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# Turbot Stock Assessment for 2001 and Recommendations for Management 

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#### Abstract

We summarise results of analysis of biological data, research survey data and observer data for turbot (Atheresthes stomias). Results from the analysis of observer data indicate that the variability in the CPUE index from the commercial fishery is mainly due to area and seasonal effects (and in some cases depth) and that these variables can change among years. The analysis of the discard data indicates that this fishery appears to be passive with little evidence of targeting behaviour. The turbot CPUE index from research surveys conducted between 1984 and 2000 shows no trend but does show cyclic fluctuation over that period. Size and age composition data show no distinct trends over time. The instantaneous total mortality rate estimates for 1980, 1998 and 2000 did not differ despite a twenty year exploitation history. We conclude that the current fishing mortality rate for turbot stocks off the West Coast of Canada is at or below the sustainable level.


## Résumé

Nous résumons les résultats de l'analyse de données biologiques, de données de relevés de recherche et de données d'observateurs pour la plie à grande bouche (Atheresthes stomias). Les résultats de l'analyse des données d'observateurs révèlent que la variabilité de l'indice CPUE issu de la pêche commerciale est principalement imputable aux effets du secteur et de la saison de pêche (et dans certains cas de la profondeur) et que ces variables peuvent changer d'une année à l'autre. L'analyse des données sur les rejets indique que cette pêche semble être opportuniste, ne manifestant qu'une faible tendance à être ciblée. L'indice CPUE de plie à grande bouche issu des relevés de recherche menés entre 1984 et 2000 ne révèle aucune tendance mais fluctue toutefois de façon cyclique pendant cette période. Les données sur la composition par taille et par âge ne montrent aucune tendance distincte au fil du temps. Les estimations du taux instantané de mortalité totale pour 1980, 1998 et 2000 ne sont pas différentes en dépit d'un dossier de pêche de 20 ans. Nous concluons que le taux actuel de mortalité par pêche imposé aux stocks de plie à grande bouche de la côte ouest du Canada se situe au ou sous le niveau durable

### 1.0 Introduction

The groundfish resource in British Columbia (B.C.) increased in importance in the late 1970s with the implementation of Extended Jurisdiction in 1977 and subsequent expansion of the domestic fleet. This prompted the first assessments of groundfish in 1977 (Westrheim 1977). Recommendations for quota management of groundfish species were not forthcoming until 1979 (Ketchen 1980). Since that time, detailed and interim assessments have been conducted annually including recommendations for catch limitations. The last detailed assessment of turbot (also known as arrowtooth flounder) was conducted in 1987 (Fargo 1988).

Turbot is an important component of the trawl fishery. Over the past 45 years most of the turbot catch has been discarded at sea. The species is soft-fleshed and proteolysis occurs in the muscle tissue of turbot soon after death producing a mush flesh. Although alternative processing methods have been developed for the species there is no large market for these fillets.

Although turbot are of limited economic importance, they are an important ecological component of the offshore ecosystem as well as the ecosystem in Hecate Strait. This is particularly relevant as investigators shift their emphasis from single species to multi-species or ecosystem assessment. Previous studies indicate that the major food item in the diet of the adults is fish although they also consume cephalopods, euphausiids and shrimp. Juvenile turbot are prey items for large pollock, and Pacific cod.

In this document, we summarise biological information and present the results of an analysis of catch-effort, survey and biological data. We use these analyses to provide advice to managers on harvest levels for the 2002/2003 fishing year.

### 2.0 Background

### 2.1 Range and stock structure

Turbot (Atheresthes stomias) ranges from Baja California to the eastern Bering Sea and is most abundant at the northern part of its range. Off B.C., turbot show a preference for a narrow range of bottom temperature of $7-8^{\circ} \mathrm{C}$. The species shows a preference for coarse sand substrate as well (Perry et al. 1994). The species occupies the waters of the continental shelf until about age 4 when juveniles join the adult portion of the population occupying the continental slope. Limited research has been carried out on the species' life history. Little is known about the stock structure of this species on the Pacific coast.

### 2.2 Niche

Turbot inhabit depths from 50 to 900 m and show little preference for bottom temperature over their range. Adults show little affinity for a particular sediment type, while juveniles prefer sand or mud substrate. Juveniles feed primarily on mobile prey, such as cumaceans, carideans, and gammarid amphipods. Adults are piscavores and cannabalistic. Their preferred prey is herring (Clupea harengus), juvenile pollock (Theagra chalcogramma) and Pacific sandlance (Ammodytes hexapterus) (Fargo et al. 1981). Diet variation results as much from food availability as it does from prey preference.

### 2.3 Growth/Lifespan

Turbot exhibit sexual dimorphism. As juveniles ( $<380 \mathrm{~mm}$ ), males and females grow at the same rate. However, after sexual maturity, females grow faster than males, and attain a larger maximum size (Figures 1 and 2). The maximum size observed for males in biological samples over the last 20 years is 750 mm . The maximum size for females over the same period is 840 mm . The maximum weights of males and females determined from biological samples are 1676 g and 3102 g , respectively. The growth in weight is also similar among the sexes until maturation (age 5). Thereafter the weight gain by females is significantly higher than for males.

Age at $50 \%$ maturity for males is 4 years and for females is 5 years (Figure 3). Recruitment to the exploitable population begins at age 5 but recruitment is not knife-edged. The maximum age recorded for this species in B.C. is 25 years, although most of the exploited population is less than 15 years of age.

### 2.4 Reproduction

Turbot are batch spawners and peak spawning occurs at depths deeper than 350 m in the fall and winter months (Rickey 1995). The mature population off Washington migrates seasonally from depths of about 183 m in summer to depths exceeding 475 m in winter (Rickey 1995). There is some evidence to suggest that the time of first spawning and the time of peak spawning vary interannually (Rickey 1995). The species produces pelagic eggs that drift with the ocean currents. There is some evidence of an extended larval period of several months however (Rickey 1995). Fecundity of the species is not known (DiCosimo 1998).

Larvae remain in the upper 100 m of the water column for about 4 weeks. Young of the year one and two year olds occupy shallower depths than the adults, while three and four year olds are generally found in deeper water with the adults. Turbot occupy separate spawning (winter) and feeding (summer) areas. They undergo a seasonal bathymetric movement from shallower water to deeper depths in the fall and winter.

### 3.0 Data Sources

### 3.1 Commercial trawl data

The Department of Fisheries and Oceans has maintained records of groundfish catch and effort data from 1954 to 1995 using a combination of voluntary skipper interviews, vessel logbooks, landings records (sales slips or validation records) and observations at the waterfront. These data are archived in a database called GFCATCH, the history of which has recently been described in detail by Rutherford (1999).

Skipper interviews and logbooks provided information on fishing areas and amount of effort, however, the catch for each species was estimated. Species composition was usually limited to the dominant species retained in the catch (Rutherford 1999). Skipper interview and logbook data were transcribed into a trip report by DFO staff. Sales slips or validation records provided accurate weights of species landed, but little information on fishing location or effort. If an offload was observed, information might be gathered that supplemented or superseded logbooks and landing records. For example, errors in species identification might be corrected. The "best" estimate of catch required synthesis of all data sources. Typically, the actual weights from landings were used to adjust the trip reports by prorating the landed weights using fishing location and catch information recorded at sea (Rutherford 1999).

### 3.2 Commercial trawl observer data 1996-2000: PacHarv database

A mandatory at-sea observer program was implemented for most Option A and some Option B trawl vessels in 1996. This includes some $90 \%$ of the trawl fleet. The observers provide information on catch locations, bridge log data and species composition (by weight). Observers also collect biological data for selected species. A relational database, PacHarvest, was developed by the slope rockfish assessment team using Microsoft Server 7.0 (Schnute et al. 1999). The database is located on the Windows NT server PacStad at the Pacific Biological Station, Nanaimo, B.C. Documentation and database shells for connecting to PacHarvest can be found on the DFO Intranet at http://pacstad/pacharvdb/Default.htm. Further details can be found on the website and in Schnute et al. (1999).

A detailed explanation of catch and effort data used in the general linear modelling section of this document is provided in Appendix 1, including a description of the data selection and grooming procedures.

### 3.3 Dockside validation

Since 1996 every trawler unloading is monitored at the port of landing. The dockside validator estimates the species composition of the landing by weight. This information is used together with observer at-sea information to resolve the species composition (by weight) of the catch. Dockside validation data for trawl is contained in the database tables B5_Validation_Headers and B6_Validation Species of the PacHarv database described above.

### 3.4 Landing statistics

Annual landing statistics for turbot are presented by PFMC statistical area in Table 1. The landings exhibit cyclic fluctuations between the mid 1950s and the mid 1980s. Between the late 1980s and 2001 landings have increased to triple the previous high.

Table 1. Turbot landings and discards by Major Statistical Area and calendar year, 1954-2001.

| Year | Major Area |  |  |  |  |  |  |  | Total | Discard |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 4B | 3C | 3D | 5A | 5B | 5C | 5D | 5E |  |  |
| 54 | 104 | 96 | - | 274 | 10 | - | 49 | - | 533 |  |
| 55 | 68 | 304 | 26 | 747 | - | - | 177 | - | 1322 |  |
| 56 | 81 | 741 | 55 | 730 | 10 | - | 507 | - | 2124 |  |
| 57 | 19 | 132 | - | 236 | - | 2 | 139 | - | 528 |  |
| 58 | 55 | 119 | - | 35 | 1 | 12 | 103 | - | 325 |  |
| 59 | 80 | 156 | 0 | 178 | 1 | - | 292 | - | 707 |  |
| 60 | 92 | 96 | 5 | 318 | 38 | 8 | 369 | - | 926 |  |
| 61 | 48 | 335 | 1 | 1089 | 2 | 4 | 638 | - | 2117 |  |
| 62 | 78 | 185 | 23 | 452 | 25 | 1 | 591 | - | 1355 |  |
| 63 | 101 | 48 | - | 385 | 9 | - | 77 | - | 620 |  |
| 64 | 133 | 29 | 4 | 457 | 9 | 4 | 195 | - | 831 |  |
| 65 | 137 | 104 | 5 | 278 | 7 | 1 | 145 | - | 677 |  |
| 66 | 248 | 51 | 1 | 133 | 1 | - | 252 | - | 686 |  |
| 67 | 86 | 65 | 0 | 1049 | 22 | - | 417 | - | 1639 |  |
| 68 | 138 | 31 | 1 | 301 | - | 85 | 386 | - | 942 |  |
| 69 | 114 | 41 | 22 | 1156 | 48 | 19 | 508 | - | 1908 |  |
| 70 | 51 | - | - | 42 | 9 | 2 | 212 | - | 316 |  |
| 71 | 11 | 4 | - | - | 2 | 0 | 20 | - | 37 |  |
| 72 | 3 | 1 | - | 27 | - | - | 202 | - | 233 |  |
| 73 | 14 | 19 | - | 37 | - | 16 | 342 | - | 428 |  |
| 74 | 4 | 17 | - | 9 | 62 | 17 | 224 | - | 333 |  |
| 75 | 18 | 83 | 2 | 147 | 58 | 6 | 457 | - | 771 |  |
| 76 | 1 | 121 | 1 | 40 | 86 | 43 | 760 | - | 1052 |  |
| 77 | 1 | 275 | 5 | 150 | 98 | 24 | 843 | 26 | 1422 |  |
| 78 | 30 | 203 | 4 | 385 | 262 | 77 | 1294 | 27 | 2282 |  |
| 79 | 5 | 40 | 23 | 138 | 52 | 182 | 1265 | 7 | 1712 |  |
| 80 | 12 | 85 | 5 | 121 | 170 | 80 | 821 | 4 | 1298 |  |
| 81 | 18 | 69 | 15 | 151 | 132 | 64 | 398 | 6 | 853 |  |
| 82 | 9 | 66 | - | 137 | 143 | 11 | 139 | - | 505 |  |
| 83 | 8 | 31 | - | 52 | 75 | 9 | 92 | 2 | 269 |  |
| 84 | 4 | 23 | - | 20 | 49 | 11 | 199 | 8 | 314 |  |
| 85 | 0 | 73 | 11 | 35 | 23 | 2 | 435 | 12 | 591 |  |
| 86 | - | 131 | 6 | 190 | 11 | 6 | 367 | 14 | 725 |  |
| 87 | - | 37 | 9 | 276 | 100 | 26 | 641 | 15 | 1104 |  |
| 88 | 3 | 31 | 30 | 116 | 45 | 3 | 120 | 19 | 367 |  |
| 89 | - | 38 | 34 | 144 | 348 | 4 | 30 | 11 | 609 |  |
| 90 | - | 325 | 379 | 527 | 521 | 12 | 781 | 16 | 2561 |  |
| 91 | 2 | 218 | 639 | 318 | 881 | 23 | 152 | 27 | 2260 |  |
| 92 | 3 | 560 | 989 | 761 | 1045 | 62 | 149 | 24 | 3593 |  |
| 93 | 3 | 410 | 1828 | 388 | 1023 | 16 | 336 | 47 | 4051 |  |
| 94 | 2 | 825 | 1425 | 280 | 1020 | 40 | 300 | 81 | 3973 |  |
| 95 | 0 | 315 | 1267 | 447 | 583 | 24 | 326 | 46 | 3008 |  |
| 96 | 0 | 458 | 2529 | 368 | 799 | 56 | 343 | 6 | 4558 | 3271 |
| 97 | 0 | 581 | 914 | 434 | 624 | 30 | 277 | 25 | 2885 | 2453 |
| 98 | 0 | 632 | 1778 | 444 | 623 | 3 | 325 | 7 | 3812 | 3284 |
| 99 | 0 | 506 | 878 | 463 | 860 | 31 | 777 | 19 | 3555 | 3818 |
| 00 | 0 | 662 | 1082 | 377 | 1644 | 10 | 466 | 14 | 4255 | 3561 |
| 01 | 0 | 2165 | 2100 | 55 | 1277 | 4 | 797 | 65 | 6464 | 1124 |

### 3.5 Biological data

Biological samples containing length, sex, maturity, and ageing information have been collected from the trawl fishery in British Columbia continuously since the mid-1940s. Samples collected on turbot, however, were too few in number to warrant analysis and we used biological data collected from research cruises for the analysis in this assessment.

