Canadian Stock Assessment Secretariat Research Document 98/66

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Secrétariat canadien pour l'évaluation des stocks Document de recherche 98/66

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# Assessment of Haddock on Eastern Georges Bank 

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#### Abstract

Haddock catches from eastern Georges Bank fluctuated around 5,000 tons from the early 1980s to 1993. Under restrictive management measures, catches declined from $6,377 \mathrm{t}$ in 1991 to a low of $2,111 \mathrm{t}$ in 1995, but increased again to $3,720 \mathrm{t}$ and $2,850 \mathrm{t}$ in 1996 and 1997 respectively. About $70 \%$ of the 1997 catch weight was comprised of haddock from the 1992 and 1993 year-classes. The trend in ages 3-8 abundance from surveys increased from 1992 to 1995 and has fluctuated since then. Surveys indicate that the 1996 year-class may be comparable to the moderate 1983, 1985, 1987 and 1992 year-classes.

Total population biomass (ages $1+$ ) has steadily increased from near historic low levels of $12,171 \mathrm{t}$ in 1993 to $28,809 \mathrm{t}$ in 1998. The recent increase, due principally to the 1992 year-class, but also supported by the 1991 and 1993 year-classes, was enhanced by increased survivorship of young haddock resulting from reduced capture of small fish in the fisheries. The abundance of the 1996 year-class was estimated at about 13 million. The incoming 1995 and 1997 year-classes appear relatively weak at about 6 and 3 million respectively. Exploitation rate for ages 4 and older has consistently been below the $\mathrm{F}_{0,1}$ target of $20 \%$ since 1995. Reduced fishing mortality in recent years has resulted in increased survival of incoming year-classes and greater abundance at older ages.

Combined Canada/USA projected yield at $\mathrm{F}_{0.1}=0.25$ in 1998 would be about $6,000 \mathrm{t}$. If fished at $\mathrm{F}_{0.1}$ in 1998, the biomass for ages 3 and older is projected to increase from $22,726 \mathrm{t}$ to $28,012 \mathrm{t}$ at the beginning of 1999. The 1996 year-class was estimated to contribute almost $30 \%$ to the 1999 age $3+$ biomass. A catch of $3,000 t$ in 1998, about what was caught in 1997, results in a negligible risk that fishing mortality rate will exceed $\mathrm{F}_{0.1}$ and that the biomass for ages 3 and older will decrease. That same yield gives risks of $10 \%$ and $35 \%$ that an increase in biomass of $10 \%$ and $20 \%$ respectively would not be achieved.


## Résumé

Les captures d'aiglefin dans la partie est du banc Georges ont été de l'ordre de 5000 t entre les années 1980 et 1993. Suite à l'imposition de mesures de gestion strictes, elles sont tombées de 6 377, en 1991, à 2111 ten 1995 pour ensuite augmenter à 3720 tet 2850 ten , respectivement, 1996 et 1997. Environ $70 \%$ des captures de 1997, en poids, étaient formées d'aiglefins des classes d'âges de 1992 et 1993. L'abondance classes d'âges 3 à 8 , déterminée par relevés, s'est accrue de 1992 à 1995 et a ensuite fluctué. Les relevés montrent que la classe de 1996 pourrait être comparable aux classes moyennes de 1983, 1985, 1987 et 1992.

La biomasse de la population totale (âges $1+$ ) a augmenté de façon constante du niveau historiquement faible de 12171 t , en 1993, à 28809 t , en 1998. L'augmentation récente, s'expliquant surtout par la classe de 1992 mais aussi par celles de 1991 et 1993, s'est vue favorisée par une meilleure survie des jeunes aiglefins s'expliquant par une réduction de la capture des poissons de petite taille. L'abondance de la classe de 1996 a été estimée à 13 millions de poissons environ. Les classes de 1995 et 1997 semblent relativement faibles comptant respectivement 6 et 3 millions d'individus. Depuis 1995 , le taux d'exploitation des âges 4 et plus a constamment été inférieur à l'objectif du $\mathrm{F}_{0,1}$ de $20 \%$. La faible mortalité par pêche des demières années a donné lieu à une augmentation de la survie des classes à venir et à une abondance accrue des classes plus âgées.

Le rendement prévu total du Canada et des États-Unis au niveau $\mathrm{F}_{0,1}=0,25$ serait de 6000 t environ en 1998. À une exploitation au niveau du $\mathrm{F}_{0,1}$ en 1998 , la biomasse des âges 3 et plus devrait augmenter pour passer de 22726 t à 28012 t au début de 1999. Il a été estimé que la classe de 1996 devrait représenter près de $30 \%$ de la biomasse des âges $3+$ en 1999. Des captures de 3000 ten 1998, soit environ le volume de 1997, devraient constituer un risque négligeable que la mortalité par pêche soit supérieure à celle du niveau $F_{0,1}$ et que la biomasse des âges 3 et plus diminue. Ce même rendement correspond à un risque de $10 \%$ ou de $35 \%$ qu'une augmentation de la biomasse respective de $10 \%$ et $20 \%$ ne soit pas atteinte.

## Introduction

Since 1990 Canada has used eastern Georges Bank, fishery statistical units 5 Zj and 5 Zm (Fig. 1), as the basis for a management unit (Gavaris 1989). Results from the previous assessment (Gavaris and Van Eeckhaute 1997) showed that biomass increased from 12,000 t in 1993 to about $24,000 \mathrm{t}$ in 1996 and declined slightly to $23,000 \mathrm{t}$ in 1997. Present biomass levels were estimated to be well below historical levels. In the 1930s to the mid 1950s when biomass was over $60,000 \mathrm{t}$ and about $40,000 \mathrm{t}$ in the late 1970s and early 1980s. Recent year-classes sizes were also well below historical levels. Since 1968, only two year-classes, the 1975 and 1978 at about 50 million each, have been comparable to average recruitment during the 1930s to the mid 1950s. The 1991, 92 and 93 year-classes were estimated at 7,15 and 10 million respectively while others since the 1987 year-class were considerably weaker. Exploitation rate declined from over $40 \%$ in 1991 and 1992 to below $20 \%$ in 1995 and 1996, resulting in increased survival of the moderate 1992 and 1993 year-classes. The 1997 Canadian quota was $3,200 \mathrm{t}$. For a catch of about this level, there was a greater than $80 \%$ chance of $3+$ biomass increasing with an expected biomass increase of about $10 \%$ and less than a $10 \%$ chance of exceeding $\mathrm{F}_{0.1}$.

In this assessment update we included the latest information from the 1997 Canadian and USA fisheries and made minor revisions to the 1996 data. Results from the Department of Fisheries and Oceans, Canada (DFO) survey in the spring of 1998 and the National Marine Fisheries Service,USA (NMFS) surveys in the spring and fall of 1997 were incorporated. We also explored some alternative assessment model formulations and evaluated the effect of the change in trawl nets used for the NMFS spring survey.

## The Fishery

## Commercial Catches

Under restrictive management measures, catches declined from $6,377 \mathrm{t}$ in 1991 to a low of $2,111 \mathrm{t}$ in 1995, but increased again to $3,720 \mathrm{t}$ and $2,850 \mathrm{t}$ in 1996 and 1997 respectively (Table 1, Fig. 2). Greater catches in the late 1970s and early 1980s, ranging up to 23,189 t in 1980, were associated with good recruitment. Substantial quantities of small fish were discarded in those years (Overholtz et al 1983). Catches subsequently declined and fluctuated at about $5,000 \mathrm{t}$ during the mid to late 1980s.

The haddock on Georges Bank have supported a commercial fishery since the early 1920s (Clark et al 1982). Catches during the 1930s to 1950s ranged between $15,000 \mathrm{t}$ and $40,000 \mathrm{t}$, averaging about $25,000 \mathrm{t}$ (Fig. 3). Catches probably attained record high levels of about $60,000 \mathrm{t}$ during the early 1960 s but since the early 1970 s catches have been lower.

The predominantly USA fishery was joined by Canadian and distant water fleets notably the USSR and Spain by the early 1960s. In 1953, the International Commission for the Northwest Atlantic Fisheries (ICNAF) implemented a minimum mesh size of 114 mm in the body and codend of towed gear. A Total Allowable Catch was introduced in 1970 by ICNAF in an attempt to curb rapidly declining abundance. Seasonal closures of haddock spawning areas were also instituted in that year as an adjunct and have been retained by Canada and the USA (Halliday 1988). Both the season and the area closed have gone through several modifications.

Following the declaration of economic zones to 200 mi by coastal states in 1977, only Canada and the USA continued haddock fisheries on Georges Bank. After the establishment of a maritime boundary in 1984 by the International Court of Justice, the Canadian and USA fisheries have been restricted to their respective jurisdictions. Canada has retained a quota regulatory system and uses ancillary measures to augment management. Fishermen now pay for access to the fishery, for dockside monitoring and contribute to the costs of at sea observer coverage. The USA has not regulated catch by quotas since 1977 but has relied on other measures (area closures, trip limits, fish size, etc.) and has recently instituted an effort regulatory system. Further details of regulatory measures since 1977 are summarized in Table 2.

As in 1995 and 1996, Canadian catches in 1997 of $2,739 \mathrm{t}$ were below the quota due to closure of the fisheries when the cod quotas were reached. During 1994 to 1997, all Canadian groundfish fisheries on Georges Bank remained closed from January to early June.

| Fishery Sector | 1993 |  | 1994 |  |  | 1995 |  |  | 1996 |  | 1997 |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: |
|  | Quota | Catch | Quota | Catch | Quota | Catch | Quota | Catch | Quota | Catch |  |  |
| Fixed gear <65' | 1508 | 1216 | 791 | 784 | 592 | 357 | 1085 | 919 | 754 | 714 |  |  |
| Mobile gear $<65^{\prime}$ | 2212 | 1646 | 1439 | 1206 | 1268 | 1175 | 2280 | 1713 | 1625 | 1451 |  |  |
| Fixed gear 65'-100' | 50 | 8 | 30 | 8 | 25 | 0 | 45 | 49 | 32 | 36 |  |  |
| Mobile gear 65'-100' | 32 | 32 | 30 | 33 | 25 | 27 | 189 | 181 | 32 | 35 |  |  |
| Vessels >100' | 1198 | 826 | 710 | 290 | 590 | 444 | 921 | 513 | 757 | 573 |  |  |
| Totals | 5000 | 3728 | 3000 | 2411 | 2500 | 2003 | 4500 | 3375 | 3200 | 2809 |  |  |

Catches are from quota reports and may not correspond exactly with statistics.
Although the number of vessels fishing with fixed gear declined in 1997, the number of vessels fishing with mobile gear and the total number of days fished by all fishing sectors was about the same as in 1996. All landings were monitored at dockside and at-sea monitoring by observers resulted in coverage of almost $10 \%$ of days fished. Discarding and misreporting have been considered negligible since 1992.

In recent years the Canadian fishery has been conducted by vessels using otter trawls, longlines, handlines and gillnets. During 1997, all vessels over 65 ft operated on enterprise allocations, otter trawlers under 65 ft operated on individual quotas, fixed gear vessels $45-65 \mathrm{ft}$ operated on self-administered individual quotas and fixed gear vessels under 45 ft operated on community quotas administered by local boards. Most haddock were caught by otter trawlers less than 65 ft (roughly equivalent to tonnage classes 1-3) and longliners less than 65 ft (Table 3). Unlike recent years, catches in June were not very high but the fishery improved in July (Table 4).

USA catches for 1997 were derived from logbooks coupled with dealer reports, as was done for 1994-96. Catches in 1997, which come almost exclusively from tonnage classes 3 and 4 otter trawlers (Table 5), remained low because the seasonal spawning area was closed to fishing during the entire year (Table 6). Effort in the USA fishery was regulated using Days-at-Sea limits. To curtail targeting of haddock, a 500 lb trip limit was introduced in 1994 and raised to $1,000 \mathrm{lb}$ in July 1996. In September 1997 the limit was raised to $1,000 \mathrm{lb}$ per day and a maximum of $10,000 \mathrm{lb}$ per trip (Table 2). The trip limit resulted in an increase in the discard rate. USA discards for 1994-1997 were estimated from dealer data and vessel trip reports at $258 \mathrm{t}, 25$ $\mathrm{t}, 41 \mathrm{t}$ and 63 t respectively. USA statistics for 1994-1997 are preliminary.

## Size and Age Composition

Sampling of the 1997 Canadian fishery for size and age composition of the catch by at sea observers and port sampling achieved coverage of all principal gears and all seasons (Table 7, Fig. 4). No sampling was available for discards of groundfish by-catch in the 1997 Canadian scallop fishery, although haddock by-catch in previous years has not been substantial. The observer samples were obtained on a set by set basis and these were pooled to the trip level to make them compatible with port samples before being combined with them. The calculations were done using the length-weight relationship which was derived from commercial fishery samples (Waiwood and Neilson 1985) ; weight $\left(\mathrm{kg}\right.$ round) $=0.0000158$ length $(\mathrm{cm})^{2.91612}$. With decreasing landings of haddock in the USA fishery, few port samples were available to characterize the size and age composition of the landings. Sea sampling for discards was limited and current reporting rates in vessel trip reports (logbooks) are inadequate to reliably estimate the quantity of haddock discards.

The size composition of catch in the Canadian fisheries for otter trawlers less than 65 ft , otter trawlers greater than 65 ft and longliners was similar, peaking at about 56 cm (Fig. 5). The gillnet catch represented a small proportion of the total catch, but consisted of larger haddock.

It has been suggested that fishery operations may not be typical when an observer is on board, therefore sampling by observers might not reveal high grading if it is occurring. Another source of data that may be examined for discarding or high grading is samples collected by surveillance officers. There are few length frequency samples from surveillance boardings of Georges Bank fisheries and the number of haddock measured per sample is low. The available data however, does not reveal any persistent patterns to suggest that high grading might be a significant problem (Appendix A).

Survey and commercial otoliths were read by L. Van Eeckhaute for DFO and by N. Munroe for NMFS. Results of intra-reader agreement tests for the DFO reader and inter-reader tests are detailed in Appendix B. No intra-reader agreement tests are available for the NMFS reader. Between reader tests for the DFO reader and the NMFS reader was poor for one of the tests with a substantial bias and, while the other test had satisfactory agreement, there was a slight bias in the opposite direction. At present, ageing discrepancies between NMFS and DFO have marginal impact on the fishery age composition as there is little catch by the USA in 5 Zj and 5 Zm . NMFS survey results are based on the NMFS readings. The cause of this bias and poor agreement should be investigated and appropriate measures taken to ensure that both agers are interpreting the otoliths in the same way.

Catch and weight at age for the commercial fishery were calculated by applying age length keys to length frequencies using the methods described in Gavaris and Gavaris (1983). About $70 \%$ of the 1997 catch weight was comprised of haddock from the 1992 and 1993 yearclasses. In contrast to pre-1994 catches, few haddock of ages 2 and 3 were caught in 1997, due in part to the type of gear used and to avoidance of areas with small fish (Fig. 6). In comparison to the age composition of the catch during earlier periods, age groups 4 to 6 and $9+$ appear to be well represented. Catches associated with the weak 1989 and 1990 year-classes are small. The strong 1962 and exceptional 1963 year-classes dominated catches in the early 1970s, therefore the haddock of ages $9+$ constituted a greater than average proportion of the catch in that period.

The updated 1996 and the new 1997 catch at age by quarter (Table 8) were used to augment the 1969-95 results reported by Gavaris and Van Eeckhaute (1997). Annual catch at age and average fishery weight at age are summarised in Tables 9 and 10. Fishery weights at age may be affected by variations in fishing practices and the time of year that catches occur. Survey weights at age were used to investigate changes in growth.

## Abundance Indices

## Commercial Catch Rate

Catch rate trends from the Canadian commercial fishery for selected trips by tonnage class 2 and 3 otter trawlers and longliners showed an increasing trend from 1993 to 1995 and remained relatively stable through 1996 and 1997 (Fig. 7). In contrast to 1995 and 1996, otter trawlers catch rates were relatively low in June of 1997 but displayed the expected increase in December. Longliner catch rates for tonnage class 2 increased markedly through season in 1997. Changes to regulations, gear modifications and varying fishing practices in recent years make comparison of catch rates from year to year difficult to interpret. Therefore, these were not used as indices of abundance.

## Research Surveys

Surveys have been conducted by the NMFS each year in the fall since 1963 and in the spring since 1968 and by DFO each year in the spring since 1986. All these surveys use a stratified random design. The strata boundaries for the NMFS spring and fall surveys are shown in Fig. 8 and those for the DFO spring survey are shown in Fig. 9. For the NMFS surveys, two vessels have been employed and there was a change in the trawl door in 1985. Conversion factors (Table 11), derived experimentally from comparative fishing, have been applied to the survey results to make the series consistent. Additionally, two trawl nets were used on the NMFS spring survey, a modified Yankee 41 during 1973-81 and a Yankee 36 in other years, but no conversion factors are available for haddock.

