207 | 0.4752 |
| Males | 353 | -12.046 | 0.0463 | 3.0729 | 0.0141 | 0.2097 | 0.1697 | 0.008 | 1.165 | 0.3412 |
| F+M | 769 | -11.950 | 0.0355 | 3.0387 | 0.0106 | 0.2680 | 0.2327 | 0.008 | 1.207 | 0.4158 |
| 4B (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 150 | -12.653 | 0.1895 | 3.2713 | 0.0502 | 0.8797 | 0.5006 | 0.16 | 1.945 | 0.9751 |
| Males | 45 | -13.105 | 0.3022 | 3.3963 | 0.0844 | 0.4700 | 0.3200 | 0.175 | 1.3 | 0.6547 |
| F+M | 195 | -12.743 | 0.1495 | 3.2952 | 0.0401 | 0.7852 | 0.4956 | 0.16 | 1.945 | 0.9019 |

Table D.3. Length-weight parameter estimates, standard errors (SE) and number of observations (n) for Walleye Pollock (females, males and combined) for all research/charter and commercial samples from midwater trawls operating in stock areas 5CDE, 5AB, 3CD, and 4B from 1978 to 2015. $\bar{W}_{\text {pred }}=$ mean weight (in kg) from the fitted data set.

| Stock | $n$ | $\ln (\mathrm{a})$ | $\begin{gathered} \text { SE } \\ \ln (a) \end{gathered}$ | b | SE(b) | $\bar{W}_{i}$ | $\begin{gathered} \hline \mathbf{S D} \\ W_{i} \end{gathered}$ | $\begin{array}{r} \min \\ W_{i} \end{array}$ | $\begin{array}{r} \max \\ W_{i} \end{array}$ | $\bar{W}_{\text {pred }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $5 C D E+5 A B+3 C D$ (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 2,021 | -11.563 | 0.0394 | 2.9714 | 0.0103 | 0.9966 | 0.6098 | 0.027 | 4.654 | 0.8456 |
| Males | 1,671 | -11.658 | 0.0411 | 3.0001 | 0.0110 | 0.8069 | 0.4869 | 0.046 | 2.875 | 0.6693 |
| F+M | 3,694 | -11.592 | 0.0283 | 2.9806 | 0.0075 | 0.9106 | 0.5654 | 0.027 | 4.654 | 0.7644 |
| $5 C D E+5 A B+3 C D$ (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 1,407 | -11.934 | 0.0657 | 3.0738 | 0.0173 | 0.9101 | 0.4903 | 0.142 | 2.709 | 1.2227 |
| Males | 1,274 | -11.777 | 0.0625 | 3.0400 | 0.0167 | 0.7874 | 0.3969 | 0.142 | 2.063 | 1.0096 |
| F+M | 2,675 | -11.819 | 0.0449 | 3.0476 | 0.0119 | 0.8526 | 0.4526 | 0.142 | 2.709 | 1.1339 |


| Stock | $n$ | $\ln (\mathrm{a})$ | $\begin{gathered} \text { SE } \\ \ln (a) \end{gathered}$ | b | SE(b) | $\bar{W}_{i}$ | $\begin{gathered} \mathbf{S D} \\ W_{i} \end{gathered}$ | min <br> $W_{i}$ | max <br> $W_{i}$ | $\bar{W}_{\text {pred }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 5CDE (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 1,567 | -11.538 | 0.0512 | 2.9617 | 0.0132 | 1.1031 | 0.5640 | 0.031 | 2.667 | 1.0320 |
| Males | 1,322 | -11.735 | 0.0488 | 3.0165 | 0.0129 | 0.8909 | 0.4750 | 0.046 | 2.875 | 0.7942 |
| F+M | 2,888 | -11.618 | 0.0350 | 2.9838 | 0.0091 | 1.0056 | 0.5352 | 0.031 | 2.875 | 0.9149 |
| 5CDE (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 479 | -10.037 | 0.1618 | 2.5916 | 0.0410 | 1.2674 | 0.3891 | 0.25 | 2.504 | 1.3768 |
| Males | 416 | -10.033 | 0.1780 | 2.5898 | 0.0455 | 1.1515 | 0.3176 | 0.247 | 2.012 | 1.1830 |
| F+M | 895 | -10.042 | 0.1184 | 2.5926 | 0.0301 | 1.2135 | 0.3621 | 0.247 | 2.504 | 1.2981 |
| 5AB (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 76 | -12.620 | 0.3905 | 3.2656 | 0.0985 | 1.5757 | 0.8348 | 0.38 | 4.654 | 0.9977 |
| Males | 39 | -12.903 | 0.8839 | 3.3392 | 0.2269 | 1.1830 | 0.4557 | 0.484 | 2.313 | 0.6349 |
| F+M | 115 | -12.659 | 0.3491 | 3.2759 | 0.0886 | 1.4425 | 0.7500 | 0.38 | 4.654 | 0.8401 |
| 5AB (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 551 | -12.236 | 0.1151 | 3.1656 | 0.0301 | 0.9524 | 0.4240 | 0.142 | 2.709 | 1.0045 |
| Males | 588 | -12.313 | 0.1209 | 3.1936 | 0.0323 | 0.7386 | 0.2755 | 0.142 | 2.063 | 0.7450 |
| F+M | 1,139 | -12.158 | 0.0798 | 3.1488 | 0.0211 | 0.8417 | 0.3710 | 0.142 | 2.709 | 0.8745 |
| 3CD (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 380 | -12.175 | 0.0767 | 3.1530 | 0.0217 | 0.4437 | 0.3372 | 0.027 | 1.486 | 0.6422 |
| Males | 311 | -12.479 | 0.0886 | 3.2508 | 0.0253 | 0.4037 | 0.2816 | 0.081 | 1.262 | 0.5475 |
| F+M | 691 | -12.268 | 0.0594 | 3.1842 | 0.0169 | 0.4257 | 0.3138 | 0.027 | 1.486 | 0.6025 |
| 3CD (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 378 | -9.850 | 0.1598 | 2.4750 | 0.0445 | 0.3944 | 0.1162 | 0.181 | 0.708 | 0.3754 |
| Males | 269 | -10.305 | 0.1602 | 2.6065 | 0.0455 | 0.3351 | 0.0937 | 0.16 | 0.593 | 0.3365 |
| F+M | 645 | -10.014 | 0.1095 | 2.5224 | 0.0308 | 0.3697 | 0.1113 | 0.16 | 0.708 | 0.3602 |
| 4B (Research Survey) |  |  |  |  |  |  |  |  |  |  |
| Females | 2,687 | -10.643 | 0.0556 | 2.6923 | 0.0155 | 0.4183 | 0.2110 | 0.041 | 1.905 | 0.5454 |
| Males | 3,289 | -10.926 | 0.0486 | 2.7655 | 0.0138 | 0.3283 | 0.1490 | 0.034 | 0.991 | 0.4533 |
| F+M | 5,979 | -10.831 | 0.0360 | 2.7414 | 0.0101 | 0.3684 | 0.1842 | 0.015 | 1.905 | 0.4886 |
| 4 B (Commercial) |  |  |  |  |  |  |  |  |  |  |
| Females | 542 | -11.841 | 0.1569 | 3.0505 | 0.0431 | 0.5021 | 0.1728 | 0.138 | 1.47 | 0.6961 |
| Males | 809 | -11.949 | 0.1193 | 3.0725 | 0.0337 | 0.3644 | 0.1222 | 0.118 | 0.87 | 0.5436 |
| F+M | 1,355 | -12.043 | 0.0916 | 3.1017 | 0.0256 | 0.4193 | 0.1596 | 0.118 | 1.47 | 0.6173 |



Figure D.1. Length-weight relationship for the North (5CDE) stock of Walleye Pollock - derived from randomly selected research survey samples, regardless of gear type. Records with absolute value of standardised residuals >3 (starting with a preliminary fit) were dropped, removing 18 observations for the combined-sex fit.


Figure D.2. Length-weight relationship for the South (5AB3CD) stock of Walleye Pollock - derived from randomly selected research survey samples, regardless of gear type. Records with absolute value of standardised residuals >3 (starting with a preliminary fit) were dropped, removing 9 observations for the combined-sex fit.


Figure D.3. Length-weight relationship for the combined North and South stocks of Walleye Pollock derived from randomly selected research survey samples, regardless of gear type. Records with absolute value of standardised residuals >3 (starting with a preliminary fit) were dropped, removing 28 observations for the combined-sex fit.

For the assessment model input, the following procedure was followed to combine the lengthweight information by sex into an interpolated unsexed length-weight relationship. This approach was used because it was felt that the sex ratio in the samples used to estimate the functions would not be representative of the population and that it was better to give equal weight to the length-weight model for each sex.

1. A combined sex weight $W_{i}$ was calculated for every length $i$ and sex $m$ or $f$, using the sex-specific length-weight parameters calculated using (D.1) and (D.2), weighting equally between the estimated weight for each sex.
2. Parameters $a$ and $b$ were estimated such that the least-squares sum (D.3) between the averaged weight and an estimated weight was minimised.

The resulting function is the average of the sex-specific length-weight functions (Figure D.4).

$$
\begin{align*}
& W_{i}=0.5\left(a_{m} L_{i}^{b_{m}}\right)+0.5\left(a_{f} L_{i}^{b_{f}}\right) \\
& \sum_{i=10}^{100}\left(W_{i}-a L_{i}^{b}\right)^{2} \tag{D.3}
\end{align*}
$$

The allometric parameters $(a, b)$ calculated using (D.2) were close to those estimated in Figure D. 1 for BC North and Figure D. 2 for BC South.

$$
W_{i}=e^{a} L_{i}^{b}\left\{\begin{array}{l}
\text { North: } a=-11.84492, b=3.03728 \\
\text { South: } a=-11.78992, b=3.02122
\end{array}\right.
$$



Figure D.4. Interpolated combined sex length-weight models used to estimate the Walford parameters used in the WAP delay-difference stock assessment model. [left panel]: North area model; [right panel]: South area model.

## D.1.2. von Bertalanffy Growth

Paired observations $i$ of length and age by sex, $\left\{L_{i s}, a_{i s}\right\}$, for $s=1,2$ (males, females) were available using the break and burn method (MacLellan 1997) for only 17 specimens from commercial trips ( 16 otoliths in 5CDE and one otolith in 5AB). Otoliths aged by surface readings were more abundant (commercial: 230 otoliths in 5CDE and 399 otoliths in 5AB). The remaining samples come from research surveys but are aged using pectoral fin rays or an unknown method. Unfortunately, pectoral fin ageing is thought to be biased, especially at older ages (MacLellan et al. 1990), because fin ray deposition slows down or ceases at ages greater than ten. A summary of available age data in the GFBioSQL database by stock and ageing method appears in Table D.4. Also note the low number of samples from which these otoliths were collected.

Table D.4. Walleye Pollock number of ages available in GFBioSQL database by ageing method (accessed 2016-12-13). Number of samples appear in parentheses.

| Stock | Otolith (surface) |  |  |  | Otolith (B\&B) |  |  |  | Pectoral fin |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Female | Male | Total | Female | Male | Total | Female | Male | Total | Female | Male | Total |
| 5CDE | $165(5)$ | $65(5)$ | $230(5)$ | $14(2)$ | $2(1)$ | $16(2)$ | 0 | 0 | 0 | 0 | 0 | 0 |
| 5AB | $210(7)$ | $189(7)$ | $399(7)$ | 0 | $1(1)$ | $1(1)$ | $305(1)$ | $144(1)$ | $449(1)$ | 0 | 0 | 0 |
| 3CD | 0 | 0 | 0 | 0 | 0 | 0 | $100(1)$ | $89(1)$ | $189(1)$ | $123(2)$ | $87(2)$ | $210(2)$ |
| 4B | 0 | 0 | 0 | 0 | 0 | 0 | $87(2)$ | $146(2)$ | $233(2)$ | $36(2)$ | $314(2)$ | $350(2)$ |

Growth was formulated as a von Bertalanffy model where lengths by sex, $L_{i s}$, for fish $i=1, \ldots, n_{s}$ are given by:

$$
\begin{equation*}
L_{i s}=L_{\infty s}\left[1-e^{-\kappa_{s}\left(a_{i s}-t_{0 S}\right)}\right]+\varepsilon_{s}, \quad \varepsilon_{s} \square N\left(0, \sigma_{s}^{2}\right) \tag{D.4}
\end{equation*}
$$

where for each sex $s$,
$L_{\infty s}=$ the average length at maximum age of an individual,
$\kappa_{s}=$ growth rate coefficient, and
$t_{0 s}=$ age at which the average size is zero.
The negative log likelihood for each sex $s$, used for minimisation is:

$$
\ell\left(L_{\infty}, \kappa, t_{0}, \sigma\right)=n \ln (\sigma)+\frac{\sum_{i}^{n}\left(L_{i}-E_{i}\right)^{2}}{2 \sigma^{2}}, \quad i=1, \ldots, n
$$

The limited DFO data (Table D.4) did not yield satisfactory growth curves (Figure D. 5 to Figure D.7) and were presumably biased by the ageing methods, which cannot resolve older ages. Therefore, we requested help from an Alaskan colleague, Martin Dorn (Research Fish Biologist, NOAA Fisheries, Sand Point, Seattle), who supplied us with 8,882 age-length pairs randomly selected from six biannual surveys conducted in the Gulf of Alaska (GoA) between 2005 and 2015. These fish had all been aged from otoliths prepared using the "break \& burn" method and he advised us to use the specimens from the Eastern GoA only ( $\mathrm{n}=847$ ) as growth varied across the GoA. We used these data to estimate a growth function for use in this BC Pollock stock assessment (Figure D.8). The resulting function adequately fit the mean weight data for the BC North stock for all three of the credible knife-edge age at full selectivity assumptions (see all three panels in Figure D.9). However, we could not use this function for the BC South stock because fish sampled from Dixon Entrance were, on average, twice as large as those sampled from more southern BC waters. This North stock is likely to belong to a larger stock that includes Dixon Entrance, northern Hecate Strait, and waters off of Southeast Alaska (Gustafson et al. 2000). Consequently, we decided to divide BC Pollock into two stocks, with the North stock encompassing Dixon Entrance and upper Hecate Strait and the South stock including all Pollock from Moresby Gully to the US-Canada border.




Figure D.5. von Bertalanffy fits to BC Walleye Pollock ages determined by surface-read otoliths.


Figure D.6. von Bertalanffy fits to BC Walleye Pollock ages determined by pectoral fin rays.


Figure D.7. von Bertalanffy fits to BC Walleye Pollock ages determined by unknown methods.


Figure D.8. von Bertalanffy fits to eastern Gulf of Alaska Walleye Pollock ages determined by brokenburnt otoliths.

For the BC South stock, the eastern GOA growth function did not provide a satisfactory fit to the mean weight data (either with knife-edge age=3 - see panel [a] Figure D. 10 or knife-edge age=4, see panel [a] Figure D.11, so we turned to other sources (Table D.1). Growth functions based on fin ray ages published by Saunders et al. (1989) for the west coast of Vancouver Island (3CD) and the Strait of Georgia (apparently derived from age-length pairs not available in the general DFO database) were tested but featured growth rate coefficients ( $\kappa$ ) that were so steep ( 0.31 and 0.91 , respectively) that neither function could fit the South mean weight data satisfactorily when knife-edge age=4 (panels [b] and [c], Figure D.11. Even with knife-edge age $=3$, neither function gave good fits to the mean weight data (panels [b] and [c], Figure D. 10. We also enquired with NOAA colleagues who were working on pelagic fish off the west coasts of Washington and Oregon who were able to provide us with sampled lengths for Pollock from their fisheries, but were unable to provide us with age-length pairs. In the end, we found a growth function published by Janusz and Horbowy (1997) for Walleye Pollock in the Sea of Okhotsk ( $L_{\infty}=50.827, \kappa=0.199, t_{0}=-1.790$, ages from otoliths), which provided satisfactory fits to the observed South mean weight data for both knife-edge age=3 (panel [d],Figure D. 10) and knife-edge age=4 (panel [d], Figure D.11.) We have no reason to believe that the Sea of Okhotsk relationship represents BC South other than the estimated growth was consistent with our mean weight data. We also noted that some authors have suggested similarities between Pollock populations in disconnected regions. For example, Thompson (1981) noted that populations of Pollock from the Strait of Georgia were more similar to those from the Bering Sea than they were to those in Dixon Entrance.

The following procedure was followed to combine the von-Bertalanffy growth model by sex into an interpolated unsexed growth model. This approach was used because it was felt that the sex ratio in the samples used to estimate the growth functions would not be representative of the population and that it was better to give equal weight to the growth model for each sex.

1. A combined sex length $L_{i}$ is calculated for every age $i$ and sex $m$ or $f$, using the sexspecific growth model parameters calculated using (D.4) and (D.5), weighting equally between the estimated length for each sex.
2. Parameters $L_{\infty}, \kappa, t_{0}$ are estimated such that the least-squares sum (D.6) between the averaged length and an estimated length is minimised.

The resulting function is the average of the sex-specific growth functions (Figure D.12). There is no evidence that these data show changes in size at age over time (Table D.5).

$$
\begin{align*}
& L_{i}=0.5\left(L_{\infty m} e^{\left(-\kappa_{m}\left[a_{i m}-t_{0 m}\right]\right)}\right)+0.5\left(L_{\infty f} e^{\left(-\kappa_{f}\left[a_{i f}-t_{0 f}\right]\right)}\right)  \tag{D.5}\\
& \sum_{i=1}^{50}\left(L_{i}-L_{\infty} e^{\left(-\kappa\left[a_{i}-t_{0}\right]\right)}\right)^{2} \quad(\mathrm{D} .6) \tag{D.6}
\end{align*}
$$





Figure D.9. Comparison of fits to the mean weight data for the BC North stock (Section D.2.2) under three knife-edge age assumptions. All three panels use the Eastern GoA growth function (Table D.6) and assume $M=0.3$. NLL=negative log likelihood provided for each fit; (a): knife-edge age $=3$ ( $N L L=0.059$ ); (b) knife-edge age=4 (NLL=-5.97); (c) knife-edge age=5 ( $N L L=2.95$ ).


Figure D.10. Comparison of fits to the mean weight data for the BC South stock (Section D.2.2) under four growth rate assumptions (Table D.6), assuming the same age $=3$ at knife-edge selectivity and $M=0.3$. NLL=negative log likelihood provided for each fit; (a): E. GoA (NLL=50.5); (b) WCVI (NLL=10.0); (c) SG (NLL=17.3); (d) OS (NLL=10.8).


Figure D.11. Comparison of fits to the mean weight data for the BC South stock (Section D.2.2) under four growth rate assumptions (Table D.6), assuming the same age $=4$ at knife-edge selectivity and $M=0.3$. NLL=negative log likelihood provided for each fit; (a): E. GoA (NLL=204); (b) WCVI (NLL=76.1); (c) SG ( $N L L=58.9$ ); (d) $O S$ ( $N L L=0.58$ ).


Figure D.12. Interpolated combined sex growth models used to estimate the Walford parameters used in the WAP delay-difference stock assessment model. [left panel]: Eastern Gulf of Alaska model; [right panel]: Okhotsk Sea model.

Table D.5. Mean length (cm) of Walleye Pollock, based on number of fish in brackets, by age and year for survey data from the Eastern Gulf of Alaska (Martin Dorn, pers. comm.).

| Age | 2005 | 2007 | 2009 | 2011 | 2013 |  | 2015 |  | Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 5.0 [1] | 7.0 [6] | --- | --- | --- | --- | 8.0 | [2] | 7.0 | [9] |
| 1 | 19.5[190] | 18.8 [40] | 19.8 [97] | 19.6 [99] | 18.6 | [103] | 19.3 | [28] | 19.4 | [557] |
| 2 | 30.7 [51] | 31.9 [28] | 33.1 [62] | 31.0 [56] | 33.3 | [8] | 32.1 | [24] | 31.8 | [229] |
| 3 | 39.7 [202] | 38.1 [73] | 42.3 [96] | 40.6 [57] | 40.8 | [39] | 38.3 | [100] | 39.9 | [567] |
| 4 | 42.9 [128] | 41.5 [36] | 48.6 [38] | 47.3 [29] | 47.7 | [41] | 42.7 | [26] | 44.5 | [298] |
| 5 | 45.1 [108] | 48.1 [16] | 53.9 [32] | 48.9 [52] | 51.6 | [30] | 50.1 | [21] | 48.3 | [259] |
| 6 | 49.8 [29] | 50.4 [10] | 55.0 [14] | 56.4 [19] | 55.2 | [14] | 54.2 | [26] | 53.3 | [112] |
| 7 | 52.8 [11] | 53.7 [7] | 57.5 [8] | 59.6 [13] | 56.0 | [5] | 55.1 | [8] | 56.0 | [52] |
| 8 | 56.3 [8] | 51.0 [1] | 58.0 [4] | --- | 57.6 | [7] | 58.3 | [8] | 57.2 | [28] |
| 9 | 52.0 [1] | --- | 57.7 [3] | 66.0 [1] | --- | --- | --- | --- | 58.2 | [5] |
| 10 | 65.0 [2] | --- | --- | --- | --- | --- | --- | --- | 65.0 | [2] |

Table D.6. Walleye Pollock growth parameters by sex from von Bertalanffy fits (D.4); estimated parameters from a combined-sex, interpolated model (D.5) for the non-DFO sources were used in the delay-difference model (shaded). DFO parameters for the coastwide population are included for comparison only. E.GoA = eastern Gulf of Alaska, WCVI = west coast Vancouver Is., SG = Strait of Georgia, OS = Okhotsk Sea, DFO =Dept. Fisheries \& Oceans Canada.

| Source | Assess | Ageing Method | Females |  |  | Males |  |  | Combined |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | $L_{\infty}$ | K | to | $L_{\infty}$ | K | to | $L_{\infty}$ | K | to |
| oA (Dorn pers.comm.) |  | Otoliths (B\&B) | 1.221 | 0.192 | -1.357 | 61.66 | 0.226 | -1.55 | 66.94 | 0.2 | . 13 |
| WCVI (Saunders et al. 1989) | So | Pectoral fins | 56.500 | 0.30 | -0.960 | 50.5 | 0.320 | -0.970 | 53.499 | 0.309 | -0.968 |
| SG (Sauders et al. 1989) | South | Pectoral fins | 46.500 | 0.835 | 0.559 | 42.500 | 0.997 | 0.590 | . 49 | 905 | 2 |
| OS (Janusz \& Horbowy (1997) | South | Otoliths | 53.300 | 0.177 | -1.930 | 48.40 | 0.23 | -1.560 | 50.8 | . 19 | -1.790 |
| DFO (GFBioSQL 2017-12-13) | Coast | Pectoral fins | 69.804 | 0.243 | -0.806 | 65.135 | 0.201 | -1.608 | 73.97 | . 18 | -1.431 |
| DFO (GFBioSQL 2017-12-13) | Coast | Otoliths (sfc) | 82.603 | 0.1 | 2.000 | 74.1 | 0.137 | 1.990 | 81.4 | 0.13 | -1.818 |

## D.1.3. Maturity

Maturity data for Walleye Pollock were obtained from GFBioSQL on Dec 13, 2016. Ages were scarce for this species and so data filtering was minimal for the maturity analysis. The following summary characterises the distribution of the 977 maturity data records used:

- stock - 3CD (385), 5AB (497), 5CDE (95)
- sex - males (405), females (572)
- trip type - non-observed commercial (394), charter (583)
- sample type - total catch (98), random(879)
- sampled catch - unsorted (583), sorted labelled as "keepers" (394)
- ageing method - unknown (210), surface otoliths (382), broken-burnt otoliths (12), pectoral fins (373)
- maturity codes - 1 (75), 2 (188), 3 (69), 4 (218) 5 (12), 6 (6), 7 (28), 10 (1), 12 (380)

In the GFBioSQL database, Walleye Pollock uses maturity convention 25 (Hake-Pollock, 7-stages), but may also have been using maturity convention 2 (Hake 1977+, 12-stages) as maturity codes 10 and 12 appeared. For the analysis, all stages 3 and higher were assumed to be mature, and a maturity ogive (Figure D.14) was fit to otolith and pectoral fin data using a double-normal model:

$$
m_{a s}=\left\{\begin{array}{cc}
e^{-\left(a-v_{s}\right)^{2} / \rho_{s t}}, & a \leq v_{s}  \tag{D.7}\\
1, & a>v_{s}
\end{array}\right.
$$

where, $m_{a s}=$ maturity at age $a$ for sex $s$ (combined),
$v_{s}=$ age of full maturity for sex $s$,
$\rho_{s L}=$ variance for the left limb of the maturity curve for sex $s$.
The ages at $50 \%$ and full maturity are estimated at 3.6 y and 4.6 y , respectively, for otolith-aged fish and 2.4 y and 3.4 y , respectively, for fin-aged fish (Figure D.13). There appears to be a consistent difference between these two ageing methods, although the fit to these sparse data can be influenced by the initial values offered to the minimization of (D.7).
All commercial data comprise sorted fish and all charter data comprise unsorted fish (see maturity data composition bullets above). Comparing maturity ogives for sorted and unsorted catches (Figure D.14) unfortunately means lumping ages determined by surface-read otoliths and broken-burnt otoliths for the sorted ogive and pectoral fins and unknown methods for the unsorted ogive. The results are presented for comparison only.
The maturity schedules presented here (Table D.7) are not used in this assessment because the knife-edge selectivity assumption ( $k=3$ or 4 y ) used by the delay-difference model assumes that maturity matches selectivity, i.e., all recruited fish are mature. This analysis shows that the estimated age at maturity for Walleye Pollock is similar to that for knife-edge recruitment investigated in the population model.
Alternative maturity ogives by length are shown in Figure D.15. Length data are more abundant than age data. The estimated length at full maturity from unsorted catch samples is 56.6 cm for
the commercial fishery and 49.7 cm for research surveys. For sorted catch samples, the length at full maturity for the commercial fishery is 57.3 cm (no sorted samples from surveys).