Biological samples for turbot were most numerous during research cruises conducted in Hecate Strait including the previously mentioned Hecate Strait survey and turbot biomass survey. In the early years of the Hecate Strait survey attention was devoted to sampling every species in the catch to provide data for multispecies stock assessment work. This permitted the collection of size composition data only. Since 1998, however, biological samples for important commercial species have included information on length, sex, stage of maturity and ageing structures.

### 3.6 Research surveys

Two surveys in Hecate Strait provide some information on turbot. In 1980 a turbot biomass survey was conducted in Hecate Strait (Fargo et al. 1981). Biological information collected on that survey included age, length, sex and maturity data. In 1984 a multispecies trawl survey was initiated in Hecate Strait (Fargo and Tyler 1990). This work was carried out as part of the Hecate Strait Project (Fargo 1986, Fargo 1989) with an objective to develop an ecological basis for mixed species stock assessment. The survey provides synoptic data that allowed the mapping of fish assemblages available to bottom trawls in that region. Although the Hecate Strait Project work halted in 1993, the multispecies survey was continued (Hand et al. 1994, Workman et al. 1996, Wilson et al. 1991). The survey now provides data on the abundance and distribution for groundfish species in the region. The survey data has also been used to document the spatial and temporal changes in species composition in Hecate Strait. The fishing gear used on the survey, a Yankee 36 bottom trawl, has remained the same since its inception. The net is equipped with a small-mesh codend liner to ensure sampling of all size/age groups.

The survey employs a systematic depth stratified design to achieve broad spatial coverage. A grid of $10 \times 10 \mathrm{~nm}$ blocks was superimposed on a chart of the region. Sampling stations within each block were allocated for each 20 m depth interval. The selection of a station within a stratum was made by the fishing master who searched each stratum for trawlable bottom. At the end of each tow, the species composition of the catch by weight is determined and length measurements were made for all species in the catch. Exceptions to this procedure occurred when the catch was $>3000 \mathrm{lbs}$. whereupon a random subsample was taken for the collection of biological data.

### 4.0 Biological Statistics

Current estimates for life history parameters for turbot are presented in Table 2. These estimates are based on port samples and research samples without reference to sample type or to the underlying representation of the samples in the catch or in the survey biomass. Because the underlying sampling structure has not been taken into account, we have not attempted to include error estimates for these parameters.

### 4.1 Length weight relationship

The expression describing the length weight relationship is:

$$
\ln W_{i}=a+b \ln L_{i},(1 \leq i \leq n)
$$

where $W_{i}$ is the weight $(\mathrm{kg})$ and $L_{i}$ is the length $(\mathrm{mm})$ of fish $i$, were determined from pooled samples for 1980 to 2000. Males rarely reach a size of 600 mm while females commonly reach that size (Figure 1). Weight at age is similar among the sexes until around 300 mm , the time of sexual maturation. Thereafter the females surpass the males in weight at length.


Figure 1. Length-weight relationships for turbot males and females. The data are pooled from survey samples collected from 1980 to 2000.

### 4.2 Length-age relationships

Von Bertalanffy growth curves (Figure 2) were fit to data for both sexes using the equation below where $l_{t}$ is length at age $t, L_{\infty}$ is the ultimate length for the population, $K$ is a growth coefficient and $t_{o}$ is the time when length would theoretically be zero.. Growth in length for turbot males slows markedly after about age seven while females continue to gain significant growth in length until slowing at about age twelve.

$$
l_{t}=L_{\infty}\left[1-e^{-K\left(t-t_{0}\right)}\right]
$$



Figure 2. Length age relationships (jitter plot) for turbot females (upper panel) and males (lower panel).
The data represent pooled survey samples from 1945 to 1999.

### 4.3 Maturity

We estimated length at maturity for the catch-age model using data obtained from samples from research cruises and the commercial fishery. Stage of maturity was determined macroscopically and fish were partitioned into one of seven maturity stages (Workman et al. 1996), two immature and five mature. Fish at stages one and two were treated as immature and fish at stages 3-7 were treated as mature. Length-maturity data were obtained from pooling the data to acquire an adequate sample size. Maturity ogives were fit to these data using a simple logistic
regression (Hosmer and Lemeshowe 1989), where the probability of a fish being mature at a given length $L, P_{L}$, is a function of the length, $L$, and the regression coefficients $\beta_{0}$ and $\beta_{l}$.

$$
P_{L}=\frac{e^{\beta_{0}+\beta_{1} L}}{1+e^{\beta_{0}+\beta_{1} L}}
$$

Males mature at a size of about 310 mm compared to females at about 380 mm . The rate of maturity at length is different among the sexes as well with males maturing faster than females.
$\mathrm{L}_{100}$, the length at which $100 \%$ of the fish are mature, is 430 mm for males and 550 mm for females (Figure 3).


Figure 3. Maturity ogives for male and female turbot. Data are pooled from research survey samples from 1980-2000.

### 4.4 Natural mortality

Pauly (1980) described a multiple regression relationship between $M$ and life history characteristics and temperature regime of 175 fish stocks. He found that $M$ varied linearly with species asymptotic length $L_{\infty}$, the growth coefficient $K$, and mean annual ocean temperature $T$

$$
\log M=0.065-0.287 \log L_{\infty}+0.604 \log K+0.513 \log T
$$

Estimates of $M$ from Pauly's method were 0.38 and 0.28 for males and females, respectively. This is well above the value of 0.2 used for Gulf of Alaska and Washington State assessments (Wilderbuer and Sample 2000). However, many of species that Pauly examined were warm water species and this may have biased his results. We propose to use $M=0.2$ for both sexes of this species as this value is consistent with the estimate that would be obtained using the relationship described by Hoenig (1983) below and the approximate maximum age, $t_{\max }$, observed for this species in B.C. waters of between $20(\mathrm{M}=0.22)$ and $24(\mathrm{M}=0.19)$.

$$
\ln M=0.984 \ln \left(t_{\max }\right)
$$

### 4.5 Size and Age composition

Size composition summaries from the 1980 turbot biomass survey are presented in Figure 4. The sample in 1980 was collected during the turbot biomass survey while the samples from 1984 to 2000 were collected during the Hecate Strait multispecies survey. Interannual variation in the proportion of juveniles ( $<380 \mathrm{~mm}$ ) and adults ( $\geq 380 \mathrm{~mm}$ ) is apparent. The proportion of juveniles was noticeably smaller for the 1980 samples than for the others. The overall size composition has remained fairly consistent since 1984. This may indicate an increase in recruitment due to 1998 and 1999 year-classes. There is no indication from these data that the size range has been truncated over time. We compared the age composition from 1980 with that from 2000 to see if juvenation had occurred (Figure 5). The proportions of younger and older fish were actually higher in the 2000 sample than in the 1980 sample. The sample sizes from samples taken from the commercial fishery samples are so small that no interpretation of them is made.

Table 2. Estimates of biological parameters for turbot caught in the trawl fishery or taken in trawl surveys off the west coast of Canada.

|  | Males |  |  |  |  |  | Females |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| K | 0.278 |  |  |  |  |  | 0.192 |  |  |  |  |  |
| $L^{\infty}$ | 471 mm |  |  |  |  |  | 617 mm |  |  |  |  |  |
| $t_{0}$ | -0.234 |  |  |  |  |  | -0.278 |  |  |  |  |  |
| $M$ | 0.2 |  |  |  |  |  | 0.2 |  |  |  |  |  |
| $w_{j}$ | $\begin{array}{r} 22 . \\ 566 . \\ 799 . \end{array}$ | 89.0 | 186.0 | 290.8 | 395.6 | 485.1 | $\begin{array}{\|r} \hline 21.1 \\ 863.1 \\ 1581.1 \\ \hline \end{array}$ | $\begin{array}{rr} 90.7 & 207.2 \\ 1020.0 & 1162.7 \\ 1651.9 & 1712.5 \\ \hline \end{array}$ |  | $\begin{array}{r} 356.8 \\ 1289.6 \end{array}$ | $\begin{array}{rr} \hline 524.2 & 695.9 \\ 1401.5 & 1497.9 \end{array}$ |  |
|  |  | 629.5 | 678.7 | 720.7 | 754.6 | 779.4 |  |  |  |  |  |  |
|  |  | 815.1 | 825.5 |  |  |  |  |  |  |  |  |  |
| $l_{j}$ | $\begin{aligned} & 136 . \\ & 408 . \\ & 459 . \end{aligned}$ | 217.7 | 279.2 | 325.7 | 361.0 | 387.7 | 134.4 <br> 464.8 <br> 569.1 | $\begin{aligned} & 218.8 \\ & 491.5 \\ & 577.5 \end{aligned}$ | 288.4 | 345.9 | 393.4 | 432.5 |
|  |  | 423.3 | 434.9 | 443.7 | 450.4 | 455.5 |  |  | 513.5 | 531.6 | 546.6 | 558.9 |
|  |  | 462.2 | 464.4 |  |  |  |  |  | 584.5 |  |  |  |


| $p_{j}$ | 0.00 | 0.00 | 0.00 | 0.03 | 0.21 | 0.58 | 0.85 | 0.95 | 0.00 | 0.00 | 0.01 | 0.10 | 0.45 | 0.79 | 0.92 | 0.97 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | 0.98 | 0.99 | 0.99 | 1.00 | 1.00 | 1.00 | 1.00 |  | 0.98 | 0.99 | 0.99 | 0.99 | 1.00 | 1.00 | 1.00 |  |
| $a$ | 0.0000127 |  |  |  |  |  | 0.00000914 |  |  |  |  |  |  |  |  |  |
| $b$ | 2.93 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

where:
$K, L_{\infty}, t_{0}$ are coefficients estimated for the von Bertalanffy growth formulation
$M$ is the instantaneous rate of natural mortality
$w_{j}$ : mean weight at age $j$
$l_{j}$ : mean length at age $j$
$p_{j}$ : mean proportion mature at age $j$
$j$ indexes age groups 1-15
$a, b$ are the length-weight coefficients


Figure 4. Size composition (cm) for turbot in Hecate Strait, 1980-2000.


Figure 5. Age composition of female turbot from research and commercial samples 1980,1997,1998 and 2000.

### 4.6 Total mortality

We estimated $Z$, the instantaneous total mortality rate from the age composition data for females for samples collected on the turbot biomass survey in 1980, samples collected on the 1998 and 2000 multispecies survey and from all research and port samples collected from 1997 - 2001. We used the method of Ricker (1975) where he used log to the base 10 of numbers at age to restrict the $y$-axis to 2.0. The facilitates the comparison of samples for different years. This calculation also requires the assumption that recruitment variation is low, which we believe is justified for this species (Rickey 1995, Wilderbuer and Sample (2000)). We used the regression relationship between the log of the numbers for fully recruited ages on age (Figure 6). To obtain estimates of $Z$ the sign of the slope from the regression was changed and then it was multipied by 2.3 (Ricker 1975). The estimates of $Z$ were all between 0.35 and 0.37 . This implies a range of 0.15 to 0.17 for $F$ assuming that $M$ is 0.2 (the estimate that is used in U.S. stock assessments). Moreover, there was no significant difference in the slopes from the regressions for any of the time periods (analysis of covariance, $\mathrm{p}<0.0001$ ).


Figure 6. Estimates of total mortality, $Z$, for female turbot from research and commercial samples, 19802001.

### 5.0 Analysis of Observer data

### 5.1 Analytical Procedure used for catch/effort data

A stepwise multiple linear regression (where data are modelled assuming lognormal variability) was used to estimate trends in abundance from CPUE data derived from the commercial catch and effort database (see Appendix 1 [Section 10.0] for how these data were generated). This approach is commonly used to analyse fisheries catch and effort data and are described in Hilborn and Walters (1992) and Quinn and Deriso (1999).

Quinn and Deriso (1999) describe a general linear model based on the lognormal distribution:

$$
U=U_{r} \prod_{i} \prod_{j} P_{i j}^{X_{i j}} e^{\varepsilon}
$$

Eq. 1
where $U$ is the observed CPUE, $U_{r}$ is the reference CPUE, $P_{i j}$ is a factor $i$ at level $j$, and $X_{i j}$ is a categorical variable which takes a value of 1 when factor $P_{i j}$ is true and 0 when it is false. $\varepsilon$ is a normal random variable with mean $=0$ and standard deviation $\sigma$.

Taking the logarithm of Eq. 1 gives the following general form for one explanatory factor:

$$
\ln U=\ln U_{r}+\sum_{i} \sum_{j} X_{i j} \ln P_{i j}+\varepsilon
$$

or Eq. 2

$$
Y=\beta_{0}+\sum_{k} \beta_{k} X_{k}+\varepsilon
$$

where the subscript $k$ in the second form of Eq. 2 combines subscripts $i$ and $j$ in the first form, $\beta_{0}$ is the intercept of $\ln (\mathrm{CPUE})$ and $\beta_{k}$ is the logged coefficient of the categorical variable for the factor under consideration.

The model described in Eq. 1and Eq. 2 is overparameterised and can take on an infinite number of solutions. The approach used to overcome this problem in this analysis was to fix one of the $\beta_{k}$ coefficients and to estimate the remainder of the coefficients relative to the fixed coefficient. Practically this is done in the regression model by dropping one coefficient (usually the first) and estimating the model with $k-1$ coefficients. The dropped coefficient will be equal to zero (in log space).

Categorical variable coefficients obtained by dropping one factor will take on different values depending on which coefficients has been dropped. Following the suggestion of Francis (1999), these coefficients are transformed to "canonical" coefficient calculated relative to the geometric mean ( $\bar{\beta}$ )of the series:

$$
\begin{equation*}
\beta_{k}^{0}=\beta_{k} / \bar{\beta} \tag{Eq. 3}
\end{equation*}
$$

As the analysis is done in $\log$ space, this is equivalent to:

$$
b_{k}^{0}=\mathrm{e}^{\left(\hat{\beta}_{k}-\bar{\beta}\right)}
$$

## Eq. 4

where $\hat{\beta}_{k}$ is the coefficient calculated for each value of the predictor variable and $\bar{\beta}$ is the mean of those coefficients, including the dropped coefficient. When this procedure is applied to the annual abundance variable ('year' or 'fishing year'), the resulting set of canonical indices is termed the "Standardised" CPUE index [ $Y_{k}^{0}$ ] in this report.