The distribution of catches for the most recent survey of each series was similar to the distribution over the previous 5 year period (Figs. 10-12). During spring, haddock of all ages are broadly distrubuted on top of the bank but ages 2 and older are more concentrated along the Northern Edge and Northeast Peak. In the fall, haddock are less prevalent on top of the bank and higher densities are found in deeper waters on the slopes of the Northern Edge and the Northeast Peak. The percent of biomass, ages $3-8$, on the Canadian side of 5 Zjm from the three surveys was summarised for the most recent years. During the NMFS fall survey almost all of the biomass occurred on the Canadian side. During the DFO spring survey, generally conducted in late February, most of the biomass was on the Canadian side although the percentage was lower in 1992 to 93 . During the NMFS spring survey, generally conducted in late March, the percentage on the Canadian side was typically lower than for the $\mathrm{DFO}^{-}$survey but these results were more variable.

| Percentage of biomass on Canadian side |  |  |  |
| :---: | :---: | :---: | :---: |
|  | Spring |  | Fall |
| Year | DFO | NMFS | NMFS |
| 1992 | 68 | 78 | 100 |
| 1993 | 67 | 43 | 99 |
| 1994 | 99 | 100 | 100 |
| 1995 | 98 | 62 | 100 |
| 1996 | 96 | 17 | 100 |
| 1997 | 92 | 93 | 100 |
| 1998 | 100 | N/A | N/A |

The trend in abundance for ages 3-8 increased from 1992 to 1995 and has fluctuated since then with no persistent pattern though all 3 surveys were low in 1997 (Fig. 13). Abundance peaked at record highs during the early 1960s. After declining to a record low in the early 1970s, it peaked again in the late 1970s, though at a lower level, and again during the mid to late 1980s at about half the level of the 1970s peak. Survey results for ages 1 and 2 indicate that the abundance of the 1996 year-class may be comparable to the moderate 1983, 1985, 1987 and 1992 year-classes (Fig. 14). These year-classes were considerably smaller than the strong 1975 and 1978 year classes and the very strong 1962 and exceptional 1963 year-classes. Age specific results from these surveys are shown in Tables 12-14.

The weights at age from the DFO spring survey, $\overline{W_{a}}$, were defined as the average of the mean weights at length, $l$, and age, $a, \overline{W_{l a}}$ (defined below), weighted by the stratified numbers in 5 Zj and 5 Zm at that length and age, $A_{l a}$ :
$\overline{W_{a}}=\frac{\sum_{l}\left(\overline{W_{l a}} A_{l a}\right)}{\sum_{I} A_{l a}}$.
$\overline{W_{l a}}$ is defined as the average of the weights, $W$, at length and age for all haddock sampled for detail measurements, $d$, where $A_{\text {lad }}$ is the total numbers sampled for detailed measurements at that length and age:
$\bar{W}_{l a}=\frac{\sum_{d} W_{l a d}}{\sum A_{l a d}}$.
With the exception of the higher than average weight at age for the 1989 and 1990 yearclasses, average weight at age from surveys in recent years did not display persistent trends (Fig. 15).

## Estimation of Stock Parameters

## Calibration of VPA

The VPA used quarterly catch at age, $C_{a, t}$, for ages $a=0,1,2 \ldots 8$, and time $t=1969.0$, $1969.25,1969.5,1969.75,1970.0 \ldots 1996.75$, where $t$ represents the beginning of the time interval during which the catch was taken. The VPA was calibrated to bottom trawl survey abundance indices, $I_{s, a, t}$, for

$$
\begin{aligned}
& s=\text { DFO spring, ages } a=1,2,3 \ldots 8, \text { time } t=1986.16,1987.16 \ldots 1998.0 \\
& s=\text { NMFS spring (Yankee } 36) \text {, ages } a=1,2,3 \ldots 8, \text { time } t=1969.29,1970.29,1971.29, \\
& 1972.29,1982.29,1983.29 \ldots 1997.29 \\
& s=\text { NMFS spring (Yankee 41), ages } a=1,2,3 \ldots 8, \text { time } t=1973.29,1974.29 \ldots 1981.29 \\
& s=\text { NMFS fall, ages } a=0,1,2 \ldots 5, \text { time } t=1969.69,1970.69 \ldots 1997.69
\end{aligned}
$$

Since forecast projections were required for the entire year 1998, the DFO spring survey in 1998 was designated as occurring at time 1998.0 instead of 1998.16. The NMFS fall survey captures young of the year and that information is included as 0 group, but older haddock appear less available during this season. Survey indices for older ages where catches were sparse and there were frequent occurrences of zero catches were not included. Zero observations for abundance indices were treated as missing data as the logarithm of zero is not defined. During years when discarding was high, survey information was used along with interviews to obtain estimates of the USA catch. This lack of complete independence between catch and survey data does not influence population estimates but may deflate variance estimates marginally.

The adaptive framework, ADAPT, (Gavaris 1988) was used to calibrate the sequential population analysis with the research survey abundance trend results. The model formulation employed assumed that the random error in the catch at age was negligible. The errors in the abundance indices were assumed independent and identically distributed after taking natural logarithms of the values. The annual natural mortality rate, $M$, was assumed constant and equal to 0.2 . A model formulation using as parameters the natural logarithm of population abundance at the beginning of the year was considered because of close to linear behavior for such a parameterization (Gavaris 1993). The following model parameters were defined:
$\sigma_{a, t^{\prime}}=\ln$ population abundance
for ages $a=1,2, \ldots 8$ at time $t^{\prime}=1998.0$,
$k_{s, a}=\ln$ calibration constants
for each abundance index source, $s$, and relevant ages, $a$.
A solution for the parameters was obtained by minimizing the sum of squared differences between the natural logarithm observed abundance indices and the natural logarithm population abundance adjusted for catchability by the calibration constants. The objective function for minimization was defined as

$$
\underset{s, a, t}{\Psi}(\hat{\theta}, \hat{\kappa})=\sum_{s, a, t}\left(\psi_{s, a, t}(\hat{\theta}, \hat{\kappa})\right)^{2}=\sum_{s, a, t}\left(\ln I_{s, a, t}-\left(\hat{\kappa}_{s, a}+\ln N_{a, t}(\hat{\theta})\right)\right)^{2}
$$

For convenience, the population abundance $N_{a, t}(\hat{\theta})$ is abbreviated by $N_{a, t}$. At time $t^{\prime}$, the population abundance was obtained directly from the parameter estimates, $N_{a, r^{\prime}}=e^{\hat{\theta}_{a,,^{\prime}}}$. For
all other times, the population abundance was computed using the virtual population analysis algorithm, which incorporates the common exponential decay model
$N_{a+\Delta t, t+\Delta t}=N_{a, t} t^{-\left(F_{a, t}+M_{a}\right) \Delta t}$.

Year was used as the unit of time, therefore ages were expressed as years and the fishing and natural mortality rates were annual instantaneous rates. The fishing mortality rate, $F_{a, i}$, exerted during the time interval $t$ to $t+\Delta t$, was obtained by solving the catch equation
$C_{a, t}=\frac{F_{a, t} \Delta t N_{a, t}\left(1-e^{-\left(F_{a, t}+M_{a}\right) \Delta t}\right)}{\left(F_{a, t}+M_{a}\right) \Delta t}$
using a Newton-Raphson algorithm. The fishing mortality rate for the oldest age in the last time interval of each year was assumed equal to the weighted average for ages fully recruited to the fishery during that time interval
$F_{8, t}=\sum_{a=4}^{7} N_{a, t} F_{a, t} / \sum_{a=4}^{7} N_{a, t}$.
The covariance matrix of the parameters was estimated using the common linear approximation (Kennedy and Gentle 1980 p.476)
$\operatorname{Cov}(\hat{\theta}, \hat{\kappa})=\hat{\sigma}^{2}\left[J^{T}(\hat{\theta}, \hat{\kappa}) J(\hat{\theta}, \hat{\kappa})\right]^{-1}$
where $\hat{\sigma}^{2}$ is the mean square residual and $J(\hat{\theta}, \hat{\kappa})$ is the Jacobian matrix. The bias of the parameters was estimated using Box's (1971) approximation, which assumes that the errors are normally distributed
$\operatorname{Bias}(\hat{\theta}, \hat{\kappa})=\frac{-\hat{\sigma}^{2}}{2}\left(\sum_{s, a, t} J_{s, a, t}(\hat{\theta}, \hat{\kappa}) J_{s, a, t}^{T}(\hat{\theta}, \hat{\kappa})\right)^{-1}\left(\sum_{s, a, t} J_{s, a, t}(\hat{\theta}, \hat{\kappa})\right) \operatorname{tr}\left[\left(\sum_{s, a, t} J_{s, a, t}(\hat{\theta}, \hat{\kappa}) J_{s, a, t}^{T}(\hat{\theta}, \hat{\kappa})\right)^{-1} H_{s, a, t}(\hat{\theta}, \hat{\kappa})\right]$
where $J_{s, a, t}(\hat{\theta}, \hat{\kappa})$ are vectors of the first derivatives for each $\psi_{s, a, t}(\hat{\theta}, \hat{\kappa})$ and $H_{s, a, t}(\hat{\theta}, \hat{\kappa})$ are the Hessian matrices of second derivatives for each $\psi_{s, a, t}(\hat{\theta}, \hat{\kappa})$.

Population quantities of interest for management advice are functions of the estimated parameters. The variance and bias of an arbitrary quantity, $\hat{\eta}=g(\hat{\theta}, \hat{\kappa})$ where $g$ is the transformation function, were estimated using the methods described in Ratkowsky (1983)

$$
\begin{aligned}
& \operatorname{Var}(\hat{\eta})=\operatorname{tr}\left[G G^{T} \operatorname{cov}(\hat{\theta}, \hat{\kappa}) \mid\right. \\
& \operatorname{Bias}(\hat{\eta})=G^{T} \operatorname{Bias}(\hat{\theta}, \hat{\kappa})+\operatorname{tr}|W \operatorname{cov}(\hat{\theta}, \hat{\kappa})| / 2
\end{aligned}
$$

where $G$ is the vector of first derivatives of $g$ with respect to parameters and $W$ is the matrix of second derivatives of $g$ with respect to parameters.

Confidence statements for management quantities of interest were obtained using the non-parametric bias corrected percentile bootstrap method. Non-parametric bootstrap techniques offer the advantage of not making any assumptions about the error distribution. The bootstrap samples are used to calculate the bootstrap replicate estimates, $\hat{\eta}^{b}$, of the quantity of interest.

The percentile method (Efron 1982) is a simple and direct way of forming an empirical cumulative frequency distribution. The probability that $\hat{\eta}$ is less than or equal to some value is defined as the proportion of bootstrap replicates, $\hat{\eta}^{b}$, less than or equal to that value.
$\hat{\Omega}(x)=\operatorname{Prob}\{\hat{\eta} \leq x\}=\frac{\#\left\{\hat{\eta}^{b} \leq x\right\}}{B}$,
where $B$ is the total number of bootstrap replicates. For conceptual and graphing purposes, it is convenient to consider the empirical cumulative frequency distribution as the set of paired values $\left(\alpha, \vec{\eta}^{b}\right)$, where $\vec{\eta}^{b}$ are the ordered bootstrap replicates and $\alpha$ are the respective probability levels equal to $\frac{1}{B}, \frac{2}{B}, \frac{3}{B}, \cdots, \frac{B}{B}$.

Frequently, the median of the bootstrap percentile density function does not equal the estimate obtained with the original data sample. The bias-corrected percentile method (Efron 1982) makes an adjustment for this type of bias. The bias-corrected percentile method can be thought of as an algorithm to replace the $\vec{\eta}^{b}$ in the paired values $\left(\alpha, \vec{\eta}^{b}\right)$ with the bias adjusted quantity $\vec{\eta}_{B C}^{b}$. The notation $\hat{\Omega}^{-1}()$ or $\Phi^{-1}()$ is used to represent the inverse distribution function, i.e. the critical value corresponding to the specified probability level. For each $\alpha$ in the paired values $\left(\alpha, \bar{\eta}^{b}\right)$, calculate the bias adjusted quantity $\vec{\eta}_{B C}^{b}$.
$\vec{\eta}_{B C}^{b}=\hat{\Omega}^{-1}\left(\Phi\left(2 z_{0}+z_{\alpha}\right)\right)$.
Here, $\Phi$ is the cumulative distribution function of a standard normal variate, $z_{\alpha}=\Phi^{-1}(\alpha)$ and $z_{0}=\Phi^{-1}(\hat{\Omega}(\hat{\eta}))$. The term $z_{0}$ achieves the bias adjustment. If the median of the bootstrap density function is equal to $\hat{\eta}$, then $\hat{\Omega}(\hat{\eta})$ will be $0.5, z_{0}$ will be zero, and $\vec{\eta}_{B C}^{b}$ will equal $\vec{\eta}^{b}$ (i.e. no bias adjustment). Note that computations are not carried out for $\alpha=\frac{B}{B}$ because $z_{\alpha}=\Phi^{-1}(\alpha=1)$ is not defined.

In previous assessments the NMFS spring survey has been treated as a continuous index for the entire time series. As noted earlier, the NMFS spring survey used a Yankee 41 trawl during 1973-81 and a Yankee 36 in other years. The limited comparative fishing done between these nets caught few haddock thereby not permiting computation of a conversion factor, but analyses by Hayes and Buxton (1992) suggested that differences were not significant. We examined this issue again by treating the NMFS spring survey data for 1973-81 as a distinct survey. We refer to this as the split model and we refer to the approach used in last year's assessment as the base.

Box plots of the age specific annual catchabilities (ln and linear) showed that those for the Yankee 41 net were higher than for the Yankee 36 net (Fig. 16). A graph of the estimated ln catchabilities showed little or no overlap in the one standard error range for most ages (Fig. 17). Although the overall mean square residual for the split model improved marginally, this is hardly detectable in the residual time series (Fig. 18).

We concluded that the evidence for a difference in catchability between the Yankee 36 and the Yankee 41 was compelling and the assessment was conducted in a manner which takes this into account. Nevertheless, the impact on terminal year population abundance was not great as the greater number of observations for the Yankee 36 trawl largely determined the average catchability (Fig. 19).

The population abundance estimates show a large relative error and substantial bias at ages 1 and 2 reflecting the variability in the abundance indices (Table 15). The average magnitude of residuals is large and though several large residuals can be identified, the respective observations do not appear influential and should not impact parameter estimates of current abundance (Fig. 20). The table below shows the average of squared residuals for each series.

| Survey | Age |  |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 |  |
| Can. Spring | - | 0.68 | 0.91 | 0.28 | 0.21 | 0.48 | 0.60 | 0.99 | 0.64 |  |
| NMFS Spring 36 | - | 0.89 | 0.52 | 0.49 | 0.34 | 0.60 | 0.94 | 1.36 | 0.64 |  |
| NMFS Spring 41 |  | 1.03 | 1.23 | 0.89 | 0.29 | 0.60 | 0.90 | 3.55 | 2.08 |  |
| NMFS Fall | 1.33 | 1.78 | 1.18 | 0.60 | 0.47 | 0.67 | - | - | - |  |

Myers and Cadigan (1995) reported that correlated errors among ages within a survey can be sufficiently large to produce model mis-specification biases in estimates of population parameters from standard assessment methods. Their simulation however, showed that maximum likelihood estimators from models which ignored correlation performed similar to those from models which incorporated correlation when the correlated errors were small, e.g. $\rho=0.15$. An estimate of the correlation among ages within a survey was computed using the standard sample estimator for the coefficient of linear correlation where the pairs of observations were the residuals from each abundance index source: $\left(e_{i, t}, e_{j, t}\right)$ for all ages $i \neq j$ and all times $t$. For the four survey sources used in this assessment, the correlation was found to be small; DFO spring survey $\hat{\rho}=0.08$, NMFS spring survey with the Yankee 36 net $\hat{\rho}=0.10$, NMFS spring survey with the Yankee 41 net $\hat{\rho}=0.18$ and NMFS fall survey $\hat{\rho}=0.19$. Accordingly, no further corrective measures were taken to account for bias from this type of model mis-specification for this stock.

Results from analyses assuming a multinomial error structure for the objective functions were examined (Appendix C). The multinomial error model, which places less reliance on small catches, resulted in a $40 \%$ lower estimate of population abundance in 1998. The model based on the logarithm of survey observations assumes that small catches are as reliable as large catches on the logarithmic scale. The results from the multinomial model were considered illustrative and further work is required to investigate the robustness of such alternate error structure models. Diagnostics to help discriminate between alternative models need to be investigated. Model misspecification is an additional source of uncertainty that has not been incorporated in the subsequent risk analyses.

A further variation in model formulation which was explored retained the standard error assumption about the logarithm of survey observations but did not make any assumptions about the fishing mortality on age 8 being equal to the average on ages 4 to 7 by estimating the abundance of all cohorts directly as parameters (Appendix C). The population abundance trends from this model, for constant catchability at ages 3 to 8 in the Canadian spring survey, were similar to those from the standard model therefore the assumption on fishing mortality for age 8 was considered appropriate.

## Retrospective Analysis of Calibrated VPA

Results from assessments for several other stocks have identified a discrepancy between past estimates of stock status and current estimates using additional data (retrospective pattern). Results for this stock indicate that this assessment does not suffer from a retrospective pattern (Fig. 21). Fig. 21 tracks successive estimates of year-class abundance at age and shows that estimates are fairly stable although there is sometimes a substantial change after the first estimate of a year-class when more data becomes available, as evidenced for the 1992 and 1996 yearclasses. There were no trends of concern in the $3+$ biomass pattern and the $4+F$ when weighted by population numbers (Fig. 22). The unweighted $4+\mathrm{F}$ pattern was not as stable as the weighted Fs, the increased variability probably due to poorly estimated Fs on small year-classes.

Projections based on the previous assessment (Gavaris and Van Eeckhaute 1997) indicated that with a Canadian quota of 3,200 t in 1997, biomass for ages 3 and older was expected to increase by about $8 \%$ and there was a very high chance that fishing mortality would be lower than $\mathrm{F}_{0.1}$. The realised biomass growth between 1997 and 1998 was about $10 \%$ and the realised fishing mortality rate in 1997 was about half of $\mathrm{F}_{0.1}$ with a Canadian catch of $2,739 \mathrm{t}$.