Figure D.13. Maturity-at-age ogives by ageing method for BC Walleye Pollock (data from Dec 13, 2016, GFBioSQL) as double-normal fits using (D.7) for both sexes combined, where maturity is defined by stages $\geq 3$. Solid blue line shows fit to otolith data; blue dashed line indicates fit to pectoral fin data.


Figure D.14. Maturity-at-age ogives by catch sorting for BC Walleye Pollock (data from Dec 13, 2016, GFBioSQL) as double-normal fits using (D.7) for both sexes combined, where maturity is defined by stages $\geq 3$. Solid blue line shows fit to otolith data; blue dashed line indicates fit to pectoral fin data.

Table D.7. Proportion of Walleye Pollock mature at each age $\left(m_{a}\right)$ up to age 10y. Maturity stages 1 and 2 describe immature fish while stages 3 to 12 are considered mature. Model fits are presented for the binomial logit (BL, comparison only) and the double normal (DN).

| Age | \# Fish | Obs. $m_{a}$ | BL $m_{a}$ | DN $m_{a}$ | Age | \# Fish | Obs. $m_{a}$ | BL $m_{a}$ | DN $m_{a}$ |
| ---: | ---: | ---: | ---: | :---: | ---: | ---: | ---: | ---: | ---: |
|  |  | Otoliths |  |  |  |  | Pectoral |  |  |
| 1 | 2 | 0 | 0.07347 | 0.00004 | 1 | 0 | 0 | 0.01474 | 0.01388 |
| 2 | 23 | 0.08696 | 0.16830 | 0.00500 | 2 | 43 | 0.23256 | 0.24427 | 0.23256 |
| 3 | 154 | 0.12987 | 0.34053 | 0.13948 | 3 | 88 | 0.88636 | 0.87475 | 0.88636 |
| 4 | 139 | 0.77698 | 0.56854 | 0.77461 | 4 | 71 | 0.98592 | 0.99342 | 1 |
| 5 | 99 | 0.97980 | 0.77078 | 1 | 5 | 43 | 1 | 0.99969 | 1 |
| 6 | 107 | 0.91589 | 0.89563 | 1 | 6 | 58 | 1 | 0.99999 | 1 |
| 7 | 51 | 0.74510 | 0.95633 | 1 | 7 | 45 | 1 | 1.00000 | 1 |
| 8 | 19 | 0.78947 | 0.98242 | 1 | 8 | 17 | 1 | 1 | 1 |
| 9 | 9 | 0.66667 | 0.99304 | 1 | 9 | 5 | 1 | 1 | 1 |
| 10 | 1 | 1 | 0.99726 | 1 | 10 | 3 | 1 | 1 | 1 |
|  |  | Unsorted |  | 0 | 0.01658 | 0.02087 | 1 | 0 |  |
| 1 | 2 | 0.16949 | 0.10180 | 0.12429 | 2 | 7 | 0.28571 | 0.76385 | 0.28038 |
| 2 | 59 | 0.39367 | 0.43247 | 0.42829 | 3 | 21 | 0.52381 | 0.78915 | 0.53463 |
| 3 | 221 | 0.3936 |  |  |  |  |  |  |  |
| 4 | 102 | 0.87255 | 0.83669 | 0.85370 | 4 | 108 | 0.82407 | 0.81241 | 0.81302 |
| 5 | 58 | 1 | 0.97179 | 1 | 5 | 84 | 0.97619 | 0.83364 | 0.98603 |
| 6 | 67 | 1 | 0.99570 | 1 | 6 | 98 | 0.90816 | 0.85291 | 1 |
| 7 | 48 | 0.97917 | 0.99936 | 1 | 7 | 48 | 0.75000 | 0.87029 | 1 |
| 8 | 18 | 1 | 0.99990 | 1 | 8 | 18 | 0.77778 | 0.88589 | 1 |
| 9 | 5 | 1 | 0.99999 | 1 | 9 | 9 | 0.66667 | 0.89983 |  |
| 10 | 3 | 1 | 1.00000 | 1 | 10 | 1 | 1 | 0.91224 | 1 |



Figure D.15. Maturity-at-length ogives for BC Walleye Pollock: (left) unsorted samples from the commercial fishery (trip types 1,4,5 combined in blue) and research surveys (trip types 2,3 combined in green); (right) sorted samples from the commercial fishery (red). Curves depict double-normal fits using (D.7) for both sexes combined, where maturity is defined by stages $\geq 3$.

## D.1.4. Natural Mortality

Although the Alaskan fisheries stock assessments use age-specific mortality rates for Walleye Pollock, the underlying assumption is that $M=0.30$ for age at full maturity (Dorn et al. 2015). Age-specific $M$ values of $0.90,0.45$, and 0.30 for ages 1,2 and $3+$, respectively, have been used in Alaskan Eastern Bering Sea catch-age models since 1982 (lanelli et al. 2015). The delay-difference model used for the BC population assumes that maturity matches selectivity, i.e., all recruited fish are mature, and by extension, all mature fish have a natural mortality rate of 0.30 , which is the value adopted by this assessment. Sensitivity runs were made with $M=0.25$ and $M=0.35$ to bracket plausible values for this parameter.
In the DFO database GFBioSQL, the maximum age is 11 years for two female and two male specimens caught at depths between 106 and 381 m in PMFC area 5CD, specifically in Moresby Gully; however the mean age is $5.0 \mathrm{y}(\mathrm{n}=1494)$ and the 0.99 quantile is 9 y . McFarlane and Beamish (1990) reported a maximum age of 28 y from the Bering Sea using a burnt otolith section method. These authors generally found that the burnt otolith method produced higher estimates of age when comparing them to ages from fin rays (Beamish and McFarlane 1995, and see Figure D.13).
The current assessment does not use catch-at-age information as the data are insufficient and potentially biased by the ageing methodology. However, natural mortality ( $M$ ) can be estimated using Quinn and Deriso (1999, p.361) based on Hoenig (1983):

$$
M=-\ln (0.01) / t_{m} \quad \text { (D.8) }
$$

where, $t_{m}=$ maximum observed age reach by $1 \%$ of the population.
Using the maximum age observed in the DFO database $t_{m}=11 \mathrm{y}, M=0.419$, which seems high compared to $M=0.30$ adopted for use in the delay-difference model. However, using $t_{m}=28 \mathrm{y}$, natural mortality $M=0.164$, which seems low for this species.

Then et al. (2015) revisited various natural mortality estimators and recommended the use of an updated Hoenig estimator based on nonlinear least squares:

$$
M_{\mathrm{est}}=4.899 t_{\max }^{-0.916}
$$

where $t_{\max }=$ maximum age. For WAP with a maximum age of 11 y (GFBio) or 28 y (McFarlane and Beamish 1990), the updated Hoenig estimator suggests that $M=0.55$ or 0.23 , respectively.

## D.1.5. Knife-edge Selectivity and Walford Plot

Dorn et al. (2012) provide a range of selectivity ogives for the GoA fisheries and surveys, with the median age selected to these commercial fisheries ranging from age 3 to age 5 (see columns 5 to 7 in Table D.8). Based on the ogives in this table, ages 3 and 4 were selected as the most likely ages to use for the age of knife-edge recruitment in the WAP delay-difference model. Knife-edge selectivity at age 5 was also run as an additional sensitivity for both stock definitions. Growth and length-weight parameters appropriate were used to prepare Walford plots (Figure D.16) which provide the growth parameter values used as input to the WAP delaydifference model. The Walford parameters are calculated from the knife-edge recruitment age to 30 y for the Section D.1.2 growth model. The Walford parameters will vary slightly with changing age assumptions at knife-edge recruitment for both growth models. Table D. 9 presents the Walford parameters used in the stock assessment for both growth models along with the mean length and mean weight associated with each of the knife-edge age at recruitment assumptions. Equilibrium mean weights assuming $M=0.3$ are also presented for comparative purposes.

Table D.8. Table 1.17 taken from Dorn et al. (2012) showing various selectivity ogives estimated for Gulf of Alaska pollock fisheries and surveys.

Table I. 17 Estimated selectivity at age for Gulf of Alaska Pollock fisheries and survey. The fisheries and surveys were modeled using double logistic functions.
\(\left.$$
\begin{array}{rccccccccc}\hline \text { Age } & \begin{array}{c}\text { POP fishery } \\
\text { (1964-71) }\end{array} & \begin{array}{c}\text { Foreign } \\
\text { (1972-81) }\end{array} & \begin{array}{c}\text { Foreign and } \\
\text { JV (1982- } \\
\text { 1988) }\end{array} & & \begin{array}{c}\text { Domestic } \\
(1989-2000)\end{array} & \begin{array}{c}\text { Domestic } \\
(2001-2006)\end{array} & \begin{array}{c}\text { Recent } \\
\text { domestic } \\
(2007-2012)\end{array} & \begin{array}{c}\text { Acoustic } \\
\text { survey }\end{array} & \begin{array}{c}\text { Bottom trawl } \\
\text { survey }\end{array}
$$ <br>
\hline 1 \& 0.000 \& 0.002 \& 0.016 \& 0.004 \& 0.029 \& 0.030 \& 0.584 \& 0.358 <br>

bottom trawl\end{array}\right]\)|  |
| :--- |
| 2 |

Table D.9. Age varying biological parameters used in the WAP delay-difference stock assessment using two growth models.

|  | Age at knife-edge recruitment |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Parameter | GoA growth model (North) |  |  | Okhotsk Sea growth model (South) |
|  | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{3}$ | $\mathbf{4}$ |
| $\alpha_{g}$ | 0.347 | 0.372 | 0.144 | 0.153 |
| $\rho_{g}$ | 0.867 | 0.856 | 0.871 | 0.861 |
| length $(\mathrm{cm})$ at $W_{k}$ | 39.1 | 44.4 | 31.2 | 34.8 |
| $W_{k}(\mathrm{~kg})$ | 0.493 | 0.727 | 0.249 | 0.344 |
| $\bar{W}_{0}(\mathrm{~kg})^{1}$ | 1.076 | 1.269 | 0.483 | 0.561 |

[^7]

Figure D.16. Walford plot for WAP using age=4y as the knife-edge recruitment assumption: [left panel]: Eastern GoA growth model (North); [right panel]: Okhotsk Sea growth model (South). Plotted points are the estimated weight-at-age from the growth model and the line is the fitted Walford plot to the points.

## D.2. MEAN WEIGHT

Data used to estimate the mean weight by year for this stock assessment were selected following the relevant guidelines in Table D.1. The biological data for WAP (downloaded from the GFBio database, Dec 13, 2016) yielded 345,429 records which were filtered as follows:

- year $=1967: 2016 \quad$ \{modern\} 344,377 records
- stock = c("3CD","5AB","5CDE")
- trip type $=c(1,4,5)$
- sample type $=c(1,2,6,7)$
- gear type $=c(1,6)$
- $\quad$ spp. category $=1$
\{modern \{oast
\{comm. incl. JV Hake\}
\{random\}
218,136 records
102,254 records
100,596 records
\{unsorted\}
98,292 records
50,998 records
or
- spp. category $=3 \quad$ \{keepers $=$ sorted $\} \quad 46,769$ records

This process resulted in 50,998 biological records for unsorted samples, all containing length data but only 266 records with weight data, and 46,769 records for sorted (keeper) samples, all but one with length data and 2,750 records with weight data. Weights, missing or otherwise, were calculated from the measured lengths using the length-weight regression (D.1) described in Section D.1.1, specifically $W=e^{a} L^{b}=e^{-11.82032} L^{3.03028}$.

## D.2.1. GLM method - unsorted vs. sorted samples (coastwide)

To remove some of the variance due to influential factors in the data, an additive lognormal model (Schnute et al. 2004) was used to adjust the annual index of fish weight for minor PMFC area:

$$
\begin{equation*}
w_{i j}=\mu+\alpha_{i}+\beta_{j}+\sigma \varepsilon_{i j m} \tag{D.10}
\end{equation*}
$$

where, $\mu=$ the overall mean;
$\alpha_{i} \quad=\quad$ year effect (with missing years)
unsorted: $i_{1}=1972, i_{N}=2016$, where $N=38$ available years,
sorted: $\quad i_{1}=1975, i_{N}=2009$, where $N=31$ available years;

| $\beta_{j}=$ | minor PMFC area effect |  |
| ---: | :--- | :--- |
|  | unsorted: $(1,2,3,4,5,6,7,8,11,12,20,21,23,35)$, |  |
|  | $\quad$ sorted: (0, 1, 2, 3, 4, 5, 8, 11, 12, 20, 21, 23); |  |
| $m$ | $=$ | number of fish weight values; |
| $\sigma$ | $=$ | standard deviation of the model; and |
| $\varepsilon_{i j m}$ | $=$ | independent residuals assumed to be standard normal $\mathrm{N}(0,1)$. |

The fitted unsorted-catch model had a residual standard error of 0.2984 on 50,947 degrees of freedom (multiple $R^{2}=0.6695$, adjusted $R^{2}=0.6691$, Figure D.17). The fitted sorted-catch model had a residual standard error of 0.3682 on 46,727 degrees of freedom (multiple $R^{2}=$ 0.4458 , adjusted $R^{2}=0.4453$, Figure D.18).

The main purpose of the GLM fit was to adjust for trend in the annual indices of weight; however, the process rendered the scale of the indices relative. To transform the relative indices back to absolute, they were multiplied by the ratio of the geometric mean of the nonstandardised annual indices ( $0.7606 \mathrm{~kg} / \mathrm{fish}$ ) to the geometric mean of the standardised (and sometimes normalized) indices ( $1 \mathrm{~kg} /$ fish ); see results in Table D.10.

$$
\begin{equation*}
w_{a i}=w_{s i}\left[\left(\prod_{i_{1}}^{i_{N}} w_{u i}\right)^{1 / N} /\left(\prod_{i_{1}}^{i_{N}} w_{s i}\right)^{1 / N}\right] \tag{D.11}
\end{equation*}
$$

where, $i_{1, \ldots, N}=$ annual index (unsorted: $N=38$ years, sorted: $N=31$ years),
$w_{u i}=$ unstandardized annual mean weights (kg/fish),
$w_{s i}=$ GLM-standardized annual mean weights (kg/fish), and
$w_{a i}=$ adjusted GLM-standardized annual mean weights (kg/fish).
The standardization removed substantial spatial effects from area, specifically PMFC minor area (Figure D.17, Figure D.18). Additional effects (e.g., season, depth, sex) were explored but their effects on the annual index series were minimal and consequently not used. Only season could have been used without losing annual indices; data for the other effects were not universally available for all years. The mean weight of unsorted fish ( $0.67989 \mathrm{~kg} / \mathrm{fish}$, Table D.10) was lower than that for sorted (kept) fish ( $1.09049 \mathrm{~kg} / \mathrm{fish}$, Table D.11). The delay-difference model assumes that signals in mean weight trend result from recruitment, not spatial movement of the fishery.

Table D.10. Annual mean weight (kg) per Walleye Pollock using samples of unsorted catch aboard commercial trawlers: $w_{u i}=$ non-standardized (non-std), $w_{s i}=$ GLM-standardized (glm-std), $w_{a i}=$ adjusted GLM-standardized (adj glm-std); numbers of fish used for annual mean calculations are also reported. Final row reports geometric mean weight of all years with data.

| Year | \# Fish | Fish wt. <br> (non-std) | Fish wt. <br> (glm-std) $)$ | Fish wt. <br> (adj glm-std) | Year | \# Fish | Fish wt. <br> (non-std) | Fish wt. <br> (glm-std) | Fish wt. <br> (adj glm-std) |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1972 | 140 | 0.87075 | 1.00025 | 0.67990 | 1998 | 152 | 0.96381 | 0.92997 | 0.63212 |
| 1973 | 2324 | 1.00374 | 0.75808 | 0.51529 | 1999 | 610 | 1.17332 | 1.09209 | 0.74232 |
| 1974 | 77 | 0.83167 | 1.29147 | 0.87784 | 2000 | 1502 | 0.83841 | 1.04592 | 0.71094 |
| 1976 | 593 | 1.36151 | 1.48094 | 1.00663 | 2001 | 3074 | 0.24238 | 0.95541 | 0.64941 |
| 1977 | 1308 | 1.27842 | 2.02779 | 1.37834 | 2002 | 293 | 0.25857 | 0.79753 | 0.54210 |
| 1978 | 6071 | 1.29251 | 1.32626 | 0.90149 | 2003 | 732 | 0.62835 | 0.80124 | 0.54463 |
| 1979 | 5350 | 1.38308 | 1.27879 | 0.86922 | 2004 | 547 | 0.73322 | 0.94263 | 0.64073 |
| 1980 | 1204 | 0.82737 | 1.49355 | 1.01520 | 2005 | 340 | 0.93635 | 0.91294 | 0.62054 |
| 1981 | 908 | 1.10047 | 0.82731 | 0.56234 | 2006 | 1440 | 0.64403 | 1.00675 | 0.68431 |


| Year | \# Fish | Fish wt. <br> (non-std) | Fish wt. <br> (glm-std) | Fish wt. <br> (adj glm-std) | Year | \# Fish | Fish wt. <br> (non-std) | Fish wt. <br> (glm-std) | Fish wt. <br> (adj glm-std) |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1985 | 255 | 1.43418 | 1.08711 | 0.73894 | 2007 | 867 | 0.76250 | 0.86537 | 0.58821 |
| 1988 | 3598 | 0.70947 | 1.51776 | 1.03166 | 2008 | 104 | 0.74438 | 0.95243 | 0.64739 |
| 1989 | 3527 | 0.43554 | 1.29587 | 0.88084 | 2009 | 470 | 0.78158 | 0.81327 | 0.55280 |
| 1990 | 719 | 0.54339 | 0.97948 | 0.66577 | 2010 | 562 | 0.75736 | 0.94711 | 0.64378 |
| 1991 | 337 | 0.48092 | 1.08684 | 0.73875 | 2011 | 592 | 0.50006 | 0.87682 | 0.59599 |
| 1992 | 3340 | 0.33721 | 1.03111 | 0.70087 | 2012 | 1190 | 0.53003 | 0.79461 | 0.54012 |
| 1993 | 1336 | 0.27738 | 0.88058 | 0.59855 | 2013 | 610 | 0.72040 | 0.78201 | 0.53155 |
| 1994 | 60 | 0.37424 | 0.81273 | 0.55244 | 2014 | 1276 | 0.49386 | 0.86495 | 0.58793 |
| 1996 | 3942 | 0.27973 | 0.87060 | 0.59177 | 2015 | 384 | 0.92280 | 0.80019 | 0.54391 |
| 1997 | 924 | 0.45531 | 0.82335 | 0.55965 | 2016 | 240 | 0.82869 | 1.00191 | 0.68102 |
| - | - | - | - | - | - | $\Sigma=$ | $\prod^{1 / N}=$ | $\prod^{1 / N}=$ | $\prod^{1 / \mathrm{N}}=$ |
|  |  |  |  |  |  | 50998 | 0.67972 | 1.00000 | 0.67972 |



Figure D.17. Normalised mean weight (kg/fish) of WAP coastwide estimated from (D.10) using data from sampling unsorted catch (original geometric mean $=0.680 \mathrm{~kg} / f \mathrm{fish}$ ). Panels from top to bottom show how the annual indices change as residual variance from each factor (in this case, only PMFC minor area) is removed. Broken lines show the index series in the panel above (using the factor accepted just prior to that depicted in the current panel).

Table D.11. Annual mean weight (kg) per Walleye Pollock from samples of sorted (coded as "keepers") catch aboard commercial trawlers coastwide: see Table D. 10 caption for details.

| Year | $\begin{array}{r} \# \\ \text { Fish } \end{array}$ | Fish wt. (non-std) | Fish wt. (glm-std) | Fish wt. (adj glm-std) | Year | $\begin{array}{r} \# \\ \text { Fish } \\ \hline \end{array}$ | Fish wt. (non-std) | Fish wt. (glm-std) | Fish wt. (adj glm-std) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1975 | 475 | 1.11523 | 1.00740 | 1.09016 | 1994 | 1997 | 1.29558 | 1.36343 | 1.47544 |
| 1976 | 1408 | 1.34492 | 0.81094 | 0.87757 | 1995 | 3871 | 1.31178 | 1.20889 | 1.30821 |
| 1977 | 699 | 1.31660 | 1.09264 | 1.18241 | 1996 | 2282 | 1.55122 | 1.25881 | 1.36223 |
| 1978 | 3106 | 1.29478 | 1.09655 | 1.18665 | 1997 | 554 | 1.19400 | 1.46935 | 1.59007 |
| 1979 | 1623 | 1.42501 | 1.10631 | 1.19720 | 1998 | 548 | 1.16599 | 1.16118 | 1.25658 |
| 1980 | 1233 | 0.97085 | 1.13920 | 1.23280 | 1999 | 2552 | 1.18719 | 0.89451 | 0.96800 |
| 1981 | 1469 | 1.12347 | 0.76335 | 0.82607 | 2000 | 1707 | 1.12938 | 0.92542 | 1.00145 |
| 1982 | 3404 | 1.20553 | 0.90398 | 0.97825 | 2001 | 2000 | 0.39071 | 0.91688 | 0.99221 |
| 1983 | 2095 | 1.24211 | 0.92093 | 0.99659 | 2002 | 768 | 0.29984 | 0.72025 | 0.77943 |
| 1984 | 868 | 1.10698 | 0.96721 | 1.04667 | 2003 | 1749 | 0.83749 | 0.79294 | 0.85808 |
| 1985 | 3296 | 1.30633 | 0.86806 | 0.93938 | 2004 | 983 | 0.53450 | 0.99397 | 1.07563 |
| 1986 | 251 | 1.50615 | 1.06892 | 1.15674 | 2005 | 750 | 0.72745 | 0.65608 | 0.70999 |
| 1990 | 1047 | 1.27638 | 1.32541 | 1.43431 | 2006 | 350 | 0.77008 | 0.85718 | 0.92760 |
| 1991 | 438 | 1.53763 | 1.14233 | 1.23618 | 2007 | 500 | 1.12009 | 0.73788 | 0.79851 |
| 1992 | 2587 | 1.39607 | 1.32153 | 1.43010 | 2009 | 200 | 1.48541 | 0.86547 | 0.93657 |
| 1993 | 1958 | 1.35347 | 1.29682 | 1.40336 | - | - | - | - | - |
| - | - | - | - | - | - | $\begin{array}{r} \Sigma= \\ 46768 \end{array}$ | $\begin{array}{r} \Pi_{1 / \mathrm{N}}= \\ 1.08216 \end{array}$ | $\begin{array}{r} \Pi^{1 / \mathrm{N}}= \\ 1.00000 \end{array}$ | $\begin{array}{r} \Pi^{1 / \mathrm{N}}= \\ 1.08216 \end{array}$ |



Figure D.18. Normalised mean weight (kg/fish) of WAP coastwide estimated from (D.10) using data from sampling sorted (keepers) catch (original geometric mean $=1.082$ kg/fish). See Figure D. 17 caption for additional detail.

## D.2.2. North vs. South (unsorted)

The same GLM standardisation (D.10) and (D.11) using PMFC minor area as a factor was applied to the North and South stocks of Walleye Pollock to derive stock-specific mean weight series. Only unsorted samples were used for these analyses. The most striking feature between the North and the South is that fish are roughly twice as large in the North with a geometric mean weight of $1.056 \mathrm{~kg} /$ fish vs. $0.521 \mathrm{~kg} / \mathrm{fish}$ in the South. The mean weight series for the North appears in Table D. 12 and Figure D. 19 and for the South in Table D. 13 and Figure D. 20.

While coastwide mean weights were used in the exploratory phase of the assessment, only mean weights from the North and South were used in the delay-difference model results in Appendix F.