The use of the canonical form allows the computation of standard errors for every coefficient, including the dropped coefficient (Appendix 2 [Section 11.0]). Ordinarily, the use of a fixed reference coefficient sets the standard error for that coefficient to zero and spreads the error associated with that coefficient to the other coefficients in the variable.

Eq. 2 can be extended to include as many factors as are thought to be reasonable, including interaction terms. A selection procedure has been developed (Vignaux 1993, Vignaux 1994; Francis 2001) to determine the relative importance of these factors in the model and to establish a stopping rule which will include only the most important factors. This procedure involves a forward stepwise fitting algorithm which generates a regression model iteratively, starting with the simplest model (one dependent and one independent variable).

The following procedure was used to fit the models, given a data set with candidate predictor variables:

1. Calculate the regression with each predictor variable against the natural $\log$ of CPUE $(\mathrm{kg} / \mathrm{hr})$.
2. Generate the AIC (Akaike Information Criterion; Akaike 1974) for each regression based on the number of model degrees of freedom. Select the predictor variable that has the lowest AIC.
3. Repeat Steps 1 and 2, accumulating the number of selected predictor variables and increasing the model degrees of freedom, until the increase in residual deviance $\left(=R^{2}\right)$ for the final iteration is less than 0.01 .

The AIC is used for predictor selection to account for variables which may have equivalent explanatory power in terms of residual deviance but add fewer degrees of freedom to the model (Francis 2001).

A direct comparison of a number of alternative estimates of annual CPUE is made by standardising all available indices relative to the geometric mean of the index series. The simplest estimate of mean annual CPUE is:

$$
\begin{equation*}
R_{j}=\frac{\sum_{k=1}^{M_{i}} C_{j k}}{\sum_{k=1}^{M_{j}} E_{j k}} \tag{Eq. 5}
\end{equation*}
$$

where $M_{j}$ is the number of records in the data set for year $j, C_{j k}$ is the catch and $E_{j k}$ is the effort associated with each record in the data set for year $j$. The series of annual abundance indices calculated in this manner is termed the "Arithmetic" CPUE index in this report and is the arithmetic mean of CPUE weighted by effort. This index can also be scaled relative to its geometric mean $(\bar{R})$ in the same manner as the canonical standardised index (Eq. 3):

$$
R_{j}^{0}=R_{j} / \bar{R}
$$

Another simple index of annual abundance based on CPUE is:


Eq. 7
where $U_{j}$ is the annual geometric mean of the CPUE observations. The resulting series of indices is termed the "Unstandardised" CPUE index in this report as it is equivalent to a GLM where the only predictor variable is the year (= abundance) term. This index can also be scaled relative to its geometric mean $(\bar{U})$ in the same manner as the canonical standardised index (Eq. 3):

$$
U_{j}^{0}=U_{j} / \bar{U}
$$

Eq. 8

### 5.2 Catch regions

A preliminary examination of the turbot catch and effort data was made to see if there was a basis on which to separate turbot catch into regions which were consistent with turbot abundance patterns. Accordingly, simple turbot CPUE was calculated for $5 \mathrm{~km}^{2}$ grids (using Eq. 3) over the entire period of available data (January 1994 to March 2001) to determine areas of high and low turbot CPUE and whether these patterns could be used to establish sensible catch regions (Figure 7; Figure 8). This was done using both the landed and discard turbot catches as it was expected that reasonable patterns should be consistent with both forms of catch.

Table 3. Proposed definitions for turbot catch regions based on the existing DFO "major", "minor" and "locality" names (Rutherford 1999). The names of the major and minor areas are taken from Appendix E, Figure 6 in Rutherford (1999).

| Proposed turbot catch region name | Major area | Minor area | Locality |
| :--- | :---: | :---: | :---: |
| West \& North Charlottes | part of 5D \& all of 5E | 2AW, IW, IE | see Table 4 |
| Upper Hecate Strait | part of 5D | part of 1E, 4, 5U, 2AE \& part of | see Table 4 |
|  |  | 5 L |  |
| Moresby Gully | 5C \& part of 5B \& 5E | 2BW, 2BE, 5L, 6 \& part of 8 | see Table 4 |
| Mitchell Gully | part of 5B | part of 8 | see Table 4 |
| Goose Island Gully | part of 5A \& 5B | parts of 8 \& 11 | see Table 4 |
| upper west coast Vancouver Island | part of 3D \& part of 5A | $27 \&$ part of 11 | see Table 4 |
| lower west coast Vancouver Island | all of 3C \& part of 3D | $21,23: 26$ | all |
| Georgia Strait undefined for turbot | 4B | all | all |

The regions of high catch rates separated by areas of low catch rates described in the previous paragraph can be approximated using the existing area and "locality" definitions which are presently available in the DFO catch and effort databases (Rutherford 1999). As these definitions have remained reasonably constant over the history of DFO catch reporting, they can be applied over the entire period. The proposed area definitions for turbot catch regions in terms of the DFO major and minor areas and the smaller localities are presented in Table 3 and Table 4.

Figure 7 and Figure 8 demonstrate that there are areas where turbot catch rates (both landed and discard) are low between regions of high catch rates. These form natural boundaries which can be used to define turbot catch regions. One such break occurs just off Brooks Peninsula on the west coast of Vancouver Island. Another break lies in a diagonal from Cape Scott to Triangle Island at the top end of Vancouver Island. The three "gullies" which characterise Queen Charlotte Sound are clearly defined in Figure 7, particularly the division between Mitchell and Moresby gullies. There is an area of high CPUE in the upper regions of Hecate Strait and another off Langara Island, with a clear break occurring in mid-Dixon Entrance (Figure 7). These patterns are clearer when the landed catches are used because the discard catch rates tend to be more ubiquitous and spread out. However, the general pattern described here is consistent for both forms of turbot catch (compare Figure 8 with Figure 7)

Table 4. Locality definitions (Rutherford 1999) used to establish the proposed turbot catch regions defined in Table 3. As the localities are not numbered uniquely, they must be identified in the context of each minor area. The names of the minor areas are taken from Appendix E, Figure 6 in Rutherford (1999). Locality numbers are those shown in Appendix E, Figures 7 to 10 in Rutherford (1999). Commas separate unique localities; a colon (' $\because$ ') indicates that all localities are taken inclusively in the range indicated. Locality ' 0 ' is "unknown".

| Minor Area <br> Name |  <br> North <br> Charlottes | Upper <br> Hecate <br> Strait | Moresby | Mitchell | Goose <br> Island | Upper <br> Vancouver <br> Island |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 E | 3,5 | $1,2,4$ |  |  |  |  |
| 2BW |  |  | $1: 5$ |  |  |  |
| 2 BE |  |  | $1: 10$ |  |  |  |
| 5 L |  | 2 | $1,3: 10$ |  |  |  |
| 6 |  |  | $1: 4$ |  |  |  |
| 8 |  |  | $6,11: 12$ | $3: 5,14: 15$ | $0: 2,7: 10,13$ |  |
| 9 |  |  |  |  | $1: 2$ |  |
| 11 |  |  |  |  | $0: 3,5,7,12$ | $4,6,10: 11$ |



Figure 7. Left panel: distribution map of turbot CPUE for retained catch only (calculated for $5 \mathrm{~km}^{2}$ grids using Eq. 3 and based on total catch and effort from January 1994 to March 2001). Right panel: same distribution map as in left panel but with DFO "locality" boundaries (see Rutherford 1999 for definitions and names) superimposed, along with proposed boundaries for turbot catch regions based on the locality boundaries and the combined CPUE patterns from the catch and discard data.


Figure 8. Left panel: distribution map of turbot CPUE for discarded catch only (calculated for $5 \mathrm{~km}^{2}$ grids using Eq. 3 and based on total discards and effort from February 1996 to March 2001). Right panel: same distribution map as in left panel but with DFO "locality" boundaries (see Rutherford 1999 for definitions and names) superimposed, along with proposed boundaries for turbot catch regions based on the locality boundaries and the combined CPUE patterns from the catch and discard data.

### 5.3 Distribution of annual landing and discard catches

The dataset used to estimate catch rates from the observer data can also be used to summarise the retained and discard catches based on the catch regions described in Section 5.2 (Table 5). This table indicates that the total coastwide catches of turbot have ranged between 7,000 and 8,500 metric tonnes between 1996/97 and 2000/01, except for 1997/98 when both the retained and discard catches dropped considerably. The coastwide split between retained and discard catches has averaged approximately $45 \%$ discarded, ranging between $41 \%$ and $51 \%$ with no apparent trend over time (Table 5). Catches and discards of turbot have also been summarised by standard fishing year (1 April-31 March) for the major DFO catch reporting regions (Table 6).

When the catches and discards are examined separately by catch region, the distribution between retained and discarded catch varies considerably between regions and between fishing years. The most striking differences are the high level of retained catch in the lower west coast of Vancouver Island (about $75 \%$ of the catch is retained) and the high rates of discard in upper west coast Vancouver Island, upper Hecate Strait and the Queen Charlotte Islands. Catches in Queen Charlotte Sound have been evenly split between retained and discarded catches. In Mitchell Gully there is a trend towards increasing levels of discard over time while there is a decreasing trend for discards in Moresby Gully. The other areas are reasonably consistent between the fishing years. Retained catches are highest in the lower west coast of Vancouver Island while discard catches are highest in upper Hecate Strait (Table 5).

Table 5. Total retained and discard catches $(\mathrm{t})$ and their distribution by standard fishing year (1 April 31 March) for the total B.C. Coast and for the seven catch regions defined in Table 3. Discard estimates prior to the 1996/97 fishing year should be regarded as incomplete as they are not based on observer estimates.

| Fishing Year | Catch Type |  |  | \% of Total |  | Catch Type |  |  | \% of Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Retained | Discard | Total | \% Ret | \% Disc | Retained | Discard | Total | \% Ret | \% Disc |
|  | Total B.C. |  |  |  |  | Lower WCVI |  |  |  |  |
| 1991/92 | 2,206 | 711 | 2,918 | 76\% | 24\% | 857 | 128 | 985 | 87\% | 13\% |
| 1992/93 | 3,592 | 709 | 4,297 | 84\% | 16\% | 1,548 | 249 | 1,796 | 86\% | 14\% |
| 1993/94 | 3,943 | 1,095 | 5,035 | 78\% | 22\% | 2,026 | 312 | 2,338 | 87\% | 13\% |
| 1994/95 | 3,933 | 773 | 4,705 | 84\% | 16\% | 2,148 | 133 | 2,281 | 94\% | 6\% |
| 1995/96 | 3,642 | 715 | 4,357 | 84\% | 16\% | 1,895 | 225 | 2,119 | 89\% | 11\% |
| 1996/97 | 5,001 | 3,476 | 8,477 | 59\% | 41\% | 3,115 | 803 | 3,919 | 79\% | 20\% |
| 1997/98 | 2,722 | 2,522 | 5,244 | 52\% | 48\% | 1,645 | 583 | 2,228 | 74\% | 26\% |
| 1998/99 | 3,938 | 3,217 | 7,156 | 55\% | 45\% | 2,188 | 523 | 2,711 | 81\% | 19\% |
| 1999/00 | 3,900 | 4,031 | 7,932 | 49\% | 51\% | 1,309 | 764 | 2,074 | 63\% | 37\% |
| 2000/01 | 4,726 | 3,440 | 8,168 | 58\% | 42\% | 2,218 | 578 | 2,796 | 79\% | 21\% |
|  | Upper WCVI |  |  |  |  | Goose Island Gully |  |  |  |  |
| 1991/92 | 95 | 49 | 143 | 66\% | 34\% | 341 | 155 | 497 | 69\% | 31\% |
| 1992/93 | 222 | 94 | 315 | 70\% | 30\% | 708 | 119 | 827 | 86\% | 14\% |
| 1993/94 | 200 | 42 | 241 | 83\% | 17\% | 487 | 136 | 622 | 78\% | 22\% |
| 1994/95 | 156 | 112 | 267 | 58\% | 42\% | 454 | 91 | 545 | 83\% | 17\% |
| 1995/96 | 121 | 37 | 158 | 77\% | 23\% | 850 | 132 | 982 | 87\% | 13\% |
| 1996/97 | 19 | 149 | 168 | 11\% | 89\% | 507 | 486 | 993 | 51\% | 49\% |
| 1997/98 | 51 | 87 | 138 | 37\% | 63\% | 599 | 671 | 1,270 | 47\% | 53\% |


| Fishing Year | Catch Type |  |  | \% of Total |  | Catch Type |  |  | \% of Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Retained Discard Total |  |  | \% Ret \% Disc |  | Retained Discard |  | Total | \% Ret | \% Disc |
| 1998/99 | 151 | 186 | 337 | 45\% | 55\% | 710 | 887 | 1,597 | 44\% | 56\% |
| 1999/00 | 33 | 134 | 167 | 20\% | 80\% | 681 | 932 | 1,613 | 42\% | 58\% |
| 2000/01 | 54 | 128 | 182 | 30\% | 70\% | 987 | 795 | 1,782 | 55\% | 45\% |
|  | Mitchell G | ully |  |  |  | Moresby | Gully |  |  |  |
| 1991/92 | 130 | 29 | 159 | 82\% | 18\% | 597 | 27 | 625 | 96\% | 4\% |
| 1992/93 | 430 | 26 | 455 | 95\% | 6\% | 478 | 16 | 494 | 97\% | 3\% |
| 1993/94 | 527 | 41 | 568 | 93\% | 7\% | 315 | 43 | 358 | 88\% | 12\% |
| 1994/95 | 483 | 20 | 503 | 96\% | 4\% | 316 | 28 | 343 | 92\% | 8\% |
| 1995/96 | 211 | 98 | 309 | 68\% | 32\% | 162 | 34 | 196 | 83\% | 17\% |
| 1996/97 | 304 | 219 | 522 | 58\% | 42\% | 685 | 575 | 1,260 | 54\% | 46\% |
| 1997/98 | 80 | 99 | 179 | 45\% | 55\% | 60 | 263 | 323 | 19\% | 81\% |
| 1998/99 | 158 | 130 | 288 | 55\% | 45\% | 387 | 425 | 813 | 48\% | 52\% |
| 1999/00 | 78 | 186 | 264 | 30\% | 70\% | 1,011 | 584 | 1,595 | 63\% | 37\% |
| 2000/01 | 30 | 187 | 218 | 14\% | 86\% | 898 | 199 | 1,098 | 82\% | 18\% |
|  | Upper He | ate Strait |  |  |  | North \& | West Char |  |  |  |
| 1991/92 | 148 | 222 | 370 | 40\% | 60\% | 38 | 101 | 139 | 27\% | 73\% |
| 1992/93 | 154 | 119 | 273 | 56\% | 44\% | 52 | 86 | 137 | 38\% | 63\% |
| 1993/94 | 335 | 293 | 627 | 53\% | 47\% | 53 | 228 | 281 | 19\% | 81\% |
| 1994/95 | 303 | 123 | 426 | 71\% | 29\% | 73 | 266 | 340 | 21\% | 78\% |
| 1995/96 | 373 | 110 | 484 | 77\% | 23\% | 30 | 79 | 109 | 28\% | 72\% |
| 1996/97 | 344 | 989 | 1,333 | 26\% | 74\% | 27 | 255 | 282 | 10\% | 90\% |
| 1997/98 | 277 | 672 | 949 | 29\% | 71\% | 10 | 147 | 157 | 6\% | 94\% |
| 1998/99 | 327 | 914 | 1,241 | 26\% | 74\% | 17 | 152 | 169 | 10\% | 90\% |
| 1999/00 | 768 | 1,283 | 2,051 | 37\% | 63\% | 20 | 148 | 168 | 12\% | 88\% |
| 2000/01 | 470 | 1,426 | 1,896 | 25\% | 75\% | 69 | 127 | 196 | 35\% | 65\% |