## Mortality Estimates from Surveys

Survey abundance estimates for this stock are very variable, therefore age specific mortality derived from them fluctuate widely. Age aggregated total mortality for adult fish has somewhat lower variablility but also fluctuates considerably. Ages 3-8 from the DFO and NMFS spring surveys and ages 2-5 from the NMFS fall survey were used to calculate mortality estimates. Applying a 3 year running median smoother to the mortality estimates helps to identify some general patterns (Fig. 23). Total mortality was relatively stable at just below 1 until about the mid 1980s when it decreased. During the late 1980s and early 1990s it increased to above 1 before declining to the current lower level below 0.5 .

## Stock Status

The results from the standard lognormal model formulation were considered appropriate on which to base the status of the stock. For each cohort, the terminal population abundance estimates from ADAPT were adjusted for bias and used to construct the history of stock status. Gavaris and Van Eeckhaute (1997) considered that this approach for bias adjustment, in the absence of an unbiased point estimator with optimal statistical properties, was preferable to using the biased point estimates (Tables 16-17). The weights at age from the DFO spring survey, $\overline{W_{a}}$, (Table 18) were used to calculate beginning of year population biomass (Table 19)

A weight of 2.4 kg , which was midway between the age 6 and 8 weight for that cohort, was used for age 7 in 1995. For 1969-85, the 1986-95 average weight at each age was used. The weight at age for ages $9+$ haddock was used for age 9 .

Total population biomass (ages 1+) has steadily increased from near historic low levels of $12,171 \mathrm{t}$ in 1993 to $28,809 \mathrm{t}$ in 1998 (Fig. 24). The recent increase, due principally to the 1992 yearclass, but also supported by the 1991 and 1993 yearclasses ( Fig .25 ), was enhanced by increased survivorship of young haddock resulting from reduced capture of small fish in the fisheries. The continuing increase is being sustained by the 1996 year-class. The biomass trend for ages 3 and older is similar. The strength of the 1996 yearclass was estimated to be about 13 million, comparable to the 1983, 1985, 1987 and 1992 year-classes, while those during 1988-90 were less than 3 million. The 1991 and 1993 yearclasses were estimated at about 7 and 10 million respectively while the incoming 1995 and 1997 yearclasses appear to be relatively weak at about 6 and 3 million respectively. Population biomass during the late 1970s and early 1980s was considerably higher, ranging to almost $50,000 \mathrm{t}$ due to recruitment of the strong 1975 and 1978 year-classes whose abundance was estimated at about 50 million. However, biomass declined rapidly in the early 1980s as subsequent recruitment was poor and these two yearclasses were fished intensely at a young age.

Exploitation rate for ages 4 and older has consistently been below the $\mathrm{F}_{0.1}$ target of 20\% $\left(\mathrm{F}_{0.1}=0.25\right)$ since 1995 (Fig. 26). Historically, exploitation rate has generally exceeded $\mathrm{F}_{0.1}$ and showed a marked increased between 1989 and 1992 to almost $50 \%$, the highest level observed. Reduced fishing mortality in recent years has resulted in increased survival of incoming yearclasses (Fig. 27). The number of haddock surviving to age 6 of the 1992 yearclass was about twice that of the equally abundant 1983 yearclass, and about the same as that of the 1975 or 1978 yearclasses which were more than 3 times as abundant. Examination of the percent age composition and the absolute abundance at age in 1998 compared to the averages over earlier time periods suggests that all ages are well represented in the population proportionally but absolute abundance at recruiting ages is considerably lower that the 1931-55 average (Fig 28).

Gains in fishable biomass may be partitioned into those associated with somatic growth of haddock which have previously recruited to the fishery and those associated with new recruitment to the fishery. We used age 2 as a convenient age of first recruitment to the fishery. It is evident (Fig. 29) that in most years since 1969, the bulk of the biomass gain was attributed to somatic growth, but in years of strong recruitment, about half the biomass gain was due to incoming recruits. Surplus production is defined as the gains in fishable biomass which are in excess of the needs to offset losses from natural mortality. When the fishery yield is less than the surplus production, there is a net increase in the population biomass. Since 1993, the surplus production has either exceeded or has been about equal to the yield and consequently, the population has sustained a net biomass growth (Fig 30).

## Fishery Reference Points

## Yield per Recruit

The yield per recruit formulation used previously was revised to assume complete fishing on the last age group (i.e. + group). A partial recruitment of $0, .05, .5,1,1 \ldots 1$, was assumed for ages $1,2,3,4,5 \ldots 16+$. Fishery weights at age for ages 1 to 16 were estimated from the 1985-97 fishery to reflect recent fishery conditions. These weights were converted to length using $L=(W / 0.0000158)^{1 / 2.91612}$ (the inverse of the relationship used to convert fishery lengths to
weights) to enable a Von Bertalanffy fit of average length and age data. The predicted lengths at age from the resulting relationship, $L_{t}=76.22\left(1-e^{-0.1925(t-(-2.509))}\right)$, were then converted back to weight using $W=0.0000158 L^{291612}$ (Fig. 31). Using these predicted weights, the estimated $F_{0.1}$ was 0.25 with a yield of 0.75 kg per age 1 recruit (Fig. 32).

## Stock and Recruitment

Exact age composition of the catch from unit areas 5 Zj and 5 Zm was not available prior to 1969 . Although total catch of haddock from unit areas 5 Zj and 5 Zm is considered reliable, only an approximate age composition of the catch could be obtained by prorating the 5 Z catch at age with the $5 \mathrm{Zjm}: 5 \mathrm{Z}$ landings ratio. Using these data however, it was possible to reconstruct an illustrative population analysis for the period between 1930 and 1955 which is suitable for comparing productivity. The results indicated that the current total biomass was less than a third of the average sustained over those two decades (Fig. 33).

The pattern of recruitment against adult biomass indicates that the chance of observing a strong year-class is significantly worse for biomass below about $40,000 \mathrm{t}$. Since 1969 , only the 1975 and 1978 year-classes have been near the average abundance of year-classes observed during that historic period (Fig. 34).

## Prognosis

Yield projections were done using the bias adjusted 1998 beginning of year population abundance estimates. The abundance of the 1998 year-class was assumed to be 6 million at age 0 . Partial recruitment to the fishery for ages 1,2 and 3 , fishery weights at age and beginning of year population weights at age were averaged over the previous 4 years for use in the 1998 forecasts (Table 20).

Combined Canada/USA projected yield at $\mathrm{F}_{0.1}=0.25$ in 1998 would be about $6,000 \mathrm{t}$ (Table 20, Fig. 35). If fished at $\mathrm{F}_{0.1}$ in 1998, the biomass for ages 3 and older is projected to increase from 22,726 t to $28,012 \mathrm{t}$ at the beginning of 1999. The 1992 and 1993 year-classes would comprise about $40 \%$ of the $3+$ biomass and $60 \%$ of the forecast yield. With the recruitment of the 1996 year-class, the adult biomass can be expected to increase further if exploitation rate is maintained at the current moderate level. The 1996 yearclass was estimated to contribute almost $30 \%$ to the 1999 age $3+$ biomass.

Uncertainty about year-class abundance generates uncertainty in forecast results. This uncertainty was expressed as risk of achieving reference targets. For example, a combined Canada/USA catch of $3,000 \mathrm{t}$ in 1998, about what was caught in 1997, results in a negligible risk that fishing mortality rate will exceed $\mathrm{F}_{0.1}$ and that the biomass for ages 3 and older will decrease (Fig. 36). That same yield gives risks of $10 \%$ and $35 \%$ that an increase in biomass of $10 \%$ and $20 \%$ respectively would not be achieved. These uncertainty calculations do not include variations in weight at age, partial recruitment to the fishery and natural mortality, or systematic errors in data reporting and model mismatch. Therefore, overall uncertainty would be greater but these results provide rough guidelines. Increasing the number of age groups contributing to the yield should lead to greater precision in the advice, reduced fluctuations in biomass caused by recruitment variability, and result in more stable yield between years. A larger spawning biomass could enhance recruitment by capitalizing on the opportunities for greater egg and larval survival
when environmental conditions are favorable.

The Georges Bank ecosystem is complex with numerous species interactions. Further, species adapt to fluctuations in abundance of both their prey and predators. These interactions were modelled by a constant natural mortality and there were no indications that this assumption was severely violated. Currently available information does not permit more complex models to be employed.

Environmental conditions on Georges Bank have varied but have not displayed extreme deviations in recent years. Although environmental conditions are thought to influence fisheries processes, convincing relationships with quantities such as recruitment, survival rates and fish catchability have not been established for this stock.

## Comparison of Results for Canadian and USA Management Units

The existence of two centers of haddock aggregation on Georges Bank with distinct spawning components has long been recognized (Fig. 37). One aggregation spawns on the Northeast Peak in the spring and migrates to the bank slopes on the Northeast Edge and Peak as the waters warm in the summer. The other component spawns around the Nantucket Shoals in the spring and migrates to the bank slopes around the Great South Channel as the waters warm in the summer. The former is referred to as the Eastern component and the latter as the Western component. There is evidence for limited but poorly quantified exchange between the two components. Haddock from the Western component are characterized by faster growth.

The USA conducts fisheries for haddock on both the Western and Eastern components. The USA defines a management unit encompassing both Eastern and Western components of the Georges Bank haddock resource, specifically NAFO Division 5Z (small amounts of haddock caught in NAFO Subarea 6 are included). Canada conducts fisheries for haddock on the Eastern component only and is concerned with regulatory measures which could be applied to it in order to achieve benefits. Accordingly, Canada defines unit areas 5 Zj and 5 Zm as a management unit.

Between 1969 and 1985, catches from 5Zjm averaged about $56 \%$ of the total catches from 5 Z , ranging between $44 \%$ and $67 \%$ (Fig. 38). During 1985 to 1996, catches from 5 Zjm were consistently above $83 \%$ of the total catches from 5 Z, averaging about $88 \%$, but in 1997 that percentage dropped to $67 \%$.

Over this period, the total biomass for the two management units showed a similar pattern (Fig. 39). The biomass in 5Z declined from 94,000 $t$ in 1980 to $17,000 \mathrm{t}$ in 1993 and has since increased to $40,000 \mathrm{t}$ in 1998. In 5Zjm, the biomass declined from 49,000 t in 1980 to $12,000 \mathrm{t}$ in 1993 and has reached about $29,000 \mathrm{t}$ in 1998. Since 1985, the biomass in 5 Zjm has consistently been over about $70 \%$ of the total 5 Z biomass.

The 1975 and 1978 year-classes were the most abundant on Georges Bank since 1969 (Fig. 40). The abundance at age 1 for these two year-classes was about 105 million and 83 million respectively for all of 5 Z and about 54 million and 52 million in 5 Zjm . Subsequent yearclasses have been considerably weaker with the strongest among them being the 1983, 1985, 1987, 1992 and 1996 year-classes. The abundance at age 1 of these year-classes was $17,15,17$, 17 and 13 million respectively for all of $5 Z$ and $15,13,15,14$ and 13 million in $5 Z \mathrm{jm}$. The 1968
through 1980 year-classes in 5 Zjm averaged about $60 \%$ of the abundance for all of 5 Z while those after 1980, with the exception of 1994 and 1995, have generally comprised over $70 \%$ of the total for 5 Z , averaging about $80 \%$.

The fishing mortality rates in 5 Zjm and in all of 5 Z showed a decline between the early and mid 1970s, followed by an increase until 1980 (Fig. 41). Between 1980 and 1990, the fishing mortality rate fluctuated between about 0.3 and 0.4. It then increased rapidly to about 0.45 in 5 Z and 0.65 in 5 Zjm by 1993 and subsequently declined to below 0.2 in both 5 Zjm and 5 Z by 1995 .

Between 1969 and 1985, the contribution to production by the Eastern and Western components was roughly equivalent, and both components appeared to have been exploited to the same degree. Since 1985 however, over $80 \%$ of the production on Georges Bank was attributed to the Eastern component. By 1997, the Eastern component increased to almost half of its biomass level observed during the late 1970s and early 1980s while Georges Bank as a whole only increased to about a third of its respective biomass level. There is some evidence that the production from the Western component is improving over the last few years. The 1994 and 1995 year-classes were estimated to be about equally represented in both components, however the 1996 year-class is poorly represented in the Western component.

## Acknowledgements

We are grateful to staff at the Northeast Fisheries Science Center in Woods Hole, MA, and in particular R.W. Brown for providing information from the USA fishery and the NMFS surveys and for their assistance in interpretation of that data. G. Donaldson and M. Showell provided samples from the Canadian fishery. We thank members of the fishing industry who spent time to discuss their experiences in the fishery. We are very appreciative of the work contributed by Vivian Haist which explores alternative model formulations.

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Table 1. Nominal catches ( t ) of haddock from unit areas 5 Zjm . For "others" it was assumed that $40 \%$ of the total 5 Z catch was in 5Zjm.

| Year | Canada | USA | Others | Total |
| :---: | :---: | :---: | :---: | :---: |
| 1969 | 3941 | 6622 | 695 | 11258 |
| 1970 | 1970 | 3153 | 357 | 5480 |
| 1971 | 1610 | 3534 | 770 | 5914 |
| 1972 | 609 | 1551 | 502 | 2662 |
| 1973 | 1565 | 1396 | 396 | 3357 |
| 1974 | 462 | 955 | 573 | $2747^{1}$ |
| 1975 | 1353 | 1705 | 29 | 3087 |
| 1976 | 1355 | 973 | 24 | 2352 |
| 1977 | 2871 | 2429 | 0 | $8266{ }^{1}$ |
| 1978 | 9968 | 4724 | 0 | $16223{ }^{1}$ |
| 1979 | 5080 | 5211 | 0 | 10291 |
| 1980 | 10017 | 5615 | 0 | $23189{ }^{1}$ |
| 1981 | 5658 | 9077 | 0 | 14735 |
| 1982 | 4872 | 6280 | 0 | 11152 |
| 1983 | 3208 | 4454 | 0 | 7662 |
| 1984 | 1463 | 5121 | 0 | 6583 |
| 1985 | 3484 | 1683 | 0 | 5167 |
| 1986 | 3415 | 2200 | 0 | 5615 |
| 1987 | 4703 | 1418 | 0 | 6121 |
| 1988 | $4046{ }^{2}$ | 1693 | 0 | 5739 |
| 1989 | 3059 | 787 | 0 | 3846 |
| 1990 | 3340 | 1189 | 0 | 4529 |
| 1991 | 5446 | 931 | 0 | 6377 |
| 1992 | 4061 | 1629 | 0 | 5690 |
| 1993 | 3727 | 421 | 0 | 4148 |
| 1994 | 2411 | 33 | 0 | $2702{ }^{3}$ |
| 1995 | 2064 | 22 | 0 | $2111^{3}$ |
| 1996 | 3643 | 36 | 0 | $3720^{3}$ |
| 1997 | 2739 | 48 | 0 | $2850{ }^{3}$ |

${ }^{1}$ Includes 757t, 2966t, 1531t and 7557t in 1974, 1977, 1978 and 1980 respectively for USA discards.
${ }^{2} 1895 \mathrm{t}$ excluded because of suspected area misreporting.
${ }^{3}$ Includes $258 \mathrm{t}, 25 \mathrm{t}, 41$ and 63 in 1994,1995, 1996 and 1997 respectively for USA discards.