Table D.12. Annual mean weight (kg) per Walleye Pollock using unsorted samples from the North stock caught aboard commercial trawlers: : $w_{u i}=$ non-standardized (non-std), $w_{s i}=$ GLM-standardized (glmstd), $w_{a i}=$ adjusted GLM-standardized (adj glm-std); number of samples and fish used for annual mean calculations are also reported. Final row reports geometric mean weight of all years with data..

| Year | $\#$ <br> Samp | \# <br> Fish | Fish wt. <br> (non-std) | Fish wt. <br> (glm-std) | Fish wt. <br> (adj glm-std) | Year | $\#$ <br> Samp | Fish wt. <br> Fish | Fish wt. <br> (non-std) <br> (glm-std) | Fish wt. <br> (adj glm-std) |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1973 | 4 | 666 | 0.44522 | 1.01932 | 1.07634 | 2003 | 2 | 125 | 1.01387 | 0.98346 | 1.03847 |
| 1974 | 1 | 77 | 0.83361 | 0.67685 | 0.71471 | 2004 | 3 | 219 | 0.85789 | 0.87875 | 0.92790 |
| 1976 | 4 | 593 | 1.36630 | 1.08168 | 1.14219 | 2005 | 3 | 193 | 1.06979 | 0.78859 | 0.83270 |
| 1977 | 10 | 738 | 1.19360 | 1.59989 | 1.68938 | 2006 | 2 | 109 | 0.82390 | 0.97472 | 1.02924 |
| 1978 | 26 | 5293 | 1.26248 | 1.10469 | 1.16649 | 2007 | 3 | 170 | 0.88071 | 0.76224 | 0.80487 |
| 1979 | 34 | 5150 | 1.40211 | 1.17811 | 1.24401 | 2008 | 1 | 51 | 1.22924 | 0.80679 | 0.85192 |
| 1980 | 6 | 1204 | 0.82942 | 1.37333 | 1.45015 | 2009 | 4 | 187 | 1.35303 | 1.03832 | 1.09641 |
| 1981 | 4 | 908 | 1.10393 | 0.76645 | 0.80933 | 2010 | 4 | 230 | 1.12128 | 1.22896 | 1.29770 |
| 1985 | 1 | 255 | 1.43928 | 1.00857 | 1.06498 | 2012 | 2 | 122 | 0.99180 | 0.97924 | 1.03401 |
| 1997 | 2 | 97 | 1.08712 | 1.41041 | 1.48931 | 2013 | 5 | 206 | 0.96015 | 0.81048 | 0.85582 |
| 1998 | 3 | 152 | 0.96654 | 0.94375 | 0.99655 | 2014 | 4 | 204 | 1.07054 | 0.85253 | 0.90022 |
| 1999 | 5 | 610 | 1.17708 | 0.95078 | 1.00397 | 2015 | 5 | 331 | 0.99307 | 1.23533 | 1.30443 |
| 2000 | 11 | 923 | 1.14732 | 1.00518 | 1.06140 | 2016 | 1 | 60 | 1.61203 | 1.09369 | 1.15487 |
| - | - | - | - | - | - | - | $\Sigma$ | $\Sigma=$ | $\prod^{1 / N}=$ | $\prod^{1 / N}=$ | $\prod^{1 / N}=$ |
|  |  |  |  |  |  |  |  |  | 150 | 18873 | 1.05594 |



Figure D.19. Normalised mean weight (kg/fish) of WAP North estimated from (D.10) using data from sampling unsorted catch (original geometric mean $=1.056 \mathrm{~kg} /$ fish). Panels from top to bottom show how the annual indices change as residual variance from each factor (in this case, only PMFC minor area) is removed. Broken lines show the index series in the panel above (using the factor accepted just prior to that depicted in the current panel).

Table D.13. Annual mean weight (kg) per Walleye Pollock using unsorted samples from the South stock caught aboard commercial trawlers: see Table D. 12 caption for details.

| Year | Samp | Fish | Fish wt. (non-std) | Fish wt. (glm-std) | Fish wt. <br> (adj glm-std) | Year | Samp | $\begin{array}{r} \# \\ \text { Fish } \\ \hline \end{array}$ | Fish wt. (non-std) | Fish wt. (glm-std) | Fish wt. (adj glm-std) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1972 | 2 | 140 | 0.86657 | 0.99684 | 0.51888 | 2002 | 2 | 293 | 0.25799 | 0.81064 | 0.42196 |
| 1973 | 9 | 1658 | 1.22141 | 1.00120 | 0.52115 | 2003 | 9 | 607 | 0.54768 | 0.81378 | 0.4235 |
| 1977 | 4 | 570 | 1.38474 | 1.99986 | 1.04098 | 2004 | 5 | 328 | 0.64893 | 0.99373 | 0.51726 |
| 1978 | 4 | 778 | 1.51574 | 1.66741 | 0.86793 | 2005 | 4 | 147 | 0.76201 | 1.02091 | 0.53141 |
| 1979 | 1 | 200 | 1.01578 | 1.76370 | 0.91805 | 2006 | 21 | 1331 | 0.62708 | 1.14122 | 0.59403 |
| 1988 | 13 | 3598 | 0.70662 | 1.71509 | 0.89274 | 2007 | 9 | 697 | 0.73123 | 0.96939 | 0.5045 |
| 1989 | 18 | 3527 | 0.43429 | 1.31448 | 0.68422 | 2008 | 1 | 53 | 0.28124 | 1.06983 | 0.5568 |
| 1990 | 4 | 719 | 0.54164 | 0.99494 | 0.51789 | 2009 | 4 | 283 | 0.40613 | 0.67862 | 0.35324 |
| 1991 | 2 | 337 | 0.47957 | 1.10330 | 0.57429 | 2010 | 6 | 332 | 0.50607 | 0.86530 | 0.4504 |
| 1992 | 50 | 3340 | 0.33660 | 1.04748 | 0.54524 | 2011 | 11 | 592 | 0.49856 | 0.84970 | 0.44229 |
| 1993 | 24 | 1336 | 0.27704 | 0.89481 | 0.46577 | 2012 | 18 | 1068 | 0.47613 | 0.90615 | 0.47167 |
| 1994 | 1 | 60 | 0.37348 | 0.82553 | 0.42971 | 2013 | 8 | 404 | 0.59731 | 0.88366 | 0.45997 |
| 1996 | 53 | 3942 | 0.27940 | 0.88310 | 0.45968 | 2014 | 17 | 1072 | 0.38372 | 0.97481 | 0.50741 |
| 1997 | 16 | 827 | 0.38079 | 0.83661 | 0.43547 | 2015 | 1 | 53 | 0.49991 | 0.85971 | 0.44750 |
| 2000 | 10 | 579 | 0.35104 | 0.93747 | 0.48797 | 2016 | 3 | 180 | 0.56772 | 0.56092 | 0.29197 |
| 2001 | 42 | 3074 | 0.24218 | 0.86350 | 0.44947 | - | - | - | - | - |  |
| - | - | - | - | - |  |  | $\begin{gathered} \Sigma= \\ 372 \end{gathered}$ | $\begin{array}{r} \Sigma= \\ 32125 \end{array}$ | $\begin{array}{r} \Pi_{1 / \mathrm{N}}= \\ 0.52052 \end{array}$ | $\begin{array}{r} \Pi_{1 / \mathrm{N}}^{=} \\ 1.00000 \end{array}$ | $\begin{array}{r} \Pi^{1 / \mathrm{N}}= \\ 0.52052 \end{array}$ |



Figure D.20. Normalised mean weight (kg/fish) of WAP South estimated from (D.10) using data from sampling unsorted catch (original geometric mean $=0.521 \mathrm{~kg} /$ fish). See Figure D. 19 caption for additional detail.

## D.3. HABITAT

Walleye Pollock is ubiquitous along the BC coast, with an estimated area of occupancy ranging from $\sim 50,200 \mathrm{~km}^{2}$ using trawl occurrence (Figure A.2) to $\sim 72,400 \mathrm{~km}^{2}$, using bathymetry limits (Figure D.21). The estimated bathymetry limits are derived from the distribution of this species captured in observer log trawl tows (bottom and midwater), which spans depths 62 to 448 m $98 \%$ of the time (Figure D.22). Species caught concurrently in coastwide observer log tows that captured at least one Walleye Pollock, herein referred to as "pollock tows" (Figure D.23), are dominated by Walleye Pollock ( $25 \%$ of total catch) and include significant amounts of Arrowtooth Flounder Atheresthes stomias (16\%), Pacific Ocean Perch Sebastes alutus (14\%), and Pacific Hake Merluccius productus (13\%).
Regional variations in depth distributions of pollock tows occur along the BC coast, and species caught concurrently also vary. In addition to the BC offshore stock assessed, we present three PMFC combinations that are typically used by managers of this species - 5CDE (Hecate Strait, Dixon Entrance, west coast Haida Gwaii), 5AB (Queen Charlotte Sound, Queen Charlotte Strait), and 3CD (west coast Vancouver Island and mouth of Juan de Fuca Strait).

The 5CDE region is dominated by shallow trawl effort (Figure D.24), which presumably indicates the targeting of flatfish in Hecate Strait. Pollock tows occur a little deeper than this and are dominated by catches of Arrowtooth Flounder (27\% of total catch), Walleye Pollock (20\%), and

Dover Sole (7\%), amongst an assemblage of rockfish, flatfish, and skates (Figure D.25, Table D.14).
Region 5AB contains three important gullies - Goose Island Gully, Mitchell's Gully, and Moresby Gully. The effort of the trawl fleet in this region appears to be limited to depths shallower than 400 m with a mode at 100 m (Figure D.26), whereas the depth of pollock tows shows a mode at $\sim 200 \mathrm{~m}$. While this region is highly important to the POP fishery, the inclusion of PMFC minor area 12 increases the catch contribution of Walleye Pollock in pollock tows Figure D. 27 to equal that of POP (both $\sim 27 \%$ of total catch). This region has the highest percentage of Walleye Pollock in depths where it is caught (Table D.14). The next greatest contributor to catch is Arrowtooth Flounder at only 10\%.

Region 3CD (excluding PMFC minor area 20) has traditionally been fished to great depths ( $>700 \mathrm{~m}$, not shown) due to favourable bathymetry but the shallow mode of the fleet effort distribution has always been the greatest (Figure D.28, which includes minor area 20). The pollock tows follow the shallow fleet effort with a peak at $\sim 150 \mathrm{~m}$. The species that occur in pollock tows appear in Figure D.29. The dominant concurrent species in this region are Pacific Hake (40\%), followed by Walleye Pollock (30\%) and Yellowtail Rockfish Sebastes flavidus (8\%). Catches of other concurrent species appear in Table D.14.


Figure D.21. Highlighted bathymetry (green) between 62 and 448 m serves as a proxy for benthic habitat for Walleye Pollock along the BC coast. Within Canada's exclusive economic zone (EEZ, blue highlighted area), the green highlighted region covers $72,419 \mathrm{~km}^{2}$. The boundaries in red delimit the PMFC areas.


Figure D.22. BC Offshore - Depth frequency of bottom tows that captured Walleye Pollock (WAP) from commercial trawl logs (1996-2007 in PacHarvest, 2007-2016 in GFFOS) in areas outside the Strait of Georgia (transparent histogram). The vertical solid lines denote the 1\% and 99\% quantiles. The black curve shows the cumulative frequency of tows that encounter WAP while the red curve shows the cumulative catch of WAP at depth (scaled from 0 to 1). The median depths of WAP encounters (inverted grey triangle) and of cumulative catch (inverted red triangle) are indicated along the upper axis. ' $N$ ' reports the total number of tows; ' $C$ ' reports the total catch ( $t$ ). The shaded histogram in the background reports the relative trawl effort on all species offshore down to 700 m .


Figure D.23. BC Offshore - Distribution of catch weights summed over the period February 1996 to January 2017 for important finfish species in bottom and midwater trawl tows that caught at least one Walleye Pollock coastwide. Tows were selected over a depth range between 62 and 448 m (the 1\% and 99\% quantile range, see Figure D.22). Relative concurrence is expressed as a percentage by species relative to the total catch weight summed over all finfish species in the specified period. Walleye Pollock is indicated in blue on the $y$-axis; other species of interest to SARA are indicated in red.


Figure D.24. 5CDE - Depth frequency of bottom and midwater trawl tows that captured Walleye Pollock (WAP) from commercial trawl logs (1996-2007 in PacHarvest, 2007-2016 in GFFOS) in PMFC major areas 5CDE (transparent histogram). The shaded histogram in the background reports the relative trawl effort on all species in 5CDE down to 700 m. Plot details appear in Figure D. 22 .


Figure D.25. 5CDE - Distribution of catch weights summed over the period February 1996 to January 2017 for important finfish species in bottom and midwater trawl tows that caught at least one Walleye Pollock off Haida Gwaii and in Dixon Entrance. Tows were selected over a depth range between 55 and 457 m (the 1\% and 99\% quantile range, see Figure D.24). Plot details appear in Figure D. 23.


Figure D.26. 5AB - Depth frequency of bottom and midwater trawl tows that captured Walleye Pollock (WAP) from commercial trawl logs (1996-2007 in PacHarvest, 2007-2016 in GFFOS) in PMFC major areas 5AB (transparent histogram). The shaded histogram in the background reports the relative trawl effort on all species in $5 A B$ down to 700 m . Plot details appear in Figure D.22.


Figure D.27. 5AB - Distribution of catch weights summed over the period February 1996 to January 2017 for important finfish species in bottom and midwater trawl tows that caught at least one Walleye Pollock in Queen Charlotte Sound and Strait. Tows were selected over a depth range between 90 and 401 m (the 1\% and 99\% quantile range, see Figure D.26). Plot details appear in Figure D. 23.


Figure D.28. 3CD - Depth frequency of bottom and midwater trawl tows that captured Walleye Pollock (WAP) from commercial trawl logs (1996-2007 in PacHarvest, 2007-2016 in GFFOS) in PMFC major areas 3CD (transparent histogram). The shaded histogram in the background reports the relative trawl effort on all species in 3CD down to 700 m . Plot details appear in Figure D.22.


Figure D.29. 3CD - Distribution of catch weights summed over the period February 1996 to January 2017 for important finfish species in bottom and midwater trawl tows that caught at least one Walleye Pollock off the west coast of Vancouver Island and at the mouth of Juan de Fuca Strait. Tows were selected over a depth range between 64 and 470 m (the 1\% and 99\% quantile range, see Figure D.28). Plot details appear in Figure D. 23.

Table D.14. Top 20 species by catch weight (sum of landed + discarded from Feb, 1996 to Jan,2016, observer logs only) that co-occur in Walleye Pollock bottom and midwater trawl tows along the BC coast, in PMFC 5DE, in PMFC 5AB + minor 12, and in PMFC 3CD + minor 20. Rockfish species of interest to COSEWIC appear in red font, target species (occurs in every tow) appears in blue font.

| Code | Species | Latin Name | Catch (t) | Catch (\%) |
| :---: | :---: | :---: | :---: | :---: |
| Coast |  |  |  |  |
| 228 | Walleye Pollock | Theragra chalcogramma | 44,553 | 24.562 |
| 602 | Arrowtooth Flounder | Atheresthes stomias | 28,435 | 15.676 |
| 396 | Pacific Ocean Perch | Sebastes alutus | 26,060 | 14.367 |
| 225 | Pacific Hake | Merluccius productus | 22,724 | 12.528 |
| 418 | Yellowtail Rockfish | Sebastes flavidus | 8,840 | 4.873 |
| 626 | Dover Sole | Microstomus pacificus | 7,073 | 3.899 |
| 440 | Yellowmouth Rockfish | Sebastes reedi | 5,887 | 3.245 |
| 405 | Silvergray Rockfish | Sebastes brevispinis | 5,723 | 3.155 |
| 222 | Pacific Cod | Gadus macrocephalus | 5,149 | 2.839 |
| 628 | English Sole | Parophrys vetulus | 2,943 | 1.623 |
| 610 | Rex Sole | Errex zachirus | 2,791 | 1.539 |
| 066 | Spotted Ratfish | Hydrolagus colliei | 2,778 | 1.532 |
| 056 | Big Skate | Raja binoculata | 2,299 | 1.268 |
| 044 | Spiny Dogfish | Squalus acanthias | 1,872 | 1.032 |
| 439 | Redstripe Rockfish | Sebastes proriger | 1,467 | 0.809 |
| 401 | Redbanded Rockfish | Sebastes babcocki | 1,399 | 0.771 |
| 394 | Rougheye Rockfish | Sebastes aleutianus | 1,157 | 0.638 |
| 467 | Lingcod | Ophiodon elongatus | 1,130 | 0.623 |
| 451 | Shortspine Thornyhead | Sebastolobus alascanus | 1,029 | 0.567 |
| 450 | Sharpchin Rockfish | Sebastes zacentrus | 888 | 0.490 |
| 5CDE |  |  |  |  |
| 602 | Arrowtooth Flounder | Atheresthes stomias | 19,014 | 26.797 |
| 228 | Walleye Pollock | Theragra chalcogramma | 14,180 | 19.985 |
| 626 | Dover Sole | Microstomus pacificus | 5,235 | 7.379 |
| 396 | Pacific Ocean Perch | Sebastes alutus | 4,575 | 6.447 |
| 222 | Pacific Cod | Gadus macrocephalus | 3,303 | 4.655 |
| 405 | Silvergray Rockfish | Sebastes brevispinis | 3,230 | 4.552 |
| 225 | Pacific Hake | Merluccius productus | 2,860 | 4.030 |
| 628 | English Sole | Parophrys vetulus | 2,689 | 3.790 |
| 066 | Spotted Ratfish | Hydrolagus colliei | 2,471 | 3.482 |
| 418 | Yellowtail Rockfish | Sebastes flavidus | 2,263 | 3.189 |
| 056 | Big Skate | Raja binoculata | 2,178 | 3.069 |
| 610 | Rex Sole | Errex zachirus | 1,987 | 2.800 |
| 394 | Rougheye Rockfish | Sebastes aleutianus | 803 | 1.131 |
| 451 | Shortspine Thornyhead | Sebastolobus alascanus | 581 | 0.818 |
| 614 | Pacific Halibut | Hippoglossus stenolepis | 577 | 0.813 |
| 044 | Spiny Dogfish | Squalus acanthias | 534 | 0.753 |
| 059 | Longnose Skate | Raja rhina | 403 | 0.567 |
| 455 | Sablefish | Anoplopoma fimbria | 400 | 0.563 |
| 401 | Redbanded Rockfish | Sebastes babcocki | 383 | 0.540 |
| 439 | Redstripe Rockfish | Sebastes proriger | 369 | 0.520 |
| 5AB |  |  |  |  |
| 228 | Walleye Pollock | Theragra chalcogramma | 21,148 | 27.355 |
| 396 | Pacific Ocean Perch | Sebastes alutus | 20,569 | 26.606 |
| 602 | Arrowtooth Flounder | Atheresthes stomias | 7,693 | 9.951 |
| 225 | Pacific Hake | Merluccius productus | 6,548 | 8.470 |
| 440 | Yellowmouth Rockfish | Sebastes reedi | 5,391 | 6.974 |
| 418 | Yellowtail Rockfish | Sebastes flavidus | 4,186 | 5.415 |
| 405 | Silvergray Rockfish | Sebastes brevispinis | 2,007 | 2.596 |


| Code | Species | Latin Name | Catch (t) | Catch (\%) |
| :---: | :--- | :--- | ---: | ---: |
| 626 | Dover Sole | Microstomus pacificus | 1,383 | 1.789 |
| 401 | Redbanded Rockfish | Sebastes babcocki | 941 | 1.217 |
| 439 | Redstripe Rockfish | Sebastes proriger | 931 | 1.204 |
| 222 | Pacific Cod | Gadus macrocephalus | 833 | 1.077 |
| 044 | Spiny Dogfish | Squalus acanthias | 683 | 0.883 |
| 610 | Rex Sole | Errex zachirus | 621 | 0.803 |
| 450 | Sharpchin Rockfish | Sebastes zacentrus | 540 | 0.698 |
| 437 | Canary Rockfish | Sebastes pinniger | 442 | 0.571 |
| 451 | Shortspine Thornyhead | Sebastolobus alascanus | 417 | 0.539 |
| 467 | Lingcod | Ophiodon elongatus | 386 | 0.499 |
| 417 | Widow Rockfish | Sebastes entomelas | 360 | 0.465 |
| 455 | Sablefish | Anoplopoma fimbria | 340 | 0.440 |
| 394 | Rougheye Rockfish | Sebastes aleutianus | 269 | 0.348 |
| $3 C D$ |  |  |  |  |
| 225 | Pacific Hake | Merluccius productus | 12,266 | 40.053 |
| 228 | Walleye Pollock | Theragra chalcogramma | 7,638 | 24.941 |
| 418 | Yellowtail Rockfish | Sebastes flavidus | 2,344 | 7.654 |
| 602 | Arrowtooth Flounder | Atheresthes stomias | 1,843 | 6.017 |
| 222 | Pacific Cod | Gadus macrocephalus | 1,029 | 3.359 |
| 396 | Pacific Ocean Perch | Sebastes alutus | 883 | 2.882 |
| 044 | Spiny Dogfish | Squalus acanthias | 644 | 2.103 |
| 626 | Dover Sole | Microstomus pacificus | 513 | 1.676 |
| 405 | Silvergray Rockfish | Sebastes brevispinis | 481 | 1.571 |
| 467 | Lingcod | Ophiodon elongatus | 385 | 1.257 |
| 437 | Canary Rockfish | Sebastes pinniger | 275 | 0.897 |
| 450 | Sharpchin Rockfish | Sebastes zacentrus | 209 | 0.682 |
| 610 | Rex Sole | Errex zachirus | 194 | 0.632 |
| 440 | Yellowmouth Rockfish | Sebastes reedi | 182 | 0.595 |
| 607 | Petrale Sole | Eopsetta jordani | 174 | 0.570 |
| 417 | Widow Rockfish | Sebastes entomelas | 173 | 0.566 |
| 439 | Redstripe Rockfish | Sebastes proriger | 166 | 0.543 |
| 066 | Spotted Ratfish | Hydrolagus colliei | 161 | 0.526 |
| 059 | Longnose Skate | Raja rhina | 152 | 0.496 |
| 455 | Sablefish | Anoplopoma fimbria | 142 | 0.465 |
|  |  |  |  |  |

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## APPENDIX E. MODEL EQUATIONS

## E.1. INTRODUCTION

The software used in this stock assessment of Walleye Pollock (WAP) is a variant of the integrated Statistical Catch Age Model (iSC $\forall \mathrm{M}$ ), developed by Steven Martell (Martell 2010) and modified by Robyn Forrest in 2015 to run a delay-difference (DD) model for Pacific Cod (Gadus macrocephalus) on the west code of Canada (Forrest et al. 2015). The DD model was written in AD Model Builder template code (Fournier et al. 2012) and was compiled for WAP using the PBSadmb package (Schnute et al. 2017).

The original $i$ SC $\forall \mathrm{M}$ code for Pacific Cod has been modified several times - in 2015 for Shortspine Thornyhead (Sebastolobus alascanus, Starr and Haigh 2017) and in 2017 for Walleye Pollock (Theragra chalcogramma):

1. The parameter $R_{0}$ (equilibrium unfished age- 0 recruits) was estimated while $\bar{R}$ (average annual recruitment) was tethered to equal $R_{0}$. In the code, this entailed forcing theta(4) $=$ theta(1). There seems to be some difference of opinion on how these parameters should be handled.
2. The analytical method for MSY implemented in the original code gave results that did not appear to be correct. We adapted code from Awatea, used to model populations of Pacific Ocean Perch (Sebastes alutus), to estimate MSY and $B_{\text {MSY }}$ through a brute-force method by projecting forward under constant $F$ policies and finding the $F$ that gave the greatest yield (MSY). The associated biomass is $B_{\mathrm{MSY}}$. This method was crude but effective.
3. The original version of the model assumed that the mean weight data vector is continuous by year for a given range. An un-interrupted series was not available for Walleye Pollock so we recoded the input routine to accept a matrix with year and mean weight. Other model components (e.g., mean weight likelihood) were re-coded accordingly. This revision allows the use of annual mean weight data with missing years.

## E.2. DELAY-DIFFERENCE MODEL

Note: The text and equations below are taken from Forrest et al. (2015), and modified as needed.
Delay difference models represent an intermediate approach between aggregated surplus production models and age-structured models. The delay-difference structure tracks the effects of recruitment, survival and growth on biomass, without requiring an age-structured framework, and can perform well as long as its major assumptions are met (Hilborn and Walters 1992). Difference equations, which allow for a time-delay between spawning and recruitment, are used to build population models in discrete time-steps (generally 1 year), in which the surviving biomass for next year is predicted from the surviving biomass from last year, after adjusting for growth and adding next year's recruitment. An advantage of delay difference models over simpler production models is that they do not assume constant recruitment over time.

The key assumptions of the delay difference model are:

- Growth in mean body weight $W_{a}$ follows the linear relationship described by the Ford-Walford equation (E.1); see Section D.1.4.

$$
\begin{equation*}
W_{a}=\alpha_{g}+\rho_{g} W_{a-1} \tag{E.1}
\end{equation*}
$$

- Knife-edge selectivity, i.e., all fish aged $k$ and older, are equally vulnerable to the fishing gear. A corollary to the assumption of knife-edge selectivity is that maturity is also knife-edge and the same as selectivity. This means that all fish in the model are mature and fully selected; and
- Mortality at age remains constant, i.e., all fish aged $k$ and older have the same mortality rate.

The delay difference model collapses all the equations needed to fully describe the population's age structure into equations for the biomass $\left(B_{t}\right)$, total numbers ( $N_{t}$ ), and survival ( $S_{t}$ ) at time $t$ :

$$
\begin{gather*}
B_{t}=S_{t-1}\left(\alpha_{g} N_{t-1}+\rho_{g} B_{t-1}\right)+w_{k} R_{t}  \tag{E.2}\\
N_{t}=S_{t-1} N_{t-1}+R_{t}  \tag{E.3}\\
S_{t}=e^{-\left(M+F_{t}\right)} \tag{E.4}
\end{gather*}
$$

where:
$k=$ the age at which fish are assumed to become fully vulnerable to fishing;
$M=$ the instantaneous natural mortality rate;
$F_{t}=$ the instantaneous fishing mortality rate at time $t$;
$\left(\alpha_{g}, \rho_{g}\right)=$ the intercept and slope of the Ford-Walford equation for all ages $\geq k$;
$w_{k}=$ the weight at age $k$; and
$R_{t}=$ is the number of recruits at time $t$ calculated by the stock-recruit function, here constrained to conform to a Beverton-Holt relationship with constants $a$ and $b$ (E.26).
We assume that recruitment to the fishery and surveys occurs at age $k \in\{3,4,5\}$ for both stocks in the various alternative cases.
A list of model parameters is given in Table E.1. Equilibrium and dynamic equations are given in Tables E. 2 and E.3, respectively. Variance parameters and likelihood components of the objective function are given in Table E.4.