Table 6. Total retained and discard catches $(\mathrm{t})$ and their distribution by standard fishing year (1 April 31 March) for the major DFO catch regions (defined in Appendix E. Fig. 6 in Rutherford 1999). Discard estimates prior to the 1996/97 fishing year should be regarded as incomplete as they are not based on observer estimates.

| Fishing <br> Year | Catch Type |  |  | \% of Total |  | Catch Type |  |  | \% of Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Retained | Discard | Total | \% Ret | \% Disc | Retained | Discard | Total | \% Ret | \% Disc |
|  | Total B.C. |  |  |  |  | 3C |  |  |  |  |
| 1991/92 | 2,206 | 711 | 2,918 | 76\% | 24\% | 273 | 68 | 341 | 80\% | 20\% |
| 1992/93 | 3,592 | 709 | 4,297 | 84\% | 16\% | 557 | 150 | 707 | 79\% | 21\% |
| 1993/94 | 3,943 | 1,095 | 5,035 | 78\% | 22\% | 398 | 73 | 470 | 85\% | 16\% |
| 1994/95 | 3,933 | 773 | 4,705 | 84\% | 16\% | 814 | 55 | 869 | 94\% | 6\% |
| 1995/96 | 3,642 | 715 | 4,357 | 84\% | 16\% | 374 | 137 | 511 | 73\% | 27\% |
| 1996/97 | 5,001 | 3,476 | 8,477 | 59\% | 41\% | 644 | 435 | 1,079 | 60\% | 40\% |
| 1997/98 | 2,722 | 2,522 | 5,244 | 52\% | 48\% | 740 | 412 | 1,152 | 64\% | 36\% |
| 1998/99 | 3,938 | 3,217 | 7,156 | 55\% | 45\% | 433 | 379 | 812 | 53\% | 47\% |
| 1999/00 | 3,900 | 4,031 | 7,932 | 49\% | 51\% | 495 | 501 | 996 | 50\% | 50\% |
| 2000/01 | 4,726 | 3,440 | 8,168 | 58\% | 42\% | 1,096 | 438 | 1,534 | 71\% | 29\% |
|  | 3D |  |  |  |  | 5A |  |  |  |  |
| 1991/92 | 646 | 83 | 728 | 89\% | 11\% | 320 | 118 | 438 | 73\% | 27\% |
| 1992/93 | 1,040 | 125 | 1,165 | 89\% | 11\% | 761 | 127 | 888 | 86\% | 14\% |
| 1993/94 | 1,806 | 261 | 2,067 | 87\% | 13\% | 401 | 120 | 521 | 77\% | 23\% |
| 1994/95 | 1,413 | 141 | 1,555 | 91\% | 9\% | 305 | 120 | 425 | 72\% | 28\% |


| Fishing Year | Catch Type |  |  | \% of Total |  | Catch Type |  |  | \% of Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Retained | Discard$116$ | $\begin{aligned} & \text { Total } \\ & 1,706 \end{aligned}$ | $\begin{array}{rr} \hline \text { \% Ret } & \text { \% Disc } \\ 93 \% & 7 \% \end{array}$ |  | $\begin{array}{r}\text { Retained } \\ \hline 499\end{array}$ | $\begin{array}{r} \text { Discard } \\ 102 \end{array}$ | $\begin{array}{r} \text { Total } \\ \hline 601 \end{array}$ | $\begin{array}{r} \text { \% Ret } \\ \hline 83 \% \end{array}$ | $\begin{array}{r} \text { \% Disc } \\ \hline 17 \% \end{array}$ |
| 1995/96 | 1,590 |  |  |  |  |  |  |  |  |  |
| 1996/97 | 2,487 | 426 | 2,913 | 85\% | 15\% | 369 | 330 | 698 | 53\% | 47\% |
| 1997/98 | 936 | 222 | 1,158 | 81\% | 19\% | 435 | 352 | 787 | 55\% | 45\% |
| 1998/99 | 1,814 | 202 | 2,016 | 90\% | 10\% | 444 | 662 | 1,106 | 40\% | 60\% |
| 1999/00 | 830 | 342 | 1,172 | 71\% | 29\% | 464 | 505 | 969 | 48\% | 52\% |
| 2000/01 | 1,158 | 216 | 1,374 | 84\% | 16\% | 378 | 393 | 771 | 49\% | 51\% |
|  | 5B |  |  |  |  | 5C |  |  |  |  |
| 1991/92 | 759 | 99 | 857 | 89\% | 12\% | 23 | 22 | 45 | 51\% | 49\% |
| 1992/93 | 949 | 87 | 1,036 | 92\% | 8\% | 78 | 15 | 93 | 84\% | 16\% |
| 1993/94 | 933 | 82 | 1,015 | 92\% | 8\% | 16 | 39 | 55 | 29\% | 71\% |
| 1994/95 | 984 | 61 | 1,045 | 94\% | 6\% | 42 | 7 | 49 | 86\% | 14\% |
| 1995/96 | 749 | 147 | 896 | 84\% | 16\% | 27 | 32 | 59 | 46\% | 54\% |
| 1996/97 | 1,075 | 816 | 1,891 | 57\% | 43\% | 55 | 234 | 289 | 19\% | 81\% |
| 1997/98 | 294 | 599 | 893 | 33\% | 67\% | 30 | 133 | 164 | 18\% | 81\% |
| 1998/99 | 891 | 767 | 1,659 | 54\% | 46\% | 14 | 150 | 164 | 9\% | 91\% |
| 1999/00 | 1,304 | 978 | 2,282 | 57\% | 43\% | 20 | 293 | 313 | 6\% | 94\% |
| 2000/01 | 1,549 | 740 | 2,289 | 68\% | 32\% | 10 | 214 | 224 | 4\% | 96\% |
|  | 5D |  |  |  |  | 5E |  |  |  |  |
| 1991/92 | 148 | 222 | 370 | 40\% | 60\% | 38 | 101 | 139 | 27\% | 73\% |
| 1992/93 | 154 | 118 | 272 | 57\% | 43\% | 52 | 86 | 137 | 38\% | 63\% |
| 1993/94 | 335 | 295 | 631 | 53\% | 47\% | 52 | 224 | 276 | 19\% | 81\% |
| 1994/95 | 301 | 123 | 424 | 71\% | 29\% | 73 | 266 | 339 | 22\% | 78\% |
| 1995/96 | 373 | 113 | 487 | 77\% | 23\% | 30 | 69 | 98 | 31\% | 70\% |
| 1996/97 | 344 | 1,001 | 1,345 | 26\% | 74\% | 27 | 235 | 262 | 10\% | 90\% |
| 1997/98 | 278 | 696 | 974 | 29\% | 71\% | 9 | 109 | 118 | 8\% | 92\% |
| 1998/99 | 332 | 935 | 1,267 | 26\% | 74\% | 12 | 124 | 135 | 9\% | 92\% |
| 1999/00 | 769 | 1,283 | 2,052 | 37\% | 63\% | 18 | 130 | 148 | 12\% | 88\% |
| 2000/01 | 468 | 1,320 | 1,788 | 26\% | 74\% | 68 | 119 | 187 | 36\% | 64\% |

### 5.4 Selection of GLM models

After some preliminary exploratory analysis, four models were investigated for trends in catch rates (Table 7). These model choices were dictated by the amount of data available after the grooming process described in Appendix 1. In particular, it was decided to combine the analysis of the retained and discard catch of turbot for Queen Charlotte Sound and for Hecate Strait because there were insufficient data available to do the retained catch analysis separately for these areas. The discard catch analysis was conducted for the same combination of areas for comparability with the retained catch analysis.

Table 7. Models investigated using general linear modelling for turbot and the time periods included in the modelling.

| Area | Retained catch | Discarded catch |
| :--- | :--- | :--- |
| West coast Vancouver Island | April 1994 - March 2001 | April 1996 - March 2001 |
| Queen Charlotte Sound ("gullies") | Combined QCS \& HS analysis: | Combined QCS \& HS analysis: <br> April 1996 - March 2001 |
| Hecate Strait | April 1994 - March 2001 | Apr\| |

The analysis on retained catch was limited to the period 1 April, 1994 to 31 March, 2001 as records of the latitude and longitude of each tow were not maintained until early in 1994. These data were crucial in this analysis as the spatial explanatory variable was usually the most important variable offered in every model investigated. The analysis on discarded catch was limited to the period 1 April, 1996 to 31 March, 2001 as reliable estimates of discarded catch were not available until observer coverage was instituted on nearly $100 \%$ of the deepwater trawl fleet.

We decided to limit the analysis to those tows which recorded a catch of turbot, either retained or discarded, depending on the model being investigated. Records without retained turbot catch were dropped because it is known that the decision to land turbot is based on economic factors, with the marginal value of this species dependent on the holding capacity of the vessel and the demand for the species by the processor. Therefore, a zero catch of retained turbot is rarely an indication of a lack of abundance and should not be treated in this way. It is more difficult to maintain a similar argument for discarded turbot catch, but it is likely that vessels which are not fishing for turbot will not always be fishing in a location or depth where turbot are likely to be caught. Distinguishing between records that should have caught turbot and did not as compared to records which were not in areas where turbot would be expected to be part of the bycatch would be a very difficult task. As well, it is possible that small amounts of discarded turbot will not always be enumerated by the observer. Future analyses may attempt this type of analysis.

### 5.5 Variables for GLM models

A large number of preliminary analyses were investigated but which will not be reported. These analyses showed that the most important variables offered to the model usually were variables which defined the area of catch. As indicated in Table 8, there were a number of such variables. When all available variables were offered to the model, the procedure outlined in Section 5.1 selected an area variable first in three of the four analyses and often selected more than one area variable before the model selection procedure ceased. When more than one area variable was selected into the model, the model fit was better (as defined in terms of the deviance explained); however, the plots of the area coefficients and their standard errors showed that many of the coefficients were poorly estimated (with very large error bars) and the correct interpretation of multiple area predictor variables is not clear. Therefore, for all the models reported in this document, only the area variable with the best explanatory power (as determined in a preliminary fit) was presented during the final fitting procedure, thus eliminating all other area variables. Furthermore, the model estimates of the abundance coefficients (year coefficients) were not very sensitive to which area variable was used or how many were included in the model.

Table 8. Area categorical variables used in the turbot general linear modelling.

| Variable | Description |
| :--- | :--- |
| 5 km | constant $5 \mathrm{~km}^{2}$ grid blocks |
| 10 km | constant $10 \mathrm{~km}^{2}$ grid blocks |
| 25 km | constant $25 \mathrm{~km}^{2}$ grid blocks |
| 40 km | constant $40 \mathrm{~km}^{2}$ grid blocks |
| lat | $0.1^{\circ}$ latitude strips (WCVI models only) |
| minor | DFO minor statistical areas |

All models forced the abundance index, defined as a standard fishing year (1 April to 31 March), as the first predictor variable. This was done because in most cases the year variable was not sufficiently important in terms of explanatory power to be introduced into the model and there is little point in pursuing these analyses for stock assessment purposes without creating an abundance index.

Each of the analyses presented in this paper has been supported by a presentation of auxiliary information on the distribution of the underlying data by standard fishing year and the distribution of $\log (\mathrm{CPUE})$ for the selected predictor variables (Appendix 3).

### 5.6 Main effects model: analysis of retained catch

### 5.6.1 West coast Vancouver Island

The analysis of retained catch for the west coast of Vancouver Island introduced four explanatory variables in addition to fishing year and explained approximately 0.36 of the residual deviance (Table 9). The plot of the abundance indices indicated that the 1996/97 to 1998/99 fishing years had the highest CPUE in the series, with considerably lower CPUE in 1994/95 and 1995/96 and an apparent small drop in the two most recent years (Figure 9). The two unstandardised sets of abundance indices follow the same broad trend but appear to be more variable than the standardised indices. Examination of the supporting data (Section 12.1) shows that there were fewer tows available in the first two years (Appendix Table 3) but that the distributions of data in those years for the selected predictor variables were similar to those seen in subsequent years (Appendix Figure 2).