Table 2. Regulatory measures implemented for the 5 Z and 5 Zjm fishery management units by the USA and Canada, respectively, from 1977, when jurisdiction was extended to 200 miles for coastal states, to the present.

|  | USA | Canada |
| :---: | :---: | :---: |
| 1977-82 | Mesh size of $51 / 8^{\prime \prime}(140 \mathrm{~mm})$, seasonal spawning closures, quotas and trip limits. |  |
| 1982-85 | All catch controls eliminated, retained closed area and mesh size regulations, implemented minimum landings size ( 43 cm ). | First 5Ze assessment in 1983. |
| 1984 Oct. | Implementation of the 'Hague' line . |  |
| 1985 | $51 / 2^{\prime \prime}$ mesh size,. <br> Areas 1 and 2 closed during February-May. |  |
| 1989 |  | Combined cod-haddock-pollock quota for 4X5Zc |
| 1990 |  | $\mathbf{5 Z j m}$ adopted as management unit. <br> For MG $<65 \mathrm{ft}$. - trip limits with a $30 \%$ by-catch of haddock to a maximum of 8 trips of $35,000 \mathrm{lbs}$ per trip between June 1 and Oct. 31 and 130 mm square mesh required. <br> Fixed gear required to use large hooks until June |
| 1991 | Established overfishing definitions for haddock. | MG < 65 ft similar to 1990 but mesh size increased to 145 mm diamond. |
| 1992 |  | Introduction of ITQs and dockside monitoring. |
| 1993 | Area 2 closure in effect from Jan 1-June30. | OT fishery permitted to operate in Jan. and Feb. Increase in use square mesh. |
| 1994 | Jan.: Expanded Area 2 closure to include June and increased extent of area. <br> Area 1 closure not in effect. <br> 500 lb trip limit. <br> Catch data obtained from mandatory log books combined with dealer reports (replaces interview system). <br> May: 6" mesh restriction. <br> Dec.: Area 1,2 closed year-round. | Spawning closure extended to Jan. 1 to May 31. <br> Fixed gear vessels must choose between 5 Z or 4 X for the period of June to September. <br> Small fish protocol. <br> Increased at sea monitoring. OT $>65$ could not begin fishng until July 1. Predominantly square mesh by end of year. |
| 1995 |  | All OT vessels using square mesh. Vessels with a history since 1990 of $25 t$ or more for 3 years of cod, haddock pollock, hake or cusk combined can participate in 5Z fishery. <br> ITQ vessel require at least 2 t of cod and 8 t of haddock quota to fish Georges. |
| 1996 | July: Additional Days-at-Sea restrictions, trip limit raised to 1000 lbs . | Fixed gear history requirement dropped. |
| 1997 | May: Additional scheduled Days-at-sea restrictions. <br> September: Trip limit raised to $1000 \mathrm{lbs} /$ day, maximum of $10,000 \mathrm{lbs} /$ trip. |  |
| 1998 | Proposed: Trip limit raised to $3000 \mathrm{lbs} /$ day, maximum of $30,000 \mathrm{lbs} /$ trip. |  |

Table 3. Canadian catch ( t ) of haddock in unit areas 5Zjm by gear category and tonnage class for principle gears.

| Year | Otter Trawl |  |  |  |  |  | Longline |  |  | Other | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Side | Stern |  |  |  |  |  |  |  |  |  |
|  |  | 2 | 3 | 4 | 5 | Total | 2 | 3 | Total |  |  |
| 1969 | 777 | 0 | 1 | 225 | 2902 | 3127 | 2 | 21 | 23 | 15 | 3941 |
| 1970 | 575 | 2 | 0 | 133 | 1179 | 1314 | 6 | 72 | 78 | 2 | 1970 |
| 1971 | 501 | 0 | 0 | 16 | 939 | 955 | 18 | 129 | 151 | 3 | 1610 |
| 1972 | 148 | 0 | 0 | 2 | 260 | 263 | 23 | 169 | 195 | 3 | 609 |
| 1973 | 633 | 0 | 0 | 60 | 766 | 826 | 23 | 80 | 105 | 0 | 1565 |
| 1974 | 27 | 0 | 6 | 8 | 332 | 346 | 29 | 59 | 88 | 1 | 462 |
| 1975 | 222 | 0 | 1 | 60 | 963 | 1024 | 25 | 81 | 107 | 0 | 1353 |
| 1976 | 217 | 0 | 2 | 59 | 905 | 967 | 48 | 108 | 156 | 15 | 1355 |
| 1977 | 370 | 92 | 243 | 18 | 2025 | 2378 | 43 | 51 | 94 | 28 | 2871 |
| 1978 | 2456 | 237 | 812 | 351 | 5639 | 7039 | 121 | 47 | 169 | 305 | 9968 |
| 1979 | 1622 | 136 | 858 | 627 | 1564 | 3185 | 190 | 80 | 271 | 2 | 5080 |
| 1980 | 1444 | 354 | 359 | 950 | 6254 | 7917 | 129 | 51 | 587 | 69 | 10017 |
| 1981 | 478 | 448 | 629 | 737 | 2344 | 4159 | 331 | 99 | 1019 | 2 | 5658 |
| 1982 | 115 | 189 | 318 | 187 | 3341 | 4045 | 497 | 187 | 712 | 0 | 4872 |
| 1983 | 106 | 615 | 431 | 107 | 1130 | 2283 | 593 | 195 | 815 | 4 | 3208 |
| 1984 | 5 | 180 | 269 | 21 | 149 | 620 | 614 | 192 | 835 | 3 | 1463 |
| 1985 | 72 | 840 | 1401 | 155 | 348 | 2745 | 562 | 33 | 626 | 41 | 3484 |
| 1986 | 51 | 829 | 1378 | 95 | 432 | 2734 | 475 | 98 | 594 | 35 | 3415 |
| 1987 | 48 | 782 | 1448 | 49 | 1241 | 3521 | 854 | 113 | 1046 | 89 | 4703 |
| $1988{ }^{1}$ | 72 | 1091 | 1456 | 186 | 398 | 3183 | 428 | 200 | 695 | 97 | 4046 |
| 1989 | 0 | 489 | 573 | 376 | 536 | 1976 | 713 | 175 | 977 | 106 | 3059 |
| 1990 | 0 | 928 | 890 | 116 | 471 | 2411 | 623 | 173 | 853 | 76 | 3340 |
| 1991 | 0 | 1610 | 1647 | 81 | 679 | 4018 | 900 | 271 | 1309 | 119 | 5446 |
| 1992 | 0 | 797 | 1084 | 56 | 645 | 2583 | 984 | 245 | 1384 | 90 | 4061 |
| 1993 | 0 | 535 | 1179 | 67 | 699 | 2490 | 794 | 156 | 1144 | 94 | 3727 |
| 1994 | 0 | 495 | 911 | 79 | 112 | 1597 | 498 | 47 | 714 | 100 | 2411 |
| 1995 | 0 | 510 | 896 | 14 | 214 | 1647 | 261 | 69 | 389 | 28 | 2064 |
| 1996 | 1 | 836 | 1405 | 166 | 270 | 2689 | 548 | 107 | 932 | 21 | 3643 |
| 1997 | 0 | 681 | 1123 | 91 | 96 | 1991 | 494 | 116 | 713 | 36 | 2739 |

${ }^{1}$ Catches of 26t, 776t, 1091 t and 2 t for side ofter trawlers and stern otter trawlers tonnage classes 2,3 and 5 respectively were excluded because
of suspected area misreporting.
Table 4. Monthly catch (t) of haddock by Canada in unit areas 5Zjm.

| Year | Jan | Feb | Mar | Apr | May | Jun | Jul | Aug | Sep | Oct | Nov | Dec | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1969 | 105 | 74 | 6 | 291 | 588 | 691 | 559 | 580 | 551 | 360 | 102 | 34 | 3941 |
| 1970 | 2 | 105 | 0 | 1 | 574 | 345 | 103 | 456 | 242 | 103 | 26 | 12 | 1970 |
| 1971 | 0 | 9 | 1 | 0 | 400 | 132 | 283 | 278 | 97 | 246 | 141 | 21 | 1610 |
| 1972 | 0 | 119 | 2 | 0 | 2 | 111 | 84 | 116 | 98 | 68 | 7 | 2 | 609 |
| 1973 | 4 | 10 | 0 | 0 | 0 | 184 | 198 | 572 | 339 | 232 | 22 | 4 | 1565 |
| 1974 | 19 | 0 | 1 | 0 | 0 | 58 | 63 | 53 | 96 | 61 | 92 | 19 | 462 |
| 1975 | 4 | 14 | 0 | 0 | 0 | 166 | 256 | 482 | 100 | 166 | 118 | 45 | 1353 |
| 1976 | 0 | 7 | 62 | 68 | 60 | 587 | 152 | 190 | 186 | 26 | 9 | 7 | 1355 |
| 1977 | 102 | 177 | 7 | 0 | 23 | 519 | 1059 | 835 | 13 | 59 | 56 | 22 | 2871 |
| 1978 | 104 | 932 | 44 | 22 | 21 | 319 | 405 | 85 | 642 | 5433 | 1962 | 0 | 9968 |
| 1979 | 123 | 898 | 400 | 175 | 69 | 1393 | 885 | 396 | 406 | 261 | 53 | 22 | 5080 |
| 1980 | 38 | 134 | 14 | 29 | 223 | 2956 | 2300 | 965 | 1411 | 1668 | 104 | 176 | 10017 |
| 1981 | 38 | 481 | 568 | 4 | 254 | 1357 | 1241 | 726 | 292 | 82 | 378 | 239 | 5658 |
| 1982 | 129 | 309 | 1 | 11 | 46 | 1060 | 769 | 682 | 585 | 837 | 398 | 44 | 4872 |
| 1983 | 32 | 67 | 29 | 47 | 60 | 1288 | 387 | 483 | 526 | 195 | 88 | 6 | 3208 |
| 1984 | 3 | 5 | 81 | 88 | 73 | 433 | 219 | 254 | 211 | 71 | 25 | 0 | 1463 |
| 1985 | 1 | 11 | 33 | 99 | 26 | 354 | 392 | 1103 | 718 | 594 | 61 | 93 | 3484 |
| 1986 | 11 | 28 | 79 | 99 | 40 | 1339 | 1059 | 369 | 233 | 139 | 12 | 8 | 3415 |
| 1987 | 24 | 26 | 138 | 70 | 12 | 1762 | 1383 | 665 | 405 | 107 | 97 | 14 | 4703 |
| $1988{ }^{1}$ | 39 | 123 | 67 | 79 | 15 | 1816 | 1360 | 315 | 130 | 65 | 13 | 24 | 4046 |
| 1989 | 32 | 94 | 48 | 7 | 20 | 1398 | 356 | 566 | 141 | 272 | 108 | 18 | 3059 |
| 1990 | 35 | 14 | 50 | 0 | 7 | 1179 | 668 | 678 | 469 | 199 | 18 | 22 | 3340 |
| 1991 | 144 | 166 | 49 | 26 | 21 | 1928 | 1004 | 705 | 566 | 576 | 123 | 137 | 5446 |
| 1992 | 118 | 205 | 97 | 152 | 36 | 1381 | 619 | 414 | 398 | 401 | 209 | 28 | 4061 |
| 1993 | 466 | 690 | 96 | 78 | 25 | 723 | 505 | 329 | 202 | 198 | 230 | 185 | 3727 |
| 1994 | 1 | 3 | 1 | 2 | 0 | 398 | 693 | 373 | 375 | 220 | 211 | 134 | 2411 |
| 1995 | 1 | 1 | 1 | 1 | 0 | 762 | 326 | 290 | 281 | 109 | 197 | 96 | 2064 |
| 1996 | 0 | 0 | 0 | 0 | 0 | 1067 | 660 | 700 | 357 | 278 | 191 | 391 | 3643 |
| 1997 | 0 | 0 | 0 | 0 | 0 | 328 | 751 | 771 | 417 | 190 | 116 | 166 | 2739 |

[^0]Table 5. USA catch ( $t$ ) of haddock (excluding discard estimates) in unit areas 5Z.jm by gear category and tonnage class. Details for 1994-1997 are not available because data is preliminary.

| Year | Otter Trawl |  |  | Other | Total |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 3 | 4 | Total |  |  |
| 1969 | 3010 | 3610 | 6621 | 0 | 6622 |
| 1970 | 1602 | 1551 | 3154 | 0 | 3153 |
| 1971 | 1760 | 1768 | 3533 | 0 | 3534 |
| 1972 | 861 | 690 | 1551 | 0 | 1551 |
| 1973 | 637 | 759 | 1396 | 0 | 1396 |
| 1974 | 443 | 512 | 955 | 0 | 955 |
| 1975 | 993 | 675 | 1668 | 36 | 1705 |
| 1976 | 671 | 302 | 972 | 2 | 973 |
| 1977 | 1721 | 700 | 2423 | 5 | 2429 |
| 1978 | 3140 | 1573 | 4713 | 11 | 4724 |
| 1979 | 3281 | 1927 | 5208 | 4 | 5211 |
| 1980 | 3654 | 2955 | 5611 | 4 | 5615 |
| 1981 | 3591 | 5408 | 9031 | 45 | 9077 |
| 1982 | 2585 | 3657 | 6242 | 37 | 6280 |
| 1983 | 1162 | 3261 | 4423 | 29 | 4454 |
| 1984 | 1854 | 3260 | 5115 | 5 | 5121 |
| 1985 | 856 | 823 | 1679 | 4 | 1683 |
| $1986$ | 985 | 1207 | 2192 | 9 | 2200 |
| 1987 | 778 | 639 | 1417 | 1 | 1418 |
| 1988 | 920 | 768 | 1688 | 6 | 1693 |
| 1989 | 359 | 419 | 780 | 6 | 787 |
| 1990 | 486 | 688 | 1178 | 4 | 1189 |
| 1991 | 400 | 517 | 918 | 13 | 931 |
| 1992 | 597 | 740 | 1337 | 292 | 1629 |
| 1993 | 142 | 191 | 333 | 88 | 421 |
| 1994 |  |  | 32 | 0 | 33 |
| 1995 |  |  | 21 | 0 | 22 |
| 1996 |  |  | 36 | 0 | 36 |
| 1997 |  |  | 48 | 0 | 48 |

Table 6. Monthly catch ( t ) of haddock (excluding discard estimates) by USA in unit areas 5Zjm. Details for 19941997 are not available because data is preliminary.

| Year | Jan | Feb | Mar | Apr | May | Jun | Jul | Aug | Sep | Oct | Nov | Dec | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1969 | 525 | 559 | 976 | 1825 | 670 | 809 | 204 | 219 | 249 | 226 | 203 | 157 | 6622 |
| 1970 | 169 | 219 | 242 | 375 | 608 | 374 | 324 | 333 | 179 | 219 | 61 | 50 | 3153 |
| 1971 | 155 | 361 | 436 | 483 | 668 | 503 | 338 | 152 | 147 | 165 | 58 | 68 | 3534 |
| 1972 | 150 | 196 | 91 | 90 | 239 | 261 | 97 | 164 | 84 | 63 | 52 | 64 | 1551 |
| 1973 | 90 | 111 | 77 | 85 | 138 | 365 | 217 | 196 | 37 | 3 | 22 | 55 | 1396 |
| 1974 | 135 | 70 | 47 | 70 | 122 | 160 | 165 | 43 | 27 | 6 | 19 | 91 | 955 |
| 1975 | 152 | 123 | 32 | 116 | 388 | 489 | 138 | 95 | 57 | 24 | 52 | 39 | 1705 |
| 1976 | 116 | 147 | 83 | 106 | 323 | 162 | 7 | 6 | 5 | 2 | 3 | 13 | 973 |
| 1977 | 75 | 211 | 121 | 154 | 374 | 372 | 434 | 191 | 73 | 52 | 146 | 226 | 2429 |
| 1978 | 336 | 437 | 263 | 584 | 752 | 750 | 467 | 221 | 245 | 426 | 194 | 49 | 4724 |
| 1979 | 274 | 329 | 352 | 548 | 766 | 816 | 588 | 659 | 224 | 202 | 281 | 172 | 5211 |
| 1980 | 632 | 1063 | 742 | 784 | 711 | 461 | 324 | 254 | 221 | 91 | 110 | 222 | 5615 |
| 1981 | 550 | 1850 | 634 | 627 | 882 | 1326 | 1233 | 873 | 321 | 284 | 242 | 255 | 9077 |
| 1982 | 425 | 754 | 502 | 347 | 718 | 1801 | 757 | 145 | 201 | 216 | 276 | 138 | 6280 |
| 1983 | 492 | 931 | 272 | 181 | 310 | 1145 | 231 | 178 | 187 | 110 | 227 | 190 | 4454 |
| 1984 | 540 | 961 | 366 | 281 | 627 | 1047 | 370 | 302 | 250 | 196 | 92 | 89 | 5121 |
| 1985 | 165 | 190 | 254 | 300 | 352 | 206 | 60 | 47 | 1 | 24 | 41 | 43 | 1683 |
| 1986 | 184 | 396 | 334 | 479 | 496 | 221 | 31 | 6 | 12 | 6 | 6 | 29 | 2200 |
| 1987 | 225 | 52 | 43 | 307 | 233 | 342 | 67 | 30 | 24 | 4 | 23 | 68 | 1418 |
| 1988 | 196 | 152 | 207 | 245 | 366 | 316 | 30 | 19 | 6 | 1 | 45 | 110 | 1693 |
| 1989 | 114 | 56 | 47 | 164 | 161 | 145 | 15 | 8 | 1 | 5 | 25 | 46 | 787 |
| 1990 | 148 | 21 | 155 | 274 | 214 | 306 | 23 | 3 | 5 | 5 | 16 | 19 | 1189 |
| 1991 | 105 | 28 | 76 | 133 | 89 | 434 | 1 | 20 | 6 | 0 | 19 | 19 | 931 |
| 1992 | 253 | 81 | 51 | 149 | 353 | 669 | 20 | 20 | 17 | 3 | 2 | 12 | 1629 |
| 1993 | 15 | 12 | 16 | 55 | 84 | 209 | 6 | 3 | 3 | 7 | 2 | 8 | 421 |
| 1994 |  |  |  |  |  |  |  |  |  |  |  |  | 33 |
| 1995 |  |  |  |  |  |  |  |  |  |  |  |  | 22 |
| 1996 |  |  |  |  |  |  |  |  |  |  |  |  | 36 |
| 1997 |  |  |  |  |  |  |  |  |  |  |  |  | 48 |

Table 7. Derivation of catch at age for the 1997 5Zjm Canadian haddock fishery.

$\mathrm{OTB}=$ Otter Trawl Bottom, OTS=Otter Trawl Side, $\mathrm{GN}=\mathrm{Gill} \operatorname{Net}, \mathrm{LL}=$ Longline, $\mathrm{HL}=$ Handline, $\mathrm{I}=$ Inshore (Tonnage Classes $<=3$ ), $\mathrm{OF}=$ Offshore (Tonnage Classes $>=4$ ).