## E.2.1. Objective function components

Variance parameters and objective function components are listed in Table E.4. The objective function $f(\theta)$ in the delay-difference model contains five major components:

1. the negative log-likelihood for the relative abundance data (E.33);
2. the negative log-likelihood for the catch data (E.35);
3. the negative log-likelihood for the mean weight data (E.37);
4. the prior distributions for model parameters, and
5. three penalty functions that:
a. constrain the estimates of annual recruitment to conform to a Beverton-Holt stock-recruit function;
b. weakly constrain the log recruitment deviations to a normal distribution; and
c. weakly constrain estimates of log fishing mortality to a normal distribution, $\mathrm{N}(\ln (0.2), 4.0)$, to prevent estimates of catch from exceeding estimated biomass.

## E.2.2. Variance components and weighting of index data

The $i$ SC $\forall \mathrm{M}$ modelling framework (Martell 2010) partitions the variance using an errors in variables approach. Total variance $\vartheta$ can be fixed or estimated, and was fixed for the Walleye Pollock delay-difference model (as it was for the Shortspine Thornyhead model, Starr and Haigh 2017). Total variance is partitioned by the model into observation and process error components (E. 27 and E.28, respectively) using the parameter $\rho$, which represents the proportion of the total variance that is due to observation error (Punt and Butterworth 1993; Deriso et al. 2007). The process error component (E.28) of the total variance is applied to the estimated recruitments as shown in equation (E.39).
Two variance control parameters in the model were fixed ( $\rho=0.1, \vartheta=2.5$ ) for Walleye Pollock. The formulae for model observation and process error, respectively, are: $\sigma=\sqrt{\rho} \cdot \sqrt{1 / \vartheta}$ and $\tau=\sqrt{1-\rho} \cdot \sqrt{1 / \vartheta}$. In the Shortspine Thornyhead assessment, sensitivities to $\sigma$ and $\tau$ were examined through runs that fixed the variance control parameters at alternative values; the sensitivities showed only small effects on model outcome.
The standard deviation used when fitting the survey and CPUE abundance index data is given in equation (E.27), with each index value weighted by the inverse of the CV associated with that index, as shown in equation (E.29). The index variance is added to the total likelihood as shown in (E.33). Five surveys and one CPUE index series were fitted in this model. The relative sampling error or CV $\left(c_{j t}\right)$ associated with each survey index value was used without adding additional survey process error $c_{j}^{\prime}$. A relative error of 0.3 was assumed for each CPUE index value. We did not attempt to alter the relative weights of the component data series (Francis 2011), instead using the observation error CVs estimated by the surveys without modification.

Note that after the assessment model was accepted, some of the calculations for the variance components: $(\sigma, \tau)=(\sqrt{\rho} \cdot \varphi, \sqrt{1-\rho} \cdot \varphi)$, appear to have been incorrectly specified as $(\sqrt{\rho} / \varphi, \sqrt{1-\rho} / \varphi)$ in the model code iscamdelaydiff.tpl (dated 2017-01-03). These potential errors appear on lines 2054 (likelihood for stock-recruitment relationship in calc_objective_function) and 2513-14 (in simulation_model function). In these instances $(\sigma, \tau)=(0.5,1.5)$ instead of the correct values $(0.2,0.6)$. We don't anticipate that these errors would change the model outcomes substantially.

## E.3. REFERENCES POINTS, PROJECTIONS AND ADVICE TO MANAGERS

Typically, advice to managers is given with respect to three reference points based on maximum sustainable yield (MSY). The provisional reference points of the DFO Precautionary Approach (DFO 2006), namely $0.4 B_{\mathrm{MSY}}$ and $0.8 B_{\mathrm{MSY}}$, comprise the primary benchmarks for advice, where $B_{\mathrm{MSY}}$ is the estimated equilibrium spawning biomass at MSY. The third reference point is $u_{\text {MSY }}$, the harvest rate at MSY, and is derived from the instantaneous fishing mortality at MSY: $u_{\text {MSY }}=\left(1-e^{-F_{M S Y}}\right)$. However, exploration on the treatment of $R_{0}$ and $\bar{R}$, i.e., (i) estimating $R_{0}$ and setting $\bar{R}=R_{0}$, (ii) estimating $\bar{R}$ and setting $R_{0}=\bar{R}$, and (iii) estimating $R_{0}$ and $\bar{R}$ independently, uncovered instabilities in the trajectory of spawning biomass, specifically in the start and end points, $B_{0}$ and $B_{2017}$, respectively. Unfortunately, these two points are critical in assessing stock status using MSY-based criteria.
For these reasons, we adopted historical reference points (HRP) to assess stock status and to provide advice to managers in the form of decision tables. As a proxy for $B_{\mathrm{MSY}}$, we use average $B_{t}\left(B_{\text {avg }}\right)$, where $t=1967, \ldots, 2016$; similarly, $u_{\text {avg }}$ acts as a proxy for $u_{\text {MSY }}$. For the limit reference point (LRP) we adopt $B_{t}$ from year $t$ in which biomass was at a minimum and
subsequently recovered to exceed $B_{\text {avg. }}$. The minimum year is determined for each MCMC sample, and the biomass from this year is denoted $B_{\min }$. Consequently, there are $1000 B_{\text {min }}$ values with a distribution of years in which they occur. The upper stock reference (USR) is simply $2 B_{\text {min }}$.
Projections were made for only 2 years due to the model's inherent uncertainty and its lack of associated age structure, starting with the biomass calculated for the start of 2017, across a range of constant catch strategies. For each strategy, projections were performed for each of the 1000 MCMC samples (resulting in posterior distributions of future spawning biomass).

Recruitments for the projections were randomly generated from lognormal recruitment deviations applied to the deterministic recruitment estimate from the Beverton-Holt stock-recruitment function, using randomly generated values of $\epsilon_{t} \sim \operatorname{Normal}\left(0, \sigma_{R}^{2}\right)$. For each of the 1000 MCMC samples a time series of $\left\{\epsilon_{t}\right\}$ was generated. For each MCMC sample, the same time series of $\left\{\epsilon_{t}\right\}$ was used for every catch strategy so that, for a given MCMC sample, all catch strategies experienced the same recruitment stochasticity in projections using (E.2).

Decision tables comprise probabilities calculated as the proportion of the 1000 MCMC samples for which $B_{t=2017, \ldots, 2019}$ is greater than a reference point ( $B_{2017}, B_{\min }, 2 B_{\mathrm{min}}, B_{\text {avg }}, u_{\text {avg }}$ ), and is expressed as $\mathrm{P}\left(B_{t}>B_{\text {ref }}\right)$ or $\mathrm{P}\left(u_{t-1}>u_{\text {ref }}\right)$.

Table E.1. Notation for the $i S C \forall M$ delay-difference model used for Walleye Pollock. The term 'log' refers to natural logarithms (base e) herein.

| Symbol | Description | North | South |
| :--- | :--- | ---: | ---: |
|  | Indices (subscripts) |  |  |
| $t$ | Model year, where $t=1980,1981, \ldots, 2015 ;$ |  |  |
|  | and $t=1980$ represents unfished equilibrium conditions |  |  |
| $j$ | Gear (catch, surveys, CPUE) |  |  |
| $g$ | Ford-Walford identifier |  |  |
|  |  |  |  |
|  | Fixed input parameters | 3 | 3 |
| $k$ | Age at knife-edge recruitment | 66.944 | 50.827 |
| $L_{\infty}$ | Theoretical maximum length (cm) | 0.2118 | 0.1991 |
| $\kappa$ | von Bertalanffy growth rate | -1.136 | -1.790 |
| $t_{0}$ | Theoretical age at length = 0 cm | $7.102 \mathrm{E}-6$ | $7.354 \mathrm{E}-6$ |
| $\alpha$ | Scaling parameter of the length-weight relationship | 3.042 | 3.030 |
| $\beta$ | Exponent of the length-weight relationship | 0.3475 | 0.1444 |
| $\alpha_{g}$ | Intercept of the Ford-Walford plot, for all ages $>k$ | 0.8668 | 0.8706 |
| $\rho_{g}$ | Slope of the Ford-Walford plot, for all ages $>k$ | 0.4929 | 0.2488 |
| $W_{k}$ | Weight at age of recruitment $k$ | $\ln (0.30)$ | $\ln (0.30)$ |
| $M$ | Natural mortality (estimated in log space) |  |  |


| Symbol | Description | North | South |
| :---: | :---: | :---: | :---: |
|  | Annual input data |  |  |
| $C_{j t}$ | Catch (metric tonnes) for gear $j=1$ (total commercial) at time $t$ |  |  |
| $W_{t}$ | Mean weight (kg) of individuals $i$ in population at time $t$ where all weights are calculated from lengths: $W_{i}=\alpha L_{i}^{\beta}$ |  |  |
| $I_{j t}$ | Indices of abundance for gear $j$ at time $t$ in BC North, where $j=2$ - GB Reed (or GIG) rockfish survey series $j=3$ - Hecate Strait Assemblage survey series $j=4$ - Hecate Strait synoptic survey series $j=5$ - West Coast Haida Gwaii synoptic survey series $j=6-$ commercial WAP North CPUE series |  |  |
| $I_{j t}$ | Indices of abundance for gear $j$ at time $t$ in BC South, where $j=2$ - GB Reed (or GIG) rockfish survey series <br> $j=3$ - West Coast Vancouver Island synoptic survey series $j=4$ - Queen Charlotte Sound synoptic survey series $j=5$ - commercial WAP South CPUE series |  |  |
| $\mathrm{c}_{j t}$ | Annual coefficients of variation (CV) for $I_{j t}$ |  |  |
|  | Time-invariant parameters |  |  |
| $R_{0}$ | Equilibrium unfished age-0 recruits (est. in log space) |  |  |
| $h$ | Steepness of the stock-recruit relationship |  |  |
| $\chi$ | Recruitment compensation ratio (CR) |  |  |
| $a$ | Slope of the stock-recruit function at the origin |  |  |
| $b$ | Scaling parameter of the stock-recruit function |  |  |
| $N_{0}$ | Equilibrium unfished number of fish |  |  |
| $B_{0}$ | Equilibrium unfished biomass (t) |  |  |
| $S_{0}$ | Equilibrium unfished survival rate |  |  |
| $\bar{W}_{0}$ | Equilibrium unfished mean weight (kg) |  |  |
| $c_{j}^{\prime}$ | Additional process error in abundance indices $I_{j t}$ for gear $j$ |  |  |
| $n_{j}$ | Number of abundance indices for gear $j$ |  |  |
|  | Time-varying parameters (at time $t$ ) |  |  |
| $\omega_{t}$ | Recruitment deviations (est. in log space) |  |  |
| $F_{t}$ | Fishing mortality (est. in log space) by the commercial fishery |  |  |
| $S_{t}$ | Survival rate |  |  |
| $N_{t}$ | Numbers of fish |  |  |
| $R_{t}$ | Recruits (1000s fish) |  |  |
| $B_{t}$ | Biomass (tonnes) |  |  |
| $\bar{W}_{t}$ | Predicted mean weight (kg) |  |  |


| Symbol | Description | North | South |
| :---: | :---: | :---: | :---: |
|  | Likelihood components |  |  |
| $\vartheta$ | Total variance of the total error | 2.5 | 2.5 |
| $\varphi$ | Precison as square root of inverse variance $\vartheta$ | $\sqrt{ } 0.4$ | $\sqrt{ } 0.4$ |
| $\rho$ | Proportion of total variance due to observation error | 0.1 | 0.1 |
| $\sigma_{O}$ | Overall standard deviation of observation residuals | 0.2 | 0.2 |
| $\sigma_{R}$ | Standard deviation of In-recruitment deviations | 0.6 | 0.6 |
| $\sigma_{W}$ | Standard deviation of mean weight | 0.15 | 0.15 |
| $\sigma_{C}$ | Standard deviation of catch |  |  |
| $\sigma_{j t}$ | Annual standard deviation of observation residuals for each survey |  |  |
| $q_{j}$ | Constant of proportionality in indices of catchability (est. in log space) |  |  |
| $d_{j t}^{2}$ | Residual log difference for $I_{j t}$ indices of abundance |  |  |
| $d_{C_{t}}^{2}$ | Residual log difference for catch data |  |  |
| $d_{W_{t}}^{2}$ | Residual log difference for mean weight data |  |  |
|  | Fishery reference points |  |  |
| MSY | Maximum sustainable yield (t) |  |  |
| $B_{\text {MSY }}$ | Long-term fixed spawning biomass at MSY |  |  |
| $F_{\text {MSY }}$ | Long-term fixed fishing mortality that produces MSY |  |  |
| $u_{\text {MSY }}$ | Long-term fixed harvest rate that produces MSY ( $1-e^{-F_{\text {MSY }}}$ ) |  |  |
| $B_{\text {avg }}$ | Average spawning biomass (t) over a specified number of years (1967-2016) |  |  |
| $B_{\text {min }}$ | Minimum annual spawning biomass ( t ) from which the biomass recovered to $B_{\text {avg }}$ |  |  |

Table E.2. Summary of equilibrium equations for the delay-difference model.

| Description | Equations |  |
| :---: | :---: | :---: |
| Initialization at equilibrium with $F=0$ |  |  |
| Unfished survival | $S_{0}=e^{-M}$ | (E.5) |
| Unfished mean weight | $\bar{W}_{0}=\frac{S_{0} \alpha_{g}+W_{k}\left(1-S_{0}\right)}{1-\rho_{g} S_{0}}$ | (E.6) |
| Unfished numbers | $N_{0}=\frac{R_{0}}{1-S_{0}}$ | (E.7) |
| Unfished biomass | $B_{0}=N_{0} \bar{W}_{0}$ | (E.8) |
| Recruitment compensation ratio (CR) | $\chi=\frac{4 h}{1-h}$ | (E.9) |
| Stock-recruit parameters | $a=\chi \frac{R_{0}}{B_{0}} ; \quad b=\frac{\chi-1}{B_{0}}$ | (E.10) |
| Initialization at equilibrium with $F_{e}>0$ |  |  |
| Survival at $F_{e}$ | $S_{e}=e^{-\left(M+F_{e}\right)}$ | (E.11) |
| Mean weight at $F_{e}$ | $\bar{W}_{e}=\frac{S_{e} \alpha_{g}+W_{k}\left(1-S_{e}\right)}{1-\rho_{g} S_{e}}$ | (E.12) |
| ${ }^{1}$ Biomass at $F_{e}$ | $B_{e}=-\frac{\left(-\bar{W}_{e}+S_{e} \alpha_{g}+S_{e} \rho_{g} \bar{W}_{e}+W_{k} a \bar{W}_{e}\right)}{b\left(-\bar{W}_{e}+S_{e} \alpha_{g}+S_{e} \rho_{g} \bar{W}_{e}\right)}$ | (E.13) |

Fisheries reference points at equilbrium fishing mortality $F_{e}$

Fishing mortalities
Years to equilibrium
Biomass
Numbers
Survival
Long-term yield
MSY
Biomass at MSY
Fishing mortality at MSY
$B_{\gamma t}=S_{\gamma, t-1} \rho_{g} B_{\gamma, t-1}+\alpha_{g} N_{\gamma, t-1}+W_{k} R_{\gamma t}$
$N_{\gamma t}=S_{\gamma, t-1} N_{\gamma, t-1}+R_{\gamma t}$
$S_{\gamma t}=e^{-(M+\gamma)}$
$\gamma=\{0.01,0.02, \ldots, 1.0\}$
$t=\{2017, \ldots, T\}$, where $T=2017+200$
$Y_{\gamma T}=\left(1-e^{-\gamma}\right) B_{\gamma T}$
$Y_{e}=\max \left\{Y_{\gamma T}\right.$
$B_{e}=B_{\gamma T}, \quad$ for $\gamma$ when $Y_{e}=Y_{\gamma T}$
$F_{e}=F_{\gamma T}, \quad$ for $\gamma$ when $Y_{e}=Y_{\gamma T}$
${ }^{1}$ Steven Martell (Sea State Inc., Seattle WA, pers. comm.)

Table E.3. Time-dynamic equations and likelihood components for the delay-difference model.
Description Equations

Time-dynamic equations
Survival rate

$$
\begin{equation*}
S_{t}=e^{-M+F_{t}} \tag{E.21}
\end{equation*}
$$

Biomass

$$
\begin{equation*}
B_{t}=S_{t-1}\left(\alpha_{g} N_{t-1}+\rho_{g} B_{t-1}\right)+W_{k} R_{t} \tag{E.22}
\end{equation*}
$$

Recruits

$$
\begin{equation*}
R_{t}=R_{0} e^{\omega_{t}-0.5 \sigma_{R}^{2}} \tag{E.23}
\end{equation*}
$$

## Predicted variables used in objective function

| Predicted catch | $\widehat{C}_{t}=B_{t} \frac{F_{t}}{\left(F_{t}+M\right)}\left(1-e^{-\left(F_{t}+M\right)}\right)$ |
| :--- | :--- |
| Predicted mean weight | $\widehat{\bar{W}}_{t}=\frac{B_{t}}{N_{t}}$ |
| Predicted recruits | $\widehat{R}_{t}=\frac{a B_{t-k+1}}{1+b B_{t-k+1}}$ |

Table E.4. Calculation of variance parameters, residuals, and likelihoods.

## Description Equations

Variance parameters (SD = standard deviation)
SD of abundance index residuals
$\sigma_{O}=\sqrt{\frac{\rho}{\vartheta}}=\sqrt{\rho} \cdot \varphi$
SD of recruitment residuals
$\sigma_{R}=\sqrt{\frac{1-\rho}{\vartheta}}=\sqrt{1-\rho} \cdot \varphi$
SD of abundance index observations
$\sigma_{j t}=\frac{I_{j t}}{v_{j t}} ; \quad$ where $v_{j t}=\frac{1}{c_{j t}}$
Indices of abundance
Residuals
$z_{j t}=\log \left(I_{j t}\right)-\log \left(q_{j}\right)+\log \left(\widehat{B}_{j t}\right)$
$\bar{z}_{j}=\frac{1}{n_{j}} \sum_{t=1}^{n_{j}} z_{j t}$
$d_{j t}=z_{j t}-\bar{z}_{j}$

Log likelihood
$L_{j t}=\log \left(\sigma_{j t}^{2}\right)+\frac{d_{j t}^{2}}{2 \sigma_{j t}^{2}}$

## Catch

Residuals
$d_{C_{t}}=\log \left(C_{t}\right)-\log \left(\widehat{C}_{t}\right)$
Log likelihood

$$
\begin{equation*}
L_{t}=\log \left(\sigma_{C}^{2}\right)+\frac{d_{C_{t}}^{2}}{2 \sigma_{C}^{2}} \tag{E.34}
\end{equation*}
$$

## Mean weight

Residuals
$d_{W_{t}}=\log \left(\bar{W}_{t}\right)-\log \left(\widehat{\bar{W}}_{t}\right)$
Log likelihood
$L_{t}=\log \left(\sigma_{W}^{2}\right)+\frac{d_{W_{t}}^{2}}{2 \sigma_{W}^{2}}$

## Recruitment

Residuals
$d_{R_{t}}=\log \left(R_{t}\right)-\log \left(\widehat{R}_{t}\right)$
Log likelihood
$L_{t}=\log \left(\sigma_{R}^{2}\right)+\frac{d_{R_{t}}^{2}}{2 \sigma_{R}^{2}}$

## E.4. REFERENCES - MODEL EQUATIONS

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## APPENDIX F. MODEL RESULTS

## F.1. INTRODUCTION

This appendix describes the results from the application of a delay-difference model (see Appendix E for equations, see Schnute 1985 for model framework) to a suite of survey data (Appendix B), CPUE series (Appendix C) and time series of mean weight estimates (Appendix D) pertaining to two biological stocks of BC Walleye Pollock (WAP). The modelling was done in two stages: initially the best fit to the data was found by finding the mode of the posterior distribution (MPD), which minimised the negative log-likelihood when fitting to the data including the prior penalties (Eqs. E.33, E.35, E. 37 and E.39). Fits to the data, model estimates and comparative negative log-likelihoods are presented for all models. A Bayesian second step was adopted, with a Markov chain Monte Carlo (MCMC) search across plausible parameter combinations initialised from the MPD fit. MCMC diagnostics are presented, along with estimated parameters and derived parameters for 12 model runs for the BC North stock and 11 model runs for the BC South stock. All final advice and major outputs are based on the MCMC results. Estimates of major quantities and advice to management (such as decision tables) appearing here are also presented in the main document.
Biological data (weights converted from lengths) from DFO sources suggested that two stocks of Walleye Pollock occur along the outer coast of BC, with fish on average twice as large in the north (mainly Dixon Entrance - $1.056 \mathrm{~kg} /$ fish) compared to those in the south (beginning in the lower parts of Hecate Strait $-0.521 \mathrm{~kg} / \mathrm{fish}$ ). This information is presented in detail in Appendix D . The BC North stock probably represents the southern extreme of a larger SE Alaska population, which may provide a rescue effect for any over-harvesting in BC (Gustafson et al. 2000).
Consequently, biomass levels for the BC North stock appear to be relatively small and variable. In contrast, the BC South stock is estimated to be much larger ( $\sim 10 x$ ) than the BC North stock, and has two apparent main population centres - Juan de Fuca Strait and Queen Charlotte Strait.

No reliable DFO data (e.g., ages by broken/burnt otoliths) exist to estimate growth models for either stock; therefore, these stock assessment models use a growth model estimated from eastern Gulf of Alaska (GoA) survey data (Martin Dorn, NOAA Fisheries, Sand Point, Seattle, pers. comm.) for the BC North stock and a published growth model based on data from the Sea of Okhotsk (Janusz and Horbowy 1997), which lies between the eastern Russian mainland and Kamchatka Peninsula, just west of the Bering Sea, for the BC South stock. The growth models in Saunders et al. (1989) for the west coast Vancouver Island and the Strait of Georgia were based on ages derived from fin ray sections, with the published growth rate coefficients not capable of matching the observed mean lengths, especially at higher ages of selectivity ( $k>3$ ), because of the extremely fast growth at young ages (see Figures D. 10 and D.11).
Model runs are reported for each stock which span a range of fixed values for $M \in\{0.25,0.30,0.35\}$ and $k \in\{3,4,5\}$. We selected these values to include the most plausible values for these important parameters. Some other sensitivities were also tried, such as excluding the GIG survey (BC North) and the commercial fishery CPUE series.

This range of model exploration was undertaken because there is substantial uncertainty in specifying the productivity of this stock (as represented by $M$ and the growth model), as well as selecting the age at full knife-edge recruitment ( $k$ ), assumptions that are mandatory when using a delay difference model. Because the available data are not informative with respect to these key model parameters, it is not possible to objectively rule out these alternate hypotheses. Initially, after covering a range of plausible values for the key parameter assumptions, we chose model
runs for advice based on a subjective ranking of the MCMC diagnostics ( $1=$ good, 2 = acceptable, $3=$ poor), selecting only those runs that ranked $\leq 2$ (using the mean ranking across both authors). The regional peer review (RPR) participants were concerned about runs with high values of $F_{\max }$, and so choices based on $F$ were first identified before considering the MCMC diagnostics:

- use model runs where median $F_{\max }$ across MCMC samples is $<2$ (three runs each in BC North and BC South);
- add model runs where median annual $F_{t}$ is $>2$ only once (one run in BC North, three in BC South);
- remove one model run from BC North with poor diagostics;
- keep two BC South model runs with diagnostics flagged as poor but ranked as acceptable by at least one of the authors.

This stock assessment adopted a "Model Averaging" (MA) approach, using the selection process above to yield three model runs for BC North and six for BC South, that represent a range of plausible hypotheses to construct a "Model Average Composite" for each stock to provide advice to managers (Sections F.2.3. and F.3.3.).

## F.1.1. Historical Reference Points

The Sustainable Fisheries Framework (SFF, DFO 2009) established provisional reference points to guide management and assess harvest in relation to sustainability. These reference points are the Limit Reference Point (LRP, limit below which biological harm occurs) of $0.4 B_{\text {MSY }}$ and the upper stock reference point (USR, limit at which management needs to consider conservation action) of $0.8 B_{\text {MSY }}$, which have not been adopted in this assessment due to concerns about the stability of estimating $B_{0}$ and $B_{2017}$ using the $i \mathrm{SC} \forall \mathrm{M}$ delay-difference model (see Appendix E for discussion). In their stead, this assessment adopted historical reference points (HRPs): $B_{\text {avg }}$ (average spawning biomass from 1967-2016) as a proxy for $B_{\mathrm{MSY}}$, and $B_{\text {min }}$ (spawning biomass in the year when the reconstructed biomass reached a minimum from which it subsequently recovered to $B_{\text {avg }}$ ) in place of $0.4 B_{\mathrm{MSY}}$. The term "spawning biomass" used in this Walleye Pollock assessment is interchangeable with "mature exploitable biomass (males and females)".

The determination of $B_{\min }$ required an algorithm that could be applied to each MCMC sample (matrix row) or MPD vector:

1. Calculate $B_{\text {avg }}$ for years spanning 1967-2016;
2. Identify a set of candidate $B_{\min }$ points using the 0.005 quantile;
3. For consecutive-year lows (if any), identify the lowest point to represent the group by the low year;
4. Remove any candidate low that occurs in the final year;
5. Using these candidates (+ the final year) as break points, create vectors of $B_{t}$ between and including the break points;
6. Determine which vectors of $B_{t}$ increase to equal or exceed $B_{\text {avg }}$;
7. Of the successful candidates in the previous step, choose the one that started from the minimum $B_{t}$ to represent $B_{\text {min }}$;
8. If none of the $B_{t}$ vectors reach $B_{\text {avg }}$, increase the quantile for candidate selection and repeat the above from step 2 until a valid $B_{\min }$ is found.