Table 9. Regression results for GLM model applied to west coast Vancouver Island retained catch of turbot. Total deviance explained was 0.36 and the first four variables were accepted in the model (indicated by *). Fishing year was forced as the first variable in model to provide an index of relative abundance. Table values are the proportion of the total residual deviance explained by each predictor variable at the specified iteration.

|  | Iteration |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ |
| Fishing year* | 0.0375 |  |  |  |  |
| $\mathbf{2 5 ~ k m}^{2}$ area grid* | 0.1962 | 0.2209 |  |  |  |
| Vessel* $^{*}$ | 0.1251 | 0.1500 | 0.2929 | 0.3610 | 0.3674 |
| Month $^{*}$ | 0.1408 | 0.1739 | 0.2831 | 0.3000 | 0.0681 |
| Depth | 0.0470 | 0.0778 | 0.2290 | 0.065 |  |
| Increase in proportion | 0.0000 | 0.1835 | 0.0719 | 0.061 |  |
| deviance explained |  |  |  |  |  |

The plots of the coefficients for the $25 \mathrm{~km}^{2}$ area variable show consistently high catch rates for areas 3 and 10 (Figure 10), which probably correspond to the areas of high retained catch rates shown in Figure 7. The monthly coefficients are uniformly high from May to October (Figure 10). Several vessels show high catch rates relative to the other vessels, indicating that these vessels may specialise in targeting on this species (Figure 10).


Figure 9. Plot of fishing year abundance estimates for WCVI retained catch regression model (only $25 \mathrm{~km}^{2}$ area predictor variable offered) with each annual index normalised relative to the geometric mean of the indices in each set. Plotted lines: standardised index from the GLM (Eq. 3); unstandardised geometric mean of CPUE (Eq. 8); annual index of the arithmetic mean CPUE (Eq. 6).


Figure 10. Canonical (Eq. 3) categorical variable coefficients for the WCVI retained catch regression model (only $25 \mathrm{~km}^{2}$ area predictor variable offered) plotted as relative CPUE.


Figure 11. Standardised residuals for the WCVI retained catch regression model (only $25 \mathrm{~km}^{2}$ area predictor variable offered).

The residual diagnostics for this model show deviations from the normal assumption at both ends of the distribution. The model shows a tendency to underestimate the observations at both the lower and upper tails and a slight bulge of overestimation in the centre of the distribution (Figure 11).

### 5.6.2 Combined Queen Charlotte Sound/Hecate Strait

The analysis of retained catch for the combined dataset for Queen Charlotte Sound and Hecate introduced only three explanatory variables in addition to fishing year and explained approximately 0.40 of the residual deviance (Table 10). Fishing year as an explanatory variable was stronger in this model than for any of the other models investigated, with fishing year entering into the model in third position when this model was fitted without forcing. The $10 \mathrm{~km}^{2}$ grid area predictor variable was the best explanatory variable, providing 0.23 of the residual deviance in the first iteration.

The trend in the abundance indices is similar to that seen in the WCVI retained catch indices, with the 1996/97 fishing year showing the highest CPUE in the series and with considerably lower CPUE in 1994/95 and 1995/96. This series shows a tendency to decline towards the most recent years (Figure 9), but the current levels are higher than those experienced in 1994/95 or 1995/96 (Figure 12). The two unstandardised sets of abundance indices show a trend of increasing catch rates, indicating that the standardisation procedure has possibly
captured some changes in fishing practices which have led to higher unstandardised catch rates but which are attributable to factors in the other explanatory variables.

Table 10. Regression results for GLM model applied to combined QCS\&HS retained catch of turbot. Total deviance explained was 0.40 and the first four variables were accepted in the model (indicated by *). Fishing year was forced as the first variable in model to provide an index of relative abundance. Table values are the proportion of the total residual deviance explained by each predictor variable at the specified iteration.

|  | Iteration |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ |  |
| Fishing year* | 0.0869 |  |  |  |  |
| ${\mathbf{1 0} \text { km }^{2 *}}^{\text {Vessel* }}$ | 0.2291 | 0.2887 |  |  |  |
| Month* $^{*}$ | 0.2043 | 0.2503 | 0.3812 | 0.3974 |  |
| Depth | 0.0267 | 0.1089 | 0.3022 | 0.3888 | 0.4025 |
| Increase in proportion | 0.0569 | 0.1376 | 0.2991 | 0.0051 |  |
| deviance explained | 0.0000 | 0.2018 | 0.0924 | 0.0162 |  |

The plot of the coefficients for the $10 \mathrm{~km}^{2}$ area variable shows high catch rates for areas 18 to 20 and 22 to 24 (Figure 13). These probably correspond to the areas of high catches rates shown in Figure 7. The monthly coefficients are similar to those estimated for the WCVI retained catch (compare Figure 13 with Figure 10), being uniformly high from May to October. As seen in the WCVI retained catch analysis, a few vessels (here about 4 to 6 vessels) show relatively high catch rates, possibly being indicative of specialisation in this species by these vessels.


Figure 12. Plot of fishing year abundance estimates for the combined QCS and Hecate St. retained catch regression model (only $10 \mathrm{~km}^{2}$ area predictor variable) with each annual index normalised relative to the geometric mean of the indices in each set. Plotted lines: standardised index from the GLM (Eq. 3); unstandardised geometric mean of CPUE (Eq. 8); annual index of the arithmetic mean CPUE (Eq. 6).


Figure 13. Canonical (Eq. 3) categorical variable coefficients for the combined QCS and Hecate St. retained catch regression model (only $10 \mathrm{~km}^{2}$ area predictor variable offered) plotted as relative CPUE.


Figure 14. Standardised residuals for the combined QCS and Hecate St. retained catch regression model (only $10 \mathrm{~km}^{2}$ area predictor variable offered).

The residual diagnostics for this model seem better than for the WCVI retained catch model, with deviations from the normal assumption at both ends of the distribution being smaller and without the bulge in the centre of the distribution (Figure 14).

### 5.7 Main effects model: analysis of discards

### 5.7.1 West coast Vancouver Island

The analysis of discarded catch for the west coast of Vancouver Island introduced five explanatory variables in addition to fishing year and explained approximately 0.17 of the residual deviance (Table 11). The predictor variable for area (in this case the $0.1^{\circ}$ latitude band variable) was much less important than for the equivalent retained catch analysis, as evidenced by its late entry into the analysis (it entered in the last significant position). The plot of the abundance indices shows almost no contrast in abundance over the five years of data (Figure 15) and only a small amount of deviance is explained by this predictor variable ( 0.0019 ; Table 11). The two unstandardised sets of abundance indices follow the same trend as the standardised index and the underlying data show very little contrast (Appendix Figure 6).

Table 11. Regression results for GLM model applied to west coast Vancouver Island discarded catch of turbot. Total deviance explained was 0.17 and the first five variables were accepted in the model (indicated by *). Fishing year was forced as the first variable in model to provide an index of relative abundance. Table values are the proportion of the total residual deviance explained by each predictor variable at the specified iteration.

|  | Iteration |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ |
| Fishing year* | 0.0019 |  |  |  |  |
| Vessel* $^{*}$ | 0.0501 | 0.0527 |  |  |  |
| Month* $^{\text {Depth* }}$ | 0.0394 | 0.0401 | 0.0991 |  |  |
| Den $^{\circ}$ Latitude band* | 0.0362 | 0.0385 | 0.0856 | 0.1534 |  |
| Increase in proportion | 0.0342 | 0.0358 | 0.0841 | 0.1371 | 0.1725 |
| deviance explained | 0.0000 | 0.0508 | 0.0463 | 0.0543 | 0.0190 |

The plot of the coefficients for month shows the same trends as do the two retained catch analysis (Figure 16). The plot of the coefficients for depth band shows above average catch rates from the $250-300 \mathrm{~m}$ band to the $450-500 \mathrm{~m}$ band (Figure 16). The plot of the $0.1^{\circ}$ latitude band area variable coefficients shows high catch rates for bands 13 to 15 (Figure 16) which correspond to latitude bands $49.4^{\circ}-49.5^{\circ}, 49.6^{\circ}-49.7^{\circ}$ and $50.3^{\circ}-50.4^{\circ}$ which correspond to the regions of high discard rates shown near mid-Vancouver Island (Figure 8). While some vessels have relatively higher catch rates than others, the contrast between the highest and lowest vessels is less for the discard fishery than for the equivalent retained turbot catch fishery (compare Figure 10 and Figure 16).


Figure 15. Plot of fishing year abundance estimates for WCVI discarded catch regression model (only $0.1^{\circ}$ latitude band predictor variable offered) with each annual index normalised relative to the geometric mean of the indices in each set. Plotted lines: standardised index from the GLM (Eq. 3); unstandardised geometric mean of CPUE (Eq. 8); annual index of the arithmetic mean CPUE (Eq. 6).


Figure 16. Canonical (Eq. 3) categorical variable coefficients for the WCVI discarded catch regression model (only $0.1^{\circ}$ latitude predictor variable offered) plotted as relative CPUE.


Figure 17. Standardised residuals for the WCVI discarded catch regression model (only $0.1^{\circ}$ latitude area predictor variable offered).

The residual diagnostics for this model are very good, with almost no deviations from the normal assumption at either end of the distribution (Figure 17).

### 5.7.2 Combined Queen Charlotte Sound/Hecate Strait

The analysis of discarded catch for the combined dataset for Queen Charlotte Sound and Hecate introduced four explanatory variables in addition to fishing year and explained approximately 0.24 of the residual deviance (Table 12). As with the WCVI discard catch analysis, fishing year was a very poor predictor of $\ln (C P U E)$, accounting for only 0.0034 of the residual deviation in the first model iteration. Unlike the WCVI discard analysis, the area predictor variable (the $40 \mathrm{~km}^{2}$ grid variable) was the strongest explanatory variable, entering the model first and accounting for 0.10 of the residual deviance.

The plot of the abundance indices is similar to those for the WCVI discarded catch, with almost no contrast in the indices over the five year period (Figure 18). The unstandardised and arithmetic indices also show little contrast. The plots of the coefficients for month and 50 m depth band are similar to the same plots for the other discard analysis (compare Figure 19 with Figure 16), with the highest catch rates occurring from May to October and in the 250-300 m to the $450-500 \mathrm{~m}$ depth bands. This pattern is also evident in the plots of the auxiliary supporting
data provided in Appendix Figure 8. The $40 \mathrm{~km}^{2}$ grid area variable shows very high catch rates for areas 15 to 17 (Figure 19).

Table 12. Regression results for GLM model applied to combined QCS\&HS discarded catch of turbot. Total deviance explained was 0.24 and the first five variables were accepted in the model (indicated by *). Fishing year was forced as the first variable in model to provide an index of relative abundance and depth was truncated at $1 / 99 \%$ of the distribution. Table values are the proportion of the total residual deviance explained by each predictor variable at the specified iteration.

|  | Iteration |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ |
| Fishing year* | 0.0034 |  |  |  |  |
| 40 km $^{2}$ grid* $^{*}$ | 0.0968 | 0.0996 |  |  |  |
| Month* $^{\text {Vessel* }}$ | 0.0705 | 0.0719 | 0.1784 |  |  |
| Depth Band* | 0.0417 | 0.0434 | 0.1298 | 0.2139 |  |
| Increase in proportion | 0.0291 | 0.0324 | 0.1204 | 0.2093 | 0.2422 |
| deviance explained | 0.000 | 0.0962 | 0.0789 | 0.0354 | 0.0283 |



Figure 18. Plot of fishing year abundance estimates for the combined QCS and Hecate St. discard catch regression model (only $40 \mathrm{~km}^{2}$ area predictor variable offered and depth truncated at $1 / 99 \%$ ) with each annual index normalised relative to the geometric mean of the indices in each set. Plotted lines: standardised index from the GLM (Eq. 3); unstandardised geometric mean of CPUE (Eq. 8); annual index of the arithmetic mean CPUE (Eq. 6).

As for the west coast of Vancouver Island discard analysis, the residual diagnostics for this model are very good, with almost no deviations from the normal assumption at either end of the distribution (Figure 20).


Figure 19. Canonical (Eq. 3) categorical variable coefficients for the combined QCS and Hecate St. discard catch regression model (only $40 \mathrm{~km}^{2}$ area predictor variable offered and depth truncated at 1/99\%) plotted as relative CPUE.


Figure 20. Standardised residuals for the combined QCS and Hecate St. discard catch regression model (only $40 \mathrm{~km}^{2}$ area predictor variable offered and depth truncated at $1 / 99 \%$ ).

### 5.8 Main effects model: summary of results

All four of the main effects models introduced similar predictor variables into the final regression (Table 13). Both retained catch models accepted the same four variables in the same order. Both retained catch models also did not include depth as an explanatory variable. The only difference between these models was that the WCVI model used a larger area grid explanator $\left(25 \mathrm{~km}^{2}\right)$ than the QCS\&HS model $\left(10 \mathrm{~km}^{2}\right)$. Both models also accounted for approximately the same proportion of the total deviance ( 0.36 and 0.40 ). The discard catch models accepted the same five explanator variables but in a different order, with the area explanator variable being considerably more important for the QCS\&HS model than for the WCVI model (Table 13). Both discard catch models used the depth categorical variable, but this variable entered the model late in the stepwise selection process. The discard catch models accounted for much less of the total deviance ( 0.17 and 0.24 ) than did the retained catch models (Table 13).

Table 13. Summary of the four main effects models performed on two catch types over two areas of the B.C. coast: order of acceptance of key predictor variables and the proportion of the total deviance explained. "NA" indicates that the variable was not accepted into the model.

| Variable | WCVI Retained | QCS\&HS Retained | WCVI Discard | QCS\&HS Discard |
| :--- | ---: | ---: | ---: | ---: |
| Fishing year | 1 | 1 | 1 | 1 |
| Area variable | 2 | 2 | 5 | 2 |
| Vessel | 3 | 3 | 2 | 4 |
| Month | 4 | 4 | 3 | 3 |
| Depth | NA | NA | 4 | 5 |
| Proportion of deviance <br> explained | 0.36 | 0.40 | 0.17 | 0.24 |



Figure 21. Overlay of the canonical (Eq. 3) coefficients for fishing year for the retained and discard catch models from each of the two area modelled: [left panel] west coast Vancouver Island; [right panel] combined Queen Charlotte Sound/Hecate Strait. The coefficients and the error bars from the two retained catch models have been standardised relative to the geometric mean for the 1996/97 to 2000/01 fishing years for comparability to the discard catch models.