Table 8. Components of catch at age numbers ( 000 's) of haddock from unit areas $5 Z \mathrm{Zjm}$ by quarter.

| Quarter | Age Group |  |  |  |  |  |  |  |  |  | Annual Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9+ | 1+ |  |
| Canadian |  |  |  |  |  |  |  |  |  |  |  |
| 1996 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |  |
| 1996.25 | 0.031 | 2.657 | 135.030 | 254.856 | 125.126 | 15.622 | 6.168 | 0.000 | 21.350 | 560.841 |  |
| 1996.5 | 0.000 | 16.206 | 234.318 | 431.968 | 177.203 | 19.791 | 5.475 | 2.584 | 34.187 | 921.733 |  |
| 1996.75 | 0.087 | 8.109 | 102.654 | 163.708 | 109.190 | 23.765 | 5.607 | 0.000 | 15.085 | 428.206 | 1910.780 |
| 1997 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |  |
| 1997.25 | 0.020 | 0.613 | 2.427 | 37.160 | 69.322 | 32.539 | 4.355 | 1.835 | 6.372 | 154.641 |  |
| 1997.5 | 0.091 | 44.041 | 45.146 | 411.285 | 311.587 | 134.305 | 3.455 | 3.409 | 21.114 | 974.433 |  |
| 1997.75 | 0.777 | 28.045 | 21.052 | 76.918 | 88.967 | 19.722 | 3.790 | 1.759 | 4.061 | 245.092 | 1374.166 |
| USA |  |  |  |  |  |  |  |  |  |  |  |
| 1996 | 0.000 | 0.007 | 0.323 | 0.936 | 1.127 | 0.312 | 0.183 | 0.157 | 0.418 | 3.463 |  |
| 1996.25 | 0.000 | 0.026 | 0.872 | 1.840 | 1.706 | 0.443 | 0.254 | 0.218 | 0.581 | 5.940 |  |
| 1996.5 | 0.000 | 0.154 | 0.980 | 0.870 | 0.349 | 0.047 | 0.037 | 0.016 | 0.021 | 2.474 |  |
| 1996.75 | 0.000 | 0.104 | 0.850 | 0.918 | 0.387 | 0.055 | 0.044 | 0.018 | 0.028 | 2.404 | 14.281 |
| 1997 | 0.000 | 0.000 | 0.000 | 0.335 | 1.183 | 0.934 | 0.148 | 0.089 | 0.276 | 2.965 |  |
| 1997.25 | 0.000 | 0.000 | 0.000 | 0.828 | 2.925 | 2.309 | 0.367 | 0.220 | 0.682 | 7.332 |  |
| 1997.5 | 0.000 | 0.016 | 0.022 | 0.923 | 2.165 | 1.634 | 0.065 | 0.092 | 0.510 | 5.427 |  |
| 1997.75 | 0.000 | 0.035 | 0.045 | 0.585 | 1.509 | 0.610 | 0.179 | 0.080 | 0.153 | 3.196 | 18.919 |


| USA Discards |  |  |  |  |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1996 | 0.024 | 0.291 | 1.199 | 1.878 | 1.363 | 0.289 | 0.070 | 0.052 | 0.249 | 5.415 |  |
| 1996.25 | 0.038 | 0.454 | 1.873 | 2.934 | 2.130 | 0.451 | 0.110 | 0.081 | 0.388 | 8.459 |  |
| 1996.5 | 0.164 | 0.618 | 1.425 | 0.927 | 0.385 | 0.030 | 0.028 | 0.005 | 0.023 | 3.605 |  |
| 1996.75 | 0.170 | 0.638 | 1.472 | 0.957 | 0.397 | 0.031 | 0.029 | 0.006 | 0.023 | 3.723 | 21.202 |
| 1997 | 0.680 | 4.321 | 6.554 | 7.914 | 2.889 | 1.117 | 0.316 | 0.194 | 0.291 | 24.275 |  |
| 1997.25 | 0.452 | 2.876 | 4.363 | 5.268 | 1.923 | 0.743 | 0.210 | 0.129 | 0.194 | 16.159 |  |
| 1997.5 | 0.391 | 1.006 | 0.670 | 0.806 | 0.313 | 0.102 | 0.003 | 0.048 | 0.061 | 3.399 |  |
| 1997.75 | 0.075 | 0.194 | 0.129 | 0.155 | 0.060 | 0.020 | 0.001 | 0.009 | 0.012 | 0.656 | 44.490 |


| Total |  |  |  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1996 | 0.024 | 0.298 | 1.522 | 2.814 | 2.490 | 0.601 | 0.253 | 0.209 | 0.667 | 8.878 |  |
| 1996.25 | 0.069 | 3.137 | 137.775 | 259.630 | 128.962 | 16.516 | 6.532 | 0.299 | 22.319 | 575.240 |  |
| 1996.5 | 0.164 | 16.978 | 236.723 | 433.765 | 177.937 | 19.868 | 5.540 | 2.605 | 34.231 | 927.812 |  |
| 1996.75 | 0.257 | 8.851 | 104.976 | 165.583 | 109.974 | 23.851 | 5.680 | 0.024 | 15.136 | 434.333 | 1946.263 |
| 1997 | 0.680 | 4.321 | 6.554 | 8.249 | 4.072 | 2.051 | 0.464 | 0.283 | 0.567 | 27.240 |  |
| 1997.25 | 0.472 | 3.489 | 6.790 | 43.256 | 74.170 | 35.592 | 4.932 | 2.184 | 7.248 | 178.132 |  |
| 1997.5 | 0.482 | 45.063 | 45.838 | 413.014 | 314.065 | 136.041 | 3.523 | 3.548 | 21.685 | 983.258 |  |
| 1997.75 | 0.852 | 28.275 | 21.226 | 77.658 | 90.536 | 20.352 | 3.970 | 1.848 | 4.226 | 248.944 | 1437.575 |

Table 9. Total commercial catch at age numbers ( 000 's) of haddock from unit areas 5 Zjm .

| Year | 1 | 2 | 3 | 4 | Age $C$ 5 | 6 | 7 | 8 | $9+$ | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1969 | 0 | 18 | 1441 | 260 | 331 | 2885 | 819 | 89 | 279 | 6123 |
| 1970 | 25 | 82 | 7 | 347 | 147 | 126 | 1140 | 364 | 189 | 2425 |
| 1971 | 0 | 1182 | 247 | 31 | 246 | 157 | 159 | 756 | 407 | 3185 |
| 1972 | 259 | 1 | 376 | 71 | 21 | 92 | 37 | 16 | 431 | 1303 |
| 1973 | 1015 | 1722 | 6 | 358 | 37 | 10 | 37 | 8 | 163 | 3358 |
| 1974 | 17 | 2105 | 247 | 0 | 31 | 3 | 0 | 29 | 57 | 2488 |
| 1975 | 0 | 270 | 1428 | 201 | 5 | 34 | 1 | 2 | 28 | 1969 |
| 1976 | 73 | 149 | 166 | 814 | 125 | 0 | 19 | 0 | 17 | 1363 |
| 1977 | 0 | 7836 | 64 | 178 | 303 | 162 | 0 | 15 | 14 | 8571 |
| 1978 | 1 | 285 | 9831 | 161 | 169 | 302 | 80 | 10 | 9 | 10848 |
| 1979 | 0 | 15 | 199 | 4250 | 362 | 201 | 215 | 43 | 14 | 5300 |
| 1980 | 3 | 17561 | 342 | 299 | 2407 | 191 | 129 | 51 | 12 | 20995 |
| 1981 | 0 | 660 | 6687 | 393 | 494 | 1234 | 119 | 33 | 7 | 9627 |
| 1982 | 0 | 713 | 1048 | 2799 | 201 | 377 | 723 | 62 | 65 | 5988 |
| 1983 | 0 | 140 | 648 | 546 | 1629 | 207 | 104 | 402 | 34 | 3710 |
| 1984 | 0 | 76 | 249 | 341 | 264 | 1120 | 186 | 165 | 314 | 2716 |
| 1985 | 0 | 2063 | 374 | 176 | 189 | 123 | 371 | 53 | 114 | 3463 |
| 1986 | 6 | 38 | 2557 | 173 | 142 | 122 | 118 | 173 | 41 | 3369 |
| 1987 | 0 | 1990 | 127 | 1515 | 96 | 56 | 82 | 68 | 108 | 4042 |
| 1988 | 4 | 51 | 2145 | 121 | 877 | 109 | 36 | 46 | 98 | 3487 |
| 1989 | 0 | 1153 | 78 | 734 | 129 | 320 | 31 | 20 | 45 | 2510 |
| 1990 | 2 | 7 | 1265 | 126 | 743 | 68 | 163 | 42 | 42 | 2457 |
| 1991 | 6 | 441 | 89 | 2041 | 88 | 389 | 72 | 145 | 61 | 3332 |
| 1992 | 7 | 230 | 311 | 127 | 1446 | 89 | 315 | 26 | 90 | 2640 |
| 1993 | 7 | 247 | 343 | 279 | 85 | 635 | 34 | 153 | 74 | 1856 |
| 1994 | 1 | 241 | 737 | 148 | 54 | 48 | 125 | 29 | 39 | 1423 |
| 1995 | 2 | 60 | 525 | 414 | 53 | 25 | 3 | 51 | 16 | 1149 |
| 1996 | 1 | 29 | 481 | 862 | 419 | 61 | 18 | 3 | 72 | 1946 |
| 1997 | 2 | 81 | 80 | 542 | 483 | 194 | 13 | 8 | 34 | 1438 |

Table 10. Average weight at age ( kg ) of haddock from the commercial fishery in unit areas 5 Zjm .

| Year | 1 | 2 | 3 | $\begin{gathered} \text { Age Gmun } \\ 4 \\ \hline \end{gathered}$ | 5 | 6 | 7 | 8 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1969 | 0.600 | 0.763 | 1.282 | 1.531 | 1.649 | 1.836 | 2.298 | 2.879 |
| 1970 | 0.721 | 1.067 | 0.812 | 1.653 | 1.886 | 2.124 | 2.199 | 2.841 |
| 1971 | 0.600 | 0.928 | 1.059 | 1.272 | 2.011 | 2.255 | 2.262 | 2.613 |
| 1972 | 0.759 | 1.000 | 1.562 | 1.750 | 2.147 | 2.505 | 2.411 | 2.514 |
| 1973 | 0.683 | 1.002 | 1.367 | 1.804 | 2.202 | 1.631 | 2.885 | 3.295 |
| 1974 | 0.600 | 0.970 | 1.418 | 1.800 | 1.984 | 3.760 | 2.700 | 3.128 |
| 1975 | 0.600 | 0.872 | 1.524 | 2.062 | 1.997 | 2.422 | 4.114 | 3.557 |
| 1976 | 0.596 | 0.956 | 1.293 | 1.857 | 2.417 | 2.700 | 2.702 | 3.000 |
| 1977 | 0.600 | 0.970 | 1.442 | 1.809 | 2.337 | 2.809 | 2.700 | 3.095 |
| 1978 | 0.619 | 1.151 | 1.433 | 2.055 | 2.623 | 2.919 | 2.972 | 2.829 |
| 1979 | 0.600 | 0.987 | 1.298 | 1.805 | 2.206 | 2.806 | 3.219 | 3.277 |
| 1980 | 0.405 | 0.892 | 1.034 | 1.705 | 2.115 | 2.593 | 3.535 | 3.608 |
| 1981 | 0.600 | 0.890 | 1.262 | 1.592 | 2.270 | 2.611 | 3.505 | 4.009 |
| 1982 | 0.600 | 0.965 | 1.363 | 1.786 | 2.327 | 2.557 | 2.958 | 3.531 |
| 1983 | 0.600 | 1.024 | 1.341 | 1.750 | 2.118 | 2.509 | 2.879 | 3.104 |
| 1984 | 0.600 | 0.876 | 1.354 | 1.838 | 2.159 | 2.605 | 2.856 | 3.134 |
| 1985 | 0.600 | 0.950 | 1.230 | 1.915 | 2.227 | 2.702 | 2.872 | 3.180 |
| 1986 | 0.452 | 0.981 | 1.352 | 1.866 | 2.367 | 2.712 | 2.969 | 3.570 |
| 1987 | 0.600 | 0.833 | 1.431 | 1.984 | 2.148 | 2.594 | 2.953 | 3.646 |
| 1988 | 0.421 | 0.974 | 1.305 | 1.708 | 2.042 | 2.350 | 3.011 | 3.305 |
| 1989 | 0.600 | 0.868 | 1.450 | 1.777 | 2.183 | 2.522 | 3.012 | 3.411 |
| 1990 | 0.639 | 0.999 | 1.419 | 1.787 | 2.141 | 2.509 | 2.807 | 3.002 |
| 1991 | 0.581 | 1.197 | 1.241 | 1.802 | 2.087 | 2.596 | 2.918 | 3.012 |
| 1992 | 0.538 | 1.163 | 1.622 | 1.654 | 2.171 | 2.491 | 2.988 | 3.388 |
| 1993 | 0.659 | 1.160 | 1.724 | 2.181 | 2.047 | 2.623 | 2.386 | 3.112 |
| 1994 | 0.405 | 1.135 | 1.661 | 2.235 | 2.639 | 2.422 | 2.831 | 3.223 |
| $1995$ | 0.797 | 1.055 | 1.511 | 2.033 | 2.550 | 2.755 | 2.908 | 3.010 |
| 1996 | 0.576 | 1.022 | 1.439 | 1.795 | 2.294 | 2.485 | 3.322 | 2.032 |
| 1997 | 0.685 | 1.216 | 1.336 | 1.747 | 2.121 | 2.476 | 3.034 | 3.367 |

Table 11. Conversion factors used in the ADAPT calibration.

|  |  | Spring |  | Fall |  |
| :--- | :--- | ---: | ---: | ---: | ---: |
| Year | Door | Vessel | Conversion | Vessel | Conversion |
| 1968 | BMV | Albatross IV | NA | Albatross IV | 1.49 |
| 1969 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1970 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1971 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1972 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1973 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1974 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1975 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1976 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1977 | BMV | Albatross IV | 1.49 | Delaware II | 1.2218 |
| 1978 | BMV | Albatross IV | 1.49 | Delaware II | 1.2218 |
| 1979 | BMV | Albatross IV | 1.49 | Delaware II | 1.2218 |
| 1980 | BMV | Albatross IV | 1.49 | Delaware II | 1.2218 |
| 1981 | BMV | Delaware II | 1.2218 | Delaware II | 1.2218 |
| 1982 | BMV | Delaware II | 1.2218 | Albatross IV | 1.49 |
| 1983 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1984 | BMV | Albatross IV | 1.49 | Albatross IV | 1.49 |
| 1985 | Polyvalent | Albatross IV | 1 | Albatross IV | 1 |
| 1986 | Polyvalent | Albatross IV | 1 | Albatross IV | 1 |
| 1987 | Polyvalent | Albatross IV | 1 | Albatross IV | 1 |
| 1988 | Polyvalent | Albatross IV | 1 | Albatross IV | 1 |
| 1989 | Polyvalent | Delaware II | 0.82 | Delaware II | 0.82 |
| 1990 | Polyvalent | Delaware II | 0.82 | Delaware II | 0.82 |
| 1991 | Polyvalent | Delaware II | 0.82 | Delaware II | 0.82 |
| 1992 | Polyvalent | Albatross IV | 1 | Albatross IV | 1 |
| 1993 | Polyvalent | Albatross IV | 1 | Delaware II | 0.82 |
| 1994 | Polyvalent | Delaware II | 0.82 | Albatross IV | 1 |
| 1995 | Polyvalent | Albatross IV | 1 | Albatross IV | 1 |
| 1996 | Polyvalent | Albatross IV | 1 | Albatross IV | 1 |
| 1997 | Polyvalent | Albatross IV | 1 | Albatross IV | 1 |

Table 12. Total estimated abundance at age numbers ( 000 's) of haddock for unit areas 5 Zjm from the Canadian spring surveys.

| Year | Age Groun |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9+ | Total |
| 1986 | 5057 | 306 | 8175 | 997 | 189 | 348 | 305 | 425 | 401 | 16205 |
| 1987 | 46 | 4286 | 929 | 3450 | 653 | 81 | 387 | 135 | 1132 | 11099 |
| 1988 | 971 | 49 | 12714 | 257 | 4345 | 274 | 244 | 130 | 686 | 19671 |
| 1989 | 48 | 6664 | 991 | 2910 | 247 | 528 | 40 | 36 | 260 | 11725 |
| 1990 | 726 | 108 | 12302 | 166 | 4465 | 299 | 1370 | 144 | 389 | 19968 |
| 1991 | 393 | 2159 | 137 | 10876 | 116 | 1899 | 119 | 507 | 225 | 16431 |
| 1992 | 1914 | 3879 | 1423 | 221 | 4810 | 18 | 1277 | 52 | 655 | 14248 |
| 1993 | 3448 | 1759 | 545 | 431 | 34 | 1186 | 19 | 281 | 147 | 7849 |
| 1994 | 4197 | 15163 | 5332 | 549 | 314 | 20 | 915 | 18 | 356 | 26864 |
| 1995 | 1231 | 3224 | 6236 | 3034 | 720 | 398 | 0 | 729 | 849 | 16422 |
| 1996 | 1477 | 2059 | 4784 | 5247 | 3391 | 326 | 246 | 20 | 698 | 18247 |
| 1997 | 1033 | 1550 | 1222 | 2742 | 2559 | 1397 | 150 | 65 | 372 | 11090 |
| 1998 | 2438 | 10893 | 4312 | 3608 | 5217 | 5029 | 2645 | 329 | 653 | 35124 |

Table 13. Total estimated abundance at age numbers ( 000 's) of haddock for unit areas 5 Zjm from the USA spring surveys. From 1973-81, a 41 Yankee trawl was used while a 36 Yankee trawl was used in other years. Conversion factors to adjust for changes in door type and survey vessel were applied.