As the assessment uses $B_{\text {avg }}$ as a recovery point from a low at $B_{\text {min }}$, there is little likelihood that a recovery to $B_{\text {avg }}$ cannot be identified. However, certain trajectories, like a constantly decreasing population from $B_{0}$, will not yield an identifiable $B_{\text {min }}$ using the above algorithm.
In this assessment, the following reference points are used to determine the probabilty of projected $B_{t}$ exceeding them:

- Current spawning biomass $B_{2017}$
- Limit Reference Point (LRP): $B_{\text {min }}$
- Upper Stock Reference (USR): $2 B_{\text {min }}$
- Average spawning biomass: $B_{\text {avg }}$ (average over the years 1967-2016)
- Average harvest rate: $u_{\text {avg }}$ (average over the years 1967-2016)


## F.2. BC NORTH STOCK

## F.2.1. Example Case - North

An example model run for $B C$ North stock is presented to show representative detail in the results. This model is based on a growth model derived from eastern Gulf of Alaska data supplied by Martin Dorn (pers. comm.) and included the following elements:

- instantaneous natural mortality $M$ fixed at 0.30 ;
- knife-edge recruitment at age $k=3 \mathrm{y}$;
- steepness $h$ beta prior (mean=0.7, SD=0.15);
- 1973-2016 standardised unsorted mean weights ( $\bar{w}=1.05594 \mathrm{~kg}$ );
- length-weight allometry: $\alpha=7.1018 \mathrm{E}-06, \beta=3.0415$;
- Brody parameters: $\alpha_{g}=0.3475 \mathrm{~kg}, \rho_{g}=0.8668, w_{k}=0.4929 \mathrm{~kg}$;
- growth parameters: $L_{\infty}=66.9436 \mathrm{~cm}, \kappa=0.211778, t_{0}=-1.13642 \mathrm{y}$;
- errors: observation $\sigma_{O}=0.2$, recruitment $\sigma_{R}=0.6$, mean weight $\sigma_{W}=0.15$;
- uniform priors on $q$ from -10 to 0;
- Walleye Pollock CPUE indices (uniform -10 to 0);
- catch series (GFFOS accessed 2016-09-12);
- equilibrium start in 1967 (use all of GIG historical, including 1995)
- equal-weight for each age class when estimating von Bertalanffy model parameters by sex, combined-sex model interpolated between sexes;
- estimate $\ln \left(R_{0}\right)$ (uniform prior) and fix $\ln (\bar{R})$ and $\ln \left(R_{\text {init }}\right)$ to $\ln \left(R_{0}\right)$;
- version iscam-delaydiff. exe: built 2017-01-03.


## F.2.1.1. MPD results - example (north)

The mode of the posterior distribution (MPD) for this model (reported in Table F.1) is estimated by minimising the objective function (components summarised in Section E.2.1). The results are presented to show the fits by the model to the observed data and are used as the starting point for the MCMC simulations. MPD fits are shown for the abundance indices (Figure F.1), the annual mean weights and annual recruitment (Figure F.2). The fits to the survey and CPUE indices are generally reasonable although the model is incapable of fitting the abrupt changes in some series.

The model is not capable of fitting the high values of mean weight that occur throughout the series (Figure F.2). Instead it fluctuates near the mean of the series, effectively ignoring four or five recurring upward shifts in the data. Recruitment events exceed the long-term mean of the series 13 times; the 1974 event is roughly seven times higher than the long-term mean. Fits to the catch data are not presented because the model is parameterised so that it always fits the catch closely.

## F.2.1.2. MCMC results - example (north)

The MCMC procedure performed 60,000,000 iterations, sampling every 50,000 to give 1200 draws ( 1000 samples after dropping the first 200, including the MPD start point, as burn-in). The 1000 samples were used to estimate parameters and quantities of interest, including stock status by year and the probabilities of being above reference points.
MCMC traces show good convergence properties (no trend with increasing sample number) for the leading estimated parameters (Figure F.3), as does a diagnostic analysis that splits the samples into three segments, checking for consistency along the length of the chain (Figure F.4). Autocorrelation appears to be minimal with some periodicty over 100 lags (Figure F.5). Pairs plots of the estimated parameters (Figure F.6) show no undesirable correlation between the two primary parameters, $\ln \left(R_{0}\right)$ and $h$, though all the $q$ parameters were highly correlated with $\ln \left(R_{0}\right)$ and with each other, as would be expected. MCMC quantiles for parameters, biomass, and status with respect to historical reference points are summarised in Table F.1.

Marginal posterior distributions, along with the corresponding priors for the estimated parameters, are shown in Figure F.7. Only the steepness parameter used an informative prior, with its posterior distribution largely reflecting the prior. This indicates that there was relatively little information in this model to inform this parameter and it is unlikely that it could be estimated without using a prior.

The plot of estimated spawning biomass (Figure F.8) shows a large increase in the mid 1970s followed by a decline to a low point in 1986. Since then, spawning biomass has fluctuated at levels below the average biomass calculated from 1967-2016. Assuming a catch policy of $1000 \mathrm{t} / \mathrm{y}$, which is close to the $5-\mathrm{y}$ average catch and lower than the current TAC of $1320 \mathrm{t} / \mathrm{y}$, the projected biomass declines precipitously under conditions of average recruitment.
This model run estimates a few strong recruitment pulses in 1974, 1996, and 2012 (Figure F.9), likely to fit drops in mean weight. The first strong recruitment occurs in absence of fishing pressure and so spawing biomass increases substantially (Figure F.8). Thereafter, strong fishing pressure dampens any benefits from strong recruitment (Figure F.10). Fishing mortality peaks in 1993 at a median $F$ value of $0.652 \mathrm{y}^{-1}$ and declines thereafter until it reaches $0.334 \mathrm{y}^{-1}$ in 2016. The median value (and the 5th and 95th percentiles in parentheses) for the estimated level of
biomass depletion $\left(B_{t} / B_{\text {avg }}\right)$ at the end of the final year of the reconstruction is $0.57(0.35,0.83)$, with the MPD value of 0.52 lying close to the median of the posterior distribution of this quantity (Figure F.10).

The use of historical reference points is illustrated in Figure F.11. Under an assumed catch policy of $1000 \mathrm{t} / \mathrm{y}, B_{2019}$ lies considerably below $B_{2017}$, and both lie below $B_{\text {avg }}$. The current year biomass $B_{2017}$ lies between the LRP and the USR and the median lies just below the median USR. While the current fishing mortality rate $F_{2016}$ is higher than both the average $F$ and the minimum $F$ that occurred in 2001, it is not as high as the maximum fishing mortlaity rates experienced over the assessment time period. The phase plot (Figure F.12) also indicates that the current stock status lies between the LRP and the USR, and the current mid-year harvest rate $u_{2016}$ lies above the average harvest rate over the time series.

Table F.1. BC North: The $5^{t h}$, $50^{\text {th }}$, and $95^{t h}$ percentiles of MCMC-derived parameter estimates and quantities from 1,000 MCMC samples for the example model run. Some fixed parameters are reported as MPD only. See Appendix E for paramater definitions. Subscripts 1-4 on q refer to the fishery-independent surveys, subscript 5 refers to the commercial trawl CPUE series. Other definitions: $B_{2017}$ - biomass at the start of 2017, $u_{2016}$ - exploitation rate (ratio of total catch to vulnerable biomass) in the middle of 2016, $B_{\text {avg }}$ - average biomass from 1967 to 2016, LRP (limit reference point) - minimum median estimated biomass in the time series, USR (upper stock reference) - twice the LRP biomass. All biomass values are in tonnes. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

|  | $5 \%$ | $50 \%$ | $95 \%$ | MPD |
| ---: | ---: | ---: | ---: | ---: |
| Parameters |  |  |  |  |
| $R_{0}$ | 565 | 831 | 1,266 | 939 |
| $h$ | 0.482 | 0.736 | 0.922 | 0.788 |
| $M$ |  |  |  | 0.3 |
| $q_{1}$ | 0.132 | 0.199 | 0.291 | 0.204 |
| $q_{2}$ | 0.0905 | 0.136 | 0.202 | 0.157 |
| $q_{3}$ | 0.169 | 0.249 | 0.349 | 0.287 |
| $q_{4}$ | 0.00451 | 0.00676 | 0.0097 | 0.00789 |
| $q_{5}$ | 0.000132 | 0.000194 | 0.000275 | 0.000225 |
| $\mathrm{HRP}-$ based |  |  |  |  |
| $B_{2017}$ | 2,145 | 4,297 | 8,045 | 3,544 |
| $B_{\text {avg }}$ | 5,349 | 7,568 | 11,033 | 6,816 |
| $\mathrm{LRP}=B_{2001}$ | 1,154 | 2,333 | 3,866 | 2,249 |
| $\mathrm{USR}=2 B_{2001}$ | 2,308 | 4,665 | 7,732 | 4,498 |
| $B_{2017} / B_{\text {avg }}$ | 0.353 | 0.573 | 0.834 | 0.52 |
| $B_{2017} / B_{2001}$ | 1.19 | 1.88 | 3.12 | 1.58 |
| $B_{2017} / 2 B_{2001}$ | 0.594 | 0.942 | 1.56 | 0.788 |
| $u_{\text {avg }}$ | 0.0964 | 0.142 | 0.207 | 0.157 |
| $u_{2016} / u_{\text {avg }}$ | 1.55 | 2.01 | 2.63 | 2.12 |



Figure F.1. BC North: MPD index fits to relative abundance indices for the example model run. Circles represent observed indices with associated CVs; squares represent the model fit. Surveys: (1) GIG (Goose Island Gully) Historical, (2) HS (Hecate Strait) Assemblage, (3) HS Synoptic, (4) WCHG (west coast Haida Gwaii) Synoptic, and (5) commercial trawl catch per unit effort of Walleye Pollock.


Figure F.2. BC North: [Top] MPD fit to the mean weight data for the example model run. Predicted mean weights are shown as a red line and observations are shown as points. Error bars on mean weight observations represent a fixed CV using $\sigma_{W}=0.15$. [Bottom] MPD recruitment in thousands of age-3 individuals in year $t$ for example model run.


Figure F.3. BC North: Trace plots for MCMC output of estimated parameters in the example model run. The MCMC run shows 1000 MCMC samples after removing 200 samples. Parameters log.ro (natural log of unfished equilibrium recruitment), $h$ (steepness), and $q$ (catchability) for the surveys outlined in Figure F.1. Grey lines show the 1000 samples for each parameter, solid lines show the cumulative median (up to that sample), and dashed lines show the cumulative 2.5 and 97.5 quantiles. Red circles are the MPD estimates.


Figure F.4. BC North: Diagnostic plot for the example model run obtained by dividing the MCMC chain of 1000 samples into three segments and over-plotting the cumulative distributions of the first segment (green), second segment (red), and final segment (blue).


Figure F.5. BC North: Autocorrelation plots for MCMC output of estimated parameters in the example model run. See Figure F. 3 for parameter descriptions.


Figure F.6. BC North: Pairs plot from the northern BC region of 1000 MCMC samples for the estimated parameters of the example model run. The diagonal shows the frequency distribution of each posterior.


Figure F.7. BC North: Prior probability distributions (blue lines) used in the example model run and the comparative posterior histograms. Parameters $q_{j}$ represent catchability of the various surveys $j$ as defined in Figure F.3. The dashed red vertical lines show the MPD estimates.

## Biomass



Figure F.8. BC North: Posterior estimates of spawning biomass (1000t) for the example model run with $95 \%$ credibility intervals in grey. The current year biomass (2017, yellow point) and projected biomass (2018-2022, red line), assuming a constant catch policy of $1000 \mathrm{t} / \mathrm{y}$, are enclosed by a $95 \%$ credibility interval shaded pink. The median posterior estimate of $B_{0}$ is shown as a green point (with $95 \%$ credibility range) to the left of the time series. The MPD estimate is shown as a blue line. The total catch is shown along the bottom as red bars, with assumed TAC catch in pink.


Figure F.9. BC North: [Top] Posterior estimates of age-3 recruits for the example model run. [Bottom] Log recruitment deviations for the example model run with $95 \%$ credibility intervals.


Figure F.10. BC North: [Top] Posterior estimates of fishing mortality for the example model run. [Bottom] Biomass depletion, i.e. $B_{t} / B_{\text {avg }}$, for the example model run with $95 \%$ credibility intervals. Also displayed on the depletion figure is the MPD estimate (blue line) and the reference points $B_{\text {min }}$ (red dashed line) and $2 B_{\text {min }}$ (green dashed line).


Figure F.11. BC North: Posterior estimates of spawning biomass- and harvest-based reference points for the example model run. [Left] The current-year biomass $B_{2017}$, the projected-year biomass $B_{2019}, B_{\text {avg }}=$ average biomass from 1967 to 2016, the limit reference biomass (or $B_{\min }$ ) = biomass in 2001, and the upper stock biomass set at $2 B_{2001}$. [Right] The current-year fishing mortality rate $F_{2016}$, the average fishing mortality rate $F_{\text {avg }}$ from 1967 to 2016, the fishing mortality rate in the year of minimum biomasss $F_{2001}$, and the maximum fishing mortality experience over the time series $F_{\max }$. Box delimiter and limits represent quantiles at 0.5 (median), 0.25 and 0.75 quantiles, respectively, and the whiskers delimit the 0.05 and 0.95 quantiles. Outliers are not shown.


Figure F.12. BC North: Phase plot through time for the example model run of the medians of the ratios $B_{t} / B_{\text {avg }}$ (the biomass in year $t$ relative to $B_{\text {avg }}$ ) and $u_{t-1} / u_{\text {avg }}$ (the exploitation rate in year $t-1$ relative to $u_{\text {avg }}$ ). Blue filled circle is the starting year 1967. Years then proceed from light grey through to dark grey with the final year 2017 as a filled purple circle with limit lines represent the 0.05 and 0.95 quantiles of the posterior distributions for the final year. Vertical dashed lines indicate the historical limit (red) and upper stock reference (green) points (see legend for values), and horizontal dotted line indicates $u_{t-1}$ at $u_{\text {avg }}$.

## F.2.2. Alternative Cases - North

A necessary component to this stock assessment was testing the sensitivity of the results and the associated advice to key uncertainties in the underlying stock assessment model. Therefore, we ran a total of 12 alternative runs (Table F.2) (including the example run above), using the derived growth model from the eastern Gulf of Alaska (see Appendix D) to test the robustness of the results to uncertainties in:

- natural mortality $(M)$;
- age at knife-edge recruitment $(k)$;
- use of the GIG historical survey series; and
- use of the CPUE index series.

We tested a range of sensible values for $M$ and $k$ because these parameters control key assumptions made by the delay-difference model, given that the data available to the model are not very informative with respect to these parameters. We also tested combinations of $M$ and $k$ for the last two alternative categories (removing index series).

Although the Alaskan stock assessments use age-specific natural mortality rates for Walleye Pollock, the underlying assumption is that $M=0.30$ for age at full maturity (Dorn et al. 2015). The delay-difference model assumes that maturity matches selectivity, i.e., all recruited fish are mature, and by extension, all mature fish will have a single natural mortality rate. Alternative runs were made with $M \in\{0.25,0.30,0.35\}$ to bracket plausible values for this parameter (Table F.2) that are consistent with the Alaskan stock assessment assumption for this species.

The delay-difference assumption of knife-edge selectivity at a specific age $k$ is a strong assumption that is difficult to test without age information from the fishery. However, if such information were available, it is likely that another form of model would have been used. Dorn et al. (2012) provide a range of selectivity ogives for the GoA fisheries and surveys, with the median age selected to these commercial fisheries ranging from age 3 to age 5 (see columns 5 to 7 in Table D.7). Based on the ogives in this table, ages 3 and 4 were selected as the most likely ages to use for the age of knife-edge recruitment in the BC North Walleye Pollock delay-difference model (Table F.2).

The parameterisation of the $i \mathrm{SC} \forall \mathrm{M}$ delay-difference model combines observation and process error into a single total variance parameter $\vartheta$. This variable is partitioned into observation and process error components through the parameter $\rho$ (see Equations E. 27 and E.28).
Experimentation with a higher value for $\rho$ when using the same model to assess Shortspine Thornyhead (Starr and Haigh 2017) showed that increasing the error term led to greater uncertainty but did not appreciably affect the overall conclusions. Consequently, it was decided to leave out sensitivities to this component of the model for this assessment. These parameters were fixed in all model runs such that the overall observation error ( $\sigma_{O}$ ) was 0.2 and the recruitment process error ( $\sigma_{R}$ ) was 0.6 , the latter being a common value used as a default for teleost finfish. A further variance component $\sigma_{W}$ sets the weight used to fit the mean weight observations, which was fixed at $\sigma_{W}=0.15$ for all model runs to ensure a strong fit to the mean weight series.
Five of the alternative runs (including the example case) used all the abundance index series and paired $k=3$ with $M \in\{0.25,0.30,0.35\}$ and $k=4$ with $M \in\{0.30,0.35\}$. Five alternative runs dropped the GIG historical survey and paired $k=3$ with $M \in\{0.25,0.30\}$ and $k=4$ with
$M \in\{0.25,0.30,0.35\}$. This was done because, strictly speaking, the GIG survey did not operate in the stock definition area for the BC North stock. However, because this series is the only set of early abundance information, two blocks of runs were made, one with and one without this survey. The final two sensitivities dropped the CPUE index series as well as the GIG survey series and paired $M=0.3$ with $k \in\{3,4\}$. Again, this was done because there is uncertainty as to whether fishery-dependent data track abundance.

All 12 alternative runs were taken to the MCMC level. Each MCMC search was started at the "best fit" MPD parameter set, run for 60 million iterations, and sampled every 50,000 for 1200 samples. The first 200 samples were dropped as burn-in to yield a total posterior sample of 1000 draws.

## F.2.2.1. MPD results - alternatives (north)

This large number of alternative runs shows that the data available to this model do not allow much discrimination between the range of hypotheses tested. While there were differences in the fits to the available biomass indices and to the mean weight data, these differences tended to be small and probably could not distinguish between hypotheses except in the most extreme cases. As an example, none of the models were able to fit the high mean weights observed in the late 1970s or near the end of the time series (Figure F.13), which may be attributable to misspecification of the growth model. While the delay-difference model uses the mean weight data to scale the overall biomass and to obtain recruitment deviation information, it is likely that only an age-structured model would have sufficient flexibility to fit the entire mean weight series. The model also did not fit high points in the HS Assemblage survey or in the CPUE times series (Figures F. 14 and F.17). However, despite these difficulties, the overall fit to most abundance indices was acceptable.

In an attempt to make such comparisons more quantitative, the negative log-likelihoods for the fits to each data component of the model are summarised for the 12 alternative runs in Table F.3. Comparisons of the whole model can only be made among those runs that shared the same estimated components:

- Cases S00 to S04 - S04 had the lowest objective function value of these five cases and offered the best model fit in the first group;
- Cases S05 to S09 - S09 offered the best fit in the second group; and
- Cases S10 to S11 - S11 offered the best fit in the third group.

The three best-fit models all featured $k=4$ and $M \in\{0.35,0.30\}$, which likely reflects the growth model used for this population of large Walleye Pollock. Case S11 also showed the best fit to the mean weight component, and most of the $k=4$ models showed better fits than those where $k=3$. This observation tended to be true for the survey components as well.

Table F.2. Summary of the analyses performed to test the sensitivity of the delay-difference model to variations in natural mortality $M$, knife-edge recruitment age $k$. All runs for the BC North stock use the eastern Gulf of Alaksa growth function (Martin Dorn, pers.comm.). The column marked 'Rank' provides a subjective ranking of the MCMCs, where 1 = good, 2 = acceptable, and $3=$ poor.

| Case | Run ID Run \# |  | $M$ | $k$ | Rank |
| :--- | :---: | ---: | ---: | ---: | ---: |
| S00 | M.30+k3 | 1 | 0.3 | 3 | 1.25 |
| Four surveys | CPUE time series |  |  |  |  |
| S01 | M.25+k3 | 16 | 0.25 | 3 | 3 |
| S02 | M.30+k4 | 4 | 0.3 | 4 | 2.5 |
| S03 | M.35+k3 | 2 | 0.35 | 3 | 1.5 |
| S04 | M.35+k4 | 12 | 0.35 | 4 | 1.5 |
| Remove the GIG survey |  |  |  |  |  |
| S05 | M.25+k3-GIG | 9 | 0.25 | 3 | 3 |
| S06 | M.25+k4-GIG | 7 | 0.25 | 4 | 2 |
| S07 | M.30+k3-GIG | 3 | 0.3 | 3 | 3 |
| S08 | M.30+k4-GIG | 5 | 0.3 | 4 | 1.5 |
| S09 | M.35+k4-GIG | 10 | 0.35 | 4 | 1.25 |
| Remove both GIG survey and CPUE |  |  |  |  |  |
| S10 | M.30+k3-GIG-CPUE | 8 | 0.3 | 3 | 2 |
| S11 | M.30+k4-GIG-CPUE | 6 | 0.3 | 4 | 2 |

Table F.3. BC North MPD negative log likelihoods from the 12 alternative cases documented in Table F. 2 for each data component used in the model.

| Case | Catch | GIG <br> Hist | HS <br> Assem | HS <br> Synop | WCHG <br> Synop | North <br> CPUE | Recruits | Mean <br> Weight | ObFn <br> Value |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| S00: M.30+k3 | -101.355 | 19.734 | 48.086 | -0.694 | 7.464 | 9.000 | 80.902 | -1.100 | 198.860 |
| S01: M.25+k3 | -101.135 | 22.087 | 47.223 | -0.354 | 7.910 | 11.075 | 82.006 | -1.465 | 204.633 |
| S02: M.30+k4 | -101.465 | 23.916 | 32.534 | -2.655 | 2.898 | 4.975 | 81.486 | -6.832 | 171.551 |
| S03: M.35+k3 | -101.532 | 17.715 | 48.824 | -0.982 | 7.038 | 7.171 | 80.259 | -0.596 | 194.482 |
| S04: M.35+k4 | -101.535 | 22.377 | 30.236 | -2.781 | 2.494 | 5.077 | 80.540 | -7.329 | 165.332 |
| S05: M.25+k3-GIG | -101.116 | - | 40.897 | -0.115 | 7.654 | 11.390 | 72.901 | -0.288 | 164.739 |
| S06: M.25+k4-GIG | -101.336 | - | 32.827 | -2.160 | 3.413 | 4.403 | 72.644 | -4.646 | 138.121 |
| S07: M.30+k3-GIG | -101.322 | - | 42.312 | -0.425 | 7.229 | 9.237 | 72.718 | 0.059 | 163.199 |
| S08: M.30+k4-GIG | -101.433 | - | 30.338 | -2.325 | 3.022 | 4.311 | 72.005 | -5.972 | 132.711 |
| S09: M.35+k4-GIG | -101.510 | - | 28.094 | -2.469 | 2.626 | 4.319 | 71.733 | -6.578 | 128.908 |
| S10: M.30+k3-GIG-CPUE | -101.389 | - | 26.508 | 1.652 | 3.586 | - | 72.920 | -6.654 | 130.290 |
| S11: M.30+k4-GIG-CPUE | -101.202 | - | 23.382 | -0.348 | 0.452 | - | 71.339 | -7.695 | 118.341 |



Figure F.13. BC North: MPD fit to the mean weight data for the 12 alternative runs. Predicted mean weights are shown as red lines and observations are shown as points. Error bars on mean weight observations represent a fixed CV using $\sigma_{W}=0.15$.


Figure F.14. BC North: MPD index fits (12 alternative runs) for the HS (Hecate Strait) assemblage survey relative abundance indices. Circles represent observed indices with associated CVs, triangles represent the model fit.


Figure F.15. BC North: MPD index fits (12 alternative runs) for the HS (Hecate Strait) synoptic survey relative abundance indices. Circles represent observed indices with associated CVs, triangles represent the model fit.


Figure F.16. BC North: MPD index fits (12 alternative runs) for the WCHG (west coast Haida Gwaii) synoptic survey relative abundance indices. Circles represent observed indices with associated CVs, triangles represent the model fit.


Figure F.17. BC North: MPD index fits (12 alternative runs) for the commercial trawl Walleye Pollock CPUE relative abundance indices. Circles represent observed indices with associated CVs, triangles represent the model fit.

## F.2.2.2. MCMC results - alternatives (north)

Median estimates for current biomass $B_{2017}$ lie below the average biomass $B_{\text {avg }}$ for all alternative cases except the two that remove the CPUE and GIG index series (Table F.4). Additionally, the year in which median spawning biomass reached a minimum is not stable across the suite of runs. The maximum fishing mortality $F$ estimated in the MCMC samples exceeded realistic levels $(F>2)$ in some years for all but three cases. This is likely to be the result of the failure of the knife-edge recruitment assumption (particularly when it is $k=4$ ), which leaves too few fish in the population in some years to support the observed level of catch. The alternative cases do not model population trajectories consistently such that the minimum biomass occurs in the same year.