The abundance trends estimated by the four main effects model are surprisingly similar for the overlapping fishing years (Figure 21). The two WCVI models show some divergence in abundance trends in the two most recent fishing years while the QCS\&HS models estimate the same relative coefficients for all five years of overlap. It can also be seen from Figure 21 that the coefficient error bars estimated by all four models substantially overlap, indicating that the data provide consistent abundance trends for each of the areas when using either the retained or discarded turbot catch data.

### 5.9 Interaction effects model

An analysis of the interaction effects for all the selected model predictor parameters (with the exception of vessel interactions) was performed for each of the four models described in the previous section. The omission of vessel interactions was forced due to limitations in the available software which precluded performing the full analysis on the large number of variables generated when using all selected first order predictor variables and the consequent second order interaction combinations based on the selected predictor variables. The interaction analysis was done by creating categorical variables which described the interactions between each pair of selected predictor variables, providing a total of three to six combinations of variable interactions, depending on the analysis. Each paired interaction term was then offered to the model after the main effects model predictor variables were used. The selection criteria provided in Section 5.1 were used as the basis for deciding which interaction pairs would be considered "significant" as traditional tests of model significance are almost always true, given the relatively large number of observations in these models.

Each of the four models selected three or four of the interaction terms using the Section 5.1 criteria, with all models showing considerable improvement in fit relative to the main effects model. These models probably have far too many variables, resulting in unusual coefficient estimates in some instances and large standard errors for the more poorly determined coefficients. However, these analyses indicate that there are complex interactions in these models which affect the interpretation of the trends that are seen in the simpler main effects models.

### 5.9.1 Interaction effects model: west coast Vancouver Island retained catch

The first interaction term entering this model is an areaXmonth interaction, which is not surprising given the strong seasonal nature of most fisheries and the necessity of fishing in exposed areas only during the more clement seasons (Table 14). Interaction terms involving fishing year follow, first with month and then with the $25 \mathrm{~km}^{2}$ area grid. There is a good improvement in the amount of deviance explained by this model, increasing from 0.36 to 0.52 (Table 14). There are changes to some of the categorical coefficients compared to the estimates from the main effects model, with substantial changes in the coefficient patterns for fishing year and month while the coefficient pattern for vessel is almost completely unchanged (compare Figure 22 with Figure 10). Model fits do not appear to be improved in the interaction model compared to the main effects model (Figure 23 and Figure 11). Pairwise plots of the fishing year coefficients for every combination of the $25 \mathrm{~km}^{2}$ grids show major differences in the fishing year coefficient patterns between some of the grids (eg. grids 3 and 8 or grids 7 and 14 or grids 11 and

14; Figure 24). There are also many cases of similar trends between areas (eg. 2 and 4 or 4 and 8; Figure 24) and instances of similar trends in all years except for the first or last year (eg. 2 and 8 or 13 and 17; Figure 24). While these differences in fishing year coefficient trends are apparent to the eye, the model selection process selected this interaction term last, with an improvement in deviance of about 0.03 (Table 14).

Table 14. Regression results for interaction effects GLM model applied to west coast Vancouver Island retained catch of turbot. Total deviance explained was 0.52 , an improvement of 0.15 over the main effects model. All three of the interaction terms offered to the model were accepted (indicated by a *). These three interaction terms represent all the available interactions from the main effects model after excluding the vessel interactions. Table values are the proportion of the total residual deviance explained by each interaction variable at the specified iteration after fitting all the main effects terms indicated in Table 9.

|  | Iteration |  |  |
| :--- | ---: | ---: | ---: |
| Interaction Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ |
| $\mathbf{2 5 ~ k m}^{\mathbf{2}}$ Xmonth* | 0.4586 |  |  |
| Fishing yearXMonth* | 0.4066 | 0.4930 |  |
| Fishing yearX25 km |  |  |  |
| Increase in proportion <br> deviance explained | 0.4154 | 0.4927 | 0.5218 |



Figure 22. Main effects coefficients and 2*standard error confidence bounds for the WCVI retained catch interaction effects model (Table 14)..


Figure 23. Standardised residuals for the WCVI retained catch interaction effects regression model (only $25 \mathrm{~km}^{2}$ area predictor variable offered).


Figure 24. Fishing year coefficients for every combination of paired $25 \mathrm{~km}^{2}$ grid for the WCVI retained catch interaction model described in Table 14.

### 5.9.2 Interaction effects model: Queen Charlotte Sound \& Hecate Strait retained catch

The order of acceptance of the interaction terms for this model is the same as for the west coast Vancouver Island retained catch model: areaXmonth, followed by fishing yearXmonth and fishing yearXarea (Table 15). This is a striking similarity given the different nature of the fisheries, possibly being evidence for stock or biological affinity. As with the other retained catch model, there is a good improvement in the amount of deviance explained by this model, increasing from 0.40 to 0.58 (Table 15). The categorical coefficients are considerably changed compared to the estimates from the main effects model, particularly to the coefficient patterns for fishing year, month and $10 \mathrm{~km}^{2}$ area grid (Figure 25 and Figure 13) As with the other retained catch model, the coefficient pattern for vessel is almost completely unchanged (compare Figure 25 with Figure 13) Model fits are similar between the interaction model and the main effects model (Figure 26and Figure 14). There are a great many pairwise plots of the fishing year coefficients for every combination of the $10 \mathrm{~km}^{2}$ grids ( 276 for 24 categories) which show similar major differences between some of the grids in the fishing year coefficient patterns as were seen for the west coast Vancouver Island retained catch interaction model (eg. grids 7 and 19 or grids 2 and 22;Figure 27). There are some cases of similar trends between areas (eg. 6 and 7 or 2 and 7) and many instances of similar trends in all years except for the first or last year (eg. 2 and 5 or 8 and 24). As with the other retained catch model, the model selection process selected this interaction term last, with an improvement in deviance of around 0.03 (Table 15).

Table 15. Regression results for interaction effects GLM model applied to the combined Queen Charlotte Sound/Hecate St. retained catch of turbot. Total deviance explained was 0.58 , an improvement of 0.18 over the main effects model. All three of the interaction terms offered to the model were accepted (indicated by a *). These three interaction terms represent all the available interactions from the main effects model after excluding the vessel interactions. Table values are the proportion of the total residual deviance explained by each interaction variable at the specified iteration after fitting all the main effects terms indicated in Table 10.

|  | Iteration |  |  |
| :--- | ---: | ---: | ---: |
| Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ |
| $\mathbf{1 0 ~ k m}^{\mathbf{2}} \mathbf{X m o n t h} \boldsymbol{*}$ | 0.5142 |  |  |
| Fishing yearXMonth* | 0.4408 | 0.5440 |  |
| Fishing yearX10 $\mathbf{~ m}^{2} *$ | 0.4503 | 0.5503 | 0.5788 |
| Increase in proportion <br> deviance explained | 0.1168 | 0.0298 | 0.0348 |



Figure 25. Main effects coefficients and 2*standard error confidence bounds for the combined QCS/HS retained catch interaction effects model (Table 15)..


Figure 26. Standardised residuals for the combined QCS/HS retained catch interaction effects regression model (only $10 \mathrm{~km}^{2}$ area predictor area variable offered).



Figure 27. Fishing year coefficients for every combination of paired $10 \mathrm{~km}^{2}$ area grid for the combined QCS\&HS retained catch interaction model described in Table 15.

### 5.9.3 Interaction effects model: west coast Vancouver Island discard catch

There are three additional interaction terms included in the two discard catch models compared to the retained catch models because the two main effects models based the discard catch included a depth effect while the retained catch models did not. However, only three of these six interaction terms were included in the model based on the criteria provided in Section 5.1 (Table 16). There were no fishing yearXarea interactions included and the only interaction term including fishing year was with month (Table 16). The improvement in the amount of deviance explained by this model is proportionately large, increasing by 0.14 from 0.17 to 0.31 (Table 16). The categorical coefficients for the main effects appear to be changed less than for the retained catch models (eg. fishing year, vessel and latitude band; compare Figure 28 and Figure 16), possibly because of the larger amount of data available compared to the retained catch models. However, there are wide confidence bounds for some of the coefficient estimates and the coefficient patterns for the month and depth effects have changed. As for the main effects model, the residual diagnostics are very good and indicate a good fit to the model assumptions concerning data distribution (Figure 29). Pairwise plots of the fishing year coefficients for every combination of month show similar fishing year coefficient patterns for many of the month pairs (eg. months 2 and 3 or months 8 and 10; Figure 30). Some months show different patterns (eg. 5 and 9 or 3 and 7; Figure 30) and many months diverge in the first or last year (eg. 3 and 12 or 4 and ; Figure 30). The model selection process selected this interaction term second, with an improvement in deviance of 0.02 (Table 16).

Table 16. Regression results for interaction effects GLM model applied to the west coast Vancouver Island discard catch of turbot. Total deviance explained was 0.31 , an improvement of 0.14 over the main effects model. Only three of the six interaction terms offered to the model were accepted (indicated by a *). The six interaction terms represent all the available interactions from the main effects model after excluding vessel interactions. Table values are the proportion of the total residual deviance explained by each interaction variable at the specified iteration after fitting all the main effects terms indicated in Table 11.

|  | Iteration |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ |
| MonthXDepth band* | 0.2533 |  |  |  |
| Fishing yearXMonth* | 0.1915 | 0.2726 |  |  |
| Latitude bandXMonth* | 0.2298 | 0.2857 | 0.3062 |  |
| Fishing yearXDepth band | 0.1861 | 0.2641 | 0.2802 | 0.3137 |
| Fishing yearXLatitude band | 0.1893 | 0.2683 | 0.284 | 0.3169 |
| Latitude bandXDepth band | 0.2031 | 0.2792 | 0.2981 | 0.3281 |
| Increase in proportion <br> deviance explained | 0.0808 | 0.0194 | 0.0336 | 0.0075 |



Figure 28. Main effects coefficients and 2*standard error confidence bounds for the west coast Vancouver Island discard catch interaction effects model (Table 16)..



Figure 29. Standardised residuals for the west coast Vancouver Island retained catch interaction effects regression model (only $0.1^{\circ}$ latitude band predictor area variable offered).
















Month 8 \& Mont 11

1

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\text { Mont } 9 \text { M Mont 12 } 12
$$


1

Fishing year

Figure 30. Fishing year coefficients for every combination of paired month (1 = April) for the WCVI discard catch interaction model described in Table 16.

### 5.9.4 Interaction effects model: Queen Charlotte Sound \& Hecate Strait discard catch

Four of the six offered interaction terms were included in the QCS/HS discard model based on the criteria provided in Section 5.1 (Table 17). Fishing yearX40 $\mathrm{km}^{2}$ area grid interaction was selected third, with an improvement in deviance explained of less than 0.015 (Table 17). The improvement in the amount of deviance explained in this model is equivalent to that seen in the west coast Vancouver Island discard catch model, increasing by 0.13 from 0.24 to 0.37 (Table 17). The categorical coefficients for the main effects are quite different for the month effect compared to main effects model (Figure 31 and Figure 19) and the standard errors for the depth band and area grid coefficients are very poorly determined (Figure 31). The residual diagnostics are very good and indicate a good fit to the model assumptions concerning the data distribution (Figure 32). Pairwise plots of the fishing year coefficients for every combination of $40 \mathrm{~km}^{2}$ area grid show similar fishing year coefficient patterns for many of the area grid pairs (eg. grids 3 and 8 or grids 15 and 16; Figure 33). Only a few grid pairs show different patterns (eg. 9 and 16 or 4 and 11 ; Figure 33) and few grid pairs diverge in the first or last year.

Table 17. Regression results for interaction effects GLM model applied to the combined Queen Charlotte Sound/Hecate Strait discard catch of turbot. Total deviance explained was 0.37 , an improvement of 0.13 over the main effects model. Only four of the six interaction terms offered to the model were accepted (indicated by a*). The six interaction terms represent all the available interactions from the main effects model after excluding the vessel interactions. Table values are the proportion of the total residual deviance explained by each interaction variable at the specified iteration after having fitted all the main effects terms indicated in Table 12.

|  | Iteration |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ |
| 40 km $^{2}$ gridXMonth* | 0.3265 |  |  |  |  |
| 40 km $^{2}$ gridXDepth band* | 0.264 | 0.3473 |  |  |  |
| Fishing yearX40 km $^{2}$ grid $^{*}$ | 0.2624 | 0.3405 | 0.361 |  |  |
| MonthXDepth band* | 0.3076 | 0.3418 | 0.3595 | 0.3729 |  |
| Fishing yearXDepth band | 0.2505 | 0.3335 | 0.354 | 0.3668 | 0.3786 |
| Fishing yearXDepth band | 0.2509 | 0.3319 | 0.352 | 0.363 | 0.375 |
| Increase in proportion <br> deviance explained | 0.0843 | 0.0208 | 0.0137 | 0.0119 | 0.0057 |



Figure 31. Main effects coefficients and 2*standard error confidence bounds for the combined QCS/HS discard catch interaction effects model (Table 17). Error bars are not shown for the 'depth band' and 40 km 2 area grid coefficients as the estimated standard errors were too large to calculate in normal space.


Figure 32. Standardised residuals for the combined QCS/HS discard catch interaction effects regression model (only $40 \mathrm{~km}^{2}$ area grid predictor variable offered).



Figure 33. Fishing year coefficients for every combination of paired $40 \mathrm{~km}^{2}$ area grid for the combined QCS\&HS discard catch interaction model described in Table 17.

### 5.9.5 Summary: GLM interaction analysis

The apparent importance of interaction effects in these four models is troubling but understandable. It is not surprising that coefficients for one categorical variable will change by differing amounts across the range of another categorical variable. What is not clear is how much this will affect the estimates of key main effect variables such as the abundance variable (eg. fishing year). One difficulty in gauging changes in these estimates is the deterioration of the quality of the estimates as indicated by the wide (or non-existent) standard errors that are estimated for some of the coefficients in these interaction models (eg. Figure 31). The credibility of these models is suspect given the poor determination of the coefficients and the likely overparameterisation of the models. It is reasonable to suppose that our understanding of interaction effects will improve as more analyses are performed and the amount of data available increases.