| Year | Age Group |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9+ | Total |
| 1969 | 17 | 35 | 614 | 235 | 523 | 3232 | 1220 | 358 | 489 | 6724 |
| 1970 | 478 | 190 | 0 | 560 | 998 | 441 | 3169 | 2507 | 769 | 9113 |
| 1971 | 0 | 655 | 261 | 0 | 144 | 102 | 58 | 1159 | 271 | 2650 |
| 1972 | 2594 | 0 | 771 | 132 | 25 | 47 | 211 | 27 | 1214 | 5019 |
| 1973 | 2455 | 5639 | 0 | 1032 | 154 | 0 | 276 | 0 | 1208 | 10763 |
| 1974 | 1323 | 20596 | 4084 | 0 | 354 | 0 | 43 | 72 | 322 | 26795 |
| 1975 | 528 | 567 | 6016 | 1063 | 0 | 218 | 127 | 45 | 208 | 8773 |
| 1976 | 8279 | 402 | 433 | 1229 | 582 | 0 | 0 | 0 | 22 | 10948 |
| 1977 | 138 | 25922 | 294 | 855 | 816 | 586 | 0 | 22 | 98 | 28730 |
| 1978 | 0 | 743 | 20859 | 641 | 880 | 1163 | 89 | 23 | 116 | 24516 |
| 1979 | 10496 | 441 | 1313 | 9764 | 475 | 72 | 445 | 42 | 9 | 23057 |
| 1980 | 4364 | 67961 | 1129 | 1117 | 5822 | 628 | 381 | 705 | 359 | 82466 |
| 1981 | 3595 | 3041 | 27694 | 2887 | 719 | 2389 | 335 | 57 | 21 | 40738 |
| 1982 | 584 | 3697 | 1649 | 7743 | 745 | 447 | 669 | 0 | 0 | 15534 |
| 1983 | 238 | 770 | 686 | 359 | 2591 | 30 | 0 | 798 | 57 | 5529 |
| 1984 | 1366 | 1415 | 996 | 1001 | 936 | 1245 | 138 | 89 | 470 | 7656 |
| 1985 | 40 | 8911 | 1396 | 674 | 1496 | 588 | 1995 | 127 | 483 | 15709 |
| 1986 | 3334 | 280 | 3597 | 246 | 210 | 333 | 235 | 560 | 159 | 8953 |
| 1987 | 122 | 5480 | 144 | 1394 | 157 | 231 | 116 | 370 | 0 | 8013 |
| 1988 | 305 | 61 | 1868 | 235 | 611 | 203 | 218 | 178 | 0 | 3678 |
| 1989 | 84 | 6665 | 619 | 1343 | 267 | 791 | 58 | 92 | 47 | 9966 |
| 1990 | 1654 | 70 | 10338 | 598 | 1042 | 110 | 182 | 0 | 0 | 13995 |
| 1991 | 740 | 2071 | 432 | 3381 | 192 | 203 | 66 | 87 | 25 | 7198 |
| 1992 | 529 | 287 | 214 | 141 | 609 | 32 | 46 | 46 | 0 | 1905 |
| 1993 | 1870 | 1116 | 197 | 232 | 195 | 717 | 77 | 35 | 43 | 4481 |
| 1994 | 1025 | 4272 | 1487 | 269 | 184 | 118 | 278 | 28 | 85 | 7745 |
| 1995 | 921 | 2307 | 4096 | 1691 | 259 | 151 | 51 | 269 | 214 | 9959 |
| 1996 | 912 | 1351 | 3772 | 3232 | 1896 | 235 | 36 | 0 | 496 | 11931 |
| 1997 | 1635 | 1226 | 380 | 595 | 470 | 343 | 24 | 44 | 20 | 4736 |

Table 14. Total estimated abundance at age numbers ( 000 's) of haddock for unit areas 5 Zjm from the USA fall surveys. Conversion factors to adjust for changes in door type and survey vessel were applied.


Table 15. Statistical properties of estimates for population abundance and survey calibration constants for haddock in unit areas 5 Zjm .

| Age | Estimate | Standard Error | Relative Error | Bias | Relative Bias |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Population Abundance |  |  |  |  |  |
| 1 | 4597 | 3255 | 0.71 | 1162 | 0.25 |
| 2 | 12135 | 5449 | 0.45 | 1248 | 0.10 |
| 3 | 4158 | 1500 | 0.36 | 275 | 0.07 |
| 4 | 2565 | 802 | 0.31 | 127 | 0.05 |
| 5 | 3504 | 1081 | 0.31 | 132 | 0.04 |
| 6 | 3866 | 1145 | 0.30 | 126 | 0.03 |
| 7 | 1101 | 392 | 0.36 | 46 | 0.04 |
| 8 | 148 | 58 | 0.39 | 8 | 0.06 |
| Survey Calibration Constants |  |  |  |  |  |
| Canadian Spring Survey |  |  |  |  |  |
| 1 | 0.182 | 0.051 | 0.284 | 0.007 | 0.037 |
| 2 | 0.447 | 0.124 | 0.277 | 0.016 | 0.036 |
| 3 | 0.852 | 0.234 | 0.275 | 0.03 | 0.035 |
| 4 | 0.762 | 0.209 | 0.274 | 0.027 | 0.035 |
| 5 | 0.913 | 0.251 | 0.275 | 0.033 | 0.037 |
| 6 | 0.714 | 0.197 | 0.276 | 0.026 | 0.037 |
| 7 | 1.005 | 0.290 | 0.288 | 0.041 | 0.041 |
| 8 | 0.974 | 0.268 | 0.275 | 0.034 | 0.035 |
| USA Spring Survey - Yankee 36-1969-72/1982-97 |  |  |  |  |  |
| 1 | 0.129 | 0.029 | 0.227 | 0.003 | 0.024 |
| 2 | 0.340 | 0.077 | 0.226 | 0.008 | 0.024 |
| 3 | 0.430 | 0.097 | 0.225 | 0.010 | 0.024 |
| 4 | 0.460 | 0.103 | 0.225 | 0.011 | 0.024 |
| 5 | 0.556 | 0.122 | 0.219 | 0.013 | 0.023 |
| 6 | 0.454 | 0.100 | 0.219 | 0.010 | 0.023 |
| 7 | 0.559 | 0.126 | 0.225 | 0.013 | 0.023 |
| 8 | 0.703 | 0.167 | 0.237 | 0.017 | 0.025 |
| USA Spring Survey - Yankee 41-1973-81 |  |  |  |  |  |
| 1 | 0.231 | 0.079 | 0.343 | 0.014 | 0.059 |
| 2 | 0.522 | 0.169 | 0.323 | 0.027 | 0.052 |
| 3 | 0.665 | 0.228 | 0.343 | 0.039 | 0.059 |
| 4 | 0.814 | 0.279 | 0.343 | 0.048 | 0.059 |
| 5 | 1.021 | 0.350 | 0.343 | 0.060 | 0.059 |
| 6 | 0.937 | 0.371 | 0.396 | 0.073 | 0.078 |
| 7 | 2.332 | 0.855 | 0.367 | 0.157 | 0.067 |
| 8 | 0.893 | 0.327 | 0.367 | 0.060 | 0.067 |
| USA Fall Survey |  |  |  |  |  |
| 0 | 0.129 | 0.024 | 0.188 | 0.002 | 0.016 |
| 1 | 0.274 | 0.053 | 0.193 | 0.005 | 0.017 |
| 2 | 0.217 | 0.041 | 0.189 | 0.004 | 0.016 |
| 3 | 0.209 | 0.039 | 0.188 | 0.003 | 0.017 |
| 4 | 0.157 | 0.031 | 0.200 | 0.003 | 0.020 |
| 5 | 0.133 | 0.025 | 0.189 | 0.002 | 0.017 |

Table 16. Beginning of year population abundance numbers ( 000 's) for haddock in unit areas 5 Zjm .

| Year | Age Group |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 1+ | 2+ | 3+ |
| 1969 | 762 | 161 | 3994 | 849 | 885 | 8401 | 2799 | 177 | 0 | 18028 | 17266 | 17105 |
| 1970 | 3342 | 624 | 115 | 1982 | 461 | 431 | 4314 | 1564 | 66 | 12899 | 9557 | 8933 |
| 1971 | 311 | 2713 | 435 | 88 | 1313 | 246 | 240 | 2509 | 955 | 8810 | 8499 | 5786 |
| 1972 | 5154 | 255 | 1126 | 135 | 45 | 855 | 61 | 56 | 1383 | 9070 | 3916 | 3661 |
| 1973 | 11029 | 3977 | 208 | 586 | 46 | 19 | 620 | 17 | 32 | 16534 | 5505 | 1528 |
| 1974 | 3144 | 8121 | 1684 | 165 | 152 | 5 | 7 | 474 | 7 | 13759 | 10615 | 2494 |
| 1975 | 3218 | 2558 | 4749 | 1162 | 135 | 98 | 2 | 5 | 363 | 12290 | 9072 | 6514 |
| 1976 | 53818 | 2634 | 1842 | 2592 | 772 | 106 | 50 | 1 | 3 | 61818 | 8000 | 5366 |
| 1977 | 5914 | 43996 | 2023 | 1360 | 1403 | 520 | 87 | 24 | 0 | 55327 | 49413 | 5417 |
| 1978 | 4210 | 4842 | 28868 | 1600 | 955 | 884 | 285 | 71 | 7 | 41722 | 37512 | 32670 |
| 1979 | 52003 | 3446 | 3690 | 14543 | 1161 | 631 | 457 | 161 | 50 | 76142 | 24139 | 20693 |
| 1980 | 6645 | 42577 | 2807 | 2839 | 8106 | 626 | 342 | 185 | 93 | 64220 | 57575 | 14998 |
| 1981 | 5136 | 5438 | 19009 | 1995 | 2057 | 4521 | 343 | 168 | 107 | 38774 | 33638 | 28200 |
| 1982 | 1711 | 4205 | 3844 | 9586 | 1285 | 1244 | 2617 | 177 | 109 | 24778 | 23067 | 18862 |
| 1983 | 2629 | 1401 | 2780 | 2204 | 5325 | 869 | 683 | 1496 | 90 | 17477 | 14848 | 13447 |
| 1984 | 14880 | 2153 | 1016 | 1685 | 1313 | 2914 | 525 | 466 | 872 | 25824 | 10944 | 8791 |
| 1985 | 1548 | 12183 | 1694 | 607 | 1073 | 842 | 1395 | 267 | 236 | 19845 | 18297 | 6114 |
| 1986 | 13218 | 1267 | 8041 | 1040 | 338 | 709 | 579 | 816 | 171 | 26179 | 12961 | 11694 |
| 1987 | 1292 | 10817 | 1003 | 4296 | 700 | 150 | 472 | 372 | 517 | 19619 | 18327 | 7510 |
| 1988 | 14981 | 1058 | 7054 | 707 | 2156 | 486 | 73 | 313 | 243 | 27071 | 12090 | 11032 |
| 1989 | 787 | 12262 | 821 | 3837 | 470 | 991 | 301 | 28 | 216 | 19713 | 18926 | 6664 |
| 1990 | 2380 | 644 | 8999 | 601 | 2478 | 269 | 525 | 220 | 5 | 16121 | 13741 | 13097 |
| 1991 | 2134 | 1947 | 521 | 6220 | 379 | 1361 | 160 | 284 | 142 | 13148 | 11014 | 9067 |
| 1992 | 7609 | 1742 | 1190 | 347 | 3238 | 231 | 764 | 67 | 102 | 15290 | 7681 | 5939 |
| 1993 | 14382 | 6224 | 1214 | 693 | 171 | 1355 | 111 | 343 | 32 | 24525 | 10143 | 3919 |
| 1994 | 9522 | 11769 | 4860 | 681 | 319 | 65 | 547 | 61 | 146 | 27970 | 18448 | 6679 |
| 1995 | 4623 | 7795 | 9409 | 3290 | 421 | 211 | 8 | 333 | 23 | 26113 | 21490 | 13695 |
| 1996 | 5907 | 3783 | 6326 | 7219 | 2314 | 296 | 150 | 4 | 226 | 26225 | 20318 | 16535 |
| 1997 | 13300 | 4836 | 3070 | 4735 | 5116 | 1507 | 186 | 106 | 1 | 32857 | 19557 | 14721 |
| 1998 | 3435 | 10887 | 3883 | 2439 | 3372 | 3740 | 1054 | 140 | 80 | 29030 | 25595 | 14708 |

Table 17. Fishing mortality rate for haddock in unit areas 5 Zjm . The rate for ages $4+$ (ages 4 to 8 ) is weighted with population numbers.

| Year | Age Group |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 4+ |
| 1969 | 0 | 0.132 | 0.5 | 0.41 | 0.519 | 0.466 | 0.382 | 0.788 | 0.451 |
| 1970 | 0.009 | 0.161 | 0.068 | 0.212 | 0.43 | 0.388 | 0.342 | 0.293 | 0.309 |
| 1971 | 0 | 0.679 | 0.971 | 0.465 | 0.23 | 1.19 | 1.252 | 0.396 | 0.405 |
| 1972 | 0.059 | 0.003 | 0.454 | 0.865 | 0.692 | 0.121 | 1.062 | 0.355 | 0.251 |
| 1973 | 0.106 | 0.659 | 0.033 | 1.149 | 2.062 | 0.837 | 0.069 | 0.748 | 0.492 |
| 1974 | 0.006 | 0.336 | 0.171 | 0 | 0.244 | 0.839 | 0.006 | 0.067 | 0.086 |
| 1975 | 0 | 0.129 | 0.405 | 0.21 | 0.037 | 0.466 | 0.898 | 0.52 | 0.208 |
| 1976 | 0.002 | 0.064 | 0.103 | 0.414 | 0.194 | 0 | 0.534 | 0 | 0.349 |
| 1977 | 0 | 0.221 | 0.034 | 0.153 | 0.262 | 0.402 | 0 | 1.06 | 0.233 |
| 1978 | 0 | 0.072 | 0.486 | 0.121 | 0.214 | 0.46 | 0.373 | 0.165 | 0.234 |
| 1979 | 0 | 0.005 | 0.062 | 0.385 | 0.418 | 0.414 | 0.705 | 0.346 | 0.395 |
| 1980 | 0 | 0.606 | 0.141 | 0.122 | 0.384 | 0.402 | 0.509 | 0.347 | 0.320 |
| 1981 | 0 | 0.147 | 0.485 | 0.24 | 0.303 | 0.347 | 0.463 | 0.235 | 0.314 |
| 1982 | 0 | 0.214 | 0.356 | 0.388 | 0.191 | 0.399 | 0.359 | 0.478 | 0.366 |
| 1983 | 0 | 0.121 | 0.301 | 0.318 | 0.403 | 0.303 | 0.184 | 0.34 | 0.352 |
| 1984 | 0 | 0.04 | 0.314 | 0.251 | 0.244 | 0.536 | 0.477 | 0.481 | 0.394 |
| 1985 | 0 | 0.215 | 0.288 | 0.386 | 0.214 | 0.174 | 0.337 | 0.243 | 0.271 |
| 1986 | 0 | 0.034 | 0.427 | 0.196 | 0.61 | 0.206 | 0.243 | 0.256 | 0.254 |
| 1987 | 0 | 0.228 | 0.149 | 0.489 | 0.165 | 0.527 | 0.211 | 0.226 | 0.405 |
| 1988 | 0 | 0.054 | 0.409 | 0.208 | 0.577 | 0.278 | 0.758 | 0.171 | 0.422 |
| 1989 | 0 | 0.109 | 0.112 | 0.237 | 0.358 | 0.435 | 0.116 | 1.597 | 0.276 |
| 1990 | 0.001 | 0.012 | 0.169 | 0.26 | 0.4 | 0.32 | 0.415 | 0.235 | 0.365 |
| 1991 | 0.003 | 0.292 | 0.207 | 0.453 | 0.298 | 0.377 | 0.671 | 0.82 | 0.447 |
| 1992 | 0.001 | 0.161 | 0.341 | 0.506 | 0.671 | 0.531 | 0.601 | 0.536 | 0.637 |
| 1993 | 0.001 | 0.047 | 0.378 | 0.576 | 0.768 | 0.706 | 0.398 | 0.651 | 0.654 |
| 1994 | 0 | 0.024 | 0.19 | 0.282 | 0.213 | 1.861 | 0.297 | 0.772 | 0.319 |
| 1995 | 0 | 0.009 | 0.065 | 0.152 | 0.152 | 0.14 | 0.488 | 0.188 | 0.155 |
| 1996 | 0 | 0.009 | 0.09 | 0.144 | 0.229 | 0.267 | 0.146 | 1.839 | 0.167 |
| 1997 | 0 | 0.019 | 0.03 | 0.139 | 0.113 | 0.157 | 0.081 | 0.087 | 0.129 |

Table 18. Average weight at age from the Canadian spring survey.