Table F. 5 shows for each alternative run the probabilities of projected biomass in two years at $1000 \mathrm{t} / \mathrm{y}$ (at the level of recent average catch) exceeding various reference points. Only runs S10 and S11 (which omit the CPUE series) show high probabilities of $B_{2019}$ exceeding the limit reference point $B_{\text {min }}$. All other scenarios are fairly pessimistic, although it should be noted that 2 -year projections using this delay-difference model are unreliable and uncertain. That is because these models, like a surplus production model and unlike an age-structured model, only project using the stock-recruitment function, which has little predictive power.

Most of the trace plots for $R_{0}$ look acceptable but a few show an increasing trend in the median while others show large-scale shifts in the trace mean (Figure F.18). Autocorrelation plots for $R_{0}$ highlight other problems (Figure F.19), with strong significant positive serial corrleation in S01, S 05 , and S07, all of which feature $k=3$. The ranks assigned to the quality of MCMCs appear in Table F.2. While there is no statistical basis for selecting among these hypotheses, these MCMC results suggest that a selection of the runs with the best diagnostics can be used to model this Pollock stock. Consequently, this stock assessment used the subjective MCMC quality rankings to construct a Model Average posterior to provide advice to managers (Section F.2.3.), with a rank $\leq 2$ used as the quality cutoff criterion.
Quantile plots compare $B_{\text {avg }}, B_{2017} / B_{\text {avg }}, u_{2016} / u_{\text {avg }}$ and $u_{\text {avg }} B_{\text {avg }}$ of the example run to alternative runs grouped by category:

- $M \in\{0.25,0.30,0.35\}$ and $k \in\{3,4\}$ using all abundance indices (Figure F.20);
- $M \in\{0.25,0.30,0.35\}$ and $k \in\{3,4\}$ after dropping the GIG survey (Figure F.21); and
- $M=0.30$ and $k \in\{3,4\}$ after dropping both the GIG survey and the CPUE series (Figure F.22).

These plots show that $M$ and $k$ interact to change the perceived size of the stock (e.g, $B_{\text {avg }}$ ), but the estimated stock status $B_{2017} / B_{\text {avg }}$ is relatively consistent across these runs (Figure F.20). Removing the GIG survey abundance index estimates a lower average biomass ( $B_{\text {avg }}$ ) with higher mean removals relative to average biomass ( $u_{\text {avg }} B_{\text {avg }}$ ); however, estimated stock status ( $B_{2017} / B \mathrm{avg}$ ) remains similar across all these runs and to that of the example case (Figure F.21). Removing the GIG survey and the commercial CPUE series does change estimated stock status (higher than that for the example case) even though average removal rates are a lot higher (Figure F.22).

Table F.4. BC North median values for select MCMC-derived parameters and quantities for 12 alternative runs from the BC North stock. The value for $B_{2019}$ is that assuming a TAC of 1000 tyy. Model senstivity details appear in Table F.2.

| Run | $h$ | $B_{\text {avg }}$ | $\frac{B_{2017}}{B_{\text {avg }}}$ | $\frac{B_{2019}}{B_{\text {avg }}}$ | $\mathrm{Yr}_{\min }$ | $\frac{B_{2017}}{B_{\min }}$ | $F_{\max }$ | $u_{\text {avg }}$ | $\frac{u_{2016}}{u_{\text {avg }}}$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| S00: M.30+k3 | 0.74 | 7,568 | 0.57 | 0.29 | 2001 | 1.9 | 0.71 | 0.14 | 2.0 |
| S01: M.25+k3 | 0.75 | 4,962 | 0.48 | 0.12 | 1986 | 2.3 | 1.7 | 0.21 | 2.0 |
| S02: M.30+k4 | 0.78 | 3,377 | 0.44 | 0.11 | 1986 | 3.1 | 19 | 0.35 | 1.7 |
| S03: M.35+k3 | 0.73 | 10,270 | 0.60 | 0.36 | 2001 | 1.9 | 0.51 | 0.11 | 2.0 |
| S04: M.35+k4 | 0.77 | 3,695 | 0.46 | 0.13 | 1986 | 3.1 | 16 | 0.33 | 1.6 |
| S05: M.25+k3-GIG | 0.78 | 4,628 | 0.54 | 0.27 | 1986 | 2.7 | 10 | 0.26 | 1.8 |
| S06: M.25+k4-GIG | 0.81 | 3,272 | 0.51 | 0.24 | 1986 | 2.9 | 19 | 0.36 | 1.6 |
| S07: M.30+k3-GIG | 0.77 | 5,325 | 0.65 | 0.42 | 1986 | 3.7 | 8.0 | 0.24 | 1.6 |
| S08: M.30+k4-GIG | 0.80 | 3,438 | 0.58 | 0.30 | 1986 | 3.5 | 19 | 0.35 | 1.5 |
| S09: M.35+k4-GIG | 0.80 | 3,725 | 0.62 | 0.36 | 1986 | 3.5 | 17 | 0.34 | 1.5 |
| S10: M.30+k3-GIG-CPUE | 0.75 | 6,986 | 1.3 | 0.90 | 2000 | 9.9 | 18 | 0.20 | 0.80 |
| S11: M.30+k4-GIG-CPUE | 0.80 | 4,248 | 1.1 | 0.74 | 1986 | 8.9 | 20 | 0.31 | 0.87 |

Table F.5. BC North: Assuming a constant catch policy of 1000 ty, the probability that $B_{2019}$ (or $u_{2018}$ ) is greater than reference points used in this assessment for 12 alternative runs. Model senstivity details appear in Table F.2. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

| Run | $\mathrm{P}\binom{B_{2019>}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019>}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| :--- | ---: | ---: | ---: | ---: | ---: |
| S00: M.30+k3 | 0 | 0.46 | 0.03 | 0 | 1 |
| S01: M.25+k3 | 0 | 0.39 | 0.12 | 0 | 1 |
| S02: M.30+k4 | 0.05 | 0.46 | 0.25 | 0 | 0.99 |
| S03: M.35+k3 | 0.01 | 0.66 | 0.06 | 0 | 1 |
| S04: M.35+k4 | 0.04 | 0.52 | 0.28 | 0 | 0.99 |
| S05: M.25+k3-GIG | 0.07 | 0.64 | 0.35 | 0.01 | 0.98 |
| S06: M.25+k4-GIG | 0.12 | 0.66 | 0.36 | 0.02 | 0.94 |
| S07: M.30+k3-GIG | 0.08 | 0.87 | 0.59 | 0.03 | 0.94 |
| S08: M.30+k4-GIG | 0.12 | 0.75 | 0.48 | 0.04 | 0.91 |
| S09: M.35+k4-GIG | 0.14 | 0.81 | 0.53 | 0.05 | 0.90 |
| S10: M.30+k3-GIG-CPUE | 0.03 | 1 | 0.99 | 0.37 | 0.22 |
| S11: M.30+k4-GIG-CPUE | 0.06 | 0.98 | 0.92 | 0.25 | 0.44 |



Figure F.18. BC North: Trace plots (12 alternative runs) for MCMC samples of $\log \left(R_{0}\right)$ (natural log of unfished equilibrium recruitment). The MCMC run had chain length 60 million and a sample taken at every $50,000^{\text {th }}$ iteration to yield 1,000 MCMC samples after a removing a burn-in of 200 samples. Grey lines show the 1000 samples for each parameter, solid lines show the cumulative median (up to that sample), and dashed lines show the cumulative 2.5 and 97.5 quantiles. Red circles are the MPD estimates.


Figure F.19. BC North: Autocorrelation plots (12 alternative runs) for MCMC samples of $\log \left(R_{0}\right)$ (natural $\log$ of unfished equilibrium recruitment). The MCMC runs had 1,000 MCMC samples each.


Figure F.20. BC North: Quantile plots comparing the example model case (S00 with $M=0.30$ and $k=3$ using the eastern Gulf of Alaska growth model) to alternative runs that vary natural mortality, where $M=$ $0.25,0.30,0.35$, and knife-edge recruitment age $k$, where $k=3,4$. Box delimiter and limits represent quantiles at 0.5 (median), 0.25 and 0.75 quantiles, respectively, and the whiskers delimit the 0.05 and 0.95 quantiles. Outliers are not shown.


Figure F.21. BC North: Quantile plots comparing the example reference case $S 00$ to alternative runs that do not use the GIG (Goose Island Gully) historical surveys. Quantile box delimiters are detailed in Figure F. 20.


Figure F.22. BC North: Quantile plots comparing the example reference case SOO to alternative runs that do not use the GIG survey series or the commercial trawl CPUE series. Quantile box delimiters are detailed in Figure F. 20.

## F.2.3. Model Average Composite - North

Three alternative BC North runs were selected for inclusion to the Model Average posterior based on the following criteria (see Tables F. 2 and F.4):

- use model runs where the median $F_{\max }$ across MCMC samples was <2;
- add model runs where median annual $F_{t}$ was $>2$ only once;
- remove model runs with poor diagnostics (rank > 2).

For the BC North stock, these criteria selected 3 out of 12 models:

- S00: M.30+k3 (rank=1.25, med. $F_{\max }=0.71$, no.yrs med. $F_{t}>2=0$ )
- S03: M.35+k3 (rank=1.50, med. $F_{\max }=0.51$, no.yrs med. $F_{t}>2=0$ )
- S10: M.30+k3-GIG-CPUE (rank=2.00, med. $F_{\max }=18.4$, no.yrs med. $F_{t}>2=1$ )

Table F. 6 gives the model-based and HRP-based quantities ( $0.05,0.50$, and 0.95 quantiles) from the model average posterior based on 3000 pooled MCMC samples. Table F. 8 gives the decision table for this Model Average Composite scenario, showing the probabilities that $B_{2019}$ will exceed various reference points. Figure F. 23 shows the stock status $B_{2017} / B_{\text {avg }}$ of the Model Average Composite scenario and the 3 models that contribute to the composite model. Finally, Tables F.9-F. 11 show the $2-\mathrm{y}$ decision tables for scenarios contributing to the composite.

Table F.6. BC North: The $5^{\text {th }}, 50^{\text {th }}$, and $95^{\text {th }}$ percentiles of MCMC-derived quantities from 3000 MCMC samples comprising the Model Average Composite scenario. Definitions: $B_{2017}$ - current year spawning biomass, $B_{\text {avg }}$ - average biomass from 1967 to 2016, $B_{\text {min }}$ - minimum biomass that acts as the LRP (and USR $=2 L R P$ ), $u_{2016}$ - harvest rate (ratio of total catch to vulnerable biomass) in the middle of 2016, and $u_{\text {avg }}$ - average harvest rate from 1967 to 2016. All biomass values are in tonnes. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

|  | $5 \%$ | $50 \%$ | $95 \%$ |
| :--- | ---: | ---: | ---: |
| Model-based |  |  |  |
| $B_{2017}$ | 2,621 | 6,185 | 13,927 |
| $B_{\text {avg }}$ | 5,634 | 7,837 | 14,626 |
| $B_{2017} / B_{\text {avg }}$ | 0.385 | 0.683 | 1.62 |
| $u_{2016}$ | 0.106 | 0.214 | 0.406 |
| $H R P-$ based |  |  |  |
| $B_{\text {min }}$ | 654 | 2,051 | 4,818 |
| $2 B_{\text {min }}$ | 1,307 | 4,101 | 9,636 |
| $B_{\min } / B_{\text {avg }}$ | 0.0921 | 0.270 | 0.388 |
| $2 B_{\min } / B_{\text {avg }}$ | 0.184 | 0.540 | 0.775 |
| $B_{2017} / B_{\text {min }}$ | 1.29 | 2.31 | 16.1 |
| $u_{\text {avg }}$ | 0.0744 | 0.150 | 0.234 |
| $u_{2016} / u_{\text {avg }}$ | 0.602 | 1.79 | 2.52 |

Table F.7. BC North: Decision table for the Model Average Composite scenario for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2018}$ and mid-year harvest rate $u_{2017}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities are the proportion of MCMC samples from 3 pooled scenarios chosen for their well-behaved MCMC diagnostics. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=0.99$,
$P\left(B_{2017}>2 B_{\min }\right)=0.62, P\left(B_{2017}>B_{\text {avg }}\right)=0.27$, and $P\left(u_{2016}>u_{\text {avg }}\right)=0.74$. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

| Catch | $\mathrm{P}\binom{B_{2018}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2018}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2018}>}{2 B_{\text {min }}}$ | $\mathrm{P}\binom{B_{2018}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2017}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.23 | 0.99 | 0.58 | 0.26 | 0 |
| 100 | 0.17 | 0.99 | 0.56 | 0.26 | 0 |
| 200 | 0.12 | 0.98 | 0.54 | 0.25 | 0 |
| 300 | 0.10 | 0.98 | 0.53 | 0.24 | 0.01 |
| 400 | 0.08 | 0.97 | 0.51 | 0.24 | 0.08 |
| 500 | 0.07 | 0.96 | 0.50 | 0.23 | 0.24 |
| 600 | 0.06 | 0.95 | 0.48 | 0.22 | 0.45 |
| 700 | 0.05 | 0.94 | 0.47 | 0.22 | 0.58 |
| 800 | 0.04 | 0.93 | 0.46 | 0.21 | 0.65 |
| 900 | 0.03 | 0.91 | 0.45 | 0.21 | 0.68 |
| 1000 | 0.03 | 0.90 | 0.43 | 0.20 | 0.70 |
| 1200 | 0.02 | 0.87 | 0.42 | 0.18 | 0.74 |
| 1400 | 0.01 | 0.84 | 0.40 | 0.18 | 0.80 |
| 1600 | 0.01 | 0.80 | 0.39 | 0.16 | 0.85 |
| 1800 | 0.01 | 0.76 | 0.38 | 0.15 | 0.90 |
| 2000 | 0.01 | 0.71 | 0.37 | 0.13 | 0.93 |
| 2500 | 0 | 0.62 | 0.35 | 0.11 | 0.98 |
| 3000 | 0 | 0.54 | 0.34 | 0.09 | 0.99 |
| 3500 | 0 | 0.48 | 0.32 | 0.07 | 1 |
| 4000 | 0 | 0.43 | 0.30 | 0.05 | 1 |
| 4500 | 0 | 0.40 | 0.28 | 0.04 | 1 |
| 5000 | 0 | 0.37 | 0.26 | 0.03 | 1 |

Table F.8. BC North: Decision table for the Model Average Composite scenario for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities are the proportion of MCMC samples from 3 pooled scenarios chosen for their well-behaved MCMC diagnostics. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=0.99$,
$P\left(B_{2017}>2 B_{\min }\right)=0.62, P\left(B_{2017}>B_{\text {avg }}\right)=0.27$, and $P\left(u_{2016}>u_{\text {avg }}\right)=0.74$. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\text {min }}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.23 | 0.99 | 0.54 | 0.24 | 0 |
| 100 | 0.17 | 0.98 | 0.51 | 0.23 | 0 |
| 200 | 0.13 | 0.96 | 0.48 | 0.22 | 0 |
| 300 | 0.09 | 0.95 | 0.46 | 0.21 | 0.03 |
| 400 | 0.07 | 0.93 | 0.44 | 0.20 | 0.17 |
| 500 | 0.05 | 0.89 | 0.42 | 0.19 | 0.41 |
| 600 | 0.04 | 0.86 | 0.40 | 0.17 | 0.58 |
| 700 | 0.04 | 0.82 | 0.39 | 0.16 | 0.65 |
| 800 | 0.02 | 0.79 | 0.38 | 0.15 | 0.69 |
| 900 | 0.02 | 0.74 | 0.37 | 0.14 | 0.71 |
| 1000 | 0.02 | 0.71 | 0.36 | 0.12 | 0.74 |
| 1200 | 0.01 | 0.63 | 0.35 | 0.11 | 0.81 |
| 1400 | 0.01 | 0.55 | 0.34 | 0.09 | 0.88 |
| 1600 | 0 | 0.51 | 0.33 | 0.07 | 0.93 |
| 1800 | 0 | 0.46 | 0.32 | 0.06 | 0.96 |
| 2000 | 0 | 0.42 | 0.30 | 0.05 | 0.98 |
| 2500 | 0 | 0.36 | 0.26 | 0.03 | 1 |
| 3000 | 0 | 0.30 | 0.22 | 0.01 | 1 |
| 3500 | 0 | 0.26 | 0.18 | 0.01 | 1 |
| 4000 | 0 | 0.22 | 0.14 | 0.01 | 1 |
| 4500 | 0 | 0.20 | 0.11 | 0 | 1 |
| 5000 | 0 | 0.17 | 0.08 | 0 | 1 |



Figure F.23. BC North: Status of the current stock $B_{2017}$ relative to $B_{\text {avg }}$ with the dashed lines showing historical reference points ( $B_{\min } / B_{\text {avg }}, 2 B_{\min } / B_{\text {avg }}$ ) that mimic DFO Precautionary Approach provisional MSY-based reference points. Stock status is shown for the Model Average Composite scenario comprising 3 pooled model runs and for each of the 3 model runs (see Table F.2) for definitions of these model runs). Boxplots show the 5, 25, 50, 75 and 95 percentiles from the MCMC results. $M=$ instantaneous natural mortality $\left(y^{-1}\right) ; k=$ age $(y)$ at knife-edge recruitment.

Table F.9. BC North: Decision table for case S00: M.30+k3 for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=0.98, P\left(B_{2017}>2 B_{\min }\right)=0.41, P\left(B_{2017}>B_{\text {avg }}\right)=0.01$, and $P\left(u_{2016}>u_{\mathrm{avg}}\right)=1$. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {min }}}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\text {min }}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.28 | 0.98 | 0.34 | 0 | 0 |
| 100 | 0.20 | 0.97 | 0.29 | 0 | 0 |
| 200 | 0.13 | 0.95 | 0.24 | 0 | 0 |
| 300 | 0.08 | 0.92 | 0.18 | 0 | 0.05 |
| 400 | 0.05 | 0.88 | 0.15 | 0 | 0.29 |
| 500 | 0.03 | 0.82 | 0.12 | 0 | 0.67 |
| 600 | 0.02 | 0.76 | 0.09 | 0.90 |  |
| 700 | 0.02 | 0.68 | 0.07 | 0.97 |  |
| 800 | 0.01 | 0.62 | 0.06 | 0.99 |  |
| 900 | 0.01 | 0.54 | 0.05 | 0 | 1 |
| 1000 | 0 | 0.46 | 0.03 | 0 | 1 |
| 1200 | 0 | 0.34 | 0.02 | 0 | 1 |
| 1400 | 0 | 0.22 | 0.01 | 0 | 1 |
| 1600 | 0 | 0.14 | 0.01 | 0 | 1 |
| 1800 | 0 | 0.09 | 0 | 0 | 1 |
| 2000 | 0 | 0.07 | 0 | 0 | 1 |
| 2500 | 0 | 0 | 0 | 0 | 0 |
| 300 | 0 | 0 | 0 | 0 | 0 |

Table F.10. BC North: Decision table for case S03: M.35+k3 for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=0.99, P\left(B_{2017}>2 B_{\min }\right)=0.45, P\left(B_{2017}>B_{\text {avg }}\right)=0.02$, and $P\left(u_{2016}>u_{\text {avg }}\right)=1$. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.16 | 0.98 | 0.27 | 0 | 0 |
| 100 | 0.12 | 0.97 | 0.24 | 0 | 0 |
| 200 | 0.09 | 0.94 | 0.22 | 0 | 0 |
| 300 | 0.07 | 0.93 | 0.20 | 0 | 0.04 |
| 400 | 0.06 | 0.90 | 0.16 | 0 | 0.20 |
| 500 | 0.04 | 0.86 | 0.14 | 0 | 0.56 |
| 600 | 0.04 | 0.82 | 0.12 | 0 | 0.83 |
| 700 | 0.03 | 0.79 | 0.10 | 0 | 0.94 |
| 800 | 0.02 | 0.76 | 0.08 | 0 | 0.98 |
| 900 | 0.02 | 0.70 | 0.06 | 0 | 1 |
| 1000 | 0.01 | 0.66 | 0.06 | 0 | 1 |
| 1200 | 0.01 | 0.56 | 0.04 | 0 | 1 |
| 1400 | 0 | 0.44 | 0.04 | 0 | 1 |
| 1600 | 0 | 0.39 | 0.03 | 0 | 1 |
| 1800 | 0 | 0.30 | 0.02 | 0 | 1 |
| 2000 | 0 | 0.24 | 0.01 | 0 | 1 |
| 2500 | 0 | 0.14 | 0 | 0 | 1 |
| 3000 | 0 | 0.09 | 0 | 0 | 1 |
| 3500 | 0 | 0.06 | 0 | 0 | 1 |
| 4000 | 0 | 0.03 | 0 | 0 | 1 |
| 4500 | 0 | 0.02 | 0 | 0 | 1 |
| 5000 | 0 | 0.02 | 0 | 0 | 1 |

Table F.11. BC North: Decision table for case S10: M.30+k3-GIG-CPUE for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)=1, P\left(B_{2017}>B_{\text {avg }}\right)=0.80$, and $P\left(u_{2016}>u_{\text {avg }}\right)=0.23$. For reference, the average catch over the last 5 years (2011-2015) is $992 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.23 | 1 | 1 | 0.73 | 0 |
| 100 | 0.19 | 1 | 1 | 0.69 | 0 |
| 200 | 0.16 | 1 | 1 | 0.66 | 0 |
| 300 | 0.13 | 1 | 1 | 0.62 | 0 |
| 400 | 0.10 | 1 | 1 | 0.60 | 0 |
| 500 | 0.08 | 1 | 1 | 0.56 | 0.01 |
| 600 | 0.07 | 1 | 1 | 0.51 | 0.02 |
| 700 | 0.06 | 1 | 1 | 0.47 | 0.04 |
| 800 | 0.05 | 1 | 0.99 | 0.44 | 0.08 |
| 900 | 0.04 | 1 | 0.99 | 0.41 | 0.13 |
| 1000 | 0.03 | 1 | 0.99 | 0.37 | 0.22 |
| 1200 | 0.02 | 1 | 0.98 | 0.32 | 0.43 |
| 1400 | 0.02 | 0.99 | 0.97 | 0.26 | 0.64 |
| 1600 | 0.02 | 0.99 | 0.96 | 0.22 | 0.79 |
| 1800 | 0.01 | 0.98 | 0.93 | 0.18 | 0.89 |
| 2000 | 0.01 | 0.97 | 0.89 | 0.14 | 0.95 |
| 2500 | 0 | 0.90 | 0.78 | 0.08 | 0.99 |
| 3000 | 0 | 0.81 | 0.66 | 0.04 | 1 |
| 3500 | 0 | 0.72 | 0.53 | 0.02 | 1 |
| 4000 | 0 | 0.63 | 0.42 | 0.02 | 1 |
| 4500 | 0 | 0.56 | 0.32 | 0.01 | 1 |
| 5000 | 0 | 0.49 | 0.25 | 0.01 | 1 |

## F.3. BC SOUTH STOCK

## F.3.1. Example Case - South

An example model run for the BC South stock is presented to show more detail in the results. This model is based on the eastern Bering Sea Okhotsk growth model from Janusz and Horbowy (1997) and included the following elements:

- instantaneous natural mortality $M$ fixed at 0.30 ;
- knife-edge recruitment at age $k=3 \mathrm{y}$;
- steepness $h$ beta prior (mean=0.7, SD=0.15);
- 1972-2016 standardised unsorted mean weights ( $\bar{w}=0.52052 \mathrm{~kg}$ );
- length-weight allometry: $\alpha=7.3536 \mathrm{E}-06, \beta=3.030278$;
- Brody parameters: $\alpha_{g}=0.14441 \mathrm{~kg}, \rho_{g}=0.87063, w_{k}=0.24875 \mathrm{~kg}$;
- growth parameters: $L_{\infty}=50.82725 \mathrm{~cm}, \kappa=0.199054, t_{0}=-1.78968 \mathrm{y}$;
- errors: observation $\sigma_{O}=0.2$, recruitment $\sigma_{R}=0.6$, mean weight $\sigma_{W}=0.15$;
- uniform priors on $q$ from -10 to 0;
- Walleye Pollock CPUE indices (uniform -10 to 0);
- catch series (GFFOS accessed 2016-09-12);
- equilibrium start in 1967 (use all of GIG historical, including 1995);
- equal-weight for each age class when estimating von Bertalanffy model parameters by sex, combined-sex model interpolated between sexes;
- estimate $\ln \left(R_{0}\right)$ (uniform prior) and fix $\ln (\bar{R})$ and $\ln \left(R_{\text {init }}\right)$ to $\ln \left(R_{0}\right)$;
- version iscam-delaydiff.exe: built 2017-01-03.


## F.3.1.1. MPD results - example (south)

The mode of the posterior distribution (MPD) for this model (reported in Table F.12) is estimated by minimising the objective function (components summarised in Section E.2.1). The results are presented to show the fits by the model to the observed data and are used as the starting point for the MCMC simulations. MPD fits are shown for the abundance indices (Figure F.24), the annual mean weights and annual recruitment (Figure F.25). The fits to the survey and CPUE indices are generally reasonable although the model is incapable of fitting the abrupt changes in some series. The model is also not capable of fitting the high values of mean weight that occur at the beginning of the series (Figure F.25). Instead it fluctuates near the mean of the series, with a general dome-shaped trend that tries to fit the high values in the 1970s and 1980s. Recruitment events exceed the long-term median of the series 10 times; the 1975 event is roughly 11 times higher than the long-term median. Fits to the catch data are not presented because the model is parameterised so that it always fits the catch closely.