### 5.9.6 GLM conclusions

The consistency in the relative biomass estimates between these four reasonably independent analyses is striking (Figure 21). This is in spite of the low explanatory power associated with these coefficients from these models, amounting to less than $4 \%$ in most of the analyses except for the QCS\&HS model ( $=9 \%$; Table 9, Table 10, Table 11and Table 12) and the lack of contrast between years. Plots of the underlying distributions of LN(CPUE) by fishing year also show the lack of contrast in this variable in the most recent five fishing years. The lack of signal in the abundance trend from these fisheries may indicate that the effect of fishing has been low on the vulnerable biomass, either because the biomass level is large relative to the exploitation rate or that recruitment and mortalities are reasonably balanced.

### 6.0 Biomass estimation using spatial data from research surveys

### 6.1 Research surveys

The distributional characteristics of turbot make it a good candidate for abundance indexing using area-swept survey methods. Biomass estimates for turbot from research surveys were available for Washington, west coast Vancouver Island, Hecate Strait, Aleutian Islands and the eastern Bering Sea (Table 18). The estimates for Washington and the west coast of Vancouver Island were produced using data from the U.S. west coast triennial trawl survey (Mark Wilkins pers. comm.). The 1980 estimate for Hecate Strait is from the turbot biomass survey conducted by Fargo et al. (1981). The estimates for Hecate Strait in 1984, 87, 89, 91, 93, $95,96,98$, and 2000 were calculated using data from the Hecate Strait multispecies trawl survey using the same derivation detailed in Fargo et al. (1990). They represent the biomass of the entire population. The biomass estimates for the eastern Bering Sea and Aleutian Islands were presented in the 2000 stock assessment for these regions by Wilderbuer and Sample (2000). These estimates represent the exploitable population.

### 6.2 Biomass estimates

The abundance of turbot increased with latitude. The lowest densities were observed off Washington state and the highest densities were found in the Bering Sea where considerably more prime habitat is available than in the other areas. Biomass estimates for the west coast of Vancouver Island ranged from $16,417 \mathrm{t}$ in 1992 to $75,424 \mathrm{t}$. in 1995. However the 1992 survey did not cover 3 of the 5 depth strata that received survey coverage in the other years. The estimates for Washington Start and the west coast of Vancouver Island show no overall tend. Biomass estimates for Hecate Strait ranged from 24,362 t in 1993 to $72,011 \mathrm{t}$ in 1989. Adults comprise about $90 \%$ of the biomass by weight. Biomass in Hecate Strait increased between 1980 and 1991 then declined from 1991 to 1998 and increased in 2000 to a level comparable to 1991.

Table 18. Turbot biomass estimates for Washington, west coast of Vancouver Island, Hecate Strait, Aleutian islands and the eastern Bering Sea, 1977-2000.
Source

| Year | Biomass ${ }^{1}$ survey | Multispecies Trawl survey ${ }^{1}$ | Triennial Survey ${ }^{1}$ | Triennial Survey | U.S.- Japan $^{2}$ | U.S.Japan |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |

Area


### 6.3 Relative CPUE

A summary of the annual CPUE index for turbot from the Hecate Strait multispecies survey is presented in Figure 34. There is no overall trend but the index fluctuates from a relatively low level in 1984 to a high level maximum in 1989 followed by a decline to 1993 and an increase to 2000. The biomass level in 2000 is comparable to that in the early 1990s. A summary of the CPUE index by 10 fathom depth interval is presented in Figure 35. Most of these plots show the 1989 peak that was seen in the overall CPUE index (Figure 34). However, there is more variation in these indices. This is due partly to the smaller sample size but also to the species preference for depth (Perry et. al. 1994).


Figure 34. Mean CPUE and $90 \%$ confidence interval for turbot from the Hecate Strait multispecies survey, 1984-2000.


Figure 35. Mean CPUE and $90 \%$ confidence interval for turbot by 10 fathom depth interval from the Hecate Strait multispecies surveys, 1984 - 2000.

### 7.0 Recommendations and Yield Options

These analyses suggest that turbot populations in all areas examined are being fished at or below sustainable levels. We conclude this for the following reasons:

- There was no detectable change in $Z$ in Hecate Strait between 1980 and 2000, implying that exploitation rates in the area covered by the Hecate Strait trawl survey have not substantially changed the age distribution of the turbot population over that period. The age compositions shown in Figure 5 show no truncation of the older age classes.
- The biomass survey index for Hecate Strait shows no long term trend despite having the longest exploitation history of any area(Figure 34).
- Assuming that catch rates are a measure of population abundance, the catch rates derived from GLM models for the west coast of Vancouver Island show no trends for retained catch (Figure 9) or discard catch (Figure 12) from 1996/97 to 2000/01. Similarly, catch rates from the Queen Charlotte Sound/Hecate Strait show no trends for retained catch (Figure 15) or discard catch (Figure 18) from 1996/97 to 2000/01.
- There is little evidence that targeting by the fleet is causing the high proportion of discards in the turbot fishery (Figure 16 and Figure 19). Only a few vessels appear to specialise in landing this species (Figure 10 and Figure 13).
- Estimates of $F$ were below the best estimate for $M(=0.2$; Section 4.4).

At the present time turbot are largely taken as bycatch in other target fisheries. This is due to a lack of markets for this species. However, over the past 2-3 years there has been a lot of interest in this species and targeting has been undertaken by a few vessels as landings opportunities arise. The analysis of the Canadian observer data indicates that most of the variation CPUE over time is due to differences in areas, time of year fished and, depth fished.

Currently off B.C., there are biomass estimates for the west coast of Vancouver Island from US Triennial Survey and for Hecate Strait from the long-term multispecies survey. However, both of these surveys fail to cover the entire known range of turbot in BC, either in terms of the depth range or area. In addition, there is a substantial population in Queen Charlotte Sound but no biomass survey.

The stock structure of turbot is poorly understood. It is thought that the population of turbot off the west coast of Vancouver Island is separate from the population in Hecate Strait because of geographic barriers between the two areas. It is not known if the population in Queen Charlotte Sound is contiguous with the population in Hecate Strait, with that off the west coast of Vancouver Island or whether it is a separate stock

Given the current information, the current level of catch is probably the minimum sustainable because fishing mortality is below the instantaneous natural mortality rate. The analyses presented in this paper indicate that current levels of removals, which include discards have been sustainable over the last 20 years in Hecate Strait) and over the last 12 years for the west coast of Vancouver Island. Evidence from the discard analysis indicates that this aspect of the fishery appears to be passive with no indication of targeting behaviour. However, managers should work with the industry to reduce the level discards in this fishery where possible.

Although it is likely that the resource could sustain increased catches in some areas, it is not possible to recommend specific higher catch limits at this time. We recommend that, if catch levels for this species are increased, a monitoring program should also put in place along with agreed responses to changes in the "performance indicators" which are generated by the monitoring program. There are many examples of similar management procedures which are operational in various fisheries on a world-wide basis. This would ensure the safe development of the fishery for this species.

### 8.0 Acknowledgements

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### 10.0 Appendix 1: GLM data preparation

### 10.1 Catch and effort data

All catch and effort data were obtained from summary tables generated from the PacHarvest and GFCatch databases held by the DFO at the Pacific Biological Station in September 2001. See Schnute et al. (2000) for a description of the PacHarvest database, including the available data fields. A description of the GFCatch database is available in Rutherford (1999).

The two datasets were amalgamated for performing the analyses on the retained catch of turbot. However, only two years could be added as records of latitudes and longitudes for the individual tows were not maintained until January 1994. The addition of these data from the GFCatch database considerably aided the analysis of the retained turbot catch as there were relatively few records in this model.

Only the PacHarvest data were used for the analysis of discarded catch because it was not until mid-February 1996 that observers were present on most vessels which caught this species. There are no reliable estimates of discarded catch prior to February 1996.

### 10.2 GLM Data preparation and grooming

Records satisfying the following conditions were kept for the analysis in this report:

- Tow start date after 31 March 1994 for retained catch and 31 March 1996 for discarded catch
- Bottom trawl type
- Areas outside the Strait of Georgia (i.e. $<>$ Major Area $=4 \mathrm{~B}$ )
- Fishing success code $<=1$ (code $0=$ unknown; code $1=$ useable)
- Valid major, minor and locality area codes
- Valid depth value
- Tows from vessels which had been in the fishery for at least four years (for discarded catch analysis) and five years (for retained catch analysis)
- Tows with valid latitude and longitude co-ordinates
- Tows with valid estimates of time set

The locations of the selected tows have been translated into UTM (Universal Transverse Mercator) co-ordinates based on the latitude and longitude for each tow in the database. The UTM grid system and its application is described in detail in Schnute et al. (2000). This system was used to generate constant sized grid boxes of $5 \mathrm{~km}^{2}, 10 \mathrm{~km}^{2}, 25 \mathrm{~km}^{2}$ and $40 \mathrm{~km}^{2}$ which each offered as predictor variables in the generalised linear analysis.

Continuous variables (time of day set, number of species in tow and number of tows by vessel in data set) were entered into the model as third order polynomials. However, exploratory analysis showed that these variables had little explanatory power in any of the models and were dropped from the final analyses. As well, plots of the data showed that the time of day set variable was incomplete prior to 1996/97, making this variable unsuitable for the retained catch models. The number of species reported in a tow also seems suspiciously low in 1994/95 and 1995/96, rendering this variable unsuitable for the retained catch models.

Depth was entered into the model as a categorical variable in 50 m depth bands. Initial analysis also showed that depth had little explanatory power in these models and the depth distributions were truncated in all models at the 5 and $95 \%$-tiles to only include depth bands with sufficient number of observations. The only exception to this was the discard catch model for the combined Queen Charlotte Sound and Hecate Strait where the depth distribution was truncated to the 1 and $99 \%$-tiles as this model has the greatest number of tows. Depth as a continuous variable was also investigated in some of the regression models. However, the fit of the data to depth when offered as a $3^{\text {rd }}$ order polynomial was generally not as good as for the categorical variable.

Fields or derived fields that were kept in the data set are described in Appendix Table 1. All variables in the final models were entered as categorical variables.

Appendix Table 1. Fields in the data set used to analyse turbot catch and effort data

| Field | Description |
| :--- | :--- |
| CPUE | Kg/hour |
| Depth | Converted into 50m bands |
| Discarded catch | Kg |
| Effort | Tow time in hours |
| Latitude | In decimal degrees based on the mid-point of the tow |
| Locality | DFO locality area description (Rutherford 1999) |
| Longitude | In decimal degrees based on the mid-point of the tow |
| Minor Area | DFO minor area description (Rutherford 1999) |
| Month | From April 1994 to March 2001 |
| Nspp | Number of species in tow |
| Pcatch | Proportion turbot catch to total catch |
| Pdiscard | Proportion of turbot discard to total discard |
| Retained catch | Kg |
| Standardised fishing year | 01 April - 31 March |
| Time | Time of day when tow set |
| Tows | Total number tows by vessel in fishery |
| UTME | Universal Transverse Mercator Easting (Schnute et al 2000) |
| UTMN | Universal Transverse Mercator Northing (Schnute et al 2000) |
| Vessel | Coded |

### 11.0 Appendix 2: Calculating Standard Errors for CPUE

## By Chris Francis, 7 June 2001

This note describes how to calculate standard errors (s.e.s) for standardised CPUE indices following the suggestion of Francis (1999). These s.e.s relate to what Francis (1999) called the canonical form for CPUE indices (in which there is no reference year).

These s.e.s have two advantages over those calculated by the method that has been used by most people at NIWA. First, an s.e. is calculated for every year (i.e., we don't have an s.e. of 0 for the reference year). Second, the s.e.s are not inflated by the uncertainty associated with the reference year (see fig. 2 of Francis 1999, which shows that the s.e.s can vary a lot with the choice of reference year, and that they are always much smaller when there is no reference year).

In the procedure described here, I am assuming that you have carried out a CPUE standardisation in the conventional way (i.e., with a reference year) and are able to obtain, from whatever software you are using, the estimated covariance matrix for your year coefficients. By the year coefficients, I mean the vector of regression coefficients associated with each year. If there are $n$ years in your data set this vector will have length $n-1$ (because the reference year is excluded). Note that these year coefficients are in log space; to get year effects (in natural space) you need to exponentiate the year coefficients. Suppose that the reference year you have chosen is the $r$ th of your $n$ years (very often $r=1$ ) and that $\mathbf{V}$ is the estimated covariance matrix for the year coefficients (so $\mathbf{V}$ is an $(n-1) \times(n-1)$ matrix).

You need to construct an $n \times(n-1)$ matrix $\mathbf{Q}$, whose $i j$ th element is given by

$$
Q_{i j}= \begin{cases}(n-1) / n & \text { if } i<r \text { and } i=j \\ (n-1) / n & \text { if } i>r \text { and } i=j+1 \\ -1 / n & \text { otherwise }\end{cases}
$$

If you are working in $S$ or Splus, you can create $\mathbf{Q}$ using the following two commands

$$
\begin{aligned}
& \mathrm{Q}<-\operatorname{matrix}(-1 / \mathrm{n}, \mathrm{n}, \mathrm{n}-1) \\
& \mathrm{Q}[-\mathrm{r},]<-\mathrm{Q}[-\mathrm{r},]+\operatorname{diag}(\operatorname{rep}(1, \mathrm{n}-1))
\end{aligned}
$$

Now calculate matrix $\mathbf{V}^{0}=(\mathbf{Q} * \mathbf{V}) * \mathbf{Q}^{\prime}$, where * represents matrix multiplication and $\mathbf{Q}^{\prime}$ is the matrix transpose of $\mathbf{Q}$. In S or Splus, $\mathrm{V} 0<-(\mathrm{Q} \% * \% \mathrm{~V}) \% * \% \mathrm{t}(\mathrm{Q})$.
$\mathbf{V}^{0}$ is the covariance matrix that you want. That is, the s.e. for the $i$ th year coefficient is $\sqrt{V_{i i}^{0}}$. (Note that $\mathbf{V}^{0}$ is an $n \times n$ matrix so this provide s.e.s for all $n$ years, including the reference year).