| Year | Age Group |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | $9+$ |  |  |  |  |  |  |
| 1986 | 0.135 | 0.452 | 0.974 | 1.445 | 3.039 | 2.843 | 3.598 | 3.373 | 3.914 |  |  |  |  |  |  |
| 1987 | 0.150 | 0.500 | 0.716 | 1.672 | 2.011 | 2.548 | 3.149 | 3.147 | 3.629 |  |  |  |  |  |  |
| 1988 | 0.097 | 0.464 | 0.931 | 1.795 | 1.816 | 1.916 | 2.721 | 3.267 | 3.869 |  |  |  |  |  |  |
| 1989 | 0.062 | 0.474 | 0.649 | 1.392 | 1.995 | 2.528 | 2.155 | 2.820 | 2.963 |  |  |  |  |  |  |
| 1990 | 0.149 | 0.527 | 0.924 | 1.185 | 1.863 | 2.072 | 2.507 | 2.819 | 3.469 |  |  |  |  |  |  |
| 1991 | 0.120 | 0.689 | 0.801 | 1.510 | 1.687 | 2.428 | 2.103 | 3.125 | 3.435 |  |  |  |  |  |  |
| 1992 | 0.122 | 0.602 | 1.118 | 1.060 | 2.078 | 2.165 | 2.709 | 2.283 | 3.443 |  |  |  |  |  |  |
| 1993 | 0.122 | 0.481 | 1.227 | 1.803 | 1.272 | 2.333 | 2.340 | 2.740 | 3.293 |  |  |  |  |  |  |
| 1994 | 0.107 | 0.469 | 1.047 | 1.621 | 1.926 | 2.154 | 3.153 | 2.688 | 3.084 |  |  |  |  |  |  |
| 1995 | 0.086 | 0.493 | 0.963 | 1.556 | 2.224 | 2.447 | 2.400 | 2.991 | 3.184 |  |  |  |  |  |  |
| 1996 | 0.139 | 0.495 | 0.919 | 1.320 | 1.932 | 2.555 | 2.899 | 2.603 | 3.588 |  |  |  |  |  |  |
| 1997 | 0.132 | 0.507 | 0.782 | 1.205 | 1.664 | 2.177 | -2.450 | 2.586 | 3.163 |  |  |  |  |  |  |
| 1998 | 0.053 | 0.542 | 0.993 | 1.139 | 1.567 | 1.975 | 2.511 | 3.589 | 3.414 |  |  |  |  |  |  |

Table 19. Beginning of year biomass for haddock in unit areas 5 Zjm .

| Year | Age Group |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | $1+$ | $2+$ | $3+$ |
| 1969 | 88 | 83 | 3735 | 1277 | 1762 | 19686 | 7511 | 518 | 0 | 34659 | 34571 | 34489 |
| 1970 | 384 | 321 | 108 | 2981 | 918 | 1010 | 11576 | 4575 | 226 | 22100 | 21715 | 21394 |
| 1971 | 36 | 1398 | 407 | 132 | 2614 | 576 | 644 | 7339 | 3274 | 16421 | 16385 | 14987 |
| 1972 | 593 | 131 | 1053 | 203 | 90 | 2004 | 164 | 164 | 4741 | 9142 | 8549 | 8418 |
| 1973 | 1269 | 2049 | 194 | 881 | 92 | 45 | 1664 | 50 | 110 | 6353 | 5084 | 3035 |
| 1974 | 362 | 4183 | 1575 | 248 | 303 | 12 | 19 | 1387 | 24 | 8112 | 7750 | 3567 |
| 1975 | 370 | 1318 | 4441 | 1748 | 269 | 230 | 5 | 15 | 1244 | 9639 | 9269 | 7951 |
| 1976 | 6191 | 1357 | 1722 | 3898 | 1537 | 248 | 134 | 3 | 10 | 15102 | 8911. | 7554 |
| 1977 | 680 | 22664 | 1892 | 2045 | 2794 | 1219 | 233 | 70 | 0 | 31597 | 30917 | 8253 |
| 1978 | 484 | 2494 | 26993 | 2406 | 1902 | 2071 | 765 | 208 | 24 | 37348 | 36864 | 34369 |
| 1979 | 5982 | 1775 | 3450 | 21873 | 2312 | 1479 | 1226 | 471 | 171 | 38740 | 32757 | 30982 |
| 1980 | 764 | 21933 | 2625 | 4270 | 16140 | 1467 | 918 | 541 | 319 | 48977 | 48212 | 26279 |
| 1981 | 591 | 2801 | 17775 | 3001 | 4096 | 10594 | 920 | 491 | 367 | 40636 | 40045 | 37244 |
| 1982 | 197 | 2166 | 3594 | 14418 | 2559 | 2915 | 7022 | 518 | 374 | 33762 | 33566 | 31399 |
| 1983 | 302 | 722 | 2599 | 3315 | 10603 | 2036 | 1833 | 4376 | 309 | 26095 | 25793 | 25071 |
| 1984 | 1712 | 1109 | 950 | 2534 | 2614 | 6828 | 1409 | 1363 | 2989 | 21509 | 19798 | 18688 |
| 1985 | 178 | 6276 | 1584 | 913 | 2136 | 1973 | 3743 | 781 | 809 | 18394 | 18216 | 11940 |
| 1986 | 1780 | 572 | 7835 | 1503 | 1027 | 2016 | 2083 | 2752 | 669 | 20238 | 18458 | 17886 |
| 1987 | 194 | 5404 | 718 | 7184 | 1408 | 382 | 1486 | 1171 | 1876 | 19823 | 19629 | 14225 |
| 1988 | 1456 | 491 | 6564 | 1269 | 3916 | 931 | 199 | 1023 | 940 | 16789 | 15333 | 14842 |
| 1989 | 49 | 5813 | 533 | 5342 | 938 | 2505 | 649 | 79 | 640 | 16548 | 16499 | 10686 |
| 1990 | 354 | 339 | 8318 | 712 | 4616 | 557 | 1316 | 620 | 17 | 16851 | 16496 | 16157 |
| 1991 | 257 | 1341 | 417 | 9395 | 640 | 3304 | 336 | 887 | 488 | 17065 | 16808 | 15467 |
| 1992 | 931 | 1049 | 1331 | 368 | 6730 | 500 | 2069 | 153 | 351 | 13482 | 12552 | 11502 |
| 1993 | 1754 | 2994 | 1490 | 1250 | 217 | 3161 | 260 | 940 | 105 | 12171 | 10417 | 7423 |
| 1994 | 1016 | 5522 | 5087 | 1104 | 614 | 140 | 1725 | 164 | 450 | 15822 | 14806 | 9284 |
| 1995 | 399 | 3846 | 9062 | 5121 | 936 | 516 | 19 | 996 | 73 | 20967 | 20569 | 16723 |
| 1996 | 819 | 1872 | 5814 | 9530 | 4470 | 756 | 435 | 10 | 811 | 24517 | 23699 | 21826 |
| 1997 | 1758 | 2450 | 2400 | 5706 | 8511 | 3281 | 456 | 274 | 3 | 24839 | 23080 | 20631 |
| 1998 | 183 | 5900 | 3856 | 2777 | 5284 | 7386 | 2647 | 502 | 273 | 28809 | 28626 | 22726 |

Table 20. Projection results for 1998 at $\mathrm{F}_{0.1}$ of haddock in unit areas 5 Zjm .

| Year | Age Group |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | $1+$ | $2+$ | $3+$ |
| Beginning of Year Population Numbers (000s) |  |  |  |  |  |  |  |  |  |  |  |  |
| 1998 | 3435 | 10887 | 3883 | 2439 | 3372 | 3740 | 1054 | 140 | 80 |  |  |  |
| 1999 | 5731 | 2813 | 8737 | 2841 | 1555 | 2150 | 2385 | 672 | 89 |  |  |  |
| Partial Recruitment to the Fishery |  |  |  |  |  |  |  |  |  |  |  |  |
| 1998 | 0.00 | 0.08 | 0.45 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |  |  |  |  |
| Fishing Mortality |  |  |  |  |  |  |  |  |  |  |  |  |
| 1998 | 0.000 | 0.020 | 0.113 | 0.250 | 0.250 | 0.250 | 0.250 | 0.250 |  |  |  |  |
| Weight at beginning of year for population ( kg ) |  |  |  |  |  |  |  |  |  |  |  |  |
| 1999 | 0.10 | 0.51 | 0.91 | 1.31 | 1.85 | 2.29 | 2.62 | 2.94 | 3.34 |  |  |  |
| Beginning of Year Projected Population Biomass (t) |  |  |  |  |  |  |  |  |  |  |  |  |
| 1999 | 588 | 1432 | 7988 | 3708 | 2851 | 4921 | 6248 | 1978 | 298 | 30032 | 29444 | 28012 |
| Projected Catch Numbers (000s) |  |  |  |  |  |  |  |  |  |  |  |  |
| 1998 | 0 | 195 | 375 | 491 | 679 | 753 | 212 | 28 |  |  |  |  |
| Average weight for catch (kg) |  |  |  |  |  |  |  |  |  |  |  |  |
| 1998 | 0.62 | 1.11 | 1.49 | 1.95 | 2.40 | 2.53 | 3.02 | 2.91 |  |  |  |  |
| Projected Yield ( () |  |  |  |  |  |  |  |  |  |  |  |  |
| 1998 | 0 | 216 | 558 | 959 | 1630 | 1908 | 642 | 82 |  | 5995 |  |  |



Fig. 1. Fisheries statistical unit areas in NAFO Subdivision 5Ze.


Fig. 2. Nominal catch of haddock in unit areas 5Zjm.


Fig. 3. Historic catch of haddock in 5 Zjm compared to recent catches.


Fig. 4. Haddock landed in 5Zjm by month and gear by the Canadian commercial fishery in 1997 with sampling levels.


Fig. 5. Length compositions of the principal Canadian 5 Zjm commercial haddock fisheries in 1997 are fairly similar but haddock caught by gillnets are somewhat larger than those caught by other gears.


Fig. 6. Age composition of the Canadian 5Zjm commercial fisheries haddock catch in 1997 compared to the average for three periods which represent different stages in the Georges Bank fishery. Ages 4 and 5 made up the bulk of the catch, a consequence of lower selection for smaller haddock than in the past and the resulting higher survival of the 1992 year-class.


Fig. 7. Catch rates for haddock from Canadian commercial fishery gadoid trips ( $90 \%$ cod, haddock and pollock) in 5Zjm for vessels which fished during 1994 and reported more than 1t of landings. A generally increasing trend is seen from 1993 to 1995 with 1996 and 1997 values similar to 1995. (LL = longline, OT = otter trawl, $\mathrm{TC}=$ tonnage class).


Fig. 8. Stratification scheme used for USA surveys. The 5Zjm management area is indicated by shading.


Fig. 9. Stratification scheme used for the Canadian survey. The 5 Zjm management area is indicated by shading..

Age 1


Age 2


Age 3+


Fig. 10. Distribution of 5Zjm haddock as observed from the DFO spring survey. The squares are shaded relative to the average catch for 1993 to 1997. The expanding symbols represent the 1998 survey catches.


Fig. 11. Distribution of 5Zjm haddock as observed from the NMFS spring survey. The squares are shaded relative to the average catch for 1992 to 1996. The expanding symbols represent the 1997 survey catches.


Fig. 12. Distribution of 5Z.jm haddock as observed from the NMFS fall survey. The squares are shaded relative to the average catch for 1992 to 1996. The expanding symbols represent the 1997 survey catches.


Fig. 13. Beginning of year biomass for ages $2+$ from the NMFS fall and ages $3+$ from the NMFS and DFO spring research surveys (adjusted by calibration constants) for haddock in unit areas 5Zjm. Square $=$ NMFS fall (October/November); circle $=$ NMFS spring (March/April); diamond = DFO spring (February/March).


Fig. 14. Beginning of year biomass for ages 0 and 1 from the NMFS fall and ages 1 and 2 from the NMFS and DFO spring research surveys (adjusted by calibration constants) for haddock in unit areas 5 Zjm . Fall values are compared to the beginning of the subsequent year.


Fig. 15. Weight at age for haddock in unit areas 5Zjm derived from the DFO spring surveys.


Fig. 16. Ln and linear age specific annual catchabilites for base, Yankee 36 ('69-‘72',82-‘97) and Yankee 41 ('73-‘81) indices from NMFS spring survey.


Fig.
17. Ln catchabilities with one standard deviation for base, Yankee 36 and Yankee 41 indices from NMFS spring survey.


Fig. 18. Residuals for base and split model from NMFS spring survey.


Fig. 19. Impact on terminal year population numbers of base versus split model.


Fig. 20a. Age by age plots of A) the observed and predicted $\ln$ abundance index versus $\ln$ population numbers, and $B$ ) residuals plotted against year for haddock in unit areas 5 Zj and 5 Zm for the DFO spring survey.


Fig. 20b. Age by age plots of $A$ ) the observed and predicted In abundance index versus $\ln$ population numbers, and $B$ ) residuals plotted against ycar for haddock in unit areas 5 Zj and 5 Zm for the NMFS spring survey. The survey was used as two separate indices as a net change occurred in the series. During 1969-72/1982-97 a Yankee 36 was used while a Yankee 41 was used from 1973-81.


Fig. 20c. Age by age plots of A) the observed and predicted $\ln$ abundance index versus $\ln$ population numbers, and $B$ ) residuals plotted against year for haddock in unit areas 5 Zj and 5 Zm for the NMFS fall survey.


Fig. 21. Successive estimates of year-class abundance as additional years of data were included in the assessment did not display any persistent trends.


Fig. 22. Retrospective estimates of biomass and fishing mortality did not display any persistent trends for over or under estimation as successive years of data were excluded in the assessment.


Fig. 23. Mortality estimates from surveys show considerable year to year variation. A 3 year running median smoother was applied to examine the trends.


Fig. 24. Beginning of year biomass for haddock in unit areas 5 Zjm .


Fig. 25. Number of age 1 recruits for haddock in unit areas 5 Zjm .


Fig. 26. Fishing mortality rate for haddock ages 4 and older in unit areas 5 Zjm .


Fig. 27. Decay of the 1992 5Zjm haddock year-class versus the 1983, 1975 and 1978 as they progress through the fishery.


Fig. 28. Comparison of age composition and absolute abundance of the 5 Zjm haddock population in 1998 to earlier periods.


Fig. 29. Amount of productivity attributible to growth of ages 2 to 85 Zjm haddock and the amount contributed by recruitment of age 2 haddock.


Fig. 30. Surplus production of 5 Zjm haddock available to the commercial fishery compared to amount actually harvested.


Fig. 31. Predicted weight at age for 5 Zjm haddock as used in yield per recruit analysis.


Fig. 32. Yield per recruit analysis for 5 Zjm haddock assuming complete fishing on last age group (i.e. + group) with the $\mathrm{F}_{0.1}$ value identified.


Fig. 33. Historic catch and biomass of haddock in 5 Zjm compared to recent catches and biomass.


Fig. 34. Relationship between mature (3+) 5 Zjm haddock biomass and recruits at age 1 from 1931 to 1955 and 1969 to 1997.


Fig. 35. Expected exploitation rate in 1998 and expected change in biomass from 1998 to 1999 for 5 Zjm haddock at various quotas.


Fig. 36. Probability of fishing mortality exceeding the $\mathrm{F}_{0.1}(=0.25)$ reference level and of the 1999 biomass being less than the 1998 biomass by $0 \%, 10 \%$ and $20 \%$ for 5 Zjm haddock at various quotas.


Fig. 37. The spawning components of the Georges Bank haddock stock are comprised of an eastern component on the Northeast Peak and a western component in the Great South Channel. Darker shading indicates higher density of aggregation on average over the year.


Fig. 38. Comparison of the catches of haddock in 5 Z and 5 Zjm and the 5 Zjm to 5 Z catch ratio.


Fig. 39. Comparison of total haddock biomass in 5 Z and 5 Zjm and the 5 Zjm to 5 Z biomass ratio.


Fig. 40. Comparison of the number of haddock recruits at age 1 in 5 Z and 5 Zjm and the ratio of 5 Zjm to 5 Z recruits.


Fig. 41. Comparison of the fishing mortality levels for haddock in 5 Z and 5 Zjm .

## Appendix A. Comparison of 5 Zjm haddock length frequencies from port and observer sampling with those from fishery officer surveillance boardings.

Haddock length frequencies from boardings made by surveillance officers in 5 Zjm were compared to those from port and observer samplers to look for evidence of discarding. The amount of data available for comparison is low. Twenty-nine surveillance samples taken from 1994 to 1997 (Tables A1 and A2) were available for comparison.

Individual length frequencies from the surveillance samples were plotted, on a percentage basis, with the combined port and observer length frequencies from the same month and gear type or the closest month/gear type category available (Fig. A1). No observer samples were used in 1994 so these comparisons are with port samples only. The number of boardings with length frequency data in any category is low and the number of haddock measured per sample is low resulting in spiky length frequencies and high sampling variability. Six comparisons were possible for each of 1994 and 1995, 1 for 1996 and 7 for 1997. Most categories had only 1 surveillance length frequency. Only 3 of 29 surveillance samples had greater than 200 measurements and 17 had less than 100 measurements. The available data show no persistent patterns to indicate a discard problem.