## F.3.1.2. MCMC results - example (south)

The MCMC procedure performed 60,000,000 iterations, sampling every 50,000 to give 1200 draws ( 1000 samples after dropping the first 200, including the MPD start point, as burn-in). The 1000 samples were used to estimate parameters and quantities of interest, including stock status by year and the probabilities of being above reference points.

MCMC traces show good convergence properties (no trend with increasing sample number) for the leading estimated parameters (Figure F.26); however, a diagnostic analysis that splits the samples into three segments, checking for consistency along the length of the chain (Figure F.27), indicates some instability over time. The autocorrelation plots (Figure F.28) confirm this with some periodicity over 100 lags and significant serial correlation for the first 5-10 lags for all parameters except $h$. Pairs plots of the estimated parameters (Figure F.29) show no undesirable correlation between the two primary parameters, $\ln \left(R_{0}\right)$ and $h$, though all the $q$ parameters were highly correlated with $\ln \left(R_{0}\right)$ and with each other, as would be expected. MCMC quantiles for parameters, biomass, and status with respect to historical reference points are summarised in Table F. 12.

Marginal posterior distributions along with the corresponding priors for the estimated parameters are shown in Figure F.30. Only the steepness parameter used an informative prior, with its posterior distribution largely reflecting the prior. This indicates that there was relatively little information in this model to inform this parameter and it is unlikely that it could be estimated without using a prior.

The plot of estimated spawning biomass (Figure F.31) shows a large increase in the mid 1970s followed by a decline to a low point in 2008. Since then, spawning biomass has increased and the median $B_{2017}$ lies very close to the average biomass $B_{\text {avg }}$. Assuming a catch policy of $3250 \mathrm{t} / \mathrm{y}$, which is almost double the current TAC of $1790 \mathrm{t} / \mathrm{y}$ in 5AB but close to the $5-\mathrm{y}$ average catch of 3256 t in the BC South region, the projected biomass declines under conditions of average recruitment.

This model run estimates a few strong recruitment pulses in 1970, 1974, and 2014 (Figure F.32). The first strong recruitment occurs in the absence of fishing pressure and so spawing biomass increases substantially (Figure F.31). Thereafter, strong fishing pressure brings the population to a minimum in 2008, but after this point fishing does not appear to dampen lacklustre recruitment (Figure F.33). Fishing mortality peaks in 2003 at a median $F$ value of $0.275 \mathrm{y}^{-1}$ and declines thereafter until it reaches $0.028 \mathrm{y}^{-1}$ in 2016. The median value (and the 5th and 95th percentiles in parentheses) for the estimated level of biomass depletion ( $B_{t} / B_{\text {avg }}$ ) at the end of the final year of the reconstruction is $1.02(0.68,1.4)$, with the MPD value of 1 lying well above the median of the posterior distribution of this quantity (Figure F.33).

The use of historical reference points is illustrated in Figure F.34. Under an assumed catch policy of $3250 \mathrm{t} / \mathrm{y}, B_{2019}$ lies lower than $B_{2017}$ and $B_{\text {avg }}$. All three biomass estimates lie above the LRP and the USR, with the median of $B_{2017}$ slightly above the median of average biomass $B_{\text {avg }}$. The current fishing mortality rate $F_{2016}$ is lower than the average $F$, the minimum $F$ that occurred in 2008, and the maximum fishing mortlaity rates experienced over the assessment time period. The phase plot (Figure F.35) confirms that the current stock status lies well above the LRP and the USR, and the current mid-year harvest rate $u_{2016}$ lies below the average harvest rate over the time series.

Table F.12. BC South: The $5^{\text {th }}, 50^{\text {th }}$, and $95^{\text {th }}$ percentiles of MCMC-derived parameter estimates and quantities from 1,000 MCMC samples for the example model run. Some fixed parameters are reported as MPD only. See Appendix E for paramater definitions. Subscripts 1-3 on q refer to the fishery-independent surveys, subscript 4 refers to the commercial trawl CPUE series. Other definitions: $B_{2017}$ - biomass at the start of 2017, $u_{2016}$ - exploitation rate (ratio of total catch to vulnerable biomass) in the middle of 2016, $B_{\text {avg }}$ - average biomass from 1967 to 2016, LRP (limit reference point) - minimum median estimated biomass in the time series, USR (upper stock reference) - twice the LRP biomass. All biomass values are in tonnes. For reference, the average catch over the last 5 years (2011-2015) is 3256 t.

|  | $5 \%$ | $50 \%$ | $95 \%$ | MPD |
| ---: | ---: | ---: | ---: | ---: |
| Parameters |  |  |  |  |
| $R_{0}$ | 7,614 | 13,738 | 36,627 | 13,347 |
| $h$ | 0.505 | 0.751 | 0.926 | 0.792 |
| $M$ |  |  |  | 0.3 |
| $q_{1}$ | 0.00562 | 0.0146 | 0.0252 | 0.0191 |
| $q_{2}$ | 0.0019 | 0.00561 | 0.0109 | 0.00779 |
| $q_{3}$ | 0.00806 | 0.0244 | 0.0468 | 0.0335 |
| $q_{4}$ | 0.00001 | 0.000031 | 0.00006 | 0.000043 |
| HRP-based |  |  |  |  |
| $B_{2017}$ | 43,736 | 91,200 | 280,228 | 67,025 |
| $B_{\text {avg }}$ | 50,723 | 89,549 | 244,104 | 66,878 |
| LRP $=B_{2008}$ | 7,296 | 16,655 | 50,664 | 10,239 |
| $\mathrm{USR}=2 B_{2008}$ | 14,593 | 33,309 | 101,328 | 20,478 |
| $B_{2017} / B_{\text {avg }}$ | 0.679 | 1.02 | 1.4 | 1 |
| $B_{2017} / B_{2008}$ | 2.59 | 5.82 | 8.11 | 6.55 |
| $B_{2017} / 2 B_{2008}$ | 1.29 | 2.91 | 4.06 | 3.27 |
| $u_{\text {avg }}$ | 0.0127 | 0.0383 | 0.0696 | 0.0511 |
| $u_{2016} / u_{\text {avg }}$ | 0.582 | 0.735 | 0.946 | 0.739 |



Figure F.24. BC South: MPD index fits to relative abundance indices for the example model run. Circles represent observed indices with associated CVs; squares represent the model fit. Surveys: (1) GIG (Goose Island Gully) Historical, (2) WCVI (west coast Vancouver Island) Synoptic, (3) QCS (Queen Charlotte Sound) Synoptic, (4) commercial trawl catch per unit effort of Walleye Pollock.


Figure F.25. BC South: [Top] MPD fit to the mean weight data for the example model run. Predicted mean weights are shown as a red line and observations are shown as points. Error bars on mean weight observations represent a fixed CV using $\sigma_{W}=0.15$. [Bottom] MPD recruitment in thousands of age-3 individuals in year $t$ for example model run.


Figure F.26. BC South: Trace plots for MCMC output of estimated parameters in the example model run. The MCMC run shows 1000 MCMC samples after removing 200 samples. Parameters log.ro (natural log of unfished equilibrium recruitment), $h$ (steepness), and $q$ (catchability) for the surveys outlined in Figure F.24. Grey lines show the 1000 samples for each parameter, solid lines show the cumulative median (up to that sample), and dashed lines show the cumulative 2.5 and 97.5 quantiles. Red circles are the MPD estimates.


Figure F.27. BC South: Diagnostic plot for the example model run obtained by dividing the MCMC chain of 1000 samples into three segments and over-plotting the cumulative distributions of the first segment (green), second segment (red), and final segment (blue).


Figure F.28. BC South: Autocorrelation plots for MCMC output of estimated parameters in the example model run. See Figure F. 26 for parameter descriptions.


Figure F.29. BC South: Pairs plot from the southern BC region of 1000 MCMC samples for the estimated parameters of the example model run. The diagonal shows the frequency distribution of each posterior.


Figure F.30. BC South: Prior probability distributions (blue lines) used in the example model run and the comparative posterior histograms. Parameters $q_{j}$ represent catchability of the various surveys $j$ as defined in Figure F.26. The dashed red vertical lines show the MPD estimates.

## Biomass



Figure F.31. BC South: Posterior estimates of spawning biomass (1000t) for the example model run with $95 \%$ credibility intervals in grey. The current year biomass (2017, yellow point) and projected biomass (2018-2022, red line), assuming a constant catch policy of $3250 t / y$, are enclosed by a $95 \%$ credibility interval shaded pink. The median posterior estimate of $B_{0}$ is shown as a green point (with $95 \%$ credibility range) to the left of the time series. The MPD estimate is shown as a blue line. The total catch is shown along the bottom as red bars, with assumed TAC catch in pink..


Figure F.32. BC South: [Top] Posterior estimates of age-3 recruits for the example model run. [Bottom] Log recruitment deviations for the example model run with $95 \%$ credibility intervals.


Figure F.33. BC South: [Top] Posterior estimates of fishing mortality for the example model run. [Bottom] Biomass depletion, i.e. $B_{t} / B_{\text {avg }}$, for the example model run with $95 \%$ credibility intervals. Also displayed on the depletion figure is the MPD estimate (blue line) and the reference points $B_{\text {min }}$ (red dashed line) and $2 B_{\text {min }}$ (green dashed line).


Figure F.34. BC South: Posterior estimates of spawning biomass- and harvest-based reference points for the example model run. [Left] The current-year biomass $B_{2017}$, the projected-year biomass $B_{2019}, B_{\text {avg }}=$ average biomass from 1967 to 2016, the limit reference biomass (or $B_{\min }$ ) = biomass in 2008, and the upper stock biomass set at $2 B_{2008}$. [Right] The current-year fishing mortality rate $F_{2016}$, the average fishing mortality rate $F_{\text {avg }}$ from 1967 to 2016, the fishing mortality rate in the year of minimum biomasss $F_{2008}$, and the maximum fishing mortality experience over the time series $F_{\max }$. Box delimiter and limits represent quantiles at 0.5 (median), 0.25 and 0.75 quantiles, respectively, and the whiskers delimit the 0.05 and 0.95 quantiles. Outliers are not shown.


Figure F.35. BC South: Phase plot through time for the example model run of the medians of the ratios $B_{t} / B_{\text {avg }}$ (the biomass in year $t$ relative to $B_{\text {avg }}$ ) and $u_{t} / u_{\text {avg }}$ (the exploitation rate in year $t$ relative to $u_{\text {avg }}$ ). Blue filled circle is the starting year 1967. Years then proceed from light grey through to dark grey with the final year 2017 as a filled orange circle with limit lines represent the 0.05 and 0.95 quantiles of the posterior distributions for the final year. Vertical dashed lines indicate the historical limit (red) and upper stock reference (green) points (see legend for values), and horizontal dotted line indicates $u_{t}$ at $u_{\text {avg }}$.

## F.3.2. Alternative Cases - South

For the BC South stock, we ran a total of 11 alternative runs (Table F.13) (including the example run above) using the growth model published by Janusz and Horbowy (1997) for the Sea of Okhotsk (off Russia's east coast and west of the Bering Sea) to test the robustness of the results to uncertainties in:

- natural mortality ( $M$ );
- age at knife-edge recruitment $(k)$;
- the use of the CPUE index series.

We tested a range of sensible values for $M$ and $k$ because these parameters control key assumptions made by the delay-difference model when the data available to the model are not very informative with respect to these parameters. We also tested combinations of $M$ and $k$ for the last alternative category (removing CPUE index series).

The delay-difference model assumes that maturity matches selectivity, i.e., all recruited fish are mature, and by extension, all mature fish have a single natural mortality rate. Alternative runs for the BC South stock were made with $M \in\{0.25,0.30,0.35\}$ to bracket plausible values for this single parameter (Table F.13). We use $k \in\{3,4,5\}$ for the BC South stock guided by ogive information in Table D. 7 of Dorn et al. (2012). We added $k=5$ to the suite of alternative values for this parameter because of the lower maximum size in this growth model. Nine of the alternative runs (including the example case) used all the abundance index series and paired $k \in\{3,4,5\}$ with $M \in\{0.25,0.30,0.35\}$. Two of the alternative runs dropped the CPUE index series and paired $M=0.3$ with $k \in\{3,4\}$.
All 11 alternative runs were taken to the MCMC level. Each MCMC search was started at the "best fit" MPD parameter set, run for 60 million iterations, and sampled every 50,000 for 1200 samples. The first 200 samples were dropped as burn-in to yield a total posterior sample of 1000 draws.

## F.3.2.1. MPD results - alternatives (south)

Visual inspection of fits from these alternative runs suggests that the data available to this model did not allow much discrimination between the range of hypotheses tested. While there were differences in the fits to the available biomass indices and to the mean weight data, these differences tended to be small and probably could not distinguish between hypotheses except in the most extreme cases. As an example, none of the models were able to fit the high mean weights observed in the late 1970s and in the late 1980s (Figure F.36), which may be attributable to misspecification of the growth model. While the delay-difference model uses the mean weight data to scale the overall biomass and to obtain recruitment deviation information, it is likely that only an age-structured model would have sufficient flexibility to fit the entire mean weight series. The model also did not fit some of the high points in the survey series or the CPUE times series (Figures F.37-F.40), but generally most index points were acceptably fitted. The model seemed to have the most difficulty fitting the WCVI survey series (Figure F.38).
In an attempt to make such comparisons more quantitative, the negative log-likelihoods for the fits to each data component of the model are summarised for the 11 alternative runs in Table F.14. Comparisons of the whole model can only be made among those runs that share the same estimated components:

- Cases S00 to S08 - S08 had the lowest objective function value of these nine cases and offered the best model fit in the first group;
- Cases S09 to S10 - S10 offered the better fit in the second group.

The two best-fit models featured high values for $M$ paired with high values of $k$, which likely reflects the growth model used for this population of smaller Walleye Pollock. Case S08 ( $M=0.35, k=5$ ) also had the best fit to the mean weight component, and most of the $k \in\{4,5\}$ models showed much better fits than those where $k=3$.

Table F.13. Summary of the analyses performed to test the sensitivity of the delay-difference model to variations in natural mortality $M$, knife-edge recruitment age $k$. All runs for the BC South stock use the Okhotsk Sea growth function (Janusz and Horbowy 1997). The column marked 'Rank' provides a subjective ranking of the MCMCs, where $1=$ good, $2=$ acceptable, and $3=$ poor.

| Case | Run ID | Run \# | M | $k$ | Rank |
| :---: | :---: | :---: | :---: | :---: | :---: |
| S00 | M.30+k3 | 4 | 0.3 | 3 | 2 |
| Sensitivity to k at M=0.30 |  |  |  |  |  |
| S01 | M.30+k4 | 5 | 0.3 | 4 | 2 |
| S02 | M.30+k5 | 15 | 0.3 | 5 | 2 |
| Sensitivity to k at M=0.25 |  |  |  |  |  |
| S03 | M.25+k3 | 11 | 0.25 | 3 | 1 |
| S04 | M.25+k4 | 9 | 0.25 | 4 | 1.5 |
| S05 | M.25+k5 | 14 | 0.25 | 5 | 2 |
| Sensitivity to $\mathbf{k}$ at $\mathbf{M}=\mathbf{0 . 3 5}$ |  |  |  |  |  |
| S06 | M.35+k3 | 12 | 0.35 | 3 | 2 |
| S07 | M.35+k4 | 13 | 0.35 | 4 | 2 |
| S08 | M.35+k5 | 16 | 0.35 | 5 | 2 |
| Removal of commercial CPUE |  |  |  |  |  |
| S09 | M.30+k3-CPUE | 10 | 0.3 | 3 | 1 |
| S10 | M.30+k4-CPUE | 8 | 0.3 | 4 | 3 |

Table F.14. BC South MPD negative log likelihoods from the 11 alternative cases documented in Table F. 13 for each data component used in the model.

| Case | Catch | GIG <br> Hist | WCVI <br> Synop | QCS <br> Synop | South <br> CPUE | Recruits | Mean <br> Weight | ObFn <br> Value |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| S00: M.30+k3 | -103.746 | 12.417 | 77.426 | 12.888 | 24.692 | 81.712 | 10.980 | 254.232 |
| S01: M.30+k4 | -103.424 | 12.807 | 67.805 | 7.540 | 20.435 | 80.728 | 0.580 | 223.310 |
| S02: M.30+k5 | -102.609 | 12.418 | 65.330 | 11.370 | 14.403 | 81.524 | -0.916 | 218.626 |
| S03: M.25+k3 | -103.645 | 14.049 | 75.907 | 13.553 | 28.263 | 80.874 | 10.145 | 256.383 |
| S04: M.25+k4 | -103.278 | 14.357 | 68.264 | 8.702 | 23.549 | 81.285 | 0.001 | 229.967 |
| S05: M.25+k5 | -102.339 | 14.152 | 65.824 | 12.677 | 17.421 | 82.400 | 0.711 | 228.370 |
| S06: M.35+k3 | -103.805 | 10.913 | 78.441 | 12.100 | 21.326 | 82.190 | 13.178 | 252.633 |
| S07: M.35+k4 | -103.521 | 11.083 | 67.216 | 6.615 | 17.741 | 80.980 | 1.440 | 218.596 |
| S08: M.35+k5 | -102.823 | 10.721 | 64.767 | 10.251 | 11.978 | 81.093 | -2.125 | 210.788 |
| S09: M.30+k3-CPUE | -103.089 | 12.146 | 35.993 | 13.843 | - | 82.603 | 22.418 | 202.150 |
| S10: M.30+k4-CPUE | -102.837 | 12.806 | 36.614 | 8.391 | - | 80.376 | -1.323 | 170.736 |



Figure F.36. BC South: MPD fit to the mean weight data for the 11 alternative runs. Predicted mean weights are shown as red lines and observations are shown as points. Error bars on mean weight observations represent a fixed CV using $\sigma_{W}=0.15$.


Figure F.37. BC South: MPD index fits (11 alternative runs) for the GIG (Goose Island Gully) historical rockfish survey relative abundance indices. Circles represent observed indices with associated CVs, triangles represent the model fit.


Figure F.38. BC South: MPD index fits (11 alternative runs) for the WCVI (west coast Vancouver Island) synoptic survey relative abundance indices. Circles represent observed indices with associated CVs, triangles represent the model fit.


Figure F.39. BC South: MPD index fits (11 alternative runs) for the QCS (Queen Charlotte Sound) synoptic survey relative abundance indices. Circles represent observed indices with associated CVs, triangles represent the model fit.


No index
No index

## Year

Figure F.40. BC South: MPD index fits (11 alternative runs) for the commercial trawl Walleye Pollock CPUE relative abundance indices. Circles represent observed indices with associated CVs, triangles represent the model fit.

## F.3.2.2. MCMC results - alternatives (south)

Median estimates for current biomass $B_{2017}$ lie at or below the average biomass $B_{\text {avg }}$ for all alternative cases(Table F.15). Unlike the alternative case for the BC North stock, the year in which median spawning biomass reached a minimum is stable across all runs, occurring in 2008. Despite this apparent stability, the maximum fishing mortality $F$ estimated in the MCMC samples exceeded realistic levels ( $F>2$ ) in some years for all but three cases. This is likely to be the result of the failure of the knife-edge recruitment assumption (particularly when it is $k=4$ or $k=5$ ), which leaves too few fish in the population in some years to support the observed level of catch. However, the alternative cases model population trajectories consistently such that the minimum biomass occurs in the same year.
Table F. 16 shows for each alternative run the probabilities of projected biomass in two years at $3250 \mathrm{t} / \mathrm{y}$ (at the level of recent average catch) exceeding various reference points. All runs except S 08 show high probabilites of $B_{2019}$ exceeding the limit reference point $B_{\min }$. All scenarios project $B_{2019}$ to fall below $B_{\text {avg }}$ and $B_{2017}$, although it should be noted that 2-year projections using this delay-difference model are unreliable and uncertain. That is because these models, like a surplus production model and unlike an age-structured model, only project using the stock-recruitment function, which has little predictive power.

Most of the trace plots for $R_{0}$ look acceptable but a few show large-scale shifts in the trace mean (Figure F.41). Autocorrelation plots for $R_{0}$ highlight other problems (Figure F.42), with high levels of serial correlation in seven of the 11 alternative runs. The runs featuring low $M$ and $k$ values show the least autocorrelation, but these are the runs with the higher negative log-likelihood fits (Table F.14). The ranks assigned to the quality of MCMCs appear in Table F.13. While there is no statistical basis for selecting among these hypotheses, these MCMC results suggest that a selection of the runs with the best diagnostics can be used to model this Pollock stock.
Consequently, this stock assessment used the subjective MCMC quality rankings to construct a Model Average posterior to provide advice to managers (Section F.3.3.), with a rank $\leq 2$ used as the quality cutoff criterion.
Quantile plots compare $B_{\text {avg }}, B_{2017} / B_{\text {avg }}, u_{2016} / u_{\text {avg }}$ and $u_{\text {avg }} B_{\text {avg }}$ of the example run to alternative runs grouped by category:

- $M=0.30$ and $k \in\{3,4,5\}$ - comparing all the medium $M$ alternatives (Figure F.43);
- $M=0.25$ and $k \in\{3,4,5\}$ - comparing all the low $M$ alternatives (Figure F.44); and
- $M=0.35$ and $k \in\{3,4,5\}$ - comparing all the high $M$ alternatives (Figure F.45).

These plots show that $M$ and $k$ interact to change the perceived size of the stock (e.g, $B_{\text {avg }}$ ), but the estimated stock status is much more consistent among the runs (Figure F.43).
Note that runs $\mathrm{S} 00, \mathrm{~S} 03$ and S 06 , all runs where $k=3$ (but include the CPUE series), have the largest estimated $B_{\text {avg }}$ and consequently the lowest $F_{\text {max }}$. All other runs estimate smaller average biomass levels and consequent maximum exploitation rates that are not credible. This is likely to be the result of the strong assumption of knife-edge recruitment that is made by the delay-difference model, resulting in insufficient biomass to accommodate catch levels in some years when $k>3$. This is a model misspecification resulting from data limitations (i.e., the lack of ageing data), leading to a cautious interpretation of these runs but probably not to outright rejection. The two runs which discard the CPUE series also estimate levels of biomass that result in unacceptable estimates for $F_{\max }$, even for runs with $k=3$.

Table F.15. BC South median values for select MCMC-derived parameters and quantities for 11 alternative runs from the BC South stock. The value for $B_{2019}$ is that assuming a TAC of 3250 t/y. Model senstivity details appear in Table F.13.

| Run | $h$ | $B_{\text {avg }}$ | $\frac{B_{2017}}{B_{\text {avg }}}$ | $\frac{B_{2019}}{B_{\text {avg }}}$ | $\mathrm{Yr}_{\min }$ | $\frac{B_{2017}}{B_{\min }}$ | $F_{\max }$ | $u_{\text {avg }}$ | $\frac{u_{2016}}{u_{\text {avg }}}$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| S00: M.30+k3 | 0.75 | 89,549 | 1.0 | 0.75 | 2008 | 5.8 | 0.28 | 0.038 | 0.73 |
| S01: M.30+k4 | 0.77 | 21,257 | 0.79 | 0.50 | 2008 | 8.0 | 18 | 0.16 | 0.82 |
| S02: M.30+k5 | 0.79 | 14,835 | 0.72 | 0.38 | 2008 | 8.3 | 18 | 0.24 | 0.82 |
| S03: M.25+k3 | 0.76 | 54,998 | 1.0 | 0.73 | 2008 | 6.4 | 0.49 | 0.062 | 0.74 |
| S04: M.25+k4 | 0.78 | 20,412 | 0.85 | 0.46 | 2008 | 8.5 | 18 | 0.16 | 0.80 |
| S05: M.25+k5 | 0.81 | 13,022 | 0.84 | 0.47 | 2008 | 8.7 | 19 | 0.24 | 0.82 |
| S06: M.35+k3 | 0.74 | 183,563 | 1.0 | 0.72 | 2008 | 5.5 | 0.12 | 0.018 | 0.74 |
| S07: M.35+k4 | 0.78 | 21,814 | 0.75 | 0.36 | 2008 | 7.6 | 14 | 0.17 | 0.82 |
| S08: M.35+k5 | 0.79 | 14,623 | 0.72 | 0.24 | 2008 | 7.9 | 20 | 0.23 | 0.85 |
| S09: M.30+k3-CPUE | 0.75 | 33,336 | 0.62 | 0.36 | 2008 | 18 | 19 | 0.14 | 0.81 |
| S10: M.30+k4-CPUE | 0.76 | 19,971 | 0.90 | 0.44 | 2008 | 17 | 18 | 0.18 | 0.69 |

Table F.16. BC South: Assuming a constant catch policy of $3250 t / y$, the probability that $B_{2019}$ (or $u_{2018}$ ) is greater than reference points used in this assessment for 11 alternative runs. Model senstivity details appear in Table F.13. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Run | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\text {min }}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| :--- | ---: | ---: | ---: | ---: | ---: |
| S00: M.30+k3 | 0 | 1 | 0.94 | 0.10 | 0.08 |
| S01: M.30+k4 | 0 | 0.97 | 0.92 | 0 | 0.58 |
| S02: M.30+k5 | 0 | 0.96 | 0.93 | 0 | 0.81 |
| S03: M.25+k3 | 0 | 1 | 0.91 | 0.09 | 0.93 |
| S04: M.25+k4 | 0 | 0.95 | 0.91 | 0 | 0.99 |
| S05: M.25+k5 | 0 | 0.98 | 0.96 | 0 | 0.72 |
| S06: M.35+k3 | 0 | 1 | 0.97 | 0.08 | 0.11 |
| S07: M.35+k4 | 0 | 0.94 | 0.88 | 0 | 1 |
| S08: M.35+k5 | 0 | 0.90 | 0.67 | 0 | 1 |
| S09: M.30+k3-CPUE | 0 | 0.94 | 0.88 | 0 | 0.99 |
| S10: M.30+k4-CPUE | 0 | 1 | 1 | 0.01 | 0.97 |



Figure F.41. BC South: Trace plots (11 alternative runs) for MCMC samples of $\log \left(R_{0}\right)$ (natural log of unfished equilibrium recruitment). The MCMC run had chain length 60 million and a sample taken at every $50,000^{\text {th }}$ iteration to yield 1,000 MCMC samples after a removing a burn-in of 200 samples. Grey lines show the 1000 samples for each parameter, solid lines show the cumulative median (up to that sample), and dashed lines show the cumulative 2.5 and 97.5 quantiles. Red circles are the MPD estimates.