### 12.0 Appendix 3: Data Summaries Supporting the GLM Analysis

The following tables and graphs summarise the data available in each dataset used for the GLM models presented in Sections 5.6 and 5.7. Each model is supported by the following data summaries consisting of two tables and five graphs:

- $\quad 1^{\text {st }}$ table: distribution of vessels, tows, catch, and effort by number years that the vessels have been in the fishery;
- $\quad 2^{\text {nd }}$ table: distribution of tows by standard fishing year and DFO minor statistical area;
- $\quad 1^{\text {st }}$ graph: box plots of the distribution of data for key variables by standard fishing year;
- $\quad 2^{\text {nd }}$ graph: box plots of the distribution of LN(CPUE) for key predictor variables used in the models.
Note that, for both graphs and the $2^{\text {nd }}$ table, the data have been restricted to vessels which had been in the fishery for at least five years (retained catch models) or for four years (discard catch models) and that records at the extremes of the depth distributions have been dropped.


### 12.1 West coast Vancouver Island retained catch analysis summary statistics

Appendix Table 2. Number of vessels in the WCVI retained catch dataset grouped by the number of years the vessel was in the fishery along with the associated number of tows, total catch in kg and total effort in hours. These latter three totals are also expressed as a descending cumulative total in percent, indicating the cumulative amount of data present in the dataset beginning at that number of years.

| Years in <br> fishery | Number <br> vessels | Total <br> tows | Cotal <br> $\mathbf{C a t c h}$ <br> $\mathbf{k g})$ | Total <br> effort <br> $(\mathbf{h r s})$ | Descending <br> cumulative <br> tows (\%) | Descending <br> cumulative <br> catch (\%) | Descending <br> cumulative <br> effort (\%) |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 19 | 164 | 160,385 | 473 | 100 | 100 | 100 |
| 2 | 20 | 645 | 932,938 | 1,561 | 98 | 99 | 98 |
| 3 | 10 | 601 | 563,382 | 1,393 | 90 | 92 | 89 |
| 4 | 6 | 489 | $1,147,583$ | 1,222 | 83 | 88 | 82 |
| 5 | 11 | 1165 | $2,247,239$ | 2,717 | 76 | 79 | 76 |
| 6 | 10 | 1634 | $2,955,348$ | 3,718 | 62 | 63 | 61 |
| 7 | 16 | 3381 | $5,579,353$ | 7,900 | 42 | 41 | 42 |

Appendix Table 3. Number of tows in the WCVI retained catch dataset by standard fishing year and DFO minor statistical area (arranged from south to north) in the dataset. Note that the tows from minor area 11 only extend to Cape Scott. The dataset has been limited to vessels which have been in the fishery for at least five years and tows outside of $5 / 95 \%$ of the depth distribution have been dropped.

| Fishing | DFO Minor Statistical Area |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Year | $\mathbf{2 1}$ | $\mathbf{2 3}$ | $\mathbf{2 4}$ | $\mathbf{2 5}$ | $\mathbf{2 6}$ | $\mathbf{2 7}$ | $\mathbf{1 1}$ | Total |
| $1994 / 95$ | 0 | 195 | 69 | 248 | 12 | 27 | 24 | 575 |
| $1995 / 96$ | 7 | 135 | 124 | 239 | 10 | 38 | 12 | 565 |
| $1996 / 97$ | 4 | 175 | 185 | 319 | 29 | 30 | 6 | 748 |
| $1997 / 98$ | 2 | 196 | 193 | 194 | 35 | 39 | 40 | 699 |
| $1998 / 99$ | 9 | 182 | 175 | 376 | 26 | 60 | 87 | 915 |
| $1999 / 2000$ | 6 | 308 | 213 | 274 | 7 | 68 | 41 | 917 |
| $2000 / 01$ | 5 | 314 | 388 | 365 | 13 | 88 | 62 | 1,235 |
| Total | 33 | 1,505 | 1,347 | 2,015 | 132 | 350 | 272 | 5,654 |



Appendix Figure 1. Distribution of data for catch (kg), time set (hours), depth (m) and day of year of tow used in the WCVI retained catch model by standard fishing year. Depth is truncated at 5 and $95 \%$ of the depth distribution and the data set confined to vessels which have been the fishery for at least five years.


Appendix Figure 2. Distribution of $\mathrm{LN}(\mathrm{CPUE})[\mathrm{kg} / \mathrm{hr}]$ for five categorical predictor variables used in the WCVI retained catch model. The width of the boxes is proportional to the number of tows in each category and the mean of $\mathrm{LN}(\mathrm{CPUE})$ is plotted as a horizontal line in all graphs.

### 12.2 Combined Queen Charlotte Sound/Hecate St retained catch analysis summary statistics

Appendix Table 4. Number of vessels in the QCS \& HS retained catch dataset grouped by the number of years the vessel was in the fishery along with the associated number of tows, total catch in kg and total effort in hours. These latter three totals are also expressed as a descending cumulative total in percent, indicating the cumulative amount of data present in the dataset beginning at that number of years.

| Years in <br> fishery | Number <br> vessels | Total <br> tows | Total <br> Catch <br> $(\mathbf{k g})$ | Total <br> effort <br> $(\mathbf{h r s})$ | Descending <br> cumulative <br> tows (\%) | Descending <br> cumulative <br> catch (\%) | Descending <br> cumulative <br> effort (\%) |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 16 | 70 | 7,812 | 153 | 100 | 100 | 100 |
| 2 | 17 | 483 | 195,426 | 1,061 | 99 | 100 | 99 |
| 3 | 15 | 756 | 375,777 | 1,636 | 93 | 98 | 93 |
| 4 | 13 | 919 | $1,445,195$ | 2,049 | 84 | 95 | 83 |
| 5 | 18 | 1596 | $1,476,905$ | 3,291 | 72 | 82 | 71 |
| 6 | 11 | 1272 | $1,828,150$ | 2,628 | 53 | 68 | 52 |
| 7 | 15 | 2971 | $5,650,248$ | 6,219 | 37 | 51 | 37 |

Appendix Table 5. Number of tows in the QCS \& HS retained catch dataset by standard fishing year and DFO minor statistical area (arranged from north to south ) in the dataset. Note that the tows from minor area 11 only extend from Cape Scott. The dataset has been limited to vessels which have been in the fishery for at least five years and tows outside of $5 / 95 \%$ of the depth distribution have been dropped.

| Fishing Year | DFO Minor Statistical Area |  |  |  |  |  |  |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2AE | 2BE | 1E | 4 | 5 | 6 | 7 | 8 | 9 | 11 |  |
| 1994/95 | 1 | 9 | 64 | 107 | 33 | 19 | 1 | 220 | 0 | 67 | 521 |
| 1995/96 | 0 | 29 | 7 | 117 | 6 | 10 | 3 | 166 | 0 | 77 | 415 |
| 1996/97 | 0 | 32 | 2 | 68 | 3 | 16 | 3 | 221 | 1 | 154 | 500 |
| 1997/98 | 0 | 7 | 9 | 44 | 17 | 23 | 2 | 248 | 0 | 318 | 668 |
| 1998/99 | 0 | 26 | 22 | 102 | 16 | 0 | 0 | 372 | 0 | 456 | 994 |
| 1999/2000 | 0 | 8 | 29 | 176 | 3 | 9 | 1 | 459 | 0 | 445 | 1,130 |
| 2000/01 | 1 | 15 | 6 | 176 | 16 | 7 | 2 | 562 | 0 | 281 | 1,066 |
| Total | 2 | 126 | 139 | 790 | 94 | 84 | 12 | 2,248 | 1 | 1,798 | 5,294 |



Appendix Figure 3. Distribution of data for catch (kg), time set (hours), depth (m) and day of year for tow used in the QCS\&HS retained catch model by fishing year. Depth is truncated at 5 and $95 \%$ of the depth distribution and the data set confined to vessels which have been the fishery for at least five years.


Appendix Figure 4. Distribution of LN(CPUE) [kg/hr] for five categorical predictor variables used in the QCS\&HS retained catch model. The width of the boxes is proportional to the number of tows in each category and the mean of LN(CPUE) is plotted as a horizontal line in all graphs.

### 12.3 West coast Vancouver Island discard catch analysis summary statistics

Appendix Table 6. Number of vessels in the WCVI discard catch dataset grouped by the number of years the vessel was in the fishery along with the associated number of tows, total catch in kg and total effort in hours. These latter three totals are also expressed as a descending cumulative total in percent, indicating the cumulative amount of data present in the dataset beginning at that number of years.

| Years in <br> fishery | Number <br> vessels | Total <br> tows | Total <br> Catch <br> (kg) | Total <br> effort <br> (hrs) | Descending <br> cumulative <br> tows (\%) | Descending <br> cumulative <br> catch (\%) | Descending <br> cumulative <br> effort (\%) |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 28 | 909 | 250,518 | 2,082 | 100 | 100 | 100 |
| 2 | 12 | 689 | 220,810 | 1,708 | 93 | 93 | 93 |
| 3 | 10 | 808 | 201,771 | 1,764 | 88 | 88 | 87 |
| 4 | 5 | 837 | 193,209 | 1,915 | 81 | 82 | 82 |
| 5 | 32 | 9709 | $2,955,346$ | 22,627 | 75 | 77 | 75 |

Appendix Table 7. Number of tows in the WCVI discard catch dataset by standard fishing year and DFO minor statistical area (arranged from south to north) in the dataset. Note that the tows from minor area 11 only extend to Cape Scott. The dataset has been limited to vessels which have been in the fishery for at least four years and tows outside of $5 / 95 \%$ of the depth distribution have been dropped.

| Fishing | DFO Minor Statistical Area |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Year | $\mathbf{2 1}$ | $\mathbf{2 3}$ | $\mathbf{2 4}$ | $\mathbf{2 5}$ | $\mathbf{2 6}$ | $\mathbf{2 7}$ | $\mathbf{1 1}$ | Total |
| $1996 / 97$ | 7 | 355 | 486 | 363 | 86 | 133 | 134 | 1,564 |
| $1997 / 98$ | 6 | 327 | 517 | 232 | 54 | 170 | 164 | 1,470 |
| $1998 / 99$ | 50 | 441 | 641 | 325 | 52 | 176 | 261 | 1,946 |
| $1999 / 2000$ | 34 | 611 | 604 | 431 | 66 | 241 | 285 | 2,272 |
| $2000 / 01$ | 15 | 512 | 685 | 435 | 65 | 254 | 303 | 2,269 |
| Total | 112 | 2,246 | 2,933 | 1,786 | 323 | 974 | 1,147 | 9,521 |



Appendix Figure 5. . Distribution of data for catch (kg), time set (hours), depth (m) and day of year for tow used in the WCVI discard catch model by fishing year. Depth is truncated at 5 and $95 \%$ of the depth distribution and the data set confined to vessels which have been the fishery for at least four years.


Appendix Figure 6. Distribution of LN(CPUE) [kg/hr] for five categorical predictor variables used in the WCVI discard catch model. The width of the boxes is proportional to the number of tows in each category and the mean of $\mathrm{LN}(\mathrm{CPUE})$ is plotted as a horizontal line in all graphs.

### 12.4 Combined Queen Charlotte Sound/Hecate St discard catch analysis summary statistics

Appendix Table 8. Number of vessels in the QCS \& HS discard catch dataset grouped by the number of years the vessel was in the fishery along with the associated number of tows, total catch in kg and total effort in hours. These latter three totals are also expressed as a descending cumulative total in percent, indicating the cumulative amount of data present in the dataset beginning at that number of years.

| Years in <br> fishery | Number <br> vessels | Total <br> tows | Total <br> Catch <br> (kg) | Total <br> effort <br> (hrs) | Descending <br> cumulative <br> tows (\%) | Descending <br> cumulative <br> catch (\%) | Descending <br> cumulative <br> effort (\%) |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 25 | 1795 | 350,061 | 3,717 | 100 | 100 | 100 |
| 2 | 14 | 1798 | 570,630 | 4,176 | 94 | 97 | 95 |
| 3 | 10 | 2072 | 719,481 | 4,446 | 89 | 92 | 88 |
| 4 | 7 | 2206 | 584,889 | 4,694 | 82 | 86 | 82 |
| 5 | 42 | 23881 | $9,409,150$ | 50,906 | 75 | 81 | 75 |

Appendix Table 9. Number of tows in the QCS \& HS discard catch dataset by standard fishing year and DFO minor statistical area (arranged from north to south ) in the dataset. Note that the tows from minor area 11 only extend from Cape Scott. The dataset has been limited to vessels which have been in the fishery for at least five years and tows outside of $5 / 95 \%$ of the depth distribution have been dropped.

| Fishing Year | DFO Minor Statistical Area |  |  |  |  |  |  |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2AE | 2BE | 1E | 4 | 5 | 6 | 7 | 8 | 9 | 11 |  |
| 1996/97 | 14 | 368 | 142 | 965 | 230 | 173 | 33 | 1,404 | 1 | 680 | 4,010 |
| 1997/98 | 17 | 248 | 104 | 898 | 282 | 228 | 18 | 1,968 | 4 | 975 | 4,742 |
| 1998/99 | 17 | 325 | 138 | 1,367 | 259 | 179 | 26 | 1,906 | 1 | 1,392 | 5,610 |
| 1999/2000 | 7 | 521 | 134 | 1,412 | 166 | 304 | 31 | 2,554 | 1 | 1,094 | 6,224 |
| 2000/01 | 9 | 326 | 122 | 1,172 | 144 | 162 | 24 | 2,217 |  | 816 | 4,992 |
| Total | 64 | 1,788 | 640 | 5,814 | 1,081 | 1,046 | 132 | 10,049 | 7 | 4,957 | 25,578 |



Appendix Figure 7. Distribution of data for catch (kg), time set (hours), depth (m) and day of year for tow used in the QCS \& HS discard catch model by fishing year. Depth is truncated at 1 and $99 \%$ of the depth distribution and the data set confined to vessels which have been the fishery for at least four years.


Appendix Figure 8. Distribution of LN(CPUE) [kg/hr] for five categorical predictor variables used in the QCS\&HS discard catch model. The width of the boxes is proportional to the number of tows in each category and the mean of $\mathrm{LN}(\mathrm{CPUE})$ is plotted as a horizontal line in all graphs.