Table A1. Details of fishery officer boardings of fishing vessels on eastern Georges Bank from 1994 to 1997 during which haddock length measurements were taken.

| Reference No. | Boarding Date | Latitude | Longitude | Gear <br> Type | Tonnage Class Code | Number Measured |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| X94-028 | 94/06/03. | 4152 | 6650 | 12 | 3 | 222 |
| X94-034 | 94/07/24. | 4154 | 6546 | 51 | 2 | 90 |
| X94-035 | 94/08/01. | 4202 | 6600 | 19 | 3 | 13 |
| X94-021 | 94/10/05. | 4202 | 6559 | 12 | 5 | 99 |
| X94-052 | 94/10/28. | 4208 | 6615 | 19 | 5 | 13 |
| X94-057 | 94/11/14. | 4211 | 6611 | 19 | 2 | 38 |
| X95-074 | 95/06/18. | 4153 | 6653 | 19 | 2 | 214 |
| X95-026 | 95/07/01. | 4217 | 6640 | 51 | 2 | 81 |
| X95-024 | 95/07/11. | 4209 | 6636 | 19 | 2 | 133 |
| X95-003 | 95/07/31. | 4207 | 6558 | 51 | 2 | 191 |
| X95-005 | 95/08/10. | 4201 | 6606 | 19 | 2 | 32 |
| X95-017 | 95/10/26. | 4207 | 6622 | 19 | 3 | 49 |
| X95-019 | 95/12/07. | 4208 | 6623 | 51 | 3 | 73 |
| X96-011 | 96/08/05. | 4208 | 6630 | 19 | 3 | 255 |
| X96-012 | 96/08/11. | 4209 | 6633 | 19 | 2 | 116 |
| X97-042 | 97/06/09. | 4201 | 6609 | 19 | 3 | 76 |
| X97-043 | 97/06/09. | 4215 | 6620 | 19 | 3 | 157 |
| X97-045 | 97/06/09. | 4210 | 6611 | 19 | 2 | 82 |
| X97-032 | 97/06/26. | 4207 | 6618 | 51 | 3 | 23 |
| X97-025 | 95/06/27. | 4203 | 6605 | 12 | 2 | 33 |
| X97-036 | 97/07/08. | 4209 | 6641 | 19 | 2 | 199 |
| X97-037 | 97/07/08. | 4209 | 6639 | 51 | 1 | 44 |
| X97-039 | 97/07/08. | 4209 | 6639 | 51 | 2 | 25 |
| X97-074 | 9/1/97 | 4206 | 6559 | 19 | 2 | 29 |
| X97-061 | 9/11/97 | 4209 | 6638 | 51 | 2 | 128 |
| X97-053 | 9/28/97 | 4208 | 6639 | 19 | 3 | 104 |
| X97-054 | 9/28/97 | 4208 | 6641 | 19 | 3 | 119 |
| X97-079 | 10/13/97 | 4206 | 6628 | 12 | 3 | 162 |
| X97-076 | 10/14/97 | 4202 | 6600 | 19 | 2 | 79 |

Table A2. Number of boardings by fishery officers of eastern Georges Bank fishing vessels for which haddock length frequecies were taken. Shading indicates that the vessels were tonnage class 4,5 or 6 . All other vessels were tonnage class 1,2 or 3 .

| Year | Gear | $\begin{gathered} \hline \text { Q2 } \\ \text { June } \\ 6 \end{gathered}$ | Q3 |  |  | Q4 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | $\begin{gathered} \text { July } \\ 7 \end{gathered}$ | $\begin{gathered} \text { Aug } \\ 8 \\ \hline \end{gathered}$ | $\begin{gathered} \text { Sept } \\ 9 \end{gathered}$ | $\begin{gathered} \text { Oct } \\ 10 \end{gathered}$ | Nov 11 | $\begin{gathered} \text { Dec } \\ 12 \end{gathered}$ |
| 1994 | 12 | 1 | 1 | 1 |  | 1 | 1 |  |
|  | 51 |  |  |  |  |  |  |  |
|  | 19 |  |  |  |  | 1 |  |  |
| 1995 | 19 | 1 | 1 | 1 |  | 1 | 1 |  |
|  | 51 |  | 2 |  |  |  |  |  |
| 1996 | 19 |  | 2 |  |  |  |  |  |
| 1997 | 19 | 4 | 1 |  | 3 | 2 |  |  |
|  | 51 | 1 | 2 |  | 1 |  |  |  |

Fig. Al. Comparison of 5 Ze haddock length frequencies from port and observer samples with those from fishery officer surveillance boardings during 1994 to 1997. In 1994 port samples only were used. Port and observer samples are combined by month and gear type and compared to individal surveillance samples which show numbers measured in brackets. (OT=otter trawl bottom; $\mathrm{N}=$ =tonnage class $1-3 ;$ OF-tonnage class $4-6 ; \mathrm{s}=$ square mesh; $\mathrm{d}=$ diamond mesh)
A. 1994



B. 1995







B. 1995 (continued)

C. 1996

D. 1997









## Appendix B. Ageing

A subsample of the 1998 DFO spring survey otoliths (N773) plus the remainder of the 1997 DFO spring survey ( N 254 ) otoliths and the Canadian commercial fishery otoliths obtained by port samplers and by the observer program were read by L. Van Eeckhaute (LVE). Numbers of otoliths examined are summarized below:

| Name | Year | Description | Total | Not aged | Ages |
| :--- | :---: | :--- | :---: | :---: | :---: |
| N254 | 1997 | DFO spring survey | 346 | 2 | 344 |
| CGS | 1997 | Canadian commercial fishery-Port Samplers | 781 | 15 | 766 |
| IOP | 1997 | Canadian commercial fishery-Observer Program | 207 | 1 | 206 |
| N773 | 1998 | DFO spring survey | 363 | 9 | 354 |

Intra-reader and inter-reader tests
Within reader tests were conducted for L.Van Eeckhaute using otoliths from the N254 survey $(\mathrm{n}=95)$ and the 1997 commercial fishery ( $\mathrm{n}=100$ ). Between reader tests with the USA reader, N.Munroe, (NM) and the Canadian reader were also completed. Fifty otoliths from the 1997 DFO survey (N254) and 50 from the 1996 NMFS fall survey (9604) were selected. Agreement matrices for the four tests are given in Figs. 1 to 4. A summary of test results follows:

| Source | Description | Test | N | \% agreement |
| :--- | :--- | :--- | :---: | :---: |
| N254 | 1997 DFO spring survey | LVE x LVE | 95 | 95 |
|  |  | LVE x NM | 50 | 63 |
| CGS ‘97 | 1997 Canadian Commercial Fishery | LVE x LVE | 100 | 95 |
| 9604 | 1996 NMFS fall survey | LVE x NM | 50 | 84 |

Intra-reader agreement is high with no bias (Figs. B1 and B3). One of the between reader tests shows low agreement and a bias with the NMFS reader ageing younger than the DFO reader (Fig. B2). The second inter-reader test shows acceptable agreement but there is a slight bias towards older ages by the NMFS reader versus the DFO reader (Fig B4).

Fig. B1. Intra-reader agreement matrix for L.Van Eeckhaute for ageing material from the 1997 DFO spring survey.

| DFO 1997 spring survey (N254) |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Second Reading | First Reading |  |  |  |  |  |  |  |  |  |  |  |  |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 10 | 12 | 15 | Omit | Total |
| 1 | 24 |  |  |  |  |  |  |  |  |  |  |  | 24 |
| 2 |  | 13 | 1 |  |  |  |  |  |  |  |  |  | 14 |
| 3 |  |  | 5 |  |  |  |  |  |  |  |  |  | 5 |
| 4 |  |  |  | 12 | 2 |  |  |  |  |  |  |  | 14 |
| 5 |  |  |  |  | 9 |  |  |  |  |  |  |  | 9 |
| 6 |  |  |  |  | 1 | 10 |  |  |  |  |  |  | 11 |
| 7 |  |  |  |  |  | 1 | 3 |  |  |  |  |  | 4 |
| 8 |  |  |  |  |  |  |  | 2 |  |  |  |  | 2 |
| 10 |  |  |  |  |  |  |  |  | 7 |  |  |  | 7 |
| 12 |  |  |  |  |  |  |  |  |  | 3 |  |  | 3 |
| 15 |  |  |  |  |  |  |  |  |  |  | 1 |  | 1 |
| Omit |  |  |  |  |  |  |  |  |  | 1 |  |  | 1 |
| Total | 24 | 13 | 6 | 12 | 12 | 11 | 3 | 2 | 7 | 4 | 1 | 0 | 95 |
| Percent agreement: 89 agreements out of 94 with 2 ages $=95 \%$ |  |  |  |  |  |  |  |  |  |  |  |  |  |

Fig. B2. Agreement matrix between N.Munroe, the NMFS reader, and L.Van Eeckhaute, the DFO reader for ageing material from the 1997 DFO spring survey.

| DFO 1997 spring survey (N254) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| L.Van <br> Eeckhaute | N.Munroe |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 12 | 15 | omit | Total |
| 1 | 11 | 1 |  |  |  |  |  |  |  |  |  |  |  | 12 |
| 2 |  | 7 |  |  |  |  |  |  |  |  |  |  |  | 7 |
| 3 |  |  | 2 |  |  |  |  |  |  |  |  |  |  | 2 |
| 4 |  |  | 1 | 7 |  |  |  |  |  |  |  |  |  | 8 |
| 5 |  |  |  | 5 | 1 |  |  |  |  |  |  |  |  | 6 |
| 6 |  |  |  | 1 | 3 | 2 |  |  |  |  |  |  |  | 6 |
| 7 |  |  |  |  |  | 1 |  |  |  |  |  |  |  | 1 |
| 8 |  |  |  |  |  |  |  |  |  |  |  |  | 1 | 1 |
| 9 |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| 10 |  |  |  |  |  |  | 1 | 1 | 2 |  |  |  | 1 | 5 |
| 12 |  |  |  |  |  |  | 1 |  |  |  |  |  |  | 1 |
| 15 |  |  |  |  |  |  |  |  |  | 1 |  |  |  | 1 |
| omit |  |  |  |  |  |  |  |  |  |  |  |  |  | 0 |
| Total | 11 | 8 | 3 | 13 | 4 | 3 | 2 | 1 | 2 | 1 | 0 | 0 | 2 | 50 |

Fig. B3. Intra-reader agreement matrix for L.Van Eeckhaute for ageing material from the Canadian commercial fishery samples.

| 1997 Canadian Commercial Samples |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Second <br> Reading | First reading |  |  |  |  |  |  |  |  |  |  |  |
|  | 2 | 3 | 4 | 5 | 6 | 7 | 9 | 10 | 12 | 13 | omit | Total |
| 2 | 12 |  |  |  |  |  |  |  |  |  |  | 12 |
| 3 | 1 | 5 |  |  |  |  |  |  |  |  |  | 6 |
| 4 |  |  | 18 |  |  |  |  |  |  |  |  | 18 |
| 5 |  |  | 1 | 26 | 1 |  |  |  |  |  | 1 | 29 |
| 6 |  |  |  |  | 11 |  |  |  |  |  |  | 11 |
| 7 |  |  |  |  |  | 5 |  |  |  |  |  | 5 |
| 9 |  |  |  |  |  |  |  |  |  |  |  |  |
| 10 |  |  |  |  |  |  | 1 | 9 |  |  |  | 10 |
| 12 |  |  |  |  |  |  |  |  | 5 |  |  | 5 |
| 13 |  |  |  |  |  |  |  |  | 1 |  |  | 1 |
| omit |  |  | 1 |  |  | 1 |  |  |  |  | 1 | 3 |
| Total | 13 | 5 | 20 | 26 | 12 | 6 | 1 | 9 | 6 | 0 | 2 | 100 |
| Percent agreement: 91 agreements out of 96 aged twice $=95 \%$ |  |  |  |  |  |  |  |  |  |  |  |  |

Fig. B4. Agreement matrix for N.Munroe and L.Van Eeckhaute for ageing material from the NMFS 1996 fall survey.

| NMFS 1996 Fall Survey (9604) |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| L. Van | N.Munroe |  |  |  |  |  |  |  |  |  |  |
| Eeckhaute | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | Omit | Total |
| 1 | 4 |  |  |  |  |  |  |  |  |  | 4 |
| 2 |  | 2 |  |  |  |  |  |  |  |  | 2 |
| 3 |  |  | 14 |  |  |  |  |  |  |  | 14 |
| 4 |  |  |  | 13 | 4 |  |  |  |  |  | 17 |
| 5 |  |  |  | 1 | 8 |  |  |  |  |  | 9 |
| 6 |  |  |  |  |  |  | 1 |  |  |  | 1 |
| 7 |  |  |  |  |  | 1 |  |  |  |  | 1 |
| 8 |  |  |  |  |  |  |  |  | 1 |  | 1 |
| 9 |  |  |  |  |  |  |  |  |  |  | 0 |
| Omit |  |  |  | 1 |  |  |  |  |  |  | 1 |
| Total | 4 | 2 | 14 | 15 | 12 | 1 | 1 | 0 | 1 |  | 50 |

Percent agreement: 41 agreements out of 49 aged by both readers $=84 \%$

# Appendix C. Alternative catch-age analyses 

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A catch-age model, employing the same model structure and objective function as used by Gavaris and van Eeckhaute(this document), was implemented using ADModel Builder Software (Otter Consulting Ltd.). This implementation, which assumes a log-normal error structure for the survey catch-at-age data and omits zero observations from the objective function, is termed the $A D A P T$ model. A second implementation of $A D A P T$, based on the same model structure but with a different objective function was also coded. In this case the objective function was based on the assumption of a multinomial distribution for the age-composition data and a log-normal distribution for the total (i.e. summed over ages) abundance in each survey. The objective functions (quantities to be minimized) for the alternate error structure models are:

ADAPT

$$
\sum_{k i j}\left(\ln I_{k j j}-\left(\ln K_{k j}\right)\right)^{2}
$$

multinomial ADAPT

$$
-\sum_{k i j} S_{k i j} \ln P_{k j i}+W \sum_{k i}\left(\ln I_{k i .}-\ln K_{k i}\right)^{2}
$$

where,
$I_{k j}$ is the survey index for survey $k$ in year $i$ for age $j$
$N_{i j}$ is the model estimate of numbers at age $j$ in year $i$

$$
\begin{aligned}
& S_{k i j}=S\left(I_{k i j} / \sum_{j} I_{k i j}\right) \\
& K_{k i j}=N_{i j} \exp \left(q_{k j}\right) \\
& K_{k i=}=\sum_{j} K_{k i j}
\end{aligned}
$$

$$
P_{k j i}=K_{k i j} / K_{k i}
$$

$S$ is a sample size weight
$W$ is the survey catch weight
$t$ is the terminal year of the analysis
The quantities, $S$ and $W$, are fixed inputs to the analysis and the quantities, $q_{k j}$ and $N_{t j}$, are fundamental model parameters estimated through the minimization.

An alternate catch-age model, where the $N$ matrix is calculated moving forward in time, was also implemented. For this model, parameters for the numbers at each age in the first year and the numbers in the first age class for each year are estimated. This frees the assumption that the terminal age fishing mortality is equal to the average fishing mortality for younger ages, but also results in imprecise estimates of absolute stock abundance. Hence, analyses were also conducted where the catchability $(q)$ for ages 3 through 8 are estimated as a single parameter for the Canadian Spring Survēy (i.e. assume flat-top or asymptotic catchability).

Table Cl. Estimates of eastern Georges Bank haddock number-at-age for the beginning of year 1998 from the analysis for the alternative model structures and error models are shown.

| model structure | objective function | numbers ( 000 's) at age |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 |
| standard | log-normal ${ }^{1}$ | 4597 | 12135 | 4158 | 2565 | 3504 | 3866 | 1101 | 148 |
| ADAPT | log-normal | 4603 | 12148 | 4163 | 2569 | 3514 | 3881 | 1107 | 149 |
|  | multinomial | 1766 | 7838 | 2526 | 1600 | 2264 | 2642 | 616 | 144 |
| forward | log-normal | 23731 | 62614 | 22327 | 14853 | 2578 | 34209 | 13580 | 3008 |
| catch-age | multinomial | 7203 | 32289 | 11297 | 8164 | 15945 | 23369 | 8415 | 2378 |
| asymptotic | log-normal | 5136 | 13571 | 4809 | 2999 | 4385 | 4861 | 1583 | 267 |
| selectivity | multinomial | 2080 | 9261 | 3082 | 2015 | 3125 | 3925 | 1066 | 276 |

${ }^{1}$ results reported in Gavaris and van Eeckhaute (1998)
Differences between the numbers-at-age reported by Gavaris and van Eeckhaute (this document) and those produced by the AdModel Builder implementation of $A D A P T$ are small, which supports that the model formulations are the same in the two analyses (Table C1). The estimates of numbers-at-age for the standard ADAPT model, but using the multinomial error structure objective function, are considerably lower than those for the log-normal objective function, particularly for younger ages. The time-series of recruitment estimates from the two analyses, indicates that the stock estimates diverge only in the most recent years (Fig. C1).


Fig. C1. Estimates of recruitment at age 0 from the alternative error structure models for 5Zjm haddock.

Estimates of the catchabilities ( $q$ ) from the "multinomial" analysis are substantially higher for two of the surveys.


Fig. C2. Estimates of survey catchability at age for eastern Georges Bank haddock from the two error structure models.

The estimated numbers-at-age in the terminal year from the forward catch-age model are substantially higher for both forms of the objective function than they are from the ADAPT formulation. The partial recruitment parameters from these runs suggest decreasing partial recruitment at higher ages (i.e. 6 to 8 ). When catchability of ages 3 through 8 in the Canadian spring survey are estimated as a single parameter, the estimates of terminal year numbers are fairly similar to those from the $A D A P T$ analysis. Also, the time-series of recruitment estimates (shown in the following figure) have high coherence between the ADAPT and forward catch-age models (Fig. C3.).
log-normal



Fig. C3. Estimates of recruitment at age 0 for eastern Georges Bank haddock from "forward" and "backward" VPA for both error structure models.

The intent of analysing the eastern Georges Bank haddock data with alternative model formulations and error structure assumptions was to explore the potential uncertainty in stock abundance estimates. The range in terminal stock estimates obtained from these analyses suggest a greater uncertainty than what is estimated for a single model. Further work is required to investigate the robustness of the alternate error structure models.


[^0]:    ${ }^{1}$ Catches of $3 t, 1846 t$ and $46 t$ for Jan., Feb., and Mar., respectively for otter trawlers were excluded because of suspected area misreporting