Figure F.42. BC South: Autocorrelation plots (11 alternative runs) for MCMC samples of $\log \left(R_{0}\right)$ (natural $\log$ of unfished equilibrium recruitment). The MCMC runs had 1,000 MCMC samples each.


Figure F.43. BC South: Quantile plots comparing the base case (S00 with $M=0.30$ and $k=3$ using the Okhotsk Sea (OS) growth model) to alternative runs (all with $M=0.30$ and using OS growth) that vary by by knife-edge recruitment age $k$, where $k \in\{3,4,5\}$. Additionally, S09 and S10 were fit after dropping the commercial CPUE data. Box delimiter and limits represent quantiles at 0.5 (median), 0.25 and 0.75 quantiles, respectively, and the whiskers delimit the 0.05 and 0.95 quantiles. Outliers are not shown.


Figure F.44. BC South: Quantile plots comparing the base case to alternative runs (all with $M=0.25$ and using OS growth) that vary by knife-edge recruitment age $k$, where $k \in\{3,4,5\}$. Quantile box delimiters are detailed in Figure F.43.


Figure F.45. BC South: Quantile plots comparing the base case to alternative runs (all with $M=0.35$ and using OS growth) that vary by knife-edge recruitment age $k$, where $k \in\{3,4,5\}$. Quantile box delimiters are detailed in Figure F.43.

## F.3.3. Model Average Composite - South

Six alternative BC South runs were selected for inclusion to the Model Average posterior based on the following criteria (see Tables F. 13 and F.15):

- use model runs where the median $F_{\max }$ across MCMC samples was $<2$;
- add model runs where median annual $F_{t}$ was $>2$ only once;
- remove model runs with poor diagnostics (rank $>2$ ).

For the BC South stock, these criteria selected 6 out of 11 models:

- S00: M.30+k3 (rank=2.0, med. $F_{\max }=0.28$, no.yrs med. $F_{t}>2=0$ )
- S01: M.30+k4 (rank=2.0, med. $F_{\max }=18.3$, no.yrs med. $F_{t}>2=1$ )
- S03: M.25+k3 (rank=1.0, med. $F_{\max }=0.49$, no.yrs med. $F_{t}>2=0$ )
- S04: M.25+k4 (rank=1.5, med. $F_{\max }=18.3$, no.yrs med. $F_{t}>2=1$ )
- S06: M.35+k3 (rank=2.0, med. $F_{\max }=0.12$, no.yrs med. $F_{t}>2=0$ )
- S07: M.35+k4 (rank=2.0, med. $F_{\max }=14.2$, no.yrs med. $F_{t}>2=1$ )

Table F. 17 gives the model-based and HRP-based quantities ( $0.05,0.50$, and 0.95 quantiles) from the model average posterior based on 6000 pooled MCMC samples. Table F. 19 gives the decision table for this Model Average Composite scenario, showing the probabilities that $B_{2019}$ will exceed various reference points. Figure F. 46 shows the stock status $B_{2017} / B_{\text {avg }}$ of the Model Average Composite scenario and the 6 models that contribute to the composite model. Finally, Tables F.20-F. 25 show the 2-y decision tables for scenarios contributing to the composite.

Table F.17. BC South: The $5^{t h}, 50^{t h}$, and $95^{t h}$ percentiles of MCMC-derived quantities from 6000 MCMC samples comprising the Model Average Composite scenario. Definitions: $B_{2017}$ - current year spawning biomass, $B_{\text {avg }}$ - average biomass from 1967 to 2016, $B_{\min }$ - minimum biomass that acts as the LRP (and $U S R=2 L R P$ ), $u_{2016}$ - harvest rate (ratio of total catch to vulnerable biomass) in the middle of 2016, and $u_{\text {avg }}$ - average harvest rate from 1967 to 2016. All biomass values are in tonnes. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

|  | $5 \%$ | $50 \%$ | $95 \%$ |
| :--- | ---: | ---: | ---: |
| Model-based |  |  |  |
| $B_{2017}$ | 12,737 | 28,923 | 317,629 |
| $B_{\text {avg }}$ | 16,938 | 33,487 | 292,976 |
| $B_{2017} / B_{\text {avg }}$ | 0.589 | 0.899 | 1.35 |
| $u_{2016}$ | 0.00787 | 0.0829 | 0.171 |
| HRP-based |  |  |  |
| $B_{\text {min }}$ | 1,543 | 6,520 | 58,110 |
| $2 B_{\text {min }}$ | 3,086 | 13,041 | 116,219 |
| $B_{\min } / B_{\text {avg }}$ | 0.0753 | 0.138 | 0.296 |
| $2 B_{\min } / B_{\text {avg }}$ | 0.150 | 0.277 | 0.593 |
| $B_{2017} / B_{\text {min }}$ | 2.33 | 6.77 | 10.8 |
| $u_{\text {avg }}$ | 0.0113 | 0.119 | 0.195 |
| $u_{2016} / u_{\text {avg }}$ | 0.589 | 0.772 | 1.00 |

Table F.18. BC South: Decision table for the Model Average Composite scenario for 5 reference points the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2018}$ and mid-year harvest rate $u_{2017}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities are the proportion of MCMC samples from 6 pooled scenarios chosen for their well-behaved MCMC diagnostics. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)$ $=0.96, P\left(B_{2017}>B_{\text {avg }}\right)=0.34$, and $P\left(u_{2016}>u_{\text {avg }}\right)=0.05$. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Catch | $\mathrm{P}\binom{B_{2018}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2018}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2018}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2018}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2017}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.05 | 1 | 0.96 | 0.20 | 0 |
| 500 | 0.03 | 1 | 0.95 | 0.19 | 0 |
| 1000 | 0.02 | 1 | 0.95 | 0.18 | 0 |
| 1500 | 0.01 | 1 | 0.95 | 0.16 | 0 |
| 1750 | 0.01 | 1 | 0.95 | 0.16 | 0.02 |
| 2000 | 0.01 | 1 | 0.95 | 0.16 | 0.07 |
| 2250 | 0.01 | 1 | 0.95 | 0.15 | 0.20 |
| 2500 | 0.01 | 1 | 0.95 | 0.15 | 0.38 |
| 2750 | 0.01 | 0.99 | 0.95 | 0.15 | 0.56 |
| 3000 | 0 | 0.99 | 0.95 | 0.14 | 0.73 |
| 3250 | 0 | 0.99 | 0.95 | 0.14 | 0.85 |
| 3500 | 0 | 0.99 | 0.95 | 0.13 | 0.93 |
| 4000 | 0 | 0.99 | 0.95 | 0.13 | 0.99 |
| 4500 | 0 | 0.99 | 0.94 | 0.12 | 1 |
| 5000 | 0 | 0.98 | 0.94 | 0.11 | 1 |
| 5500 | 0 | 0.98 | 0.94 | 0.11 | 1 |
| 6000 | 0 | 0.98 | 0.93 | 0.10 | 1 |
| 6500 | 0 | 0.97 | 0.93 | 0.10 | 1 |
| 7000 | 0 | 0.97 | 0.93 | 0.10 | 1 |
| 8000 | 0 | 0.97 | 0.91 | 0.09 | 1 |
| 9000 | 0 | 0.96 | 0.87 | 0.09 | 1 |
| 10000 | 0 | 0.94 | 0.82 | 0.08 | 1 |

Table F.19. BC South: Decision table for the Model Average Composite scenario for 5 reference points the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities are the proportion of MCMC samples from 6 pooled scenarios chosen for their well-behaved MCMC diagnostics. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)$ $=0.96, P\left(B_{2017}>B_{\text {avg }}\right)=0.34$, and $P\left(u_{2016}>u_{\text {avg }}\right)=0.05$. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.04 | 1 | 0.95 | 0.09 | 0 |
| 500 | 0.02 | 1 | 0.94 | 0.08 | 0 |
| 1000 | 0.01 | 1 | 0.94 | 0.07 | 0 |
| 1500 | 0 | 0.99 | 0.94 | 0.06 | 0.04 |
| 1750 | 0 | 0.99 | 0.94 | 0.05 | 0.15 |
| 2000 | 0 | 0.99 | 0.93 | 0.05 | 0.32 |
| 2250 | 0 | 0.98 | 0.93 | 0.05 | 0.52 |
| 2500 | 0 | 0.98 | 0.93 | 0.05 | 0.70 |
| 2750 | 0 | 0.98 | 0.93 | 0.04 | 0.84 |
| 3000 | 0 | 0.98 | 0.92 | 0.04 | 0.93 |
| 3250 | 0 | 0.97 | 0.92 | 0.04 | 0.97 |
| 3500 | 0 | 0.97 | 0.91 | 0.04 | 0.99 |
| 4000 | 0 | 0.96 | 0.89 | 0.04 | 1 |
| 4500 | 0 | 0.95 | 0.84 | 0.03 | 1 |
| 5000 | 0 | 0.94 | 0.79 | 0.03 | 1 |
| 5500 | 0 | 0.91 | 0.73 | 0.03 | 1 |
| 6000 | 0 | 0.87 | 0.67 | 0.03 | 1 |
| 6500 | 0 | 0.81 | 0.63 | 0.02 | 1 |
| 7000 | 0 | 0.76 | 0.58 | 0.02 | 1 |
| 8000 | 0 | 0.69 | 0.52 | 0.02 | 1 |
| 9000 | 0 | 0.63 | 0.48 | 0.02 | 1 |
| 10000 | 0 | 0.59 | 0.46 | 0.02 | 1 |



Figure F.46. BC South: Status of the current stock $B_{2017}$ relative to $B_{\text {avg }}$ with the dashed lines showing historical reference points ( $B_{\min } / B_{\text {avg }}, 2 B_{\min } / B_{\text {avg }}$ ) that mimic DFO Precautionary Approach provisional MSY-based reference points, and $0.4 B_{\text {avg }}$ as a proxy for $0.4 B_{0}$. Stock status is shown for the Model Average Composite scenario comprising 6 pooled model runs and for each of the 6 model runs (see Table F.13) for definitions of these model runs). Boxplots show the 5, 25, 50, 75 and 95 percentiles from the MCMC results. $M=$ instantaneous natural mortality $\left(y^{-1}\right) ; k=$ age $(y)$ at knife-edge recruitment.

Table F.20. BC South: Decision table for case S00: M.30+k3 for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)=0.99, P\left(B_{2017}>B_{\text {avg }}\right)=0.53$, and $P\left(u_{2016}>u_{\mathrm{avg}}\right)=0.02$. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0 | 1 | 0.96 | 0.12 | 0 |
| 500 | 0 | 1 | 0.95 | 0.11 | 0 |
| 1000 | 0 | 1 | 0.95 | 0.10 | 0 |
| 1500 | 0 | 1 | 0.94 | 0.10 | 0 |
| 1750 | 0 | 1 | 0.94 | 0.10 | 0.01 |
| 2000 | 0 | 1 | 0.94 | 0.10 | 0.08 |
| 2250 | 0 | 1 | 0.94 | 0.09 | 0.24 |
| 250 | 0 | 1 | 0.93 | 0.08 | 0.49 |
| 2750 | 0 | 1 | 0.93 | 0.08 | 0.71 |
| 3000 | 0 | 1 | 0.93 | 0.08 | 0.86 |
| 3250 | 0 | 1 | 0.93 | 0.08 | 0.95 |
| 3500 | 0 | 1 | 0.93 | 0.07 | 0.99 |
| 4000 | 0 | 1 | 0.92 | 0.07 | 1 |
| 4500 | 0 | 1 | 0.92 | 0.06 | 1 |
| 5000 | 0 | 0.99 | 0.92 | 0.06 | 1 |
| 5500 | 0 | 0.99 | 0.91 | 0.05 | 1 |
| 6000 | 0 | 0.99 | 0.91 | 0.05 | 1 |
| 6500 | 0 | 0.99 | 0.90 | 0.04 | 1 |
| 7000 | 0 | 0.99 | 0.90 | 0.04 | 1 |
| 8000 | 0 | 0.99 | 0.89 | 0.04 | 1 |
| 900 | 0 | 0.98 | 0.88 | 0.03 | 1 |
| 10000 | 0 | 0.97 | 0.86 | 0.03 | 1 |

Table F.21. BC South: Decision table for case S01: M.30+k4 for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)=0.94, P\left(B_{2017}>B_{\text {avg }}\right)=0.14$, and $P\left(u_{2016}>u_{\mathrm{avg}}\right)=0.10$. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.07 | 1 | 0.92 | 0.03 | 0 |
| 500 | 0.03 | 1 | 0.92 | 0.02 | 0 |
| 1000 | 0.02 | 0.99 | 0.92 | 0.01 | 0 |
| 1500 | 0.01 | 0.98 | 0.92 | 0.01 | 0.06 |
| 1750 | 0 | 0.98 | 0.92 | 0 | 0.29 |
| 2000 | 0 | 0.97 | 0.92 | 0 | 0.58 |
| 2250 | 0 | 0.97 | 0.92 | 0 | 0.81 |
| 250 | 0 | 0.96 | 0.92 | 0 | 0.92 |
| 2750 | 0 | 0.96 | 0.92 | 0 | 0.97 |
| 3000 | 0 | 0.95 | 0.91 | 0 | 0.99 |
| 3250 | 0 | 0.94 | 0.90 | 0 | 1 |
| 3500 | 0 | 0.94 | 0.89 | 0 | 1 |
| 4000 | 0 | 0.92 | 0.85 | 0 | 1 |
| 4500 | 0 | 0.90 | 0.78 | 0 | 1 |
| 5000 | 0 | 0.88 | 0.66 | 0 | 1 |
| 5500 | 0 | 0.82 | 0.53 | 0 | 1 |
| 6000 | 0 | 0.74 | 0.41 | 0 | 1 |
| 6500 | 0 | 0.62 | 0.32 | 0 | 1 |
| 7000 | 0 | 0.53 | 0.24 | 0 | 1 |
| 8000 | 0 | 0.38 | 0.12 | 0 | 1 |
| 9000 | 0 | 0.28 | 0.07 | 0 | 1 |
| 10000 | 0 | 0.21 | 0.04 | 0 | 1 |

Table F.22. BC South: Decision table for case S03: M.25+k3 for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)=0.98, P\left(B_{2017}>B_{\text {avg }}\right)=0.50$, and $P\left(u_{2016}>u_{\mathrm{avg}}\right)=0.02$. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0 | 1 | 0.95 | 0.20 | 0 |
| 500 | 0 | 1 | 0.95 | 0.18 | 0 |
| 1000 | 0 | 1 | 0.94 | 0.15 | 0 |
| 1500 | 0 | 1 | 0.93 | 0.14 | 0 |
| 1750 | 0 | 1 | 0.93 | 0.13 | 0 |
| 2000 | 0 | 1 | 0.93 | 0.12 | 0.05 |
| 2250 | 0 | 1 | 0.93 | 0.11 | 0.19 |
| 250 | 0 | 1 | 0.92 | 0.11 | 0.43 |
| 2750 | 0 | 1 | 0.92 | 0.10 | 0.67 |
| 3000 | 0 | 1 | 0.91 | 0.10 | 0.83 |
| 3250 | 0 | 1 | 0.91 | 0.09 | 0.93 |
| 3500 | 0 | 1 | 0.91 | 0.09 | 0.97 |
| 4000 | 0 | 1 | 0.90 | 0.08 | 0.99 |
| 4500 | 0 | 1 | 0.90 | 0.07 | 1 |
| 5000 | 0 | 0.99 | 0.90 | 0.06 | 1 |
| 5500 | 0 | 0.99 | 0.89 | 0.05 | 1 |
| 6000 | 0 | 0.99 | 0.88 | 0.05 | 1 |
| 6500 | 0 | 0.99 | 0.88 | 0.04 | 1 |
| 7000 | 0 | 0.98 | 0.87 | 0.04 | 1 |
| 8000 | 0 | 0.97 | 0.86 | 0.03 | 1 |
| 9000 | 0 | 0.95 | 0.84 | 0.02 | 1 |
| 10000 | 0 | 0.93 | 0.80 | 0.02 | 1 |

Table F.23. BC South: Decision table for case S04: M.25+k4 for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)=0.94, P\left(B_{2017}>B_{\text {avg }}\right)=0.24$, and $P\left(u_{2016}>u_{\mathrm{avg}}\right)=0.07$. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.07 | 1 | 0.93 | 0.08 | 0 |
| 500 | 0.03 | 1 | 0.93 | 0.06 | 0 |
| 1000 | 0.01 | 0.99 | 0.92 | 0.04 | 0 |
| 1500 | 0.01 | 0.98 | 0.92 | 0.02 | 0.03 |
| 1750 | 0 | 0.98 | 0.92 | 0.01 | 0.17 |
| 2000 | 0 | 0.98 | 0.92 | 0.01 | 0.44 |
| 2250 | 0 | 0.97 | 0.92 | 0.01 | 0.66 |
| 250 | 0 | 0.97 | 0.92 | 0 | 0.83 |
| 2750 | 0 | 0.97 | 0.92 | 0 | 0.94 |
| 3000 | 0 | 0.96 | 0.92 | 0 | 0.98 |
| 3250 | 0 | 0.95 | 0.91 | 0 | 0.99 |
| 3500 | 0 | 0.95 | 0.91 | 0 | 1 |
| 4000 | 0 | 0.94 | 0.89 | 0 | 1 |
| 4500 | 0 | 0.93 | 0.83 | 0 | 1 |
| 5000 | 0 | 0.90 | 0.74 | 0 | 1 |
| 5500 | 0 | 0.86 | 0.63 | 0 | 1 |
| 6000 | 0 | 0.79 | 0.52 | 0 | 1 |
| 6500 | 0 | 0.68 | 0.43 | 0 | 1 |
| 7000 | 0 | 0.58 | 0.33 | 0 | 1 |
| 8000 | 0 | 0.43 | 0.16 | 0 | 1 |
| 9000 | 0 | 0.27 | 0.07 | 0 | 1 |
| 10000 | 0 | 0.18 | 0.03 | 0 | 1 |

Table F.24. BC South: Decision table for case S06: M.35+k3 for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\min }$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)=1, P\left(B_{2017}>B_{\text {avg }}\right)=0.53$, and $P\left(u_{2016}>u_{\mathrm{avg}}\right)=0.01$. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0 | 1 | 0.98 | 0.09 | 0 |
| 500 | 0 | 1 | 0.97 | 0.09 | 0 |
| 1000 | 0 | 1 | 0.97 | 0.09 | 0 |
| 1500 | 0 | 1 | 0.97 | 0.08 | 0 |
| 1750 | 0 | 1 | 0.97 | 0.08 | 0.02 |
| 2000 | 0 | 1 | 0.97 | 0.08 | 0.11 |
| 2250 | 0 | 1 | 0.97 | 0.08 | 0.36 |
| 250 | 0 | 1 | 0.97 | 0.07 | 0.60 |
| 2750 | 0 | 1 | 0.97 | 0.07 | 0.79 |
| 3000 | 0 | 1 | 0.97 | 0.07 | 0.90 |
| 3250 | 0 | 1 | 0.97 | 0.07 | 0.95 |
| 3500 | 0 | 1 | 0.97 | 0.07 | 0.98 |
| 4000 | 0 | 1 | 0.97 | 0.06 | 1 |
| 4500 | 0 | 1 | 0.97 | 0.06 | 1 |
| 5000 | 0 | 1 | 0.97 | 0.06 | 1 |
| 5500 | 0 | 1 | 0.96 | 0.06 | 1 |
| 6000 | 0 | 1 | 0.96 | 0.06 | 1 |
| 6500 | 0 | 1 | 0.96 | 0.06 | 1 |
| 7000 | 0 | 1 | 0.96 | 0.05 | 1 |
| 8000 | 0 | 1 | 0.96 | 0.05 | 1 |
| 9000 | 0 | 1 | 0.96 | 0.05 | 1 |
| 10000 | 0 | 1 | 0.95 | 0.04 | 1 |

Table F.25. BC South: Decision table for case S07: M.35+k4 for 5 reference points - the current year spawning biomass, the limit reference point $B_{\min }$, the upper stock reference $2 B_{\mathrm{min}}$, the average spawning stock biomass from 1967 to 2016, and the average harvest rate over the same time period - for projection-year biomass $B_{2019}$ and mid-year harvest rate $u_{2018}$ for a range of constant catch strategies (in tonnes). Each value is the probability that projected biomass or harvest rate is greater than the indicated reference point. The probabilities that current-year spawning biomass (or harvest rate) is greater than the reference points are: $P\left(B_{2017}>B_{\min }\right)=1, P\left(B_{2017}>2 B_{\min }\right)=0.94, P\left(B_{2017}>B_{\text {avg }}\right)=0.11$, and $P\left(u_{2016}>u_{\mathrm{avg}}\right)=0.09$. For reference, the average catch over the last 5 years (2011-2015) is $3256 t$.

| Catch | $\mathrm{P}\binom{B_{2019}>}{B_{2017}}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{2 B_{\min }}$ | $\mathrm{P}\binom{B_{2019}>}{B_{\text {avg }}}$ | $\mathrm{P}\binom{u_{2018}>}{u_{\text {avg }}}$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 0.07 | 1 | 0.93 | 0.02 | 0 |
| 500 | 0.03 | 0.99 | 0.93 | 0.01 | 0 |
| 1000 | 0.02 | 0.99 | 0.93 | 0.01 | 0 |
| 1500 | 0.01 | 0.98 | 0.93 | 0.01 | 0.13 |
| 1750 | 0.01 | 0.97 | 0.93 | 0 | 0.38 |
| 2000 | 0.01 | 0.96 | 0.93 | 0 | 0.66 |
| 2250 | 0.01 | 0.96 | 0.93 | 0 | 0.84 |
| 250 | 0 | 0.95 | 0.92 | 0 | 0.93 |
| 2750 | 0 | 0.95 | 0.91 | 0 | 0.98 |
| 3000 | 0 | 0.94 | 0.90 | 0 | 0.99 |
| 3250 | 0 | 0.94 | 0.88 | 0 | 1 |
| 3500 | 0 | 0.94 | 0.86 | 0 | 1 |
| 4000 | 0 | 0.92 | 0.79 | 0 | 1 |
| 4500 | 0 | 0.89 | 0.68 | 0 | 1 |
| 5000 | 0 | 0.85 | 0.57 | 0 | 1 |
| 5500 | 0 | 0.77 | 0.46 | 0 | 1 |
| 6000 | 0 | 0.69 | 0.36 | 0 | 1 |
| 6500 | 0 | 0.59 | 0.27 | 0 | 1 |
| 7000 | 0 | 0.50 | 0.19 | 0 | 1 |
| 8000 | 0 | 0.37 | 0.12 | 0 | 1 |
| 9000 | 0 | 0.29 | 0.07 | 0 | 1 |
| 1000 | 0 | 0.24 | 0.05 | 0 | 1 |

## F.4. REFERENCES - MODEL RESULTS

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[^0]:    ${ }^{1}$ Saunders, M.W. and Andrews, W. 1998. Walleye pollock stock assessment for 1997 and recommended yield options for 1998. PSARC Working Paper G:97-7, 18 pp, Pacific Biological Station, DFO, Nanaimo BC, Canada.

[^1]:    ${ }^{2}$ Strictly speaking, this survey did not operate in the stock definition area for the BC North stock. However, because this series is the only set of early abundance information, two blocks of stock assessment runs were made for this stock, one with and one without this survey.

[^2]:    ${ }^{3}$ MacLellan, S.E., Gillespie, D., Janz, S., Charles, K., Little, D. and Rankin, J. 1990. Age determination of various freshwater fish species being analyzed for dioxin and furan studies in B.C., 1989. Unpub. tech. rep., Pacific Biological Station, DFO, Nanaimo BC, Canada.

[^3]:    ${ }^{4}$ Saunders, M.W. and Andrews, W. 1998. Walleye Pollock stock assessment for 1997 and recommended yield options for 1998. PSARC Working Paper G:97-7, 18 pp., Department of Fisheries and Oceans, Pacific Biological Station.

[^4]:    *See Table A. 2 for management actions indicated by note letter.

[^5]:    ${ }^{5}$ Starr, P.J., Kronlund, A.R., Workman, G., Olsen, N. and Fargo, J. 2006. Rock sole (Lepidopsetta spp.) in BC, Canada: Stock assessment for 2005 and advice to managers for 2006/2007. Pacific Scientific Advice Review Committee (PSARC) Working Paper, unpublished manuscript, Pacific Biological Station, DFO, Nanaimo BC, Canada.

[^6]:    ${ }^{1}$ valid biomass estimation tows only
    ${ }^{2}$ includes tows not used for biomass estimation

[^7]:    ${ }^{1}$ assumes $M=0.3$ for comparative purposes.

